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# Statistical Aspects of North Atlantic Basin Tropical Cyclones During the Weather Satellite Era, 1960-2013: Part 2 

Robert M. Wilson

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## TABLE OF CONTENTS

1. INTRODUCTION ..... 1
2. RESULTS AND DISCUSSION ..... 2
2.1 Oceanic Niño Index and Southern Oscillation Index ..... 4
2.2 Atlantic Multidecadal Oscillation Index ..... 7
2.3 Quasi-Biennial Oscillation Index ..... 8
2.4 North Atlantic Oscillation Index ..... 9
2.5 Armagh Surface Air Temperature ..... 12
2.6 Global Land-Ocean Temperature Index and Mauna Loa Carbon Dioxide Index ..... 12
2.7 Correlative Behavior of the Climate-Related Factors ..... 14
2.8 Inferred Statistical Relationships Between Tropical Cyclone Parametric
Values and Selected Climate-Related Factors ..... 24
3. SUMMARY ..... 65
REFERENCES ..... 70

## LIST OF FIGURES

1. Variation of (a) <ONI> and (b) number of months when ENL and LNL conditions prevailed for the interval 1960-2013 ..... 4
2. Variation of <SOI> for the interval 1960-2013 ..... 5
3. Scatterplot of $<$ SOI $>$ versus $<\mathrm{ONI}>$ ..... 7
4. Variation of <AMO> for the interval 1960-2013 ..... 8
5. Variation of <QBO> for the interval 1960-2013 ..... 9
6. Variation of <NAO> CPC for the interval 1960-2013 ..... 10
7. Variation of <NAO> CRU for the interval 1960-2013 ..... 10
8. Scatterplot of < NAO> CRU versus < NAO> CPC ..... 11
9. Variation of <ASAT> for the interval 1960-2013 ..... 12
10. Variation of (a) <GLOTI> and (b) <MLCO2> for the interval 1960-2013 ..... 13
11. Scatterplot of <GLOTI> versus < MLCO2> ..... 14
12. Scatterplot of <NAO> CPC versus <AMO> ..... 17
13. Scatterplot of <NAO> CRU versus <AMO> ..... 18
14. Scatterplot of <ASAT> versus <AMO> ..... 19
15. Scatterplot of <GLOTI> versus <AMO> ..... 20
16. Scatterplot of <ASAT> versus <GLOTI> ..... 21
17. Scatterplot of <AMO> versus < MLCO2> ..... 22
18. Scatterplot of <ASAT> versus <MLCO2> ..... 23
19. Scatterplot of NTC versus <AMO> ..... 28

## LIST OF FIGURES (Continued)

20. Scatterplot of NH versus < AMO > ..... 29
21. Scatterplot of NMH versus < AMO> ..... 30
22. Scatterplot of <LAT> versus <AMO> ..... 31
23. Scatterplot of <LONG> versus < AMO> ..... 32
24. Scatterplot of LP versus < AMO> ..... 33
25. Scatterplot of total ACE versus < AMO> ..... 34
26. Scatterplot of HISACE versus <AMO> ..... 35
27. Scatterplot of total PDI versus < AMO> ..... 36
28. Scatterplot of HISPDI versus < AMO> ..... 37
29. Scatterplot of <NSD> versus <AMO> ..... 38
30. Scatterplot of total NSD versus < AMO> ..... 39
31. Scatterplot of total NHD versus <AMO> ..... 40
32. Scatterplot of total NMHD versus <AMO> ..... 41
33. Scatterplot of NTCA versus <AMO> ..... 42
34. Scatterplots of (a) NTC, (b) NMH, and (c) <LP> versus <SOI> ..... 43
35. Scatterplots of (a) total ACE and (b) LISNSD versus <SOI> ..... 45
36. Scatterplots of (a) total NSD, (b) total NHD, and (c) NTCA versus <SOI> ..... 47
37. Scatterplot of LOS versus <MLCO2> ..... 48
38. Scatterplot of NTC versus < MLCO2> ..... 49
39. Scatterplot of <LAT> versus <MLCO2> ..... 50
40. Scatterplot of total NSD versus <MLCO2> ..... 51

## LIST OF FIGURES (Continued)

41. Scatterplots of (a) total NSD and (b) NTCA versus < GLOTI> ..... 53
42. Scatterplots of (a) LSD and (b) NTC versus <ONI> ..... 55
43. Scatterplots of (a) NH, (b) <NSD> and (c) total NSD versus <NAO> CRU ..... 57
44. Scatterplots of (a) NH and (b) total NSD versus <NAO> CPC ..... 59
45. Scatterplot of NTC versus <ASAT> ..... 60

## LIST OF TABLES

1. Summary of selected climate-related parametric values, 1960-2013 ........................... 2
2. Summary of climate interrelational statistics, 1960-2013 ............................................ 15
3. Summary of tropical cyclone statistics against climate factors (single-variate fits), 1960-201325
4. Comparison of observed and predicted NTC values based on selected fits ................. 62
5. Summary of estimates for the North Atlantic basin tropical cyclone parameters (rounded to nearest whole number) for the 2014 hurricane season based on the means for the interval 1995-2013 for <AMO> and the projected value of <MLCO2> for 201467
6. Summary of estimates for the North Atlantic basin tropical cyclone parameters based on whether the year 2014 is classified as ENY or NENY (rounded to nearest whole number)68

## LIST OF ABBREVIATIONS, ACRONYMS, DESIGNATORS, AND SYMBOLS

| ACE | accumulated cyclone energy |
| :---: | :---: |
| <ACE> | mean seasonal ACE |
| AMO | Atlantic Multidecadal Oscillation (index) |
| <AMO> | mean seasonal AMO |
| ASAT | Armagh surface air temperature |
| <ASAT> | mean seasonal ASAT |
| $\mathrm{CO}_{2}$ | carbon dioxide |
| CPC | Climate Prediction Center |
| CRU | Climate Research Unit |
| DOY | day of year |
| ENL | El Niño-like |
| ENY | El Niño year |
| FSD | first storm day |
| GLOTI | Global Land-Ocean Temperature Index |
| <GLOTI> | mean seasonal GLOTI |
| HISACE | highest individual storm ACE |
| HISPDI | highest individual storm PDI |
| <LAT> | mean seasonal latitude |
| LISNSD | longest individual storm NSD |
| LNL | La Niña-like |

LIST OF ABBREVIATIONS, ACRONYMS, DESIGNATORS, AND SYMBOLS (Continued)

| LNY | La Niña year |
| :---: | :---: |
| <LONG> | mean seasonal longitude |
| LOS | length of season |
| LP | lowest pressure |
| <LP> | mean seasonal LP |
| LSD | last storm day |
| MLCO2 | Mauna Loa $\mathrm{CO}_{2}$ (index) |
| <MLCO2> | mean seasonal MLCO2 |
| NAO | North Atlantic Oscillation (index) |
| < NAO> | mean seasonal NAO |
| NENM | number of ENL months |
| NENY | non-El Niño year |
| NH | number of hurricanes |
| NHD | number of hurricane days |
| NLNM | number of LNL months |
| NMH | number of major hurricanes |
| NMHD | number of major hurricane days |
| NNM | number of neutral months |
| NSD | number of storm days |
| <NSD> | mean seasonal NSD |
| NTC | number of tropical cyclones |

LIST OF ABBREVIATIONS, ACRONYMS, DESIGNATORS, AND SYMBOLS (Continued)

| NTCA | Net Tropical Cyclone Activity |
| :--- | :--- |
| NUSLFH | number of United States land-falling hurricanes |
| <NUSLFH $>$ | mean seasonal NUSLFH |
| ONI | Oceanic Niño Index |
| <ONI $>$ | mean yearly ONI |
| PDI | meaner Dissipation Index |
| <PDI> | peak wind speed |
| PWS | mean seasonal PWS |
| <PWS $>$ | Quasi-Biennial Oscillation (index) |
| QBO | mean seasonal QBO |
| $<$ QBO $>$ | mean yearly SOI |
| SOI | sea surface temperature Oscillation Index |
| <SOI $>$ | thermohaline circulation |
| SST | Technical Publication |
| THC | TP |

## NOMENCLATURE

| $a$ | $y$-intercept |
| :---: | :---: |
| $b$ | slope |
| cl | confidence level |
| $n a$ | number above median |
| $n b$ | number below median |
| $n r a$ | number of positive runs |
| $P$ | probability |
| $R_{y 12}$ | sample coefficient of multiple correlation |
| $r$ | coefficient of correlation |
| $r^{2}$ | coefficient of determination |
| $S_{y 12}$ | sample coefficient of multiple standard error of estimate |
| $s d$ | standard deviation |
| se | standard error of estimate |
| $t$ | $t$-statistic for independent samples |
| $x$ | independent variable |
| $y$ | dependent variable |
| $z$ | normal deviate for the sample results |

## TECHNICAL PUBLICATION

# STATISTICAL ASPECTS OF NORTH ATLANTIC BASIN TROPICAL CYCLONES DURING THE WEATHER SATELLITE ERA, 1960-2013: PART 2 

## 1. INTRODUCTION

This Technical Publication (TP) is part 2 of a two-part study of the North Atlantic basin tropical cyclones that occurred during the weather satellite era, 1960-2013. In particular, this TP examines the inferred statistical relationships between 25 tropical cyclone parameters and 9 specific climate-related factors, including the (1) Oceanic Niño Index (ONI), (2) Southern Oscillation Index (SOI), (3) Atlantic Multidecadal Oscillation (AMO) index, (4) Quasi-Biennial Oscillation (QBO) index, (5) North Atlantic Oscillation (NAO) index of the Climate Prediction Center (CPC), (6) NAO index of the Climate Research Unit (CRU), (7) Armagh surface air temperature (ASAT), (8) Global Land-Ocean Temperature Index (GLOTI), and (9) Mauna Loa carbon dioxide ( $\mathrm{CO}_{2}$ ) (MLCO2) index. Part 1 of this two-part study examined the statistical aspects of the 25 tropical cyclone parameters (e.g., frequencies, peak wind speed (PWS), accumulated cyclone energy (ACE), etc.) and provided the results of statistical testing (i.e., runs-testing, the $t$-statistic for independent samples, and Poisson distributions). Also, the study gave predictions for the frequencies of the number of tropical cyclones (NTC), number of hurricanes (NH), number of major hurricanes (NMH), and number of United States land-falling hurricanes (NUSLFH) expected for the 2014 season, based on the statistics of the overall interval 1960-2013, the subinterval 1995-2013, and whether the year 2014 would be either an El Niño year (ENY) or a non-El Niño year (NENY). ${ }^{1}$

## 2. RESULTS AND DISCUSSION

Table 1 provides a detailed listing of the yearly values for each of the climate-related factors that will be discussed in the following subsections. Also given in the table are the extreme monthly values per year (highs and lows) for ONI and SOI, as well as the statistics for the overall interval 1960-2013, the two subintervals 1960-1994 and 1995-2013, and the results of statistical testing.

Table 1. Summary of selected climate-related parametric values, 1960-2013.

| Year | NENM | NNM | NLNM | <ONI> | ONI <br> (High/Low) | <SOl> | SOI <br> (High/Low) | <AMO> | <QBO> | <NAO> CPC | <NAO> CRU | <ASAT> | <GLOTI> | <MLCO2> |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1960 | - | 12 | - | 0.00 | 0.32/-0.36 | 3.8 | 7.8/-2.3 | 0.237 | -7.76 | -0.56 | -0.30 | 9.44 | -0.04 | 316.91 |
| 1961 | 1 | 11 | - | 0.01 | 0.54/-0.38 | 0.8 | 13.8/-20.9 | 0.101 | 4.42 | 0.01 | 1.05 | 9.58 | 0.06 | 317.64 |
| 1962 | - | 11 | 1 | -0.24 | 0.01/-0.52 | 5.4 | 17.0/-1.4 | 0.074 | $-7.73$ | -0.38 | -0.13 | 8.76 | 0.05 | 318.45 |
| 1963 | 6 | 5 | 1 | 0.64 | 1.43/-0.50 | -2.0 | 9.4/-12.9 | 0.006 | -8.75 | -0.56 | -0.39 | 8.57 | 0.07 | 318.99 |
| 1964 | 2 | 2 | 8 | -0.38 | 1.05/-0.86 | 6.3 | 14.3/-4.0 | -0.094 | 2.38 | -0.18 | 0.24 | 9.49 | -0.20 | 319.62 |
| 1965 | 8 | 3 | 1 | 0.91 | 2.04/-0.58 | -16.8 | 2.9/-22.6 | -0.158 | -11.93 | -0.22 | -0.23 | 8.82 | -0.11 | 320.04 |
| 1966 | 4 | 8 | - | 0.37 | 1.29/-0.25 | -4.2 | 4.0/-13.9 | 0.006 | 1.29 | -0.49 | -0.22 | 9.38 | -0.05 | 321.38 |
| 1967 | - | 11 | 1 | -0.23 | 0.12/-0.58 | 3.2 | 14.6/-5.5 | -0.096 | 0.93 | 0.34 | 0.56 | 9.40 | -0.01 | 322.16 |
| 1968 | 3 | 6 | 3 | 0.07 | 0.97/-0.87 | 3.0 | 14.7/-3.4 | -0.165 | -14.41 | -1.04 | -0.62 | 9.32 | -0.05 | 323.04 |
| 1969 | 11 | 1 | - | 0.83 | 1.29/0.26 | -5.4 | 3.71-13.5 | 0.011 | 4.81 | -0.19 | -0.44 | 8.93 | 0.07 | 324.62 |
| 1970 | 1 | 5 | 6 | -0.23 | 0.61/-0.98 | 3.9 | 19.7/-10.7 | -0.102 | -13.76 | -0.35 | 0.18 | 9.29 | 0.03 | 325.68 |
| 1971 | - | - | 12 | -0.90 | -0.63/-1.42 | 11.0 | 22.6/2.1 | -0.311 | 4.80 | -0.04 | -0.55 | 9.72 | -0.06 | 326.32 |
| 1972 | 8 | 3 | 1 | 0.94 | 2.27/-0.67 | -7.4 | 8.2/-18.6 | -0.353 | -8.10 | 0.49 | -0.04 | 8.74 | 0.01 | 327.45 |
| 1973 | 3 | 1 | 8 | -0.57 | 1.91/-2.08 | 7.3 | 31.6/-13.5 | -0.215 | 4.40 | -0.16 | -0.09 | 9.33 | 0.15 | 329.68 |
| 1974 | - | 2 | 10 | -0.84 | -0.30/-1.99 | 9.9 | 20.8/-1.4 | -0.420 | -13.42 | 0.11 | 0.59 | 8.94 | -0.07 | 330.18 |
| 1975 | - | 2 | 10 | -1.04 | -0.33/-1.76 | 13.6 | 22.5/-4.9 | -0.298 | 1.25 | -0.16 | 0.05 | 9.70 | -0.01 | 331.08 |
| 1976 | 3 | 5 | 4 | -0.08 | 0.88/-1.78 | 1.1 | 13.2/-13.0 | -0.363 | 0.85 | 0.18 | -0.07 | 9.34 | -0.12 | 332.05 |
| 1977 | 6 | 6 | - | 0.52 | 0.83/0.08 | -9.9 | 7.7/-17.7 | -0.189 | -11.27 | -0.42 | -0.21 | 8.92 | 0.15 | 333.78 |
| 1978 | 1 | 11 | - | -0.08 | 0.81/-0.47 | -1.7 | 16.3/-24.4 | -0.179 | 6.81 | 0.28 | 0.21 | 9.21 | 0.06 | 335.41 |
| 1979 | 3 | 9 | - | 0.23 | 0.56/-0.10 | -1.9 | 6.71-8.2 | -0.110 | -13.16 | 0.05 | 0.19 | 8.35 | 0.12 | 336.78 |
| 1980 | 2 | 10 | - | 0.21 | 0.59/-0.09 | -3.1 | 3.2/-12.9 | -0.018 | 5.27 | -0.55 | -0.37 | 9.11 | 0.22 | 338.68 |
| 1981 | - | 9 | 3 | -0.35 | -0.12/-0.69 | 1.8 | 11.5/-16.6 | -0.075 | -1.91 | -0.27 | -0.09 | 9.09 | 0.28 | 340.10 |
| 1982 | 8 | 4 | - | 0.91 | 2.31/-0.11 | -13.1 | 9.4/-31.1 | -0.211 | -5.74 | 0.36 | 0.67 | 9.44 | 0.09 | 341.44 |
| 1983 | 6 | 2 | 4 | 0.42 | 2.25/-1.09 | -8.3 | 9.9/-33.3 | -0.069 | 1.15 | 0.30 | 0.34 | 9.77 | 0.27 | 343.03 |
| 1984 | - | 7 | 5 | -0.48 | -0.08/-1.23 | -0.1 | 5.8/-8.7 | -0.206 | -17.02 | 0.23 | 0.26 | 9.29 | 0.12 | 344.58 |
| 1985 | - | 5 | 7 | -0.61 | -0.26/-1.09 | 0.9 | 14.4/-9.6 | -0.265 | 9.63 | -0.28 | -0.47 | 8.70 | 0.08 | 346.04 |
| 1986 | 4 | 7 | 1 | 0.26 | 1.17/-0.52 | -10.7 | 10.7/-13.9 | -0.273 | -0.98 | 0.43 | 0.56 | 8.57 | 0.15 | 347.39 |
| 1987 | 12 | - | - | 1.29 | 1.72/0.93 | -13.1 | -1.4/-24.4 | 0.069 | -6.25 | -0.22 | -0.51 | 9.07 | 0.28 | 349.16 |
| 1988 | 1 | 3 | 8 | -0.82 | 0.94/-1.89 | 7.8 | 21.0/-5.0 | -0.002 | 1.56 | -0.08 | -0.32 | 9.66 | 0.34 | 351.56 |
| 1989 | - | 7 | 5 | -0.63 | 0.01/-1.85 | 6.8 | 21.0/-6.3 | -0.080 | -10.70 | 0.68 | 0.57 | 10.07 | 0.24 | 353.07 |
| 1990 | - | 12 | - | 0.27 | 0.41/0.07 | -2.2 | 13.1/-17.3 | -0.035 | 7.38 | 0.61 | 1.23 | 9.94 | 0.39 | 354.35 |
| 1991 | 7 | 5 | - | 0.65 | 1.55/0.02 | -8.8 | 5.1/-19.3 | -0.129 | -3.24 | 0.25 | 0.34 | 9.42 | 0.38 | 355.57 |

Table 1. Summary of selected climate-related parametric values, 1960-2013 (Continued).

| Year | NENM | NNM | NLNM | <ONI> | ONI (High/Low) | <SOI> | SOI (High/Low) | <AMO> | <QBO> | <NAO> CPC | <NAO> CRU | <ASAT> | <GLOTl> | <MLCO2> |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1992 | 6 | 6 | - | 0.58 | 1.65/-0.45 | -10.4 | 1.4/-25.4 | -0.216 | -7.61 | 0.54 | 1.11 | 9.45 | 0.19 | 356.38 |
| 1993 | 2 | 10 | - | 0.31 | 0.81/0.04 | -9.5 | 1.6/-21.1 | -0.207 | 3.90 | 0.16 | 0.12 | 9.27 | 0.20 | 357.07 |
| 1994 | 3 | 9 | - | 0.48 | 1.26/0.05 | -11.9 | 0.6/-22.8 | -0.173 | -13.67 | 0.54 | 0.51 | 9.38 | 0.28 | 358.82 |
| 1995 | 3 | 4 | 5 | -0.08 | 1.04/-0.96 | -1.8 | 4.2/-16.2 | 0.139 | 7.82 | -0.09 | -0.61 | 10.23 | 0.43 | 360.80 |
| 1996 | - | 8 | 4 | -0.43 | -0.18/-0.85 | 14.0 | 13.9/-0.1 | -0.054 | -16.06 | -0.25 | -1.01 | 9.23 | 0.33 | 362.59 |
| 1997 | 8 | 3 | 1 | 1.11 | 2.38/-0.58 | -11.7 | 13.3/-24.1 | 0.056 | 7.86 | -0.20 | -0.18 | 10.32 | 0.45 | 363.71 |
| 1998 | 5 | 1 | 6 | -0.03 | 2.30/-1.57 | -1.1 | 14.61-28.5 | 0.377 | -12.83 | -0.54 | 0.26 | 10.09 | 0.61 | 366.65 |
| 1999 | - | - | 12 | -1.19 | -0.87/-1.72 | 16.3 | 18.5/-0.4 | 0.123 | 10.17 | 0.33 | 0.05 | 10.18 | 0.40 | 368.33 |
| 2000 | - | 1 | 11 | -0.88 | -0.46/-1.78 | 7.8 | 22.4/-5.5 | 0.034 | -6.49 | 0.16 | 0.04 | 9.93 | 0.40 | 369.52 |
| 2001 | - | 10 | 2 | -0.26 | 0.10/-0.74 | 0.5 | 11.9/-9.1 | 0.118 | -15.24 | -0.24 | -0.45 | 9.57 | 0.52 | 371.13 |
| 2002 | 8 | 4 | - | 0.65 | 1.39/-0.18 | -6.1 | 7.71-14.6 | 0.071 | 8.72 | 0.00 | -0.04 | 10.20 | 0.61 | 373.22 |
| 2003 | 4 | 8 | - | 0.34 | 1.01/-0.49 | -3.1 | 9.8/-12.0 | 0.237 | -13.81 | 0.03 | -0.16 | 10.03 | 0.60 | 375.77 |
| 2004 | 6 | 6 | - | 0.44 | 0.76/0.08 | -4.8 | 13.1/-15.4 | 0.213 | 6.60 | 0.16 | 0.01 | 10.21 | 0.52 | 377.49 |
| 2005 | 1 | 9 | 2 | 0.08 | 0.64/-0.88 | -3.6 | 10.9-14.5 | 0.298 | -17.13 | -0.31 | -0.25 | 10.25 | 0.65 | 379.80 |
| 2006 | 4 | 5 | 3 | 0.13 | 1.18/-1.01 | -1.9 | 15.2/-15.9 | 0.273 | 4.47 | -0.31 | -0.20 | 10.42 | 0.59 | 381.90 |
| 2007 | 1 | 6 | 5 | -0.45 | 0.72/-1.30 | -0.6 | 14.4/-7.3 | 0.148 | -15.91 | 0.11 | -0.38 | 10.60 | 0.62 | 383.76 |
| 2008 | - | 5 | 7 | -0.69 | -0.12--1.58 | 10.2 | 21.3/-4.3 | 0.146 | 6.98 | -0.45 | -0.73 | 9.76 | 0.49 | 385.59 |
| 2009 | 6 | 3 | 3 | 0.37 | 1.78/-0.94 | -0.2 | 14.8/-14.7 | 0.047 | -3.49 | -0.32 | -0.42 | 9.84 | 0.59 | 387.37 |
| 2010 | 4 | 2 | 6 | -0.33 | 1.57/-1.53 | 9.8 | 27.1/-14.5 | 0.358 | -7.95 | -1.29 | -2.19 | 8.72 | 0.66 | 389.85 |
| 2011 | - | 4 | 8 | -0.73 | -0.05/-1.59 | 13.3 | 25.1/0.2 | 0.110 | 1.96 | 0.20 | 0.64 | 10.28 | 0.55 | 391.63 |
| 2012 | 2 | 8 | 2 | -0.06 | 0.75/-0.84 | -0.8 | 9.4/-10.4 | 0.222 | -20.79 | -0.53 | -0.63 | 9.70 | 0.57 | 393.82 |
| 2013 | - | 10 | 2 | -0.34 | 0.00/-0.78 | 4.0 | 13.9/-3.6 | 0.176 | 9.06 | 0.15 | 0.59 | 9.84 | 0.60 | 396.48 |
| 1960-2013 ( $n=54$ ) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| sum |  |  | 178 |  |  |  |  |  |  |  |  |  |  |  |
| mean | 3 | 6 | 3 | - | 0.81/-0.80 | -0.3 | 12.6/-12.8 | -0.026 | -3.38 | -0.07 | -0.04 | 9.50 | 0.25 | 350.52 |
| sd | 3 | 3 | 4 | 0.59 | 0.84/0.68 | 7.9 | 7.218 .5 | 0.193 | 8.88 | 0.40 | 0.57 | 0.54 | 0.25 | 23.90 |
| high | - | - | - | 1.29 | 2.38/0.93 | 16.3 | 31.6/2.1 | 0.377 | 10.17 | 0.68 | 1.11 | 10.60 | 0.66 | 396.48 |
| low | - | - | - | -1.19 | -0.87/-2.08 | -16.8 | -1.4/-33.3 | -0.420 | -20.79 | -1.29 | -2.19 | 8.35 | -0.20 | 316.91 |
| med | - | - | - | -0.02 | 0.81/-0.74 | -0.7 | 13.2-13.0 | -0.027 | -1.91 | -0.08 | -0.07 | 9.44 | 0.24 | 349.16 |
| na | - | - | - | 30 | - | 28 | - |  | 27 | 27 | 27 | 28 | 27 | 27 |
| nb | - | - | - | 24 | - | 26 | - |  | 27 | 27 | 27 | 26 | 27 | 27 |
| nra | - | - | - | 11 | - | 12 | - | 7 | 23 | 14 | 15 | 10 | 4 | 1 |
| $z$ | - | - | - | -1.57 | - | -1.08 | - |  | 4.86 | - | 0.54 | 2.16 | -5.40 | -7.02 |
| 1960-1994 ( $n=35$ ) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| sum | 111 |  | 99 |  |  |  |  |  |  |  |  |  |  |  |
| mean | 3 | 6 | 3 | 0.07 | 0.85/-0.68 | -1.5 | 11.4/-13.7 | -0.129 | -3.33 | -0.02 | 0.11 | 9.24 | 0.10 | 335.67 |
| sd | 3 | 4 | 4 | 0.60 | 0.81/0.73 | 7.8 | 7.6/8.8 | 0.144 | 7.60 | 0.40 | 0.48 | 0.40 | 0.15 | 13.42 |
| 1995-2013 ( $n=19$ ) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| sum | 52 | 97 | 79 |  |  |  |  |  |  |  |  |  |  |  |
| mean | 3 | 5 | 4 | -0.12 | 0.73/-1.01 | 2.1 | 14.8-11.1 | 0.163 | -3.48 | -0.18 | -0.30 | 9.97 | 0.54 | 377.86 |
| sd | 3 | 3 | 4 | 0.56 | 0.91/0.53 | 7.7 | 5.9/7.8 | 0.113 | 11.09 | 0.38 | 0.62 | 0.44 | 0.09 | 11.11 |
| $t$ test (1960-1994) and (1995-2013) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $t$ | - | 1.0 |  | 1.14 | 0.5/1.7 | -1.6 | -1.7/-1.1 | -7.642 | 0.06 | 1.43 | 2.70 | -6.18 | -11.67 | -11.69 |

### 2.1 Oceanic Niño Index and Southern Oscillation Index

Figure 1 displays the variation of (a) the mean yearly ONI (<ONI $>$ ) and (b) the number of months during the year when El Niño-like (ENL) and La Niña-like (LNL) conditions prevailed, based on the ONI monthly values for the overall interval 1960-2013. The $<\mathrm{ONI}>$ is simply the average of the 12 monthly values (January-December) contained in the listing of ONI values. ${ }^{2,3}$ The monthly values of the ONI are the anomalous deviations from the climate-adjusted monthly averages of the sea surface temperature (SST) in the Niño 3.4 region of the Pacific Ocean (i.e., the region bounded by $\pm 5^{\circ}$ latitude about the equator and $\pm 25^{\circ}$ longitude about long. $145^{\circ} \mathrm{W}$.) based on centered 30 -year base periods every 5 years and calculated by the National Oceanic and Atmospheric Administration/National Weather Service/CPC. The number of ENL (thick line) and LNL (thin line) months, as used in this study, is simply the number of months during the year when ONI $\geq 0.5^{\circ} \mathrm{C}$ and $\leq-0.5^{\circ} \mathrm{C}$, respectively. As an example, for the year $2013,<\mathrm{ONI}>=-0.34{ }^{\circ} \mathrm{C}$ and the number of ENL months (NENM) was 0 , the number of neutral months (NNM) was 10, and the number of LNL months (NLNM) was 2 (January and February).



Figure 1. Variation of (a) <ONI> and (b) number of months when ENL and LNL conditions prevailed for the interval 1960-2013.

The mean $<\mathrm{ONI}>$ for the overall interval $1960-2013$ is $0.00^{\circ} \mathrm{C}$, having a standard deviation $s d=0.59{ }^{\circ} \mathrm{C}$, median $=-0.02^{\circ} \mathrm{C}$, and extremes of $1.29^{\circ} \mathrm{C}$ in 1987 (an ENY) and $-1.19^{\circ} \mathrm{C}$ in 1999 (a NENY, best described as a La Niña year (LNY)). Positive-valued <ONI> occurred in 27 years and negative-valued <ONI> also occurred in 27 years. Runs-testing suggests that <ONI> varies randomly, having a normal deviate for the sample results $z=-1.57$. Likewise, a comparison of the means for the two subintervals 1960-1994 and 1995-2013 suggests that the $t$-statistic for independent samples measures 1.14 , inferring that the difference in the means for the two subintervals is not statistically important (even though the mean for the more recent subinterval is about $0.19^{\circ} \mathrm{C}$ cooler than the mean for the earlier subinterval). The large positive spikes reflect ENL warm conditions synonymous with the occurrence of traditional El Niño events, and the large negative dips reflect LNL cool conditions synonymous with the occurrence of traditional La Niña events (i.e., at least 5 consecutive months with monthly ONI values meeting the conventional El Niño or La Niña definitions).

Figure 2 shows the variation of the mean yearly SOI ( $<\mathrm{SOI}>$ ). Like the ONI, the SOI gives an indication as to the development and intensity of El Niño and La Niña events. It is calculated by the Australian Bureau of Meteorology as the pressure difference between Tahiti, French Polynesia and Darwin, Australia, based on means and standard deviations calculated over the period 1933 to 1992 inclusive. ${ }^{4,5}$ Sustained negative monthly values of the SOI below about -8 often indicate the occurrence of an El Niño event, while sustained positive monthly values of the SOI above 8 often indicate the occurrence of a La Niña event; hence, the SOI and ONI vary inversely, with positive (negative) SOI being associated with negative (positive) ONI.


Figure 2. Variation of $\langle$ SOI $\rangle$ for the interval 1960-2013.

The mean $<$ SOI $>$ for the overall interval 1960-2013 is -0.3 , having $s d=7.9$, a median $=-0.7$, and extremes of 16.3 in 1999 (an LNY) and -16.8 in 1965 (an ENY). For the year 2013, <SOI> measured 4.0. Positive-valued $<$ SOI $>$ occurred in 24 years, and negative-valued $<$ SOI $>$ occurred in 30 years. Runs-testing suggests that $<$ SOI $>$ varies randomly having $z=-1.08$, and a comparison of the means for the two subintervals 1960-1994 and 1995-2013 yields $t=-1.6$, inferring that the difference in the means is not statistically important (though to the eye, the observed variation hints of being cyclic, with warm ENL conditions prevailing in the late 1970s to the late 2000s and cooler LNL conditions prevailing in the early 1960s and now after about 2007).

Figure 3 depicts the scatterplot of <SOI> versus <ONI>. Plainly, the two parameters vary inversely (the diagonal line) as $y=-0.238-12.004 x$, where $y$ is $<$ SOI $\rangle$ and $x$ is $\langle\mathrm{ONI}\rangle$, having a coefficient of correlation $r=-0.896$, a coefficient of determination $r^{2}=0.802$ (meaning that about $80 \%$ of the variance in $<$ SOI $>$ can be explained by the variation in $<\mathrm{ONI}>$ alone, and vice versa), a standard error of estimate $s e=3.530$, and a confidence level $c l » 99.9 \%$. Identified in the scatterplot are the years of parametric extremes for <SOI> (1965 and 1999) and <ONI> (1987 and 1999). Also shown in the scatterplot is the result of Fisher's exact test for $2 \times 2$ contingency tables (the vertical and horizontal lines are the parametric medians), which indicates that the probability of obtaining the observed result, or one more suggestive of a departure from independence (chance), is $P=1.6 \times 10^{-5 \%} \%$. Hence, if the year 2014 happens to be an ENY, one clearly should expect the $<\mathrm{ONI}>$ and $<\mathrm{SOI}>$ values to lie in the lower-right quadrant of the scatterplot; on the other hand, if the year 2014 happens to be a NENY, one should expect the <ONI> and <SOI> values to lie in the upper-left quadrant of the scatterplot.


Figure 3. Scatterplot of $<\mathrm{SOI}\rangle$ versus $<\mathrm{ONI}\rangle$.

### 2.2 Atlantic Multidecadal Oscillation Index

Figure 4 displays the yearly variation of the mean seasonal AMO (<AMO>) for the overall interval 1960-2013. The AMO is defined as a fluctuation in the SST in the North Atlantic Ocean between the equator and lat. $70^{\circ} \mathrm{N} .{ }^{6-11}$ It appears to have a cycle length or period of about 70 years, fluctuating between warm (positive) and cool (negative) phases, with this pattern possibly being associated with variations in the strength of the Atlantic thermohaline circulation (THC), a density-driven, global circulation pattern that involves the movement of warm salty equatorial surface waters to higher latitudes and the subsequent cooling and sinking of these waters into the deep ocean (also called the Atlantic Meridional Overturning Circulation). As such, the warm phase of the AMO seems to represent faster THC, while the cool phase seems to represent slower THC.


Figure 4. Variation of <AMO> for the interval 1960-2013.

The mean <AMO> measures $-0.026^{\circ} \mathrm{C}$ for the overall interval 1960-2013, having $s d=0.193{ }^{\circ} \mathrm{C}$ and a median $=-0.054^{\circ} \mathrm{C}$. Runs-testing suggests that the variation of $\langle\mathrm{AMO}\rangle$ is nonrandom, having $z=-3.78$ and $c l>99.9 \%$. Comparison of the means for the two subintervals shows that the mean for the more recent subinterval is nearly $0.3^{\circ} \mathrm{C}$ warmer than the mean for the earlier subinterval, with $t=-7.64$ and $\mathrm{cl}>99.9 \%$. During the overall interval 1960-2013, there have been 25 warm years and 29 cool years. Inspection of figure 4 indicates that the $<A M O>$ switched from the warm phase to the cool phase about the mid-1960s, remaining in the cool phase until about the mid-1990s, when it switched back again to the warm phase where it has remained for the about the last 20 years. The peak yearly <AMO> occurred in $1998\left(0.377^{\circ} \mathrm{C}\right)$, although another peak occurred in 2010 $\left(0.358^{\circ} \mathrm{C}\right)$. One anticipates that the warm phase will continue for at least several years or more. (For the year 2013, <AMO> measured $0.176^{\circ} \mathrm{C}$.)

### 2.3 Quasi-Biennial Oscillation Index

The QBO refers to the quasi-periodic oscillation of the equatorial zonal winds (in $\mathrm{ms}^{-1}$ ) in the tropical stratosphere having a mean period of about 28 months (range of 20 to 36 months) that was first discovered in the 1950s. ${ }^{12-17}$ The winds alternate between the stronger easterlies (negative values) and the somewhat weaker westerlies (positive values), first developing at the top of the lower stratosphere and then slowly propagating downwards at the rate of about 1 km per month until dissipating in the tropical tropopause.

Figure 5 depicts the yearly variation of the mean seasonal $\mathrm{QBO}(<\mathrm{QBO}\rangle)$ for the overall interval 1960-2013. Its mean measures about -3.38 , having $s d=8.88$ and a median $=-1.91$. Runstesting reveals, as expected, that the variation of $<\mathrm{QBO}>$ values is nonrandom ( $z=4.86, c l>99.9^{\circ} \%$ ), although the comparison of the means for the two subintervals suggests that the difference in the means is not statistically important $(t=0.06)$. Clearly, the easterlies (mean $=-10.59)$ are about twice as strong as the westerlies (mean $=4.98$ ). The strongest easterly occurred in $2012(-20.79)$, while the strongest westerly occurred in 1999 (10.17). (For the year 2013, <QBO> measured 9.06. Because back-to-back positive-valued <QBO> has only occurred twice (1966-1967 and 1975-1976) during the overall interval 1960-2013, one strongly suspects that the year 2014 will be of negative value. For the 10 negative-valued years in the subinterval 1995-2013, it has averaged about -13 , having $s d=5.4$ and extremes of -3.49 and -20.79 .


Figure 5. Variation of $\langle\mathrm{QBO}>$ for the interval 1960-2013.

### 2.4 North Atlantic Oscillation Index

Like the SOI, the NAO is one based on changes in surface air pressure between two widely separated locations (typically, Iceland and a location in the subtropical Atlantic basin: the Azores, Portugal or Gibraltar). Also, like the SOI, the NAO fluctuates between positive and negative values, where the positive phase is associated with a stronger-than-usual subtropical high pressure center and a deeper-than-usual Icelandic low, and the negative phase is associated with the opposite behavior. ${ }^{18-21}$

Figure 6 shows the yearly variation of the mean seasonal NAO ( $<\mathrm{NAO}>$ ) as computed by the CPC for the overall interval 1960-2013. Its mean measures -0.07 , having $s d=0.04$ and a median $=-0.08$. Runs-testing reveals that the variation of $\langle\mathrm{NAO}\rangle \mathrm{CPC}$ is random $(z=0.00)$, and comparison of the means for the two subintervals suggests that the difference in the means is not statistically important ( $t=1.43$ ). The strongest positive yearly value occurred in 1989 (0.68), and the strongest negative yearly value occurred in $2010(-1.29)$, although the year 1968 also had a strong negative value as well ( -1.04 ). (For the year 2013, <NAO $>$ CPC measured 0.15.)


Figure 6. Variation of $\langle\mathrm{NAO}\rangle \mathrm{CPC}$ for the interval 1960-2013.

Figure 7 depicts the yearly variation of the $<\mathrm{NAO}>$ as computed by the CRU for the overall interval 1960-2013. Its mean measures -0.04 , having $s d=0.57$ and a median $=-0.07$. Like $<\mathrm{NAO}>$ CPC , runs-testing of the $<\mathrm{NAO}>\mathrm{CRU}$ values suggests that its variation appears random $(z=0.54)$, but unlike $<\mathrm{NAO}>\mathrm{CPC}$, comparison of the means for the two subintervals suggests that the difference in the means is statistically important ( $t=2.70, \mathrm{cl}>99 \%$ ). The strongest positive yearly value occurred in 1990 (1.23), although there have been at least two other strong positive peaks ( 1.05 in 1961 and 1.11 in 1992), and the strongest negative value occurred in $2010(-2.19)$, which is also the strongest negative value ever recorded in the nearly 190-year record. (For the year 2013, <NAO> CRU measured 0.59.)


Figure 7. Variation of <NAO> CRU for the interval 1960-2013.

Figure 8 displays the scatterplot of $<\mathrm{NAO}>\mathrm{CRU}$ versus $<\mathrm{NAO}>\mathrm{CPC}$. The inferred linear regression is $y=0.044+1.079 x$, having $r=0.758, r^{2}=0.574$, $s e=0.372$, and $c l>99.9 \%$, where $y$ is $<$ NAO $>$ CRU and $x$ is <NAO $>$ CPC. Based on Fisher's exact test for $2 \times 2$ contingency tables, the probability of obtaining the observed result, or one more suggestive a of a departure from independence (chance), is $P=1.6 \times 10^{-5 \%}$


Figure 8. Scatterplot of $<\mathrm{NAO}>\mathrm{CRU}$ versus $<\mathrm{NAO}>\mathrm{CPC}$.

### 2.5 Armagh Surface Air Temperature

The Armagh Observatory in Northern Ireland (lat. $54^{\circ} 21.2^{\prime} \mathrm{N} .$, long. $6^{\circ} 38.9^{\prime} \mathrm{W}$.) is situated about 64 m above mean sea level at the top of a small hill in an estate of natural woodland and parkland that measures about 7 ha . Mean surface air temperature has been continuously measured there since 1844 using maximum and minimum thermometers. ${ }^{22-30}$

Figure 9 depicts the yearly variation of mean seasonal ASAT (<ASAT>) for the overall interval 1960-2013. Its mean measures $9.50{ }^{\circ} \mathrm{C}$, having $s d=0.54^{\circ} \mathrm{C}$ and a median $=9.44^{\circ} \mathrm{C}$. Runstesting suggests that the variation of the $<$ ASAT $>$ values appears nonrandom ( $z=2.16, c l>95 \%$ ), and a comparison of the means for the two subintervals suggests that the difference in the means is statistically important ( $t=-6.18$, $c l>99.9 \%$ ). The highest (warmest) yearly value of $<$ ASAT $>$ occurred in $2007\left(10.60{ }^{\circ} \mathrm{C}\right)$, and the lowest (coolest) yearly value of <ASAT> occurred in $1979\left(8.35^{\circ} \mathrm{C}\right)$. For the more recent subinterval 1995-2013, only the years 1996 and 2010 had <ASAT> lower than the median ( 9.23 and $8.72{ }^{\circ} \mathrm{C}$, respectively). (For the year 2013, $<$ ASAT $>$ measured $9.84{ }^{\circ} \mathrm{C}$.)


Figure 9. Variation of <ASAT> for the interval 1960-2013.

### 2.6 Global Land-Ocean Temperature Index and Mauna Loa Carbon Dioxide Index

Figure 10 shows the yearly variation of (a) mean seasonal GLOTI (<GLOTI $>$ ) and (b) mean seasonal MLCO2 (<MLCO2>) for the overall interval 1960-2013. The GLOTI is a measure of the anomaly in global land-ocean temperatures relative to the interval 1951-1980, where the data are taken from the Global Historical Climate Network, version 3, using elimination of outliers and homogeneity adjustment. ${ }^{27,31-34}$ The MLCO2 index is a measure of the amount of atmospheric concentration of $\mathrm{CO}_{2}$ as measured from the Mauna Loa Observatory in Hawaii, located in a barren lava field of an active volcano at lat. $19^{\circ} 32^{\prime} \mathrm{N}$. and long. $155^{\circ} 35^{\prime} \mathrm{W}$. and at an altitude of 3,397 m above mean sea level. ${ }^{35-41}$


Figure 10. Variation of (a) <GLOTI> and (b) <MLCO2> for the interval 1960-2013.

For $<$ GLOTI $>$, its mean measures $0.25^{\circ} \mathrm{C}$, having $s d=0.25^{\circ} \mathrm{C}$ and a median $=0.24^{\circ} \mathrm{C}$. Runstesting suggests that the variation of $<$ GLOTI $>$ is nonrandom ( $z=-5.40, c l>99.9 \%$ ). In fact, every year post-1976 has been of positive value, and every year post-1994 has had a value larger than the mean or median. Furthermore, a comparison of the means for the two subintervals suggests that the difference in the means is statistically very important $(t=-11.67, c l » 99.9 \%)$. The warmest anomaly to date occurred in 2010 , measuring $0.66{ }^{\circ} \mathrm{C}$, and the coolest anomaly occurred in 1964 , measuring $-0.02{ }^{\circ} \mathrm{C}$. (For the year 2013, <GLOTI> measured $0.60^{\circ} \mathrm{C}$.)

The <MLCO2> has been continuously rising from one year to the next, with the lowest value occurring in 1960 ( 316.91 ppm ) and the highest value occurring in 2013 ( 396.48 ppm ), rising at an average rate of about $1.47 \mathrm{ppm}_{\mathrm{yr}}{ }^{-1}$. However, the rate of rise actually is rising faster than a simple linear increase, as a straight-edge along the figure plainly shows. Clearly, runs-testing reveals that the distribution of $<\mathrm{MLCO} 2>$ is nonrandom $(z=-7.02, c l>99.9 \%)$, as is also demonstrated from the comparison of the means for the two subintervals ( $t=-11.69$, cl»99.9\%). (In May 2014, the monthly value of MLCO2 reached $401.88 \mathrm{ppm} .{ }^{42}$ However, the yearly average probably will not exceed 400 ppm until the year 2015, due to seasonal effects.)

Figure 11 shows the scatterplot of $<$ GLOTI $>$ versus $<\mathrm{MLCO} 2>$, identifying the years of parametric extremes (1960, 1964, 2010, and 2013). The inferred linear regression is $y=-3.1258+0.0096 x$, having $r=0.932, r^{2}=0.869$, $s e=0.1120$, and $c l>99.9 \%$. Based on Fisher's exact test for $2 \times 2$ contingency tables, the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=6.4 \times 10^{-9} \%$. Hence, one expects $<$ GLOTI $>$ to continue to increase with the passage of time due to the continuing increase in the greenhouse gas of $\mathrm{CO}_{2}$ (with the 2014 yearly values lying in the upper-right quadrant of the scatterplot).


Figure 11. Scatterplot of <GLOTI> versus <MLCO2>.

### 2.7 Correlative Behavior of the Climate-Related Factors

Table 2 provides a listing of the inferred correlative statistics between the various climaterelated factors. As an example, one finds that the only inferred correlation found to be statistically important $(\mathrm{cl}>98 \%)$ is the one between $<\mathrm{ONI}>$ and $<\mathrm{SOI}>$. All other inferred correlations between either $<\mathrm{ONI}>$ or $<$ SOI $>$ and the other climate-related factors are found not to be statistically important $(\mathrm{cl}<90 \%)$. Also, one finds that $<\mathrm{QBO}>$ is not strongly correlated with any of the other climaterelated factors, whereas $<\mathrm{NAO}>\mathrm{CPC}$ and $<\mathrm{NAO}>\mathrm{CRU}$ are inferred to strongly correlate not only with each other, but also with <AMO>. <ASAT> is inferred to correlate strongly with <AMO>, <GLOTI>, and <MLCO2>; <GLOTI> with <AMO>, <ASAT> , and <MLCO2>; and <MLCO2> with <AMO>, <ASAT>, and <GLOTI>.

Table 2. Summary of climate interrelational statistics, 1960-2013.

| Parameter | a | b | $r$ | $r^{2}$ | se | cl |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| <ONl> vs. <SOl> | -0.016 | -0.067 | -0.896 | 0.802 | 0.262 | >99.9\%* |
| <ONI> vs. <AMO> | 0.002 | 0.030 | 0.010 | 0.000 | 0.593 | <90\% |
| <ONI> vs. <QBO> | -0.015 | -0.005 | -0.071 | 0.005 | 0.591 | <90\% |
| <ONI> vs. <NAO> CPC | 0.005 | 0.045 | 0.031 | 0.001 | 0.592 | <90\% |
| <ONI> vs. <NAO> CRU | 0.004 | 0.073 | 0.071 | 0.005 | 0.591 | <90\% |
| <ONI> vs. <ASAT> | 2.051 | -0.216 | -0.199 | 0.040 | 0.581 | <90\% |
| <ONI> vs. <GLOTI> | 0.010 | -0.034 | -0.014 | 0.000 | 0.593 | <90\% |
| <ONI> vs. <MLCO2> | 1.007 | -0.003 | -0.117 | 0.014 | 0.588 | <90\% |
| <SOl> vs. <ONI> | -0.238 | -12.004 | -0.896 | 0.802 | 3.530 | >99.9\%* |
| <SOI> vs. <AMO> | -0.141 | 4.296 | 0.106 | 0.011 | 7.898 | <90\% |
| <SOI> vs. <QBO> | 0.020 | 0.081 | 0.091 | 0.008 | 7.909 | <90\% |
| <SOl> vs. <NAO> CPC | -0.536 | -3.833 | -0.193 | 0.037 | 7.792 | <90\% |
| <SOI> vs. <NAO> CRU | -0.357 | -2.895 | -0.208 | 0.043 | 7.768 | <90\% |
| <SOI> vs. <ASAT> | -22.991 | 2.394 | 0.165 | 0.027 | 7.833 | <90\% |
| <SOI> vs. <GLOTI> | -0.413 | 0.624 | 0.020 | 0.000 | 7.940 | <90\% |
| <SOl> vs. <MLCO2> | -10.149 | 0.028 | 0.086 | 0.007 | 7.912 | <90\% |
| <AMO> vs. <ONI> | -0.026 | 0.003 | 0.010 | 0.000 | 0.195 | <90\% |
| <AMO> vs. <SOI> | -0.026 | 0.003 | 0.106 | 0.011 | 0.194 | <90\% |
| <AMO> vs. <QBO> | -0.029 | -0.001 | -0.033 | 0.001 | 0.195 | <90\% |
| <AMO> vs. <NAO> CPC | -0.041 | -0.199 | -0.408 | 0.167 | 0.178 | >99.5\%* |
| <AMO> vs. <NAO> CRU | -0.030 | -0.111 | -0.325 | 0.105 | 0.185 | >98\%** |
| <AMO> vs. <ASAT> | -1.740 | 0.180 | 0.505 | 0.255 | 0.168 | >99.9\%* |
| <AMO> vs. <GLOTI> | -0.172 | 0.574 | 0.733 | 0.537 | 0.133 | >99.9\%* |
| <AMO> vs. <MLCO2> | -1.704 | 0.005 | 0.591 | 0.349 | 0.162 | >99.9\%** |
| <QBO> vs. <ONI> | -3.381 | -1.080 | -0.071 | 0.005 | 8.941 | <90\% |
| <QBO> vs. <SOl> | -3.356 | 0.103 | 0.091 | 0.008 | 8.927 | <90\% |
| <QBO> vs. <AMO> | -3.421 | -1.495 | -0.033 | 0.001 | 8.959 | <90\% |
| <QBO> vs. <NAO> CPC | -3.089 | 3.978 | 0.178 | 0.032 | 8.821 | <90\% |
| <QBO> vs. <NAO> CRU | -3.304 | 2.178 | 0.139 | 0.019 | 8.877 | <90\% |
| <QBO> vs. <ASAT> | -35.536 | 3.386 | 0.207 | 0.043 | 8.771 | <90\% |
| <QBO> vs. <GLOTl> | -3.250 | -0.521 | -0.015 | 0.000 | 8.963 | <90\% |
| <QBO> vs. <MLCO2> | -2.748 | -0.002 | -0.005 | 0.000 | 8.951 | <90\% |
| <NAO> CPC vs. <ONI> | -0.074 | 0.021 | 0.031 | 0.001 | 0.401 | <90\% |
| <NAO> CPC vs. <SOl> | -0.076 | -0.010 | -0.193 | 0.037 | 0.393 | <90\% |
| <NAO> CPC vs. <AMO> | -0.096 | -0.838 | -0.408 | 0.167 | 0.366 | >99.5\%* |
| <NAO> CPC vs. <QBO> | -0.047 | 0.008 | 0.178 | 0.032 | 0.394 | <90\% |
| <NAO> CPC vs. <NAO> CRU | -0.055 | 0.532 | 0.758 | 0.574 | 0.262 | >99.9\%* |
| <NAO> CPC vs. <ASAT> | -1.391 | 0.139 | 0.189 | 0.036 | 0.394 | <90\% |
| <NAO> CPC vs. <GLOTI> | -0.045 | -0.112 | -0.070 | 0.005 | 0.400 | <90\% |
| <NAO> CPC vs. <MLCO2> | -0.180 | 0.000 | 0.018 | 0.000 | 0.391 | <90\% |

[^0]Table 2. Summary of climate interrelational statistics, 1960-2013 (Continued).

| Parameter | a | b | $r$ | $r^{2}$ | se | cl |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| <NAO> CRU vs. <ONI> | -0.036 | 0.068 | 0.071 | 0.005 | 0.568 | <90\% |
| <NAO> CRU vs. <SOl> | -0.040 | -0.015 | -0.208 | 0.043 | 0.558 | <90\% |
| <NAO> CRU vs. <AMO> | -0.061 | -0.949 | -0.325 | 0.105 | 0.540 | >98\%* |
| <NAO> CRU vs. <QBO> | -0.006 | 0.009 | 0.139 | 0.019 | 0.565 | <90\% |
| <NAO> CRU vs. <NAO> CPC | 0.044 | 1.079 | 0.758 | 0.574 | 0.372 | >99.9\%* |
| <NAO> CRU vs. <ASAT> | -1.418 | 0.146 | 0.140 | 0.019 | 0.565 | <90\% |
| <NAO> CRU vs. <GLOTI> | 0.086 | -0.477 | -0.209 | 0.044 | 0.558 | <90\% |
| <NAO> CRU vs. <MLCO2> | 1.501 | -0.004 | -0.185 | 0.034 | 0.566 | <90\% |
| <ASAT> vs. <ONI> | 9.498 | -0.184 | -0.199 | 0.040 | 0.533 | <90\% |
| <ASAT> vs. <SOI> | 9.500 | 0.011 | 0.165 | 0.027 | 0.542 | <90\% |
| <ASAT> vs. <AMO> | 9.535 | 1.416 | 0.505 | 0.255 | 0.467 | >99.9\%* |
| <ASAT> vs. <QBO> | 9.540 | 0.013 | 0.207 | 0.043 | 0.547 | <90\% |
| <ASAT> vs. <NAO> CPC | 9.516 | 0.258 | 0.189 | 0.036 | 0.541 | <90\% |
| <ASAT> vs. <NAO> CRU | 9.502 | 0.134 | 0.140 | 0.019 | 0.544 | <90\% |
| <ASAT> vs. <GLOTI> | 9.157 | 1.338 | 0.611 | 0.373 | 0.430 | >99.9\%* |
| <ASAT> vs. <MLCO2> | 4.7949 | 0.0134 | 0.592 | 0.350 | 0.4998 | >99.9\%* |
| <GLOTI> vs. <ONI> | 0.255 | -0.006 | -0.014 | 0.000 | 0.249 | <90\% |
| <GLOTI> vs. <SOI> | 0.255 | 0.001 | 0.020 | 0.000 | 0.250 | <90\% |
| <GLOTI> vs. <AMO> | 0.279 | 0.937 | 0.733 | 0.537 | 0.170 | >99.9\%** |
| <GLOTI> vs. <QBO> | 0.253 | 0.000 | -0.015 | 0.000 | 0.251 | <90\% |
| <GLOTI> vs. <NAO> CPC | 0.251 | -0.044 | -0.070 | 0.005 | 0.249 | <90\% |
| <GLOTI> vs. <NAO> CRU | 0.251 | -0.091 | -0.209 | 0.044 | 0.244 | <90\% |
| <GLOTI> vs. <ASAT> | -2.391 | 0.279 | 0.611 | 0.373 | 0.195 | >99.9\%** |
| <GLOTI> vs. <MLCO2> | -3.1258 | 0.0096 | 0.932 | 0.869 | 0.1120 | >99.9\%* |
| <MLCO2> vs. <ONI> | 350.523 | -4.756 | -0.117 | 0.014 | 23.963 | <90\% |
| <MLCO2> vs. <SO1> | 350.583 | 0.261 | 0.086 | 0.007 | 24.039 | <90\% |
| <MLCO2> vs. <AMO> | 352.431 | 73.015 | 0.591 | 0.349 | 19.465 | >99.9\%* |
| <MLCO2> vs. <QBO> | 350.472 | -0.013 | -0.005 | 0.000 | 24.134 | <90\% |
| <MLCO2> vs. <NAO> CPC | 350.597 | 1.093 | 0.018 | 0.000 | 24.127 | <90\% |
| <MLCO2> vs. <NAO> CRU | 350.237 | -7.836 | -0.185 | 0.034 | 23.708 | <90\% |
| <MLCO2> vs. <ASAT> | 102.7349 | 26.0894 | 0.592 | 0.350 | 19.455 | >99.9\%** |
| <MLCO2> vs. <GLOTI> | 327.5853 | 90.0584 | 0.932 | 0.8685 | 8.7483 | >99.9\%** |

*Inferred correlations having cl > 98\%.

Figures 12 and 13 display the scatterplots of $<\mathrm{NAO}>\mathrm{CPC}$ and $<\mathrm{NAO}>\mathrm{CRU}$ versus $<\mathrm{AMO}>$, respectively, and identifies the years of extreme values. Both measures of <NAO> are inferred to correlate inversely against <AMO>, such that negative (positive) values of $<\mathrm{NAO}>$ usually are associated with positive (negative) values of $<\mathrm{AMO}>$, and vice versa. Because $<\mathrm{AMO}>$ is expected to continue to be of positive value for the year 2014 , one expects $<\mathrm{NAO}>$, whether the CPC or CRU values, to probably be of negative value (about twice as likely as being that of positive value).


Figure 12. Scatterplot of $<\mathrm{NAO}>\mathrm{CPC}$ versus $<\mathrm{AMO}>$.


Figure 13. Scatterplot of $<\mathrm{NAO}>\mathrm{CRU}$ versus $<\mathrm{AMO}>$.

Figure 14 shows the scatterplot of <ASAT> versus <AMO>, identifying the years of extreme values. Clearly, <ASAT> tends to vary directly with $<\mathrm{AMO}>$, having $r=0.505, r^{2}=0.255$, $s e=0.467^{\circ} \mathrm{C}$, and $\mathrm{cl}>99.9^{\%}$. Based on Fisher's exact test for $2 \times 2$ contingency tables, the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.12 \%$. Hence, a positive-valued <AMO> for the year 2014 suggests that <ASAT> probably will be $\geq 9.44^{\circ} \mathrm{C}$, possibly considerably greater.


Figure 14. Scatterplot of <ASAT> versus <AMO>.

Figure 15 depicts the scatterplot of <GLOTI> versus <AMO>, identifying the years of extreme values. The inferred regression has $r=0.733$ and $r^{2}=0.537$, suggesting that about $54 \%$ of the variance in <GLOTI> can be explained by the variation in <AMO>, at least during the interval 1960-2013. Based on Fisher's exact test for $2 \times 2$ contingency tables, the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.045 \%$. Clearly, since <AMO> is expected to be of positive value in the year 2014, <GLOTI> likewise is expected to be of positive value $\left(\geq 0.24^{\circ} \mathrm{C}\right.$, in the upper-right quadrant).


Figure 15. Scatterplot of <GLOTI> versus <AMO>.

Figure 16 displays the scatterplot of <ASAT> versus <GLOTI>, identifying the years of extreme values. The inferred regression has $r=0.611$ and $r^{2}=0.323$, suggesting that about one-third of the variance in <ASAT> can be explained by the variation in <GLOTI>. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.015 \%$. Because $<$ GLOTI $>$ undoubtedly will be of positive value in the year 2014 , probably $\geq 0.6^{\circ} \mathrm{C}$, $<$ ASAT $>$ is expected to be $\geq 9.44{ }^{\circ} \mathrm{C}$, probably $\geq 9.53^{\circ} \mathrm{C}$ (in the upper-right quadrant).


Figure 16. Scatterplot of <ASAT> versus < GLOTI>.

Figure 17 shows the scatterplot of <AMO> versus <MLCO2>, identifying the years of extreme values. The inferred regression has $r=0.591$ and $r^{2}=0.349$, suggesting that about one-third of the variance in <AMO> can be explained by the variation in $<\mathrm{MLCO} 2>$, at least during the interval 1960-2013. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.12 \%$. Because <MLCO2> continues to grow year-by-year, undoubtedly, its value for the year 2014 will be greater than its value for the year $2013(=396.48 \mathrm{ppm})$ suggesting that <AMO> will be of positive value for the year 2014, having a yearly value in the upper-right quadrant.


Figure 17. Scatterplot of $<\mathrm{AMO}>$ versus $<\mathrm{MLCO} 2>$.

Figure 18 depicts the scatterplot of <ASAT> versus <MLCO2>, identifying the years of extreme values. The inferred regression has $r=0.592$ and $r^{2}=0.350$, suggesting that about one-third of the variance in <ASAT> can be explained by the variation in $<\mathrm{MLCO} 2>$, at least during the interval 1960-2013. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.015 \%$. Because <MLCO2> continues to grow year-by-year, undoubtedly, its value for the year 2014 will be greater than its value for the year $2013(=396.48 \mathrm{ppm})$ suggesting that <ASAT> will have a yearly value in the upper-right quadrant $\left(\geq 9.44^{\circ} \mathrm{C}\right)$.


Figure 18. Scatterplot of <ASAT> versus <MLCO2>.

### 2.8 Inferred Statistical Relationships Between Tropical Cyclone Parametric Values and Selected Climate-Related Factors

Table 3 provides a summary of the statistics of the 25 tropical cyclone parameters identified in reference 1 against the 9 specific climate-related factors investigated in this TP. The 25 tropical cyclone parameters include:

- First storm day (FSD).
- Last storm day (LSD).
- Length of season (LOS).
- NTC.
- NH.
- NMH.
- NUSLFH.
- Mean seasonal latitude (<LAT>).
- Mean seasonal longitude (<LONG>).
- PWS.
- Mean seasonal PWS (<PWS $>$ ).
- Lowest pressure (LP).
- Mean seasonal LP (<LP>).
- Mean seasonal ACE (<ACE $>$ ).
- Total ACE.
- Highest individual storm ACE (HISACE).
- Mean seasonal Power Dissipation Index (PDI) (<PDI $>$ ).
- Total PDI.
- Highest individual storm PDI (HISPDI).
- Mean seasonal number of storm days (NSD) (<NSD>).
- Longest individual storm NSD (LISNSD).
- Total NSD.
- Total number of hurricane days (NHD).
- Total number of major hurricane days (NMHD).
- Net Tropical Cyclone Activity (NTCA).

From the table, one finds that none of the tropical cyclone parameters correlates strongly against $<\mathrm{QBO}>$. The best climate-related factor that correlates with the most tropical cyclone parameters is $<\mathrm{AMO}>$, having 15 correlations with $c l>98 \%$ and an additional 5 correlations with $c l>95 \%$. The second best climate-related factor is <SOI>, having 8 correlations with $c l>98 \%$ and an additional 5 correlations with $\mathrm{cl}>95 \%$. In the following subsections, each of the scatterplots having $\mathrm{cl}>98 \%$ is displayed, first those against $<\mathrm{AMO}>$, then those against $<\mathrm{SOI}>,<\mathrm{MLCO} 2>$, $<\mathrm{GLOTI}>,<\mathrm{ONI}>,<\mathrm{NAO}>\mathrm{CRU},<\mathrm{NAO}>\mathrm{CPC}$, and $<\mathrm{ASAT}>$. Furthermore, estimates are made for the selected tropical cyclone parameters based on the mean values of the specific climate-related factors for the more recent subinterval 1995-2013. When $c l>98 \%$ (from the linear regression analysis) and $P<5 \%$ (from Fisher's exact test), only then is the estimate considered reliable, presuming of course the accuracy for the projected value of the specific climate-related factor.
Table 3. Summary of tropical cyclone statistics against climate factors (single-variate fits), 1960-2013.

|  | <ONI> |  |  |  |  |  | <SO1> |  |  |  |  |  | <AMO> |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Parameter | a | b | $r$ | $r^{2}$ | se | cl | a | b | $r$ | $r^{2}$ | se | cl | a | b | $r$ | $r^{2}$ | se | cl |
| FSD | 179.666 | 0.747 | 0.013 | 0.000 | 33.593 | <90\% | 179.682 | 0.062 | 0.015 | 0.000 | 33.595 | <90\% | 179.573 | -3.579 | -0.021 | 0.000 | 33.589 | <90\% |
| LSD | 311.204 | -14.770 | -0.365 | 0.133 | 22.343 | >99\%** | 311.398 | 0.838 | 0.277 | 0.077 | 23.052 | >95\%* | 311.739 | 21.137 | 0.172 | 0.030 | 23.640 | <90\% |
| LOS | 132.539 | -15.517 | -0.215 | 0.046 | 41.690 | <90\% | 132.715 | 0.775 | 0.144 | 0.021 | 42.246 | <90\% | 133.167 | 24.717 | 0.113 | 0.013 | 42.418 | <90\% |
| NTC | 11.392 | -2.562 | -0.333 | 0.111 | 4.304 | >98\%** | 11.440 | 0.200 | 0.348 | 0.121 | 4.278 | >99\%** | 11.728 | 12.929 | 0.553 | 0.306 | 3.800 | >99.9\%** |
| NH | 6.168 | -1.177 | -0.249 | 0.062 | 2.718 | >90\% | 6.193 | 0.103 | 0.292 | 0.085 | 2.684 | >95\%** | 6.352 | 7.082 | 0.493 | 0.243 | 2.442 | >99.9\%** |
| NMH | 2.446 | -0.936 | -0.296 | 0.087 | 1.794 | >95\%* | 2.468 | 0.094 | 0.397 | 0.157 | 1.724 | >99.5\%** | 2.560 | 4.411 | 0.459 | 0.211 | 1.668 | >99.9\%** |
| NUSLFH | 1.445 | -0.321 | -0.127 | 0.016 | 1.490 | <90\% | 1.449 | 0.017 | 0.092 | 0.008 | 1.496 | <90\% | 1.471 | 0.997 | 0.130 | 0.017 | 1.489 | <90\% |
| <LAT> | 22.537 | 1.314 | 0.233 | 0.054 | 3.250 | >90\% | 22.507 | -0.127 | -0.301 | 0.091 | 3.184 | >95\%* | 22.298 | -9.176 | -0.536 | 0.288 | 2.821 | >99.9\%** |
| <LONG> | 63.459 | 0.166 | 0.017 | 0.000 | 5.818 | <90\% | 63.442 | -0.067 | -0.091 | 0.008 | 5.795 | <90\% | 63.185 | -10.456 | -0.351 | 0.123 | 5.448 | >99\%** |
| PWS | 124.264 | -3.318 | -0.089 | 0.008 | 21.912 | <90\% | 124.320 | 0.241 | 0.087 | 0.008 | 21.910 | <90\% | 125.080 | 31.305 | 0.278 | 0.077 | 21.134 | >95\%** |
| <PWS> | 72.936 | -3.470 | -0.236 | 0.056 | 8.461 | >90\% | 73.010 | 0.311 | 0.284 | 0.081 | 8.351 | >95\%* | 73.125 | 7.389 | 0.166 | 0.027 | 7.863 | <90\% |
| LP | 937.158 | 6.774 | 0.181 | 0.033 | 21.825 | <90\% | 937.047 | -0.472 | -0.169 | 0.029 | 21.873 | <90\% | 935.992 | -44.795 | -0.394 | 0.155 | 20.397 | $>99.5 \%$ ** |
| <LP> | 980.753 | 3.433 | 0.262 | 0.069 | 7.478 | >90\% | 980.677 | -0.317 | -0.325 | 0.106 | 7.331 | >98\%** | 980.446 | -11.863 | -0.299 | 0.089 | 7.421 | >95\%* |
| <ACE> | 8.605 | -0.788 | -0.125 | 0.016 | 3.703 | <90\% | 8.629 | 0.098 | 0.208 | 0.043 | 3.650 | <90\% | 8.741 | 5.243 | 0.274 | 0.075 | 3.590 | >95\%* |
| Total ACE | 100.507 | -25.304 | -0.251 | 0.063 | 57.816 | >90\% | 101.087 | 2.416 | 0.321 | 0.103 | 56.564 | >98\%** | 104.820 | 165.740 | 0.542 | 0.294 | 50.193 | >99.9\%** |
| HISACE | 32.746 | -2.978 | -0.087 | 0.008 | 20.125 | <90\% | 30.989 | 0.386 | 0.188 | 0.035 | 15.967 | <90\% | 31.754 | 32.914 | 0.395 | 0.156 | 14.933 | $>99.5 \%$ ** |
| <PDI> | 6.949 | -0.803 | -0.122 | 0.015 | 3.884 | <90\% | 6.973 | 0.099 | 0.201 | 0.040 | 3.833 | <90\% | 7.110 | 6.181 | 0.309 | 0.095 | 3.722 | >95\%* |
| Total PDI | 81.871 | -20.496 | -0.216 | 0.047 | 54.938 | <90\% | 82.358 | 2.024 | 0.286 | 0.082 | 53.919 | >95\%** | 85.900 | 154.648 | 0.537 | 0.288 | 47.466 | >99.9\%** |
| HISPDI | 30.799 | -1.963 | -0.057 | 0.003 | 20.338 | <90\% | 30.890 | 0.371 | 0.145 | 0.021 | 20.243 | <90\% | 31.922 | 42.935 | 0.412 | 0.170 | 18.565 | $>99.8 \%$ ** |
| <NSD> | 4.942 | -0.310 | -0.135 | 0.018 | 1.347 | <90\% | 4.952 | 0.045 | 0.261 | 0.068 | 1.314 | >90\% | 5.007 | 2.528 | 0.363 | 0.132 | 1.269 | >99\%** |
| LISNSD | 11.751 | -0.456 | -0.072 | 0.005 | 3.768 | <90\% | 11.785 | 0.103 | 0.217 | 0.047 | 3.635 | $>99.8 \%$ ** | 11.907 | 5.638 | 0.292 | 0.085 | 3.631 | >95\%* |
| Total NSD | 58.154 | -15.196 | -0.275 | 0.076 | 31.490 | >95\%* | 58.495 | 1.495 | 0.363 | 0.131 | 30.519 | >99\%** | 60.702 | 98.628 | 0.588 | 0.346 | 26.480 | >99.9\%** |
| Total NHD | 23.712 | -6.135 | -0.246 | 0.061 | 14.304 | <90\% | 23.854 | 0.593 | 0.319 | 0.102 | 13.987 | >98\%** | 24.553 | 32.390 | 0.429 | 0.184 | 13.334 | >99.8\%** |
| Total NMHD | 5.576 | -1.355 | -0.136 | 0.019 | 5.838 | <90\% | 5.615 | 0.162 | 0.218 | 0.048 | 5.751 | <90\% | 5.964 | 14.861 | 0.493 | 0.243 | 5.128 | >99.9\%** |
| NTCA | 109.639 | -28.679 | -0.278 | 0.077 | 58.840 | >95\%* | 110.310 | 2.791 | 0.362 | 0.131 | 57.094 | >99\%** | 114.271 | 178.073 | 0.568 | 0.323 | 50.409 | >99.9\%** |

*Inferred correlation is statistically significant at $c \mid>95 \%$.
**Inferred correlation is statistically significant at $c l>98 \%$
Table 3. Summary of tropical cyclone statistics against climate factors (single-variate fits), 1960-2013 (Continued).

|  | <QBO> |  |  |  |  |  | <NAO>CPC |  |  |  |  |  | <NAO>CRU |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Parameter | a | b | $r$ | $r^{2}$ | se | cl | a | b | $r$ | $r^{2}$ | se | cl | a | b | $r$ | $r^{2}$ | se | cl |
| FSD | 181.838 | 0.642 | 0.171 | 0.029 | 33.101 | <90\% | 179.546 | -1.631 | -0.019 | 0.000 | 33.592 | <90\% | 179.751 | 2.362 | 0.040 | 0.002 | 33.570 | <90\% |
| LSD | 309.893 | $-0.382$ | -0.143 | 0.020 | 23.750 | <90\% | 311.408 | 3.022 | 0.050 | 0.003 | 23.963 | <90\% | 311.052 | -3.716 | -0.088 | 0.008 | 23.903 | <90\% |
| LOS | 129.055 | -1.024 | -0.215 | 0.046 | 41.694 | <90\% | 132.861 | 4.653 | 0.044 | 0.002 | 42.652 | <90\% | 132.301 | -6.077 | -0.081 | 0.007 | 42.552 | <90\% |
| NTC | 11.312 | $-0.023$ | -0.044 | 0.002 | 4.558 | <90\% | 11.223 | -2.246 | -0.197 | 0.039 | 4.474 | <90\% | 11.303 | -2.396 | $-0.300$ | 0.090 | 4.354 | >95\%* |
| NH | 6.088 | $-0.023$ | -0.074 | 0.005 | 2.800 | <90\% | 5.992 | -2.372 | -0.339 | 0.115 | 2.640 | >98\%** | 6.100 | -1.872 | -0.381 | 0.145 | 2.595 | <99.5\%** |
| NMH | 2.542 | 0.029 | 0.138 | 0.019 | 1.860 | <90\% | 2.364 | -1.089 | -0.232 | 0.054 | 1.826 | >90\% | 2.414 | -0.840 | -0.255 | 0.065 | 1.816 | >90\% |
| NUSLFH | 1.493 | 0.014 | 0.085 | 0.007 | 1.496 | <90\% | 1.435 | -0.133 | -0.035 | 0.001 | 1.501 | <90\% | 1.435 | -0.251 | -0.095 | 0.009 | 1.496 | <90\% |
| <LAT> | 22.556 | 0.005 | 0.014 | 0.000 | 3.340 | <90\% | 22.615 | 1.033 | 0.124 | 0.015 | 3.316 | <90\% | 22.577 | 1.077 | 0.184 | 0.034 | 3.285 | <90\% |
| <LONG> | 63.271 | $-0.056$ | -0.086 | 0.007 | 5.790 | <90\% | 63.440 | -0.262 | -0.018 | 0.000 | 5.818 | <90\% | 63.447 | -0.346 | -0.034 | 0.001 | 5.814 | <90\% |
| PWS | 123.948 | -0.092 | -0.038 | 0.001 | 21.986 | <90\% | 123.772 | -6.609 | -0.120 | 0.015 | 21.841 | <90\% | 124.113 | -4.091 | -0.106 | 0.011 | 21.877 | <90\% |
| <PWS> | 73.068 | 0.040 | 0.042 | 0.002 | 8.696 | <90\% | 72.632 | -4.064 | -0.187 | 0.035 | 8.555 | <90\% | 72.909 | -0.637 | -0.042 | 0.002 | 8.699 | <90\% |
| LP | 937.733 | 0.167 | 0.068 | 0.005 | 22.110 | <90\% | 938.038 | 11.827 | 0.214 | 0.046 | 21.689 | <90\% | 937.460 | 7.946 | 0.204 | 0.042 | 21.722 | <90\% |
| <LP> | 980.611 | -0.043 | -0.050 | 0.003 | 7.807 | <90\% | 981.055 | 4.040 | 0.209 | 0.044 | 7.605 | <90\% | 980.799 | 1.174 | 0.086 | 0.007 | 7.748 | <90\% |
| <ACE> | 8.734 | 0.039 | 0.093 | 0.009 | 3.718 | <90\% | 8.522 | -1.109 | -0.119 | 0.014 | 3.706 | <90\% | 8.602 | -0.061 | -0.009 | 0.000 | 3.732 | <90\% |
| Total ACE | 101.624 | 0.340 | 0.051 | 0.003 | 59.652 | <90\% | 98.251 | -30.157 | -0.202 | 0.041 | 58.493 | <90\% | 99.661 | -22.762 | -0.217 | 0.047 | 58.299 | <90\% |
| HISACE | 31.177 | 0.085 | 0.047 | 0.002 | 16.242 | <90\% | 30.404 | -6.600 | -0.163 | 0.026 | 16.042 | <90\% | 30.679 | -5.938 | -0.208 | 0.043 | 15.901 | <90\% |
| <PDI> | 7.053 | 0.031 | 0.071 | 0.005 | 3.903 | <90\% | 6.857 | -1.232 | -0.126 | 0.016 | 3.882 | <90\% | 6.947 | -0.038 | -0.006 | 0.000 | 3.913 | <90\% |
| Total PDI | 82.720 | 0.259 | 0.041 | 0.002 | 56.217 | <90\% | 79.952 | -25.681 | -0.183 | 0.033 | 55.314 | <90\% | 81.206 | -17.850 | $-0.181$ | 0.033 | 55.335 | <90\% |
| HISPDI | 30.868 | 0.021 | 0.009 | 0.000 | 20.370 | <90\% | 30.206 | -8.008 | -0.158 | 0.025 | 20.117 | <90\% | 30.562 | -6.564 | -0.184 | 0.034 | 20.024 | <90\% |
| <NSD> | 4.980 | 0.011 | 0.076 | 0.006 | 1.354 | <90\% | 4.871 | -0.950 | -0.280 | 0.078 | 1.306 | >95\%* | 4.913 | -0.754 | -0.316 | 0.100 | 1.291 | >98\%** |
| LISNSD | 11.959 | 0.059 | 0.140 | 0.020 | 3.740 | <90\% | 11.680 | -1.067 | -0.113 | 0.013 | 3.753 | <90\% | 11.713 | -1.304 | -0.197 | 0.039 | 3.703 | <90\% |
| Total NSD | 58.130 | 0.004 | 0.001 | 0.000 | 32.747 | <90\% | 56.072 | -27.726 | -0.339 | 0.115 | 30.803 | >98\%** | 57.203 | -25.537 | -0.445 | 0.198 | 29.326 | >99.9\%** |
| Total NHD | 24.199 | 0.147 | 0.089 | 0.008 | 14.702 | <90\% | 23.101 | -8.175 | -0.222 | 0.049 | 14.391 | <90\% | 23.502 | -5.654 | -0.219 | 0.048 | 14.402 | <90\% |
| $\begin{array}{\|l\|} \hline \text { Total } \\ \text { NMHD } \end{array}$ | 5.677 | 0.030 | 0.046 | 0.002 | 5.886 | <90\% | 5.437 | -1.858 | -0.126 | 0.016 | 5.846 | <90\% | 5.536 | -1.071 | -0.104 | 0.011 | 5.861 | <90\% |
| NTCA | 110.633 | 0.305 | 0.045 | 0.002 | 61.186 | <90\% | 106.689 | -39.526 | -0.259 | 0.067 | 59.160 | >90\% | 108.467 | -31.741 | -0.296 | 0.087 | 58.507 | >95\%* |

${ }_{*}^{*}$ *Inferred correlation is statistically significant at $c l>95 \%$.
Table 3. Summary of tropical cyclone statistics against climate factors (single-variate fits), 1960-2013 (Continued).

|  | <ASAT> |  |  |  |  |  | <GLOTI> |  |  |  |  |  | <MLCO2> |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Parameter | a | b | $r$ | $r^{2}$ | se | cl | a | b | $r$ | $r^{2}$ | se | cl | a | b | $r$ | $r^{2}$ | se | cl |
| FSD | 281.414 | -10.713 | -0.174 | 0.030 | 33.077 | <90\% | 183.324 | -14.362 | -0.107 | 0.011 | 33.404 | <90 | 287.671 | -0.308 | -0.221 | 0.049 | 32.636 | <90\% |
| LSD | 223.568 | 9.225 | 0.210 | 0.044 | 23.482 | <90\% | 304.156 | 27.607 | 0.287 | 0.083 | 22.980 | >95\%* | 209.712 | 0.289 | 0.291 | 0.085 | 24.149 | >95\%* |
| LOS | -56.846 | 19.939 | 0.256 | 0.065 | 41.268 | >90\% | 121.832 | 41.968 | 0.245 | 0.060 | 41.387 | >90\% | -76.959 | 0.598 | 0.338 | 0.114 | 39.956 | >98\%** |
| NTC | -22.591 | 3.578 | 0.429 | 0.184 | 4.119 | >99.8\%** | 8.815 | 10.107 | 0.553 | 0.306 | 3.803 | >99.9\%** | -25.399 | 0.105 | 0.555 | 0.308 | 3.770 | >99.9\%** |
| NH | -2.676 | 0.931 | 0.182 | 0.033 | 2.760 | <90\% | 5.215 | 3.738 | 0.333 | 0.111 | 2.646 | >98\% | -4.478 | 0.030 | 0.261 | 0.068 | 2.859 | >90\% |
| NMH | -5.958 | 0.885 | 0.258 | 0.066 | 1.812 | >90\% | 1.983 | 1.810 | 0.241 | 0.058 | 1.823 | >90\% | -2.980 | 0.015 | 0.199 | 0.040 | 1.953 | <90\% |
| NUSLFH | -1.614 | 0.322 | 0.117 | 0.014 | 1.492 | <90\% | 1.447 | -0.010 | -0.002 | 0.000 | 1.502 | <90\% | 0.616 | 0.002 | 0.038 | 0.001 | 1.564 | <90\% |
| <LAT> | 32.251 | -1.023 | -0.167 | 0.028 | 3.307 | <90\% | 23.702 | -4.566 | -0.341 | 0.116 | 3.139 | >98\%** | 38.425 | -0.045 | -0.327 | 0.107 | 2.708 | >99\%** |
| <LONG> | 77.213 | -1.448 | -0.136 | 0.019 | 5.756 | <90\% | 65.133 | -6.575 | -0.282 | 0.080 | 5.584 | >95\%* | 85.754 | -0.064 | -0.264 | 0.070 | 6.372 | >90\% |
| PWS | 84.665 | 4.169 | 0.104 | 0.011 | 21.881 | <90\% | 122.866 | 5.471 | 0.062 | 0.004 | 21.959 | <90\% | 138.463 | -0.041 | -0.044 | 0.002 | 22.465 | <90\% |
| <PWS> | 88.026 | -1.589 | -0.100 | 0.010 | 8.650 | <90\% | 74.734 | -7.078 | -0.203 | 0.041 | 8.525 | <90\% | 109.771 | -0.105 | -0.291 | 0.085 | 8.169 | >95\%* |
| LP | 1,033.173 | -10.109 | -0.249 | 0.062 | 21.561 | >90\% | 942.603 | -21.350 | -0.240 | 0.058 | 21.544 | >90\% | 981.652 | -0.127 | -0.138 | 0.019 | 22.642 | <90\% |
| <LP> | 989.712 | -0.943 | -0.067 | 0.004 | 7.843 | <90\% | 980.053 | 2.766 | 0.089 | 0.008 | 7.728 | <90\% | 962.171 | 0.053 | 0.165 | 0.027 | 8.225 | <90\% |
| <ACE> | 6.323 | 0.240 | 0.035 | 0.001 | 3.732 | <90\% | 9.219 | -2.417 | -0.162 | 0.026 | 3.684 | <90\% | 20.601 | -0.034 | -0.221 | 0.049 | 3.541 | <90\% |
| Total ACE | -223.135 | 34.073 | 0.312 | 0.097 | 56.749 | >95\%* | 84.033 | 64.567 | 0.270 | 0.073 | 57.513 | >95\%* | -85.405 | 0.530 | 0.214 | 0.046 | 58.438 | <90\% |
| HISACE | -15.568 | 4.892 | 0.165 | 0.027 | 16.034 | <90\% | 30.294 | 2.342 | 0.036 | 0.001 | 16.249 | <90\% | 32.379 | -0.004 | -0.006 | 0.000 | 16.172 | <90\% |
| <PDI> | 3.275 | 0.387 | 0.054 | 0.003 | 3.905 | <90\% | 7.498 | -2.159 | -0.138 | 0.019 | 3.876 | <90\% | 18.702 | -0.034 | -0.207 | 0.043 | 3.979 | <90\% |
| Total PDI | -206.843 | 30.395 | 0.295 | 0.087 | 53.769 | >95\%* | 68.335 | 53.054 | 0.235 | 0.055 | 54.683 | >90\% | -53.696 | 0.387 | 0.166 | 0.027 | 55.401 | <90\% |
| HISPDI | -22.487 | 5.610 | 0.151 | 0.023 | 20.141 | <90\% | 29.481 | 5.164 | 0.063 | 0.004 | 20.331 | <90\% | 31.348 | -0.002 | -0.002 | 0.000 | 20.488 | <90\% |
| <NSD> | 5.627 | -0.072 | -0.029 | 0.001 | 1.356 | <90\% | 4.818 | 0.485 | 0.089 | 0.008 | 1.354 | <90\% | 3.706 | 0.004 | 0.062 | 0.004 | 0.993 | <90\% |
| LISNSD | -1.586 | 1.405 | 0.204 | 0.041 | 3.701 | <90\% | 11.451 | 1.212 | 0.080 | 0.006 | 3.765 | <90\% | 5.564 | 0.018 | 0.113 | 0.013 | 3.562 | <90\% |
| Total NSD | -80.718 | 14.618 | 0.244 | 0.060 | 31.756 | >90\% | 42.629 | 60.822 | 0.464 | 0.215 | 29.013 | >99.9\%** | -154.552 | 0.607 | 0.447 | 0.200 | 29.193 | >99.9\%** |
| Total NHD | -30.148 | 5.670 | 0.210 | 0.044 | 14.431 | <90\% | 21.963 | 6.836 | 0.116 | 0.013 | 14.660 | <90\% | 11.831 | 0.034 | 0.055 | 0.003 | 14.699 | <90\% |
| Total NMHD | -21.322 | 2.832 | 0.263 | 0.069 | 5.686 | >90\% | 4.532 | 4.092 | 0.173 | 0.030 | 5.804 | <90\% | $-2.574$ | 0.023 | 0.095 | 0.009 | 5.909 | <90\% |
| NTCA | -204.500 | 33.072 | 0.295 | 0.087 | 58.516 | >95\%* | 89.455 | 79.122 | 0.323 | 0.104 | 57.973 | >98\%** | -127.784 | 0.677 | 0.267 | 0.071 | 59.110 | >95\%* |

[^1]
### 2.8.1 Inferred Correlations Against <AMO>

Figure 19 shows the scatterplot for NTC versus <AMO>, identifying the years of extreme values. The inferred regression is given approximately as $y=11.7+12.9 x$, having $r=0.55, r^{2}=0.31$, $s e=3.8$, and $\mathrm{cl}>99.9 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.6 \%$. Hence, there is reason to believe, based on the expected value for <AMO> in the year 2014 (i.e., being of positive value), that NTC will lie in the upper-right quadrant of the scatterplot (i.e., NTC $\geq 11$ ). The mean value for <AMO> during the more recent interval 1995-2013 measures 0.163 . Using this value for $<\mathrm{AMO}>$ for the year 2014 , one infers $\mathrm{NTC}=13.8 \pm 3.8$ (the $\pm 1$ se prediction interval), or $\mathrm{NTC} \geq 10$.


Figure 19. Scatterplot of NTC versus <AMO>.

Figure 20 depicts the scatterplot for NH versus <AMO>, identifying the years of extreme values. The inferred regression is given approximately as $y=6.4+7.1 x$, having $r=0.49, r^{2}=0.24$, $s e=2.4$, and $c l>99.9 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=5.0 \%$. Hence, based on the expected value for <AMO> in the year 2014 (i.e., being of positive value), NH is expected to lie in the upper-right quadrant of the scatterplot (i.e., $\mathrm{NH} \geq 6$ ). The mean value for <AMO> during the more recent interval 1995-2013 measures 0.163. Using this value for <AMO> for the year 2014 , one infers $\mathrm{NH}=7.6 \pm 2.4$, or $\mathrm{NH} \geq 5$ (rounded to the nearest whole number).


Figure 20. Scatterplot of NH versus < AMO>.

Figure 21 displays the scatterplot for NMH versus < AMO>, identifying the years of extreme values. The inferred regression is given approximately as $y=2.6+4.4 x$, having $r=0.46, r^{2}=0.21$, $s e=1.7$, and $c l>99.9 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.5 \%$. Hence, based on the expected value for <AMO> in the year 2014 (i.e., being of positive value), NMH is expected to lie in the upper-right quadrant of the scatterplot (i.e., $\mathrm{NMH} \geq 2$ ). The mean value for <AMO> during the more recent interval 1995-2013 measures 0.163 . Using this value for <AMO> for the year 2014 , one infers $\mathrm{NMH}=3.3 \pm 1.7$, or $\mathrm{NMH} \geq 2$ (rounded to the nearest whole number).


Figure 21. Scatterplot of NMH versus <AMO>.

Figure 22 shows the scatterplot for <LAT> versus <AMO>, identifying the years of extreme values. The inferred regression is given approximately as $y=22.3-9.2 x$, having $r=-0.54, r^{2}=0.29$, $s e=2.8$, and $c l>99.9 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.7 \%$. Hence, based on the expected value for <AMO> in the year 2014 (i.e., being of positive value), $<\mathrm{LAT}>$ is expected to lie in the lower-right quadrant of the scatterplot (i.e., $<\mathrm{LAT}><22^{\circ} \mathrm{N}$.). The mean value for $<$ AMO $>$ during the more recent interval 1995-2013 measures 0.163 . Using this value for $<$ AMO $>$ for the year 2014, one infers $<$ LAT $>=20.8 \pm 2.8^{\circ} \mathrm{N}$., or $<\mathrm{LAT}><23.6^{\circ} \mathrm{N}$.


Figure 22. Scatterplot of $<\mathrm{LAT}>$ versus $<\mathrm{AMO}>$.

Figure 23 depicts the scatterplot for <LONG> versus <AMO>, identifying the years of extreme values. The inferred regression is given approximately as $y=63.2-10.5 x$, having $r=-0.35$, $r^{2}=0.12$, $s e=5.4$, and $c l>99 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=29.3 \%$. The mean value for <AMO> during the more recent interval 1995-2013 measures 0.163 . Using this value for $<\mathrm{AMO}>$ for the year 2014 , one infers $<\mathrm{LONG}>=61.5 \pm 5.4^{\circ} \mathrm{W}$., or $<\mathrm{LONG}><67^{\circ} \mathrm{W}$.


Figure 23. Scatterplot of $\langle\mathrm{LONG}\rangle$ versus $<\mathrm{AMO}\rangle$.

Figure 24 displays the scatterplot for LP versus <AMO>, identifying the years of extreme values. The inferred regression is given approximately as $y=936.0-44.8 x$, having $r=-0.39, r^{2}=0.16$, $s e=20.4$, and $c l>99.5 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=2.8 \%$. The mean value for $<\mathrm{AMO}>$ during the more recent interval 1995-2013 measures 0.163 . Using this value for $<A M O>$ for the year 2014 , one infers LP $=928.7 \pm 20.4 \mathrm{mb}$, or LP $<949.1 \mathrm{mb}$.


Figure 24. Scatterplot of LP versus < AMO>.

Figure 25 shows the scatterplot for total ACE versus <AMO>, identifying the years of extreme values. The inferred regression is given approximately as $y=104.8+165.7 x$, having $r=0.54$, $r^{2}=0.29, s e=50.2$, and $c l>99.9 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.7 \%$. The mean value for $<\mathrm{AMO}>$ during the more recent interval 1995-2013 measures 0.163. Using this value for $<\mathrm{AMO}>$ for the year 2014, one infers total $\mathrm{ACE}=131.8 \pm 50.2$, or total $A C E \geq 81.6$ (units are $10^{4} \mathrm{kt}^{2}$ ).


Figure 25. Scatterplot of total ACE versus < AMO>.

Figure 26 depicts the scatterplot for HISACE versus <AMO> , identifying the years of extreme values. The inferred regression is given approximately as $y=31.8+32.9 x$, having $r=0.40, r^{2}=0.16$, $s e=14.9$, and $c l>99.5 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=1.4 \%$. The mean value for <AMO> during the more recent interval 1995-2013 measures 0.163. Using this value for $<$ AMO $>$ for the year 2014, one infers HISACE $=37.1 \pm 14.9$, or HISACE $\geq 22.2$ (units are $10^{4} \mathrm{kt}^{2}$ ).


Figure 26. Scatterplot of HISACE versus < AMO>.

Figure 27 displays the scatterplot for total PDI versus <AMO>, identifying the years of extreme values. The inferred regression is given approximately as $y=85.9+154.6 x$, having $r=0.54$, $r^{2}=0.29, s e=47.5$, and $c l>99.9 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.045 \%$. The mean value for $<\mathrm{AMO}>$ during the more recent interval 1995-2013 measures 0.163 . Using this value for $\langle\mathrm{AMO}>$ for the year 2014, one infers total PDI $=111.1 \pm 47.5$, or total PDI $\geq 63.6$ (units are $10^{6} \mathrm{kt}^{3}$ ).


Figure 27. Scatterplot of total PDI versus <AMO>.

Figure 28 shows the scatterplot for HISPDI versus <AMO>, identifying the years of extreme values. The inferred regression is given approximately as $y=31.9+42.9 x$, having $r=0.41, r^{2}=0.17$, $s e=18.6$, and $\mathrm{cl}>99.8 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=1.4 \%$. The mean value for <AMO> during the more recent interval 1995-2013 measures 0.163. Using this value for $<$ AMO $>$ for the year 2014 , one infers HISPDI $=38.9 \pm 18.6$, or HISPDI $\geq 20.3$ (units are $10^{6} \mathrm{kt}^{3}$ ).


Figure 28. Scatterplot of HISPDI versus <AMO>.

Figure 29 depicts the scatterplot for < NSD $>$ versus $<\mathrm{AMO}>$, identifying the years of extreme values. The inferred regression is given approximately as $y=5.0+2.5 x$, having $r=0.36, r^{2}=0.13$, $s e=1.3$, and $c l>99 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=29.3 \%$. The mean value for <AMO> during the more recent interval 1995-2013 measures 0.163 . Using this value for $<\mathrm{AMO}>$ for the year 2014 , one infers $<$ NSD $>=5.4 \pm 1.3$ days, or $<$ NSD $>\geq 4.1$ days.


Figure 29. Scatterplot of $<$ NSD $>$ versus $<\mathrm{AMO}>$.

Figure 30 displays the scatterplot for total NSD versus <AMO> , identifying the years of extreme values. The inferred regression is given approximately as $y=60.7+98.6 x$, having $r=0.59$, $r^{2}=0.35$, $s e=26.5$, and $c l>99.9 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.3 \%$. The mean value for $<\mathrm{AMO}>$ during the more recent interval 1995-2013 measures 0.163. Using this value for <AMO> for the year 2014, one infers total NSD $=76.8 \pm 26.5$ days, or total NSD $\geq 50.3$ days.


Figure 30. Scatterplot of total NSD versus <AMO>.

Figure 31 shows the scatterplot for total NHD versus <AMO>, identifying the years of extreme values. The inferred regression is given approximately as $y=24.6+32.4 x$, having $r=0.43$, $r^{2}=0.18, s e=13.3$, and $c l>99.8 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=5.1 \%$. The mean value for <AMO> during the more recent interval 1995-2013 measures 0.163. Using this value for <AMO> for the year 2014, one infers total NHD $=29.8 \pm 13.3$ days, or total NHD $\geq 16.5$ days.


Figure 31. Scatterplot of total NHD versus <AMO>.

Figure 32 depicts the scatterplot for total NMHD versus <AMO> , identifying the years of extreme values. The inferred regression is given approximately as $y=6.0+14.9 x$, having $r=0.49$, $r^{2}=0.24$, $s e=5.1$, and $c l>99.9 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.3 \%$. The mean value for $<\mathrm{AMO}>$ during the more recent interval 1995-2013 measures 0.163. Using this value for <AMO> for the year 2014, one infers total NMHD $=8.4 \pm 5.1$ days, or total NMHD $\geq 3.3$ days.


Figure 32. Scatterplot of total NMHD versus <AMO>.

Figure 33 displays the scatterplot for NTCA versus <AMO> , identifying the years of extreme values. The inferred regression is given approximately as $y=114.3+178.1 x$, having $r=0.57, r^{2}=0.32$, $s e=50.4$, and $\mathrm{cl}>99.9 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.3 \%$. The mean value for <AMO> during the more recent interval 1995-2013 measures 0.163 . Using this value for $<\mathrm{AMO}>$ for the year 2014 , one infers NTCA $=143.3 \pm 50.4$ days, or NTCA $\geq 92.9 \%$.


Figure 33. Scatterplot of NTCA versus < AMO>.

### 2.8.2 Inferred Correlations Against <SOI>

Figure 34 shows the scatterplots of (a) NTC, (b) NMH, and (c) <LP> versus <SOI>, identifying the years of extreme values in each panel. Concerning NTC versus $<\mathrm{SOI}>$, the inferred regression is given approximately as $y=11.4+0.2 x$, having $r=0.35, r^{2}=0.12$, $s e=4.3$, and $c l>99 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=8.5 \%$. The mean value for $<$ SOI $>$ during the more recent interval 1995-2013 measures 2.1. Using this value for <SOI> for the year 2014 , one infers NTC $=11.9 \pm 4.3$, or NTC $\geq 8$ (rounded to the nearest whole number).

(a)


$$
\begin{aligned}
& \left\{\begin{array}{l}
y=2.468+0.094 x \\
r=0.397, r^{2}=0.157 \\
s e=1.724, c l>99.5 \%
\end{array}\right. \\
& \begin{array}{|c|c|}
\hline 14 & 20 \\
\hline 13 & 7 \\
\hline
\end{array} \Rightarrow P=7.9 \%
\end{aligned}
$$

(b)


$$
\begin{aligned}
& \left\{\begin{array}{l}
y=980.677-0.317 x \\
r=-0.325, r^{2}=0.106 \\
s e=7.331, c l>98 \%
\end{array}\right. \\
& \begin{array}{|l|l|}
\hline 16 & 12 \\
\hline 11 & 15 \\
\hline
\end{array} \Rightarrow P=20.7 \%
\end{aligned}
$$

Figure 34. Scatterplots of (a) NTC, (b) NMH, and (c) $<$ LP $>$ versus $<$ SOI $>$.

Concerning NMH versus <SOI>, the inferred regression is given approximately as $y=2.5+0.1 x$, having $r=0.40, r^{2}=0.16$, $s e=1.7$, and $c l>99.5 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=7.9 \%$. The mean value for $<$ SOI $>$ during the more recent interval 1995-2013 measures 2.1. Using this value for <SOI> for the year 2014, one infers $\mathrm{NMH}=2.7 \pm 1.7$, or $\mathrm{NMH} \geq 1$.

Concerning <LP> versus <SOI>, the inferred regression is given approximately as $y=980.7-0.3 x$, having $r=-0.33, r^{2}=0.11$, se $=7.3$, and $c l>98 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=20.7 \%$. The mean value for $<$ SOI $>$ during the more recent interval 1995-2013 measures 2.1. Using this value for <SOI> for the year 2014, one infers $<\mathrm{LP}\rangle=980.0 \pm 7.3 \mathrm{mb}$, or $\langle\mathrm{LP}\rangle<987.3 \mathrm{mb}$.

Figure 35 depicts the scatterplots of (a) total ACE and (b) LISNSD versus <SOI>, identifying the years of extreme values in each panel. Concerning total ACE versus <SOI>, the inferred regression is given approximately as $y=101.1+2.4 x$, having $r=0.32, r^{2}=0.10, s e=56.6$, and $c l>98 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=20.7 \%$. The mean value for <SOI> during the more recent interval 1995-2013 measures 2.1. Using this value for <SOI> for the year 2014, one infers total $\mathrm{ACE}=106.2 \pm 56.6$, or total $\mathrm{ACE} \geq 49.6$ (units in $10^{4} \mathrm{kt}^{2}$ ).

Concerning LISNSD versus $<$ SOI $>$, the inferred regression is given approximately as $y=11.8+0.1 x$, having $r=0.22, r^{2}=0.05, s e=3.6$, and $c l>99.8 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=39.2 \%$. The mean value for $<$ SOI $>$ during the more recent interval 1995-2013 measures 2.1. Using this value for <SOI $>$ for the year 2014, one infers LISNSD $=12.0 \pm 3.6$ days, or LISNSD $\geq 8.4$ days.


Figure 35. Scatterplots of (a) total ACE and (b) LISNSD versus $<$ SOI $>$.

Figure 36 displays the scatterplots of (a) total NSD, (b) total NHD, and (c) NTCA versus $<$ SOI $>$, identifying the years of extreme values in each panel. Concerning total NSD versus <SOI>, the inferred regression is given approximately as $y=58.5+1.5 x$, having $r=0.36, r^{2}=0.13$, se $=30.5$, and $\mathrm{cl}>99 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=50.0 \%$. The mean value for $<$ SOI $>$ during the more recent interval 1995-2013 measures 2.1. Using this value for $<$ SOI $>$ for the year 2014 , one infers total NSD $=61.6 \pm 30.5$ days, or total NSD $\geq 31.1$ days.

Concerning total NHD versus $<$ SOI $>$, the inferred regression is given approximately as $y=23.9+0.6 x$, having $r=0.32, r^{2}=0.10, s e=14.0$, and $c l>98 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=50.0 \%$. The mean value for $<$ SOI $>$ during the more recent interval 1995-2013 measures 2.1. Using this value for <SOI> for the year 2014, one infers total $\mathrm{NHD}=25.1 \pm 14.0$ days, or total $\mathrm{NHD} \geq 11.1$ days.

Concerning NTCA versus <SOI>, the inferred regression is given approximately as $y=110.3+2.8 x$, having $r=0.36, r^{2}=0.13$, $s e=57.1$, and $c l>99 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=29.3 \%$. The mean value for $<$ SOI $>$ during the more recent interval 1995-2013 measures 2.1. Using this value for <SOI> for the year 2014, one infers NTCA $=116.2 \pm 57.1 \%$, or NTCA $\geq 59.1 \%$.




Figure 36. Scatterplots of (a) total NSD, (b) total NHD, and (c) NTCA versus <SOI>.

### 2.8.3 Inferred Correlations Against < MLCO2>

Figure 37 shows the scatterplot of LOS versus <MLCO2>, identifying the years of extreme values. The inferred regression is given approximately as $y=-77.0+0.60 x$, having $r=0.34, r^{2}=0.11$, $s e=40.0$, and $c l>98 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=13.8 \%$. Because the atmospheric concentration of $\mathrm{CO}_{2}$ increases each year, it makes no sense to apply the mean of $<\mathrm{MLCO} 2>$ for the more recent interval ( 377.86 ppm ). Instead, it is better to use the average rate of increase during the more recent interval for determining the likely level of <MLCO2> for the year 2014. Hence, for the year 2014, one expects <MLCO2> to measure about $396.48+1.98 \pm 0.48 \mathrm{ppm}$, or about $398.46 \pm 0.48 \mathrm{ppm}$. Using the value $<\mathrm{MLCO} 2>=398.46 \mathrm{ppm}$ for the year 2014, one infers LOS $=161.3 \pm 40.0$ days, or $\operatorname{LOS} \geq 121$ days (rounded to the nearest whole day).


Figure 37. Scatterplot of LOS versus $<\mathrm{MLCO} 2>$.

Figure 38 depicts the scatterplot of NTC versus < MLCO2>, identifying the years of extreme values. The inferred regression is given approximately as $y=-25.4+0.11 x$, having $r=0.56, r^{2}=0.31$, $s e=3.8$, and $c l>99.9 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=5.0 \%$. Using the value of $\langle\mathrm{MLCO} 2\rangle=398.46 \mathrm{ppm}$ for the year 2014 , one infers NTC $=16.4 \pm 3.8$, or $\mathrm{NTC} \geq 13$ (rounded to the nearest whole number).


Figure 38. Scatterplot of NTC versus $<\mathrm{MLCO} 2>$.

Figure 39 displays the scatterplot of <LAT> versus <MLCO2>, identifying the years of extreme values. The inferred regression is given approximately as $y=38.4-0.05 x$, having $r=-0.33$, $r^{2}=0.11$, se $=2.7$, and $c l>99 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.1 \%$. Using the value of $<\mathrm{MLCO} 2>=398.46 \mathrm{ppm}$ for the year 2014, one infers $<$ LAT $>=20.5 \pm 2.7^{\circ} \mathrm{N}$., or $<$ LAT $><23.2^{\circ} \mathrm{N}$.


Figure 39. Scatterplot of $<$ LAT $\rangle$ versus $<$ MLCO2 $>$.

Figure 40 shows the scatterplot of total NSD versus <MLCO2>, identifying the years of extreme values. The inferred regression is given approximately as $y=-154.6+0.61 x$, having $r=0.45$, $r^{2}=0.20, s e=29.2$, and $c l>99.9 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=5.1 \%$. Using the value of $\langle\mathrm{MLCO} 2\rangle=398.46 \mathrm{ppm}$ for the year 2014, one infers total $\mathrm{NSD}=87.3 \pm 29.2$ days, or total $\mathrm{NSD} \geq 58$ days (rounded to the nearest whole day).


Figure 40. Scatterplot of total NSD versus $<$ MLCO2 $>$.

### 2.8.4 Inferred Correlations Against <GLOTI>

Figure 41 depicts the scatterplots of (a) total NSD and (b) NTCA versus <GLOTI>, identifying the years of extreme values. Concerning total NSD versus <GLOTI>, the inferred regression is given approximately as $y=42.6+60.8 x$, having $r=0.46, r^{2}=0.22$, $s e=29.0$, and $c l>99.9 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=1.4 \%$. Using the value of $<$ GLOTI $>=0.54{ }^{\circ} \mathrm{C}$ for the year 2014 (i.e., the mean value for the more recent interval 1995-2013), one infers total NSD $=75.5 \pm 29.0$ days, or total NSD $\geq 46$ days (rounded to the nearest whole day).

Concerning NTCA versus <GLOTI>, the inferred regression is given approximately as $y=89.5+79.1 x$, having $r=0.32, r^{2}=0.10, s e=58.0$, and $c l>98 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=1.4 \%$. Using the value of $<\mathrm{GLOTI}\rangle=0.54{ }^{\circ} \mathrm{C}$ for the year 2014 , one infers NTCA $=132.2 \pm 58.0 \%$, or NTCA $\geq 74.2 \%$.


Figure 41. Scatterplots of (a) total NSD and (b) NTCA versus <GLOTI>.

### 2.8.5 Inferred Correlations Against < ONI>

Figure 42 displays the scatterplots of (a) LSD and (b) NTC versus <ONI>, identifying the years of extreme values. Concerning LSD versus $<\mathrm{ONI}>$, the inferred regression is given approximately as $y=311.2-14.8 x$, having $r=-0.37, r^{2}=0.13$, $s e=22.3$, and $c l>99 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=5.1 \%$. Using the value of $\angle \mathrm{ONI}\rangle=-0.12{ }^{\circ} \mathrm{C}$ for the year 2014 (i.e., the mean value for the more recent interval 1995-2013), one infers LSD $=312.0 \pm 22.3$ (day of year (DOY)), or $\mathrm{LSD} \geq 291(\mathrm{DOY}=$ on or after about October 18, 2014).

Concerning NTC versus <ONI>, the inferred regression is given approximately as $y=11.4-2.6 x$, having $r=-0.33, r^{2}=0.11$, $s e=4.3$, and $c l>98 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.1 \%$. Using the value of $\langle\mathrm{ONI}\rangle=-0.12^{\circ} \mathrm{C}$ for the year 2014 , one infers $\mathrm{NTC}=11.7 \pm 4.3$, or $\mathrm{NTC} \geq 7$ (rounded to the nearest whole number). (For the overall interval 1960-2013, there were 14 years classified as ENY. The mean NTC during these years equals 8.7, having $s d=3.8$ and extremes of 4 to 18.)


Figure 42. Scatterplots of (a) LSD and (b) NTC versus $<\mathrm{ONI}>$.

### 2.8.6 Inferred Correlations Against < NAO> CRU

Figure 43 shows the scatterplots of (a) NH, (b) <NSD>, and (c) total NSD versus <NAO> CRU, identifying the years of extreme values in each subpanel. Concerning NH versus <NAO> CRU, the inferred regression is given approximately as $y=6.1-1.9 x$, having $r=-0.38, r^{2}=0.15$, se $=2.6$, and $c l>99.5 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=29.3 \%$. Using the value of $<\mathrm{NAO}>\mathrm{CRU}=-0.30$ for the year 2014 , one infers $\mathrm{NH}=6.7 \pm 2.6$, or $\mathrm{NH}<9$.

Concerning <NSD> versus <NAO> CRU, the inferred regression is given approximately as $y=4.9-0.8 x$, having $r=-0.32, r^{2}=0.10, s e=1.3$, and $c l>98 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=50.0 \%$. Using the value of $<\mathrm{NAO}>\mathrm{CRU}=-0.30$ for the year 2014, one infers $<$ NSD $>=5.1 \pm 1.3$, or $<$ NSD $><6$ days (rounded to the nearest whole day).

Concerning total NSD versus $<$ NAO $>$ CRU, the inferred regression is given approximately as $y=57.2-25.5 x$, having $r=-0.45, r^{2}=0.20, s e=29.3$, and $c l>99.9 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=29.3 \%$. Using the value of $\angle \mathrm{NAO}>\mathrm{CRU}=-0.30$ for the year 2014, one infers total NSD $=64.9 \pm 29.3$, or total NSD $<94$ days (rounded to the nearest whole day).



Figure 43. Scatterplots of (a) NH, (b) <NSD $>$, and (c) total NSD versus $<$ NAO $>$ CRU.

### 2.8.7 Inferred Correlations Against < NAO> CPC

Figure 44 depicts the scatterplots of (a) NH and (b) total NSD versus <NAO> CPC. Concerning NH versus $<\mathrm{NAO}>\mathrm{CPC}$, the inferred regression is given approximately as $y=6.0-2.4 x$, having $r=-0.34, r^{2}=0.12$, $s e=2.6$, and $c l>98 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=29.3 \%$. Using the value of $\langle\mathrm{NAO}\rangle \mathrm{CPC}=-0.18$ for the year 2014, one infers $\mathrm{NH}=6.4 \pm 2.6$, or $\mathrm{NH}<9$.

Concerning total NSD versus $<\mathrm{NAO}>\mathrm{CPC}$, the inferred regression is given approximately as $y=56.1-27.7 x$, having $r=-0.34, r^{2}=0.12$, $s e=30.8$, and $c l>98 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=29.3 \%$. Using the value of $\angle \mathrm{NAO}>\mathrm{CPC}=-0.18$ for the year 2014, one infers total NSD $=61.1 \pm 30.8$, or total $\mathrm{NSD}<92$ days (rounded to the nearest whole day).


Figure 44. Scatterplots of (a) NH and (b) total NSD versus <NAO> CPC.

### 2.8.8 Inferred Correlation Against < ASAT>

Figure 45 displays the scatterplot of NTC versus <ASAT>. The inferred correlation is given approximately as $y=-22.6+3.6 x$, having $r=0.43, r^{2}=0.18$, $s e=4.1$, and $c l>99.8 \%$. Fisher's exact test for $2 \times 2$ contingency tables suggests that the probability of obtaining the observed result, or one more suggestive of a departure from independence, is $P=0.6 \%$. Using the value of $\angle A S A T>=9.97^{\circ} \mathrm{C}$ for the year 2014, one infers NTC $=13.1 \pm 4.1$, or $\mathrm{NTC} \geq 9$.


Figure 45. Scatterplot of NTC versus <ASAT>.

### 2.8.9 Inferred Correlations for NTC Using Single-Variate, Bivariate, and Trivariate Fits Based on Selected Climate-Related Factors

Table 4 provides a comparison of the observed yearly NTC with those based on selected inferred single-variate fits ( $<\mathrm{AMO}>,<\mathrm{SOI}>$, and $<\mathrm{MLCO} 2>$ ), bivariate fits ( $<\mathrm{AMO}>$ and $<$ SOI $>$, and <AMO> and <MLCO2>), and on a specific trivariate fit (<AMO>, <SOI>, and <MLCO2>). From the table, one finds that estimates made using <AMO> or $<$ MLCO2 $>$, both having $r=0.55$ and $s e=3.8$, appear to be slightly more reliable than the estimates using $<\mathrm{SOI}>(r=0.35$ and $s e=4.3)$. Some improvement is found when using the bivariate fits ( $R_{y 12}=0.62$ and $S_{y 12}=3.6$ ), with the trivariate fit providing the best improvement ( $R_{y 12}=0.65$ and $S_{y 12}=3.5$ ). For the single-variate fits, the fit having the best $\pm 1$ difference between estimated and observed NTC, surprisingly, appears to be the one using <MLCO2> (having an estimate of $\pm 1$ for 20 of the 54 years, or about $37 \%$ success). Accepting a broader error range ( $\pm 3$ ), one finds that $<\mathrm{AMO}>$ is the better estimator for NTC ( 39 of 54 years, or $72 \%$ ). For the bivariate fits, the one using $<\mathrm{AMO}>$ and $<$ SOI $>$ appears to be the slightly better estimator of NTC, having $\pm 1$ for 22 of 54 years ( $41 \%$ ) and $\pm 3$ for 41 of 54 years ( $76 \%$ ). The trivariate fit has a $\pm 1$ error range for 23 of 54 years ( $43 \%$ ) and a $\pm 3$ error range for 43 of 54 years $(80 \%)$. Based on $<$ AMO $>$, the estimates for NTC were poor $(\geq \pm 5)$ for 11 of the 54 years, including the years 1960, 1962, 1969, 1971, 1974, 1983, 1987, 1995, 2005, 2006, and 2011. Based on the bivariate fit using <AMO> and <SOI>, the estimates for NTC were poor for 9 of the 54 years, including the years 1960, 1962, 1969, 1973,1983, 1995, 2005, 2006, and 2012. Based on the trivariate fit using <AMO>, <SOI>, and <MLCO2>, the estimates for NTC were poor for 9 of the 54 years, including the years 1960, 1962, 1969, 1973, 1983, 1995, 2005, 2006, and 2009. (Only NTC was examined using bivariate and trivariate fits as a test case simply to determine the amount of improvement one might expect by incorporating more than one variable for estimating a tropical cyclone parameter.)

Table 4. Comparison of observed and predicted NTC values based on selected fits.

| Year | Observed | $P(<A M O>)$ | $P(<S O 1>)$ | $P(<M L C O 2>)$ | $P(<A M O\rangle,<S O 1>)$ | $\begin{aligned} & P(<A M O> \\ & <M L C O 2>) \end{aligned}$ | $\begin{gathered} \hline P(<A M O>,<S O 1>, \\ <M L C O 2>) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1960 | 7 | 15 | 12 | 8 | 15 | 11 | 11 |
| 1961 | 11 | 13 | 12 | 8 | 13 | 10 | 10 |
| 1962 | 5 | 13 | 13 | 8 | 14 | 10 | 11 |
| 1963 | 9 | 12 | 11 | 8 | 11 | 10 | 9 |
| 1964 | 12 | 11 | 13 | 8 | 12 | 9 | 10 |
| 1965 | 6 | 10 | 8 | 8 | 7 | 8 | 7 |
| 1966 | 11 | 12 | 11 | 8 | 11 | 10 | 9 |
| 1967 | 8 | 10 | 12 | 8 | 11 | 9 | 9 |
| 1968 | 8 | 10 | 12 | 9 | 10 | 8 | 9 |
| 1969 | 18 | 12 | 10 | 9 | 11 | 10 | 9 |
| 1970 | 10 | 10 | 12 | 9 | 14 | 9 | 11 |
| 1971 | 13 | 8 | 14 | 9 | 10 | 7 | 9 |
| 1972 | 7 | 7 | 10 | 9 | 6 | 7 | 6 |
| 1973 | 8 | 9 | 13 | 9 | 16 | 8 | 13 |
| 1974 | 11 | 6 | 13 | 9 | 8 | 7 | 8 |
| 1975 | 9 | 8 | 14 | 9 | 8 | 8 | 8 |
| 1976 | 10 | 7 | 12 | 9 | 7 | 7 | 7 |
| 1977 | 6 | 9 | 9 | 10 | 8 | 9 | 8 |
| 1978 | 12 | 9 | 11 | 10 | 9 | 9 | 9 |
| 1979 | 9 | 10 | 11 | 10 | 10 | 10 | 10 |
| 1980 | 11 | 11 | 11 | 10 | 11 | 11 | 10 |
| 1981 | 12 | 11 | 12 | 10 | 11 | 10 | 10 |
| 1982 | 6 | 9 | 9 | 10 | 7 | 9 | 8 |
| 1983 | 4 | 11 | 10 | 11 | 10 | 11 | 10 |
| 1984 | 13 | 9 | 11 | 11 | 9 | 10 | 9 |
| 1985 | 11 | 8 | 12 | 11 | 9 | 10 | 10 |
| 1986 | 6 | 8 | 9 | 11 | 7 | 9 | 8 |
| 1987 | 7 | 13 | 9 | 11 | 10 | 12 | 10 |
| 1988 | 12 | 12 | 13 | 12 | 13 | 12 | 12 |
| 1989 | 11 | 11 | 13 | 12 | 12 | 11 | 11 |
| 1990 | 14 | 11 | 11 | 12 | 11 | 12 | 11 |
| 1991 | 8 | 10 | 10 | 12 | 9 | 11 | 10 |
| 1992 | 7 | 9 | 9 | 12 | 7 | 10 | 9 |
| 1993 | 8 | 9 | 10 | 12 | 8 | 10 | 10 |
| 1994 | 7 | 9 | 9 | 12 | 8 | 11 | 10 |
| 1995 | 19 | 14 | 11 | 12 | 13 | 13 | 13 |
| 1996 | 13 | 11 | 14 | 13 | 13 | 12 | 13 |
| 1997 | 8 | 12 | 9 | 13 | 10 | 13 | 11 |
| 1998 | 14 | 17 | 11 | 13 | 16 | 16 | 15 |
| 1999 | 12 | 13 | 15 | 13 | 16 | 14 | 15 |
| 2000 | 15 | 12 | 13 | 13 | 13 | 13 | 14 |

Table 4. Comparison of observed and predicted NTC values based on selected fits (Continued).

| Year | Observed | $P(<A M O>)$ | $P(<S O \mid>)$ | $P(<M L C O 2>)$ | P(<AMO>,<SOI>) | P(<AMO>, <br> <MLCO2>) | $\begin{gathered} P(<A M O>,<S O 1> \\ \text { <MLCO2> }) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2001 | 15 | 13 | 12 | 14 | 13 | 14 | 14 |
| 2002 | 12 | 13 | 10 | 14 | 12 | 14 | 13 |
| 2003 | 16 | 15 | 11 | 14 | 14 | 15 | 15 |
| 2004 | 15 | 14 | 10 | 14 | 14 | 15 | 15 |
| 2005 | 28 | 16 | 11 | 14 | 15 | 16 | 16 |
| 2006 | 10 | 15 | 11 | 15 | 15 | 16 | 16 |
| 2007 | 15 | 14 | 11 | 15 | 13 | 15 | 15 |
| 2008 | 16 | 14 | 13 | 15 | 15 | 15 | 16 |
| 2009 | 9 | 12 | 11 | 15 | 12 | 14 | 14 |
| 2010 | 19 | 16 | 13 | 16 | 18 | 17 | 18 |
| 2011 | 19 | 13 | 14 | 16 | 15 | 15 | 16 |
| 2012 | 19 | 15 | 11 | 16 | 14 | 16 | 16 |
| 2013 | 14 | 14 | 12 | 16 | 15 | 16 | 17 |
| $\pm 1$ |  | 17 | 12 | 20 | 22 | 19 | 23 |
| $\pm 2$ |  | 28 | 28 | 31 | 33 | 30 | 32 |
| $\pm 3$ |  | 39 | 38 | 36 | 41 | 40 | 43 |
| Range | 4-28 | 6-17 | 8-15 | 8-16 | 6-18 | 7-17 | 6-17 |

$P(<A M O>)=11.728+12.929<A M O>, r=0.553, s e=3.8$.
$P(<$ SOl $>)=11.440+0.200<$ SOl $>, r=0.348, s e=4.3$.
$P(<$ MLCO2 $>)=-25.399+0.105<$ MLCO2>, $r=0.555$, $s e=3.8$.
$P(<\mathrm{AMO}>,<\mathrm{SOl}>)=11.750+12.225<\mathrm{AMO}>+0.168<\mathrm{SOl}>, R_{y 12}=0.625, S_{y 12}=3.6$.
$P(<\mathrm{AMO}>,<\mathrm{MLCO} \gg)=-11.535+8.098<\mathrm{AMO}>+0.066<\mathrm{MLCO} 2>, R_{y 12}=0.620, S_{y 12}=3.6$.
$P(<\mathrm{AMO}>,<\mathrm{SOl}>,<\mathrm{MLCO} 2>)=-20.025+0.600 P(<\mathrm{AMO}>,<\mathrm{SOl}>)+0.070<\mathrm{MLCO} 2>, R_{y 12}=0.653, S_{y 12}=3.5$.

## 3. SUMMARY

This is the second part of a two-part study of the 615 tropical cyclones that occurred in the North Atlantic basin during the weather satellite era, 1960-2013. Part 1 investigated the statistics of some 25 parameters associated with the tropical cyclones (e.g., NTC, NSD, PWS, ACE, etc.). Part 2 has examined the statistics of 9 specific climate-related parameters, both in relation to each other and in relation to the aforementioned 25 parameters discussed in part 1.

Regarding the intercorrelational behavior of the climate-related factors, perhaps surprisingly, one finds that $<\mathrm{ONI}>$ and $<\mathrm{SOI}>$, factors associated with the determination of the phasing for the El Niño Southern Oscillation phenomenon, are found to correlate strongly ( $r=-0.896$ and $c l>99.9 \%$ ) only against each other and not against any of the other climate-related factors. On the other hand, <NAO>, <ASAT>, and <GLOTI> are all found to correlate strongly against <AMO>, while $<\mathrm{QBO}>$ is found not to correlate strongly against any of the other climate-related factors ( $|r|<0.21$ and $\mathrm{cl}<90 \%$ ). Also, <AMO>, <ASAT>, and <GLOTI> are found to correlate strongly against $<$ MLCO2>, especially $<$ GLOTI $>(r=0.932$ and $c l>99.9 \%$ ).

Regarding the correlational behavior of the 25 tropical cyclone parameters against the 9 climate-related factors, one finds that more of the tropical cyclone parameters correlate strongly against <AMO> than any other factor, with 15 of the 25 tropical cyclone parameters correlating very strongly against <AMO> having $c l>98 \%$ and an additional 5 tropical cyclone parameters correlating strongly against <AMO> having $c l>95 \%$. Only FSD, LSD, LOS, NUSLFH, and <PWS> fail to correlate strongly against <AMO>. The second best climate-related factor is <SOI>, with 8 tropical cyclone parameters correlating very strongly against $<$ SOI $>$ having $c l>98 \%$ and an additional 5 parameters correlating strongly against <SOI> having cl>95\%.

Interestingly, none of the climate-related factors appear to correlate with the FSD. Hence, there appears to be no way for accurately predicting ahead of time using any of the 9 climate-related factors as to when to expect the FSD of a hurricane season. On the other hand, the LSD does appear to be related to $<\mathrm{ONI}>,<\mathrm{SOI}>,<\mathrm{GLOTI}>$, and $<\mathrm{MLCO} 2>$, but in contrasting ways. For example, a positive (negative) <ONI> associated with a negative (positive) <SOI> suggests that the LSD will occur sooner (later) rather than later (sooner) in the hurricane season, while increased global warming (positive $<$ GLOTI $>$ ) and increased atmospheric concentration of $\mathrm{CO}_{2}(<\mathrm{MLCO} 2>$ ) suggests that the LSD will occur later rather than earlier during the hurricane season, with the correlation against $<\mathrm{ONI}>$ being the slightly stronger correlation ( $c l>99 \%$ ). Regarding the LOS, it has been established that once the FSD is known, an estimate can be made for the LOS. Of the 9 climaterelated factors, only < MLCO2> appears to correlate strongly with the LOS (cl>98\%). Hence, given the increasing atmospheric concentration of $\mathrm{CO}_{2}$, one expects the LOS to slowly increase in length over time.

For estimating the NTC, the climate-related factors <AMO>, <ASAT>, <GLOTI>, and $<$ MLCO2> all provide some degree of being able to reliably estimate it. Since all of these climaterelated factors are now of positive value, this would seem to indicate a strong preference for the NTC to be of average to higher-than-average number for the current hurricane season. (The long-term average is about 11 tropical cyclones per season during the overall interval 1960-2013, although since 1995, the average has been higher, about 15 tropical cyclones per season with the fewest number over the past 19 years being 8 in 1997.)

Regarding the NH and NMH, they are found to correlate more strongly against <AMO> than the other climate-related factors. Hence, as with NTC, because the <AMO> is presently in its warm (positive value) phase and is expected to remain so for another decade or so, one expects the NH and NMH probably to be of average to higher-than-average number during the current hurricane season (the long-term average of NH and NMH is 6 and 2, respectively, while being 8 and 3, respectively, over the past 19 years.)

Regarding the mean seasonal NUSLFH (<NUSLFH $>$ ), none of the climate-related factors provide any indication as to how many will occur. The <NUSLFH> appears random, ranging in number from 0 to 6 per yearly season.

Regarding the total NSD, total NHD, and total NMHD, again, because the $<\mathrm{AMO}>$ is in its warm (positive values) phase, one expects values for these tropical cyclone parameters during the current hurricane season to probably be average to higher-than-average number (the long-term averages for these parameters are about 58, 24, and 6 days, respectively, and about 82,30 , and 8 days, respectively, for the past 19 years). For the NTCA (which is based on NTC, NH, NMH, total NSD, total NHD, and total NMHD), one expects activity to be average to higher than average. (The longterm average of NTCA for the overall interval 1960-2013 is about $110 \%$, while the average over the past 19 years has been about $148 \%$.)

Recall that table 3 provides the inferred linear regressions for the 25 tropical cyclone parameters against the 9 specific climate-related factors. Below is table 5 , which gives a summary of the estimates for the tropical cyclone parameters on the basis of averages for the interval 1995-2013 using $<\mathrm{AMO}>$ and $<$ SOI $>$ as the climate forcing agents and on the basis of the projected value of <MLCO2> for the year 2014. Likewise, table 6 is included, which gives a summary of the estimates for the tropical cyclone parameters on the basis of whether the year 2014 will be classified as either an ENY or a NENY.

Table 5. Summary of estimates for the North Atlantic basin tropical cyclone parameters (rounded to nearest whole number) for the 2014 hurricane season based on the means for the interval 1995-2013 for <AMO> and the projected value of $\langle\mathrm{MLCO} 2\rangle$ for 2014.

| Parameter | 1960-2013 <br> Average | 1995-2013 <br> Average | $P(<$ AMO> $)$ | $P(<$ SOI>) | $P(<$ MLCO2>) |
| :--- | :---: | :---: | :---: | :---: | :---: |
| FSD | $180(33)$ | $170(31)$ | $179(34)$ | $180(34)$ | $165(33)$ |
| LSD | $311(24)$ | $320(24)$ | $315(24)$ | $313(24)$ | $325(24)$ |
| LOS | $133(42)$ | $152(43)$ | $137(42)$ | $134(42)$ | $161(40)$ |
| NTC | $11(5)$ | $15(5)$ | $14(4)$ | $12(4)$ | $16(4)$ |
| NH | $6(3)$ | $8(3)$ | $8(2)$ | $6(3)$ | $7(3)$ |
| NMH | $2(2)$ | $3(2)$ | $3(2)$ | $3(2)$ | $3(2)$ |
| NUSLFH | $1(1)$ | $2(2)$ | $2(1)$ | $1(1)$ | $1(2)$ |
| <LAT> | $22.5(3.3)$ | $20.9(2.8)$ | $21(3)$ | $22(3)$ | $20(3)$ |
| <LONG> | $63.5(5.8)$ | $62.0(4.6)$ | $61(5)$ | $63(6)$ | $60(6)$ |
| PWS | $124(22)$ | $127(20)$ | $130(21)$ | $125(22)$ | $122(22)$ |
| <PWS> | $72.9(8.6)$ | $71.9(9.6)$ | $74(8)$ | $74(8)$ | $68(8)$ |
| LP | $937(22)$ | $930(21)$ | $929(20)$ | $936(22)$ | $931(23)$ |
| <LP> | $980.8(7.7)$ | $980.1(7.7)$ | $979(7)$ | $980(7)$ | $983(8)$ |
| <ACE> | $8.6(3.7)$ | $8.8(3.5)$ | $10(4)$ | $9(4)$ | $7(4)$ |
| Total ACE | $100.5(59.2)$ | $134.6(64.8)$ | $132(50)$ | $106(57)$ | $126(58)$ |
| HISACE | $30.9(16.1)$ | $35.9(15.7)$ | $37(15)$ | $32(16)$ | $31(16)$ |
| <PDI> | $6.9(3.9)$ | $7.2(3.5)$ | $8(4)$ | $7(4)$ | $5(4)$ |
| Total PDI | $81.8(55.7)$ | $110.8(61.9)$ | $111(48)$ | $87(54)$ | $101(55)$ |
| HISPDI | $30.8(20.2)$ | $36.6(21.0)$ | $39(19)$ | $32(20)$ | $31(20)$ |
| <NSD> | $4.94(1.35)$ | $5.31(1.66)$ | $5(1)$ | $5(1)$ | $5(1)$ |
| LISNSD | $11.76(3.74)$ | $13.09(3.45)$ | $13(4)$ | $12(4)$ | $13(4)$ |
| Total NSD | $58.12(32.44)$ | $82.05(38.60)$ | $77(26)$ | $62(31)$ | $87(29)$ |
| Total NHD | $23.70(14.62)$ | $29.87(16.13)$ | $30(13)$ | $25(14)$ | $25(15)$ |
| Total NMHD | $5.57(5.84)$ | $8.08(6.37)$ | $8(5)$ | $6(6)$ | $7(6)$ |
| NTCA | $109.6(60.7)$ | $148.3(65.6)$ | $143(50)$ | $116(57)$ | $142(59)$ |

Table 6. Summary of estimates for the North Atlantic basin tropical cyclone parameters based on whether the year 2014 is classified as ENY or NENY (rounded to nearest whole number).

| Parameter | ENY(14) | NENY(40) | $\boldsymbol{t}$ |
| :--- | :---: | :---: | :---: |
| FSD | $190(38)$ | $176(31)$ | 1.4 |
| LSD | $299(18)$ | $315(24)$ | -2.3 |
| LOS | $110(41)$ | $140(40)$ | -2.4 |
| NTC | $9(4)$ | $12(4)$ | -2.4 |
| NH | $5(3)$ | $7(3)$ | -2.1 |
| NMH | $2(2)$ | $3(2)$ | -1.6 |
| NUSLFH | $1(1)$ | $2(2)$ | -1.8 |
| <LAT> | $25(4)$ | $22(3)$ | 2.0 |
| <LONG> | $65(7)$ | $63(5)$ | 1.2 |
| PWS | $125(21)$ | $124(22)$ | 0.1 |
| <PWS> | $71(9)$ | $73(9)$ | -0.7 |
| LP | $940(19)$ | $936(23)$ | 0.6 |
| <LP> | $983(7)$ | $980(8)$ | 1.2 |
| <ACE> | $8(4)$ | $9(4)$ | -0.8 |
| Total ACE | $72(60)$ | $110(56)$ | -2.1 |
| HISACE | $27(18)$ | $32(15)$ | -1.0 |
| <PDI> | $6(4)$ | $7(4)$ | -0.8 |
| Total PDI | $59(57)$ | $90(54)$ | -1.8 |
| HISPDI | $26(22)$ | $32(20)$ | -0.9 |
| <NSD> | $4(1)$ | $5(1)$ | -3.2 |
| LISNSD | $11(4)$ | $12(3)$ | -1.0 |
| Total NSD | $41(25)$ | $64(33)$ | -2.4 |
| Total NHD | $16(15)$ | $26(14)$ | -2.3 |
| Total NMHD | $4(6)$ | $6(6)$ | -1.1 |
| NTCA | $81(59)$ | $120(59)$ | -2.1 |

Concerning table 6, one observes that an ENY, on average, tends to have a later-occurring FSD, a sooner-occurring LSD, and a shorter LOS than a NENY, with the differences in means being statistically important ( $\mathrm{cl}>95 \%$ ) for LSD and LOS. Likewise, an ENY, on average, tends to have fewer NTC and NH than a NENY. Other statistically important differences include a higher <LAT>, lower total ACE, fewer <NSD>, fewer total NSD, fewer total NHD, and smaller NTCA than a NENY.

According to the June 5, 2014, diagnostic discussion, the CPC and the International Research Institute for Climate and Society have reported that there is a $70 \%$ chance of an El Niño developing during the Northern Hemisphere summer during the year 2014, increasing to an $80 \%$ chance during the fall and winter. As to whether or not the year will be classified as ENY or not, it is dependent simply upon how soon ENL conditions manifest themselves and how strong the event becomes. Through the first four months of the year, the ONI has been of negative value, with its values indicative of El Niño-neutral conditions. Should neutral conditions continue through July, then the year 2014 would, by the definition of ENY employed in this TP, necessarily have to be classified as a NENY, even if an El Niño event should actually develop during the latter portion of the year. During the 54 years spanning 1960-2013, there have been 14 years that had ENL conditions persisting at least 6 months within the year, or about one ENY every 4 years. The longest span between ENYs during the weather satellite era has been 5 years. Since the last ENY occurred in 2009, clearly one anticipates the year 2014 to be an ENY.

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[^0]:    *Inferred correlations having $\mathrm{cl}>98 \%$.

[^1]:    *Inferred correlation is statistically significant at $c l>95 \%$.
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