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SEGO: A COMPUTER CODE FOR UNFOLDING EXPERIMENTAL
PULSE-HEIGHT DISTRIBUTIONS

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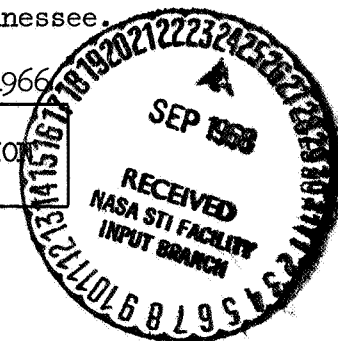
A mathematical procedure is described for unfolding scintillation pulse-height spectra to obtain an estimate of the incident energy spectra. Two methods have been developed, both giving equivalent numerical answers. The first method, known as "Low Speed Sego," reduces the number of arithmetic operations for an N channel spectrum from the order of N^3 for conventional matrix inversion methods to the order of N^2 and simplifies the problem of constructing a model for the response of the spectrometer. The second method, known as "High Speed Sego," further reduces the number of arithmetic operations to the order of $10N$ instead of N^2 . High Speed Sego makes use of the same ideas as Low Speed Sego and, in addition, transforms the response matrix of the spectrometer in such a way as to produce a large number of zero elements. This transformation is only applicable if the "Hyodo" response function can be represented by straight-line segments plus peaks (or by parabolic segments in a version under development known as "Double High Speed Sego"). (NOTE: This report was prepared primarily as a description of Low-Speed Sego, since the high-speed modifications have been described elsewhere; however, an annotated version of the high-speed program is included here, along with the results obtained when the code was applied to the standard spectra measured by R. Heath with a 3 x 3 in. NaI(Tl) spectrometer.)

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Introduction[†]

The problem of correcting experimental distributions for instrumental smearing and aberrations has long been solved for two cases. The first is when the smearing is symmetrical and the true distribution has an inherent or natural line width comparable to or broader than the instrumental smearing. The second is when the true distribution is known to be discrete with no continuous components. The first problem was essentially solved by Eddington¹ and the second by Gauss.²

In the present problem of gamma scintillation spectra, neither of the two above methods are applicable. The natural gamma-ray line width is negligible compared to the instrumental line width, and the spectrum consists of a continuous component and discrete components. Furthermore, the instrumental response is not symmetric but contains multiple peaks, tails, etc. Mathematically, the problem may be expressed as an integral equation. The desired solution must be obtained by statistical estimation, since the experimental pulse-height distributions are measured with a statistical uncertainty.

Traditionally, such problems in gamma-ray spectroscopy have been dealt with by ad hoc methods. One approach is to set up a large matrix equation and obtain the solution by a matrix inversion technique, but some procedure must be used to keep small experimental errors in the pulse-height distribution from being amplified in the result. Various techniques which have been used include: grouping the pulse-height data into rather large bins, suppressing negative components by means of special solution techniques, etc. Although these methods yield useful results, they have several difficulties:

[†] This report is primarily a description of Low Speed Sego, which was the code used to analyze the pulse height spectra representing the gamma rays given off by the Bulk Shielding Reactor II (see G. T. Chapman and W. R. Burrus, "Spectrum of Gamma Rays Emitted by a Stainless-Steel-Clad Pool-Type Reactor (BSR-II)," ORNL-TM 1284, 1968); a more detailed description of High Speed Sego can be found in an article by M. H. Young and W. R. Burrus ("A Digital Filter for Unfolding Pulse-Height Distributions") published in Nucl. Instr. Methods 62, 82(1968).

- a) A hundred or more pulse-height channels can be effectively used with modern scintillation spectrometers. Storing and inverting a matrix of this order demands the largest available computers so that frequently an ad hoc reduction of size is made resulting in loss of detail.
- b) Usually the estimate of statistical uncertainty in the final results is nonexistent or not rigorous. If negative components are set to zero, for example, a valid confidence interval might still include nonzero values. This subtlety is usually ignored.
- c) A major part of the overall problem is obtaining a catalog of calibration response functions. Although this is a straightforward technical problem, much effort must be expended in suitably interpolating the calibration functions and in obtaining absolute calibrations.

The method used for analysis of our data alleviates many of the traditional problems. It has the following features:

- a) The computation does not require that any matrix be stored. The amount of overall storage is of the order of $10n$ locations (where n is the number of pulse-height channels) so that the method is practicable on small computers.
- b) The algorithm requires only the solution to a triangular set of equations so that the number of computations is modest, being proportional to n^2 .
- c) A rigorous confidence interval for the resulting gamma spectrum is obtained.
- d) The interpolation of the calibration spectra is handled automatically, and only a minimal of calibration information must be supplied to the program.

The "Sego" Unfolding Code

The Sego code transforms the pulse-height distribution (expressed in counts $\text{channel}^{-1} \text{sec}^{-1}$) into a gamma-photon distribution (expressed in photons $\text{MeV}^{-1} \text{sec}^{-1} \text{steradian}^{-1}$). The code corrects for the variable efficiency of the spectrometer and unfolds the Compton tails and escape peak from the pulse-height distribution. The method is superficially the same as that used by Zobel³ and rediscovered many times since. In this description, we shall adopt a somewhat different point of view than that taken by Zobel.

The response of a scintillation spectrometer can be described by the integral equation

$$\int_0^{\infty} A_i(E') \phi(E') dE' = b_i + e_i \quad i = 1, 2, \dots, n \quad (1)$$

where

b_i is the observed number of counts recorded in the i th channel of a multichannel analyzer which responds to pulses from V_i to V_{i+1} in height.

e_i is the random error in the counts in the i th channel due to statistical fluctuations. We assume that an estimate of e_i can be obtained from the data (subtraction of backgrounds can be taken properly into account in the usual manner).

$A_i(E')$ is the response function of the spectrometer; i.e., the probability that a photon source of unit intensity will produce a count in the i th channel.

$\phi(E') dE'$ is the number of photons in the incident spectrum between an energy of E' and $E' + dE'$.

A two dimensional sketch of $A_i(E')$ is given in Fig. 1. Note that slices through this response surface parallel to the discrete channel number axis, $A_i(E')$ vs. i , are the conventional pulse-height distributions due to monoenergetic sources, but that slice through the response surface parallel to the energy axis, $A_i(E')$ vs. E' for fixed i , are "efficiency functions" which give the efficiency of a single channel for detection of a gamma photon.

Unfolding

As is well known, it is difficult to unfold the response of a spectrometer exactly without introducing spurious features into the desired result. Instead, what is often done (sometimes without saying so) is to leave some residual smearing in the final result so that the final gamma spectra calculated still contains some of the instrumental width. The major problem of the gamma scintillation spectrometer is not due to the instrumental smearing but is due to the

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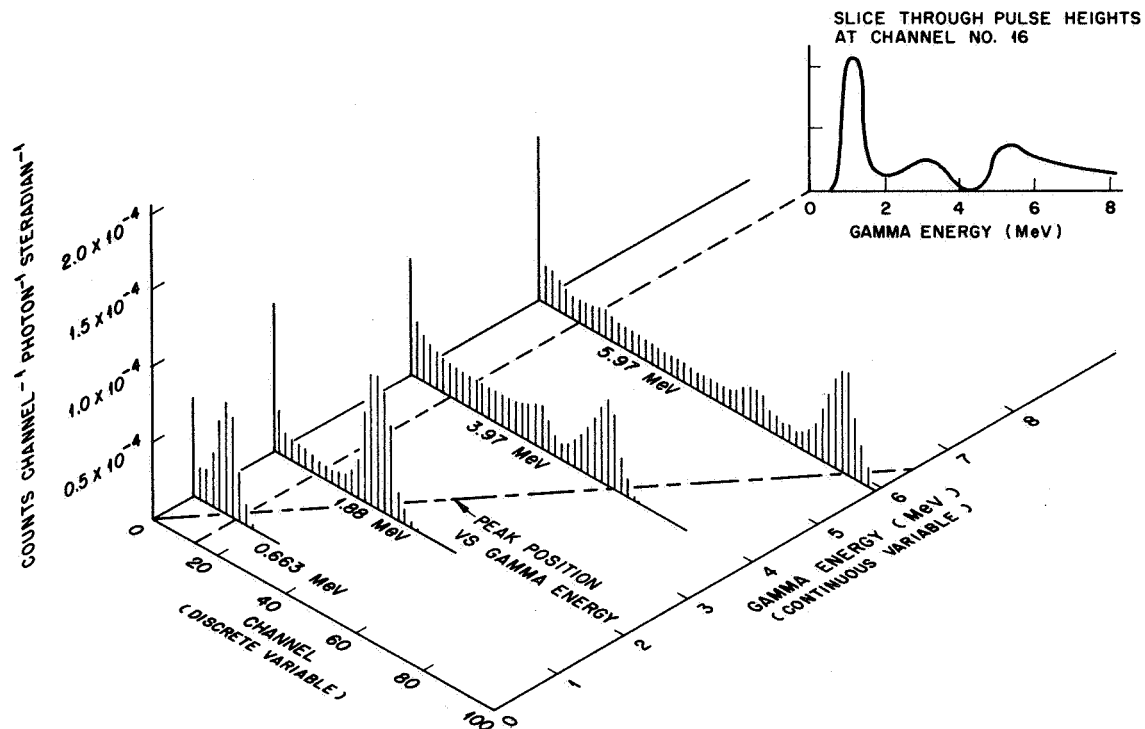


Fig. 1. Several Slices Through the Response Surface $A_i(E')$ for the 9 x 12 in. NaI(Tl) Spectrometer with Collimator.

fact that the monoenergetic response $A_i(E')$ has spurious peaks due to escape of pair production annihilation photons, etc. and a tail due to Compton scattered photons. Our approach is to admit from the beginning that we will accept the finite instrumental line width, and will try to remove the tails and spurious escape peak from the result. Let us denote the final estimated spectrum which we wish to compute by $\tilde{\varphi}(E)$. This estimated spectrum is related to the true spectrum $\varphi(E)$ by the relation

$$\tilde{\varphi}(E) = \int_0^{\infty} S(E, E') \varphi(E') dE' . \quad (2)$$

Our problem now is that we are given a statistically uncertain b_i from the experiment, and we wish to compute $\tilde{\varphi}(E)$. We may in principle compute $\tilde{\varphi}(E)$ at many values of E arbitrarily selected. But because of the nonnegativity of $\varphi(E)$, Eq. (2) implies that $\tilde{\varphi}$ will be varying at least as slowly as $S(E, E')$. It is thus adequate to compute $\tilde{\varphi}(E)$ at energies E_k , $k = 1, 2, \dots, n$ where the values of E_k are sufficiently close (say 2 to 3 per resolution width). We may think of the values of $\tilde{\varphi}(E_k)$ as the response of a conceptual pulse-height analyzer which has the response surface sketched in Fig. 2. This $W(E_k, E')$ response surface is similar to the $A_i(E')$ response surface sketched in Fig. 1 except that there is only one peak whose position is linearly related to gamma energy and whose efficiency is always constant.

The principal problem in obtaining the desired solution is finding a set of coefficients U_{ik} , $k = 1, 2, \dots, n$ for each of the n response functions $A_i(E')$ such that

$$A_i(E') = \sum_{k=1}^n U_{ik} S(E_k, E') . \quad (3)$$

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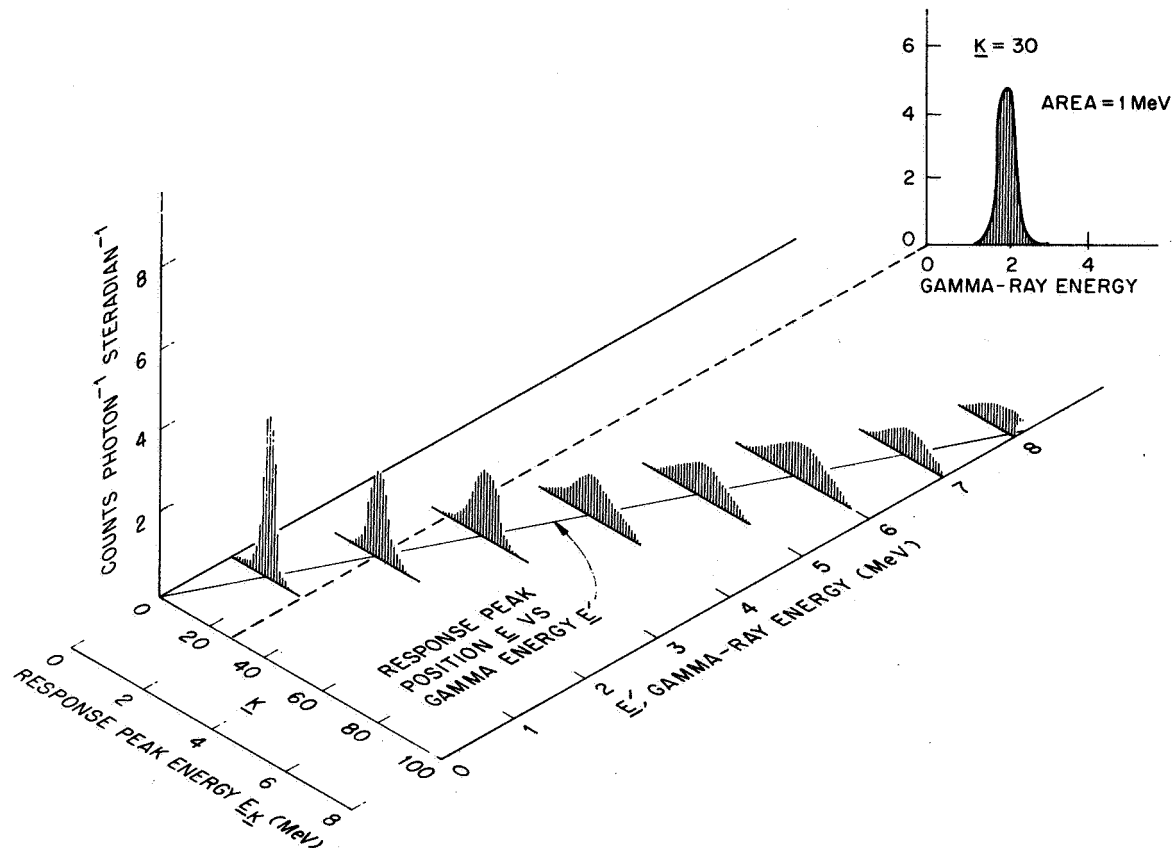


Fig. 2. Several Slices Through the Response Surface $W_k(E')$ for an "Ideal" Gamma-Ray Spectrometer.

We can achieve a rather large simplification if we accept the restriction that we will take the E_k , $k = 1, 2, \dots, n$ to be the energies at which the main peaks of $A_i(E')$ vs. E' are centered.

In illustration, Fig. 3 shows a combination of $S(E_k, E')$ functions which adds up to a particular $A_i(E')$ function. By inspection of Fig. 3 we can see that it is not necessary to use any negative coefficients and that it is not necessary to use any $S(E_k, E')$ with E_k less than the center of the peak in the $A_i(E')$ function. If, for example, E_{70} is taken as the peak of the response function $A_{70}(E')$ as shown, the peak of $A_{70}(E')$ can be fit by a single $S(E_k, E')$ function of appropriate width. It is then necessary to add together a number of other $S(E_k, E')$ functions for higher values of E_k to match the upper "tail" and any other features the response function might display.

This fitting procedure is not necessarily unique, and it may well be possible to find several different sets of nonnegative coefficients U_{ik} which will give a satisfactory fit with the experimental error in measuring $A_i(E')$. Fortunately, the nonuniqueness of the U_{ik} does not result in a nonuniqueness of the result so long as Eq. (3) is satisfied.

In order to obtain a finite set of simultaneous algebraic equations for $\tilde{\varphi}(E_k)$, substitute Eq. (3) into Eq. (1) to give:

$$\begin{aligned} \int_0^\infty A_i(E') \varphi(E') dE &= b_i + e_i \quad i = 1, 2, \dots, n \\ &= \int_0^\infty \left[\sum_{k=1}^n U_{ik} S(E_k, E') \right] \varphi(E') dE' \\ &= \sum_{k=1}^n U_{ik} \left[\int_0^\infty S(E_k, E') \varphi(E') dE' \right] \end{aligned} \quad (4)$$

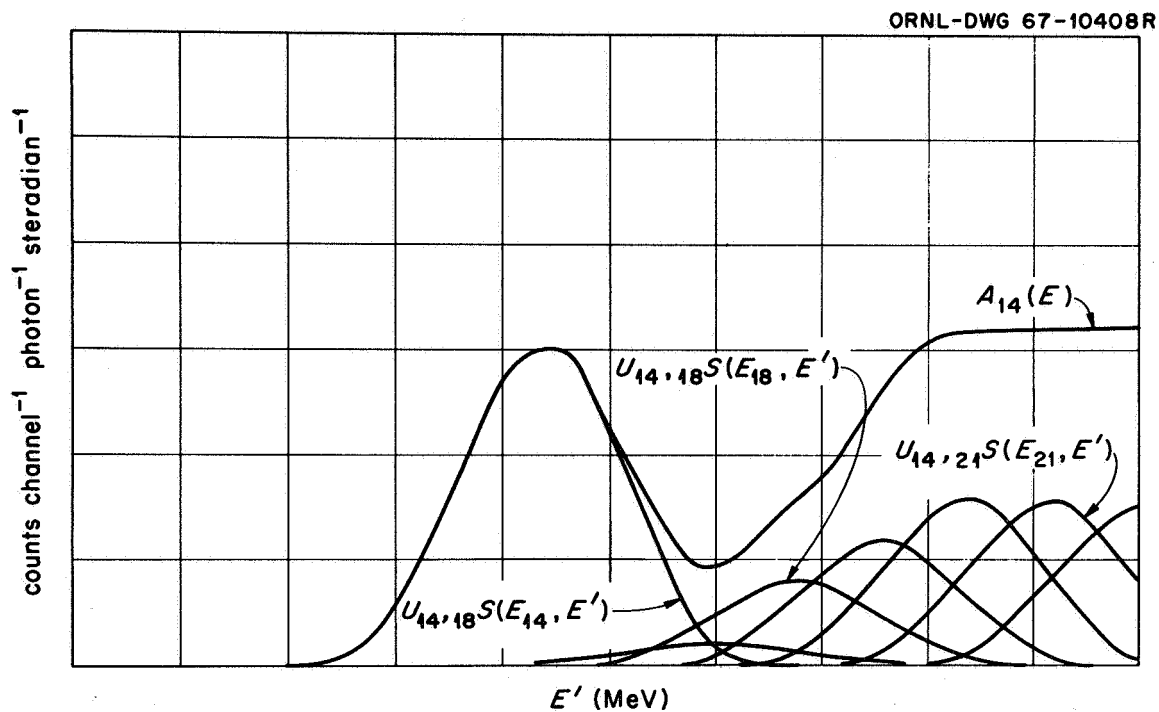


Fig. 3. An Illustration of the Combination of Several $S(E_k, E')$ Functions Which Add Up, with Weight Factors U_{ik} , to Give an $A_i(E)$ Response Function. For illustration, $i = 14$ is shown.

$$= \sum_{k=1}^n U_{ik} \tilde{\varphi}(E_k) .$$

Note that we have interchanged the order of integration and summation and have made use of the definition of $\tilde{\varphi}(E)$ given by Eq. (2).

The estimated spectrum $\tilde{\varphi}(E)$ may be obtained directly from a solution of Eq. (4) using the measured b_i for an estimate of the right-hand side. Since it is not necessary to use any E_k less than the center of the peak of $A_i(E')$ in Eq. (4), the equations given by Eq. (4) have the triangular form

$$\begin{aligned} U_{1,1} \tilde{\varphi}(E_1) + U_{1,2} \tilde{\varphi}(E_2) + \dots + U_{1,n-1} \tilde{\varphi}(E_{n-1}) + U_{1,n} \tilde{\varphi}(E_n) &\approx b_1 \\ U_{2,2} \tilde{\varphi}(E_2) + \dots + U_{2,n-1} \tilde{\varphi}(E_{n-1}) + U_{2,n} \tilde{\varphi}(E_n) &\approx b_2 \\ \dots &\dots \dots \\ U_{n-1,n-1} \tilde{\varphi}(E_{n-1}) + U_{n-1,n} \tilde{\varphi}(E_n) &\approx b_{n-1} \\ U_{n,n} \tilde{\varphi}(E_n) &\approx b_n . \end{aligned} \tag{5}$$

By solving first for $\tilde{\varphi}(E_n)$ and then working backward toward $\tilde{\varphi}(E_1)$, we can minimize the amount of arithmetic required. This method of back substitution gives the solution simply as

$$\tilde{\varphi}(E_n) \approx (1/U_{n,n})b_n$$

$$\tilde{\varphi}(E_{n-1}) \approx (1/U_{n-1,n-1})[b_{n-1} - U_{n-1,n}\tilde{\varphi}(E_n)]$$

...

(6)

$$\tilde{\varphi}(E_j) \approx (1/U_{jj})\left[b_j - \sum_{k=j+1}^n U_{jk}\tilde{\varphi}(E_k)\right]$$

...

$$\tilde{\varphi}(E_1) = (1/U_{1,1})\left[b_1 - \sum_{k=2}^n U_{1k}\tilde{\varphi}(E_k)\right]$$

Statistical Error

Since we use the measured b_i as an estimate for the true $b_i + e_i$, the resultant estimate of $\tilde{\varphi}(E_j)$ will have some statistical uncertainty. In order to compute the standard error of the $\tilde{\varphi}(E_j)$, we could first express all the solutions of Eq. (5) in the inverse form:

$$\tilde{\varphi}(E_n) = F_{n,n}b_n$$

$$\tilde{\varphi}(E_{n-1}) = F_{n-1,n-1}b_{n-1} + F_{n-1,n}b_n$$

...

$$\tilde{\varphi}(E_1) = F_{1,1}b_1 + \dots + F_{1,n}b_n$$

Then we could determine the estimated standard error in $\varphi(E_j)$ from the usual formula for the error in a sum

$$\sigma [\tilde{\varphi}(E_j)] = \sqrt{\sum_{i=1}^n F_{ji}^2 \sigma^2[b_i]} \quad (8)$$

But this calculation requires that the coefficients F_{ji} be available. Since we do not need the F_{ji} coefficients explicitly to obtain the solution $\tilde{\varphi}(E_j)$, it would be desirable to obtain the standard error of $\tilde{\varphi}(E_j)$ directly from the back solution scheme also, without the necessity of computing the inverse coefficients F_{ji} .

There does not seem to be an easy straightforward way of doing this. However a rather nice stochastic method is available based upon generating a set of random perturbations, d_i , by means of a numerical pseudo random number generator. This stochastic method allows the desired standard deviations to be obtained with many fewer arithmetic operations than would be required to evaluate the inverse coefficients F_{ji} . Furthermore, no extra computer storage is required. Suppose that we are given an experimentally measured pulse-height distribution b_i , $i = 1, 2, \dots, n$, and the corresponding standard deviations s_i , $i = 1, 2, \dots, n$. We can stochastically perturb the given b_i by adding to each term a randomly selected "normal deviate" chosen from the Gaussian distribution with mean 0 and standard deviation s_i . In this way we obtain a perturbed pulse-height distribution b'_i , $i = 1, 2, \dots, n$. Starting again with the experimental distribution b_i , $i = 1, 2, \dots, n$, we can independently select another set of perturbations to obtain b''_i , $i = 1, 2, \dots, n$, etc. as many times, r , as is desired.

Equation (5) is solved with each of the perturbed pulse-height distributions to give solutions $\tilde{\varphi}'(E_j)$, $j = 1, 2, \dots, n$; $\tilde{\varphi}''(E_j)$, $j = 1, 2, \dots, n$, etc. The standard deviation in $\tilde{\varphi}(E_j)$ can then be estimated from

$$\text{std}[\tilde{\varphi}(E_j)] = \sqrt{\frac{1}{r-1} \sum_{p=1}^r [\tilde{\varphi}^{(p)}(E_j) - \overline{\tilde{\varphi}(E_j)}]^2}$$

where $\overline{\tilde{\varphi}(E_j)}$ is the average of $\tilde{\varphi}'(E_j)$, $\tilde{\varphi}''(E_j)$, ..., $\tilde{\varphi}^{(r)}(E_j)$. As the number of perturbations r increases, the stochastic estimate of the desired standard deviation approaches that given by Eq. (8). The ratio between the result of Eq. (9) and Eq. (8) is given by $\sqrt{r/(r-1)}$ so that a choice of r anywhere from 3 to 10 is usually ample.

U Coefficient Method

Part of the *raison d'être* of the SEGO method is that a goodly portion of the work of unfolding is done in advance in the determination of the U_{ik} coefficients in Eq. (3). If one has many monoenergetic calibration sources, one may interpolate by suitable functions between the calibration sources and then determine the coefficients U_{ik} by a least squares fitting procedure. However, the problem of interpolating between available calibration energies is a formidable problem which we would like to simplify if possible. In the present case, we have made use of the physical characteristics of the response functions to reduce the amount of information needed to construct the U_{ik} coefficients.

The procedure we have used is to "parameterize" a simple piecewise continuous plus discrete approximation to the pulse-height distributions.

These parameters are listed in tabular form at convenient energies, and linear interpolation is used for intermediate values. Figure 4 shows the piecewise plus discrete representation of a typical pulse-height distribution which shows two escape peaks, a tail, and a backscatter peak. Since the position of the escape peaks and the "edge" of the tail can be accurately computed from simple formulas, sufficient parameters to characterize this distribution are

TOTAL EFFICIENCY
FIRST ESCAPE PEAK FRACTION
SECOND ESCAPE PEAK FRACTION
TAIL FRACTION
BACKSCATTER PEAK FRACTION
BACKSCATTER PEAK LOCATION
BACKSCATTER PEAK WIDTH.

The close resemblance between the U_{ik} coefficients and the pulse-height distributions is responsible for the convenience of the response surface representation. The primary difference seen in Fig. 4 is that the features of the parameterized distribution are more abrupt, with the total absorption and escape peaks being replaced by a single "spike" of about the same area. Similar the tail with a gradual "roll off" is replaced by a trapezoid with an abrupt fall. Finally the backscatter peak is replaced with a trapezoid of about the same width of the backscatter peak.

One may imagine that there is also a U_{ik} response surface which resembles the $A_i(E')$ response surface but which in general has more abrupt features. When the U_{ik} is smeared with the $S(E, E')$ function according to Eq. (3), the effect is to round off the abrupt edges so that the smeared U_{ik} surface will be identical to the $A_i(E')$ surface. It must be borne in mind that the smearing is not done along the channel number axis, but along the energy axis. Thus one can not apply the smoothing to a parameterized pulse-height distribution

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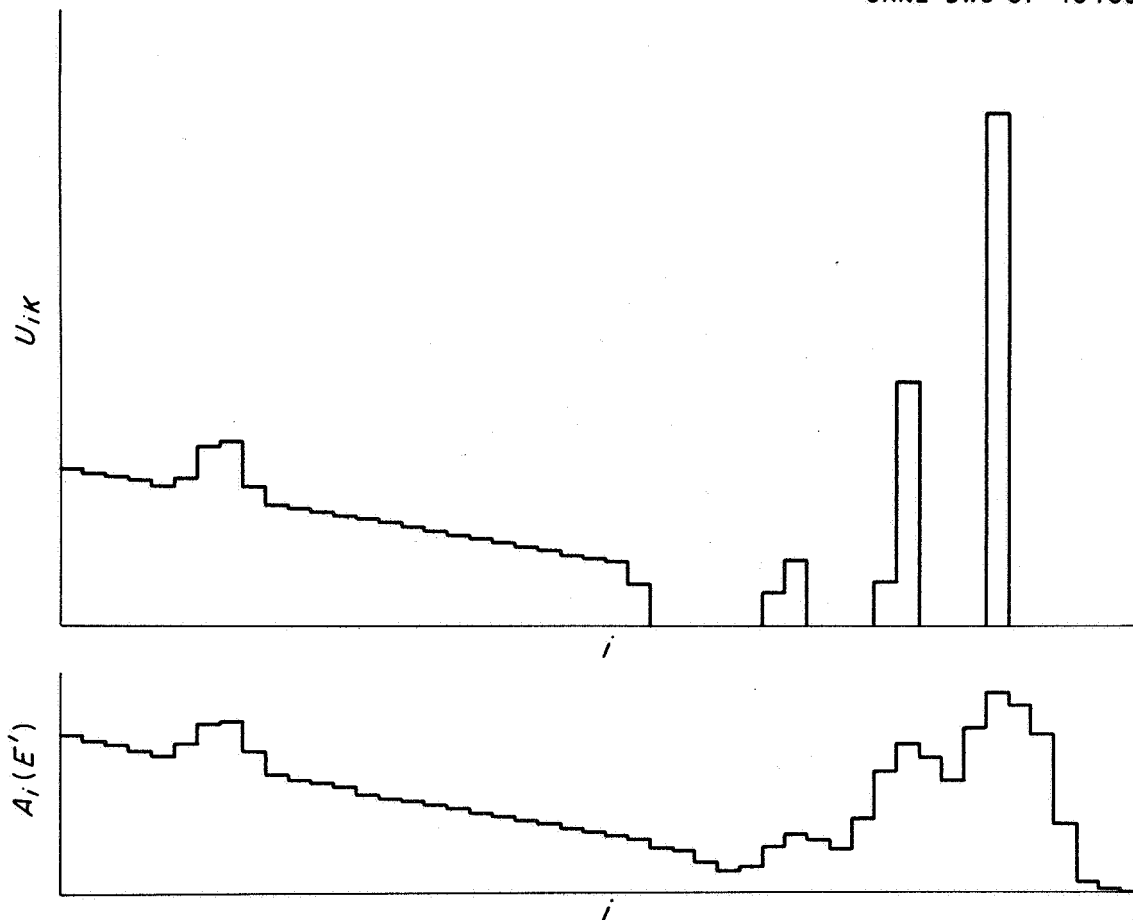


Fig. 4. Comparison of the Response Function $A_i(E')$, for a Particular Value of E' , and the Discrete U_{ik} Representation for the Corresponding Energy E_k . The graphs show two escape peaks, a tail, and a backscatter peak in addition to the main total absorption peak at the right.

directly - the pulse-height distributions are slices of the response surface in the wrong direction. Instead one must parameterize a family of calibration pulse-height distributions, and then apply the $S(E, E')$ smearing by interpolating the parameterized distributions along the E' direction.

A little thought shows that the replacement of a peak by a "spike" is plausible if the location of the peak is proportional to energy, but not if the location of the peak is fixed (or almost fixed like the backscatter peak). In the case of a fixed peak, it is necessary to parameterize the width of the peak also.

In practice, a good method seems to be to guess the parameters from inspection of the calibration spectra, by approximately matching areas of the various parameterized features with the corresponding features in the mono-energetic calibration pulse-height distributions. Then the calibration spectra are unfolded by the method. Errors in the parameterization can then be corrected iteratively. The close correspondence of the parameters to physical features of the pulse-height distribution makes it easy to "tune out" an objectionable bump in the final result by varying the value of the parameter. Of course the total efficiency of the parameterized functions and the "natural" functions must be kept the same in order to insure that the absolute magnitude of the result is correct.

If one wished to obtain the parameters from first principles without experimental "adjustment", one could construct a model of the entire "natural" response function, and then use a least squares fitting procedure to obtain the best match between the natural response function and the smeared parameterized response function. But this would be failing completely to make use of the

spirit and advance of the method which was formulated precisely to make such tedious fitting unnecessary.

The final parameters adopted for a 9 by 12 in. collimated NaI(Tl) crystal are given in Table 1. These parameters were based on matching calibration spectra from ^{24}Na , ^{88}Y , ^{16}N , and ^{137}Cs sources placed outside the collimator and along its axis. The absolute efficiencies are based on geometry and total gamma-ray cross sections of NaI, modified by a collimator correction. Back-scatter and second escape peak fractions were negligible and thus are not included.

Table 1. Parameters for 9 x 12 in.
Collimated NaI(Tl) Crystal

Energy (MeV)	Peak/Total Ratio	Tail Edge (MeV)	First Escape Fraction	$\epsilon(E)$, Efficiency (counts-cm ² -ster/photon)
$\times 10^{-5}$				
0	1.00	0	0	3.27
0.25	0.927	0.124	0	3.45
0.50	0.889	0.331	0	3.77
0.75	0.864	0.559	0	4.07
1.00	0.845	0.796	0	4.37
1.25	0.830	1.04	0.0008	4.51
1.50	0.820	1.28	0.0024	4.80
1.75	0.814	1.53	0.0046	4.94
2.00	0.809	1.77	0.0082	5.07
2.25	0.808	2.02	0.0116	5.14
2.50	0.808	2.27	0.0151	5.20
2.75	0.806	2.52	0.0184	5.23
3.00	0.804	2.76	0.0213	5.27
3.25	0.801	3.01	0.0248	5.28
3.50	0.800	3.26	0.0279	5.29
3.75	0.800	3.51	0.0306	5.29
4.00	0.800	3.76	0.0334	5.28
4.25	0.800	4.01	0.0360	5.27
4.50	0.802	4.26	0.0385	5.26
4.75	0.804	4.51	0.0408	5.24
5.00	0.805	4.76	0.0435	5.24
5.25	0.808	5.01	0.0452	5.22
5.50	0.811	5.26	0.0475	5.19
5.75	0.814	5.51	0.0496	5.17
6.00	0.817	5.76	0.0515	5.15
6.25	0.821	6.00	0.0535	5.14
6.50	0.824	6.25	0.0556	5.12
6.75	0.827	6.50	0.0576	5.11
7.00	0.830	6.75	0.0592	5.09
7.25	0.834	7.00	0.0611	5.07
7.50	0.836	7.25	0.0628	5.06
7.75	0.839	7.50	0.0644	5.04
8.00	0.842	7.75	0.0658	5.02
8.25	0.844	8.00	0.0678	5.01
8.50	0.846	8.25	0.0692	5.00
8.75	0.848	8.50	0.0708	4.99
9.00	0.850	8.75	0.0722	4.97
9.25	0.852	9.00	0.0737	4.96
9.50	0.854	9.25	0.0750	4.94
9.75	0.855	9.50	0.0765	4.93
10.00	0.856	9.75	0.0780	4.92
10.25	0.857	10.0	0.0792	4.90
10.50	0.858	10.3	0.0805	4.89
10.75	0.860	10.5	0.0819	4.88
11.00	0.861	10.6	0.0831	4.87
11.25	0.862	11.0	0.0844	4.85
11.50	0.864	11.3	0.0858	4.84
11.75	0.865	11.5	0.0870	4.83
12.00	0.866	11.7	0.0881	4.82

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Appendix I. Main FORTRAN Subroutines for Low Speed SEGO

The two main FORTRAN routines are attached as an APPENDIX. The main routine which performs the back substitution is called SEGO.

The A(I,J) routine calculates the values of the response as they are needed by SEGO. The calling program must have already prepared an appropriate list of PTOT (peak to total ratio), EDGE (Compton edge location in channels), SLOPE (slope of the Compton tail), and TAIL (the average height of the Compton tail). Escape peaks are not included in this version but their addition is trivial.

```

SUBROUTINE SEG0(C,SC,G,N)
C
C SUBROUTINE SEG0 IS A SIMPLE BACK SUBSTITUTION ROUTINE....
C BUT 9 ADDITIONAL SOLUTIONS ARE COMPUTED BY CALCULATING PERTURBED
C DATA.... BERFN IS AN INVERSE ERROR FUNCTION ROUTINE FOR
C GENERATING NORMAL RANDOM DEVIATES.....
C A(I,J) IS A SUBROUTINE WHICH RETURNS THE RESPONSE MATRIX ELEMENTS
C
DIMENSION C(402),SC(402),G(402,10)
C BACK SUBSTITUTE FOR 10 PERTURBED SOLUTIONS
DO 29 I = 1,N
    I = N+1
    G(I,1) = C(I)
    DO 16 L = 2,10
        G(I,L) = C(I) + SC(I) * BERFN(RANF(0)=0.5)
        IF (I=N) 18,22,22
    18 III = I+1
        DO 20 K = III,N
            AIK = A(I,K)
            DO 20 L = 1,10
                G(I,L) = G(I,L) + AIK*G(K,L)
                AII = 1.0/A(I,I)
            20 29 L = L+1,10
            G(I,L) = G(I,L) * AII
    29 RETURN
END

```

```

C      FUNCTION A(I,J)
C      THIS SUBROUTINE GENERATES THE HYDRO-RESPONSE MATRIX. THE
C      EFFICIENCY CORRECTION IS DEFERRED SO THAT THE RESULTING SPECTRA
C      MAY BE COMPARED DIRECTLY WITH THE RAW COUNT DATA, C(I),...
C      THE RESULTING SPECTRA SHOULD LATER BE DIVIDED BY THE EFFICIENCY,
C      THAT IS EFF(J),....
C
C      COMMON PTOT(402),EDGE(402),SLOPE(402),EFF(402),TAIL(402)
C
C      30 IF(I-J)34,32,33
C      32 A=PTOT(J)
C      33 GO TO 35
C      33 WRITE (51,933)
C      933 FORMAT(29H0YOU CANT GET THERE FROM HERE)
C      GO TO 88
C      34 A=0.0
C      35 FI=1
C      IF(EDGE(J)+1.0-FI)88,88,38
C      38 DIF=EDGE(J)-FI
C      TILT=1.0*SLOPE(J)*((2.0-FI-EDGE(J))/EDGE(J))
C      IF(DIF)40,40,41
C      40 A=TILT*((1.0-DIF)*TAIL(J)/EDGE(J))+A
C      GO TO 88
C      41 A=TILT*(TAIL(J)/EDGE(J))+A
C      88 RETURN
C      END

```

Appendix II. High Speed SEGO

The following pages give an annotated version of the program described in an article accepted for publication in Nuclear Instruments and Methods (see M. H. Young and W. R. Burrus, "A Digital Filter for Unfolding Pulse-Height Distributions").

SEGO - EFN SOURCE STATEMENT - IFN(S) -

```

C
C #####
C * SEGO UNFOLDING FILTER. CALLS THE SUBROUTINE REBIN *
C * IF LINEARITY CORRECTION IS DESIRED. *
C #####
C
C PROGRAM SEGO
C DIMENSION B(400),XRUN(4),PLOT(61),EF(21),ES(21),PT(21),
1 EC(21),BK(21),X(400,4),VU(21),VL(21),VH(21),ID(32)
COMMON B
C
C *****
C1 * READ IN THE PLOTTING SYMBOLS, THE RESPONSE PARAMETERS*
C * (TOTAL EFFICIENCY, PEAK TO TOTAL RATIO, FIRST ESCAPE *
C * PEAK TO TOTAL FRACTION, SECOND ESCAPE PEAK TO TOTAL *
C * FRACTION, BACK SCATTERING PEAK TO TOTAL FRACTION, *
C * VALLEY-FILL FUNCTION UPPER EDGE IN MEV, VALLEY-FILL *
C * FUNCTION LENGTH IN MEV, AND VALLEY-FILL FUNCTION TO *
C * TOTAL FRACTION), NO. OF CHANNELS, CHANNEL WIDTH (MEV/*
C * CH), ZERO ENERGY CHANNEL, CHANNEL-COMBINATION FACTOR *
C * (USE EVERY CHANNEL IF CHFT#1.,USE BINS OF TWO GROUPED*
C * CHANNELS IF CHFT#2.), REBIN BYPASS COMMAND, IDENTIFI-*
C * CATION OF THE SPECTRUM, AND THE SPECTRUM (COUNTS/CH.)*
C *****
C
1 READ(5,100)PLOT,SYM1,SYM2
100 FORMAT(6)A1,2A1)
2 READ(5,101)(EF(M),PT(M),ES(M),EC(M),BK(M),VU(M),VL(M)
1 ,VH(M),M#1,21)
101 FORMAT(7XE9.2,F7.4,7X,6F7.4)
WRITE (6,109)
109 FORMAT(1H1)
DO 999 JJ#2,21
VL(JJ)#.333*VL(JJ)
999 VH(JJ)#1.5*VH(JJ)
WRITE(6,99)(EF(M),PT(M),ES(M),EC(M),BK(M),VH(M),VU(M)
1 ,VL(M),M#1,21)
99 FORMAT(7XE9.2,7F10.4)
3 READ (5,107) N,G,ZO,CHFT,NRB,ID
107 FORMAT(I3,3F10.5,I7,32A1)
IF (N) 45,45,4
4 READ (5,103) (B(M),M#1,N)
WRITE (6,109)
WRITE(6,103) (B(M),M#1,N)
103 FORMAT(12X,10F6.0)
ESP#.511/G
BW#8./CHFT
KH#BW/2.
IF(NRB)6,6,5
C
C *****
C2 * NBR#1 TO REBIN, NBR#0 FOR LINEAR GAIN (BYPASS REBIN). *
C *****
C
5 CALL REBIN(N,G,ZO,CHFT)

```

SEGN - EFN SOURCE STATEMENT - IFN(S) -

```

WRITE(6,103) (R(M),M#1,N)
C
C *****
C3 * PLOTTING SCALE FACTOR DETERMINATION. *
C *****
C
6 BIGST # 0.
DO 8 I#1,N
FJ#I
ENERGY#(FJ-Z0)*G
IE#2,*ENERGY
Z#2,*ENERGY-FLOAT(IE)
EFF#(1.-Z)*EF(IE+1)+Z*EF(IE+2)
DIAG#(1.-Z)*PT(IE+1)+Z*PT(IE+2)
TERM # B(I)/(EFF*DIAG)
IF(BIGST - TERM) 7,7,8
7 BIGST # TERM
8 CONTINUE
S#50./BIGST
WRITE(6,102) N,G,Z0,S,CHFT,NRB
102 FORMAT(1X,I3,2F10.5,E10.3,F10.2,I7)
C
C *****
C4 * PERTURBING THE SPECTRUM 4 TIMES BY A RANDOM NUMBER *
C * SCHEME. *
C *****
C
NRAN#78125
DO 10 I#1,N
X(I,1)#B(I)
SQ#SQRT(R(I)+0.5)
DO 10 L#2,4
NRAN#MOD(NRAN*78125,32768)
RAN#3.0517578E-5*FLOAT(NRAN)-0.5
T#SQRT(-(ALOG((0.5-ABS(RAN))*2)))
BERF#T-(2.30753+.27061*T)/(1.+T*(.99229+.04481*T))
BERF#SIGN(BERF,RAN)
10 X(I,L)#B(I)+SQ*BERF
C
C *****
C4.5* OBTAIN DIFFERENTIAL SPECTRUM. *
C *****
C
DO 15 L#1,4
X(N+1,L)#0.
DO 15 I#1,N
15 X(I,L)#X(I,L)-X(I+1,L)
C
C *****
C5 * INITIATING OF THE MAIN LOOP, *
C *****
C
XRUN(1)#0.
XRUN(2)#0.
XRUN(3)#0.
XRUN(4)#0.

```

SEGN - EFN SOURCE STATEMENT - IFN(S) -

```

J#N+1
WRITE (6,109)
WRITE (6,104) ID,S
104 FORMAT(5X29HRUNSUM RSPL0T ERACC ENERGY,3X,32A1,3X,
1 9HP-SCALE #,E11.4)
C
C *****
C6 * THE MAIN LOOP. *
C *****
C
DO 40 JJ#1,N
JJ#1
AUP#-1.E30
ALO#1.E30
FJ#J
ENERGY#(FJ-Z0)*G
IE#2.0*ENERGY
Z#2.0*ENERGY-FLOAT(IE)
Y#1.-Z
IE1#IE+1
IE2#IE+2
C
C *****
C7 * CALCULATE THE JTH SHAPE PARAMETERS BY INTERPOLATION. *
C *****
C
EFF#Y*FF(IE1)+Z*EF(IE2)
ESC1#Y*ES(IE1)+Z*ES(IE2)
ESC2#Y*EC(IE1)+Z*EC(IE2)
BKSP#Y*BK(IE1)+Z*BK(IE2)
DIAG#Y*PT(IE1)+Z*PT(IE2)
VFUP#Y*VU(IE1)+Z*VU(IE2)
VFWE#Y*VL(IE1)+Z*VL(IE2)
VFAF#Y*VH(IE1)+Z*VH(IE2)
C
C *****
C8 * CALCULATE THE VALLEY FILLED FUNC. POSITION AND HEIGHT*
C *****
C
FUP#VFUP/G+Z0+0.5
VFW#VFWE/G
FLO#FUP-VFW
KU#FUP
KL#FLO
FU#FUP-FLOAT(KU)
FL#FLO-FLOAT(KL)
VFHT#VFAF/VFW
C
C *****
C9 * CALCULATE THE COMPTON EDGE POSITION AND TAIL HEIGHT. *
C *****
C
EDGE#ENERGY/(1.+.2555/ENERGY)
EGG#EDGE/G
TAIL#(1.-DIAG-ESC1-ESC2-BKSP-VFAF)/EGG

```

SFGD - EFN SOURCE STATEMENT - IFN(S) -

```

FKEDG#EGG+Z0+0.5
KEDG#FKEDG
FTAIL#FKEDG-FLOAT(KEDG)
C
C *****
C10 * CALCULATE THE POSITIONS OF THE BACK SCATTERING AND *
C * THE FIRST AND SECOND ESCAPE PEAKS. *
C *****
C
KB#FJ-EGG+0.5
K1#KB+KH
K2#KA-KH
KA#FJ-ESP+0.5
KC#FJ-2.*ESP+0.5
C
C *****
C11 * THE SUBTRACTION LOOP FOR THE 4 PERTURBED SPECTRA. *
C *****
C
DU 35 1#1,4
ANS#X(J,L)/DIAG
X(J,L)#ANS
C
C *****
C12 * CALCULATE THE RUNNING SUM OF THE UNSCRAMBLLED SPECTRUM*
C * AND THE UPPER AND LOWER RUNNING LIMITS. *
C *****
C
XRUN(L)#XRUN(L)+ANS/EFF
AUP#AMAX1(AUP,XRUN(L))
ALN#AMIN1(ALN,XRUN(L))
C
C *****
C13 * THE SUBTRACTION SCHEME OF THE DIFFERENTIAL SPECTRUM. *
C *****
C CALCULATE THE (J-1)TH X.
IF(JJJ)35,35,16
16 X(JJJ,L)#X(JJJ,L)+ANS*DIAG
C REMOVE THE COMPTON TAIL.
IF(KEDG-1)25,25,17
17 X(KEDG-1,L)#X(KEDG-1,L)-ANS*TAIL*(1.-FTAIL)
X(KEDG,L)#X(KEDG,L)-ANS*TAIL*FTAIL
C REMOVE THE FIRST ESCAPE PEAK.
IF(KA-1)23,23,20
20 X(KA-1,L)#X(KA-1,L)+ANS*ESC1
X(KA,L)#X(KA,L)-ANS*ESC1
C REMOVE THE SECOND ESCAPE PEAK.
IF(KC-1)23,23,22
22 X(KC-1,L)#X(KC-1,L)+ANS*ESC2
X(KC,L)#X(KC,L)-ANS*ESC2
C REMOVE THE BACK SCATTERING PEAK.
23 IF(KH)25,25,24
24 X(K1,L)#X(K1,L)-ANS*BKSP/BW
IF(K2)25,25,18
18 X(K2,L)#X(K2,L)+ANS*AKSP/BW
C REMOVE THE VALLEY FILLED FUNCTION.

```

SFGN - EFN SOURCE STATEMENT - IFN(S) -

```

25 IF(KU-1)35,35,26
26 X(KU-1,L)#X(KU-1,L)-ANS*VFHT*(1.-FU)
   X(KU,L)#X(KU,L)-ANS*VFHT*FU
27 IF(KL-1)35,35,28
28 X(KL-1,L)#X(KL-1,L)+ANS*VFHT*FL
   X(KL,L)#X(KL,L)+ANS*VFHT*(1.-FL)
35 CONTINUE
C
C *****
C14 * EVALUATE THE ERROR FROM THE UPPER AND THE LOWER *
C * LIMITS OF THE PERTURBED SPECTRA. *
C *****
C
36 XRS#XRUN(1)*S
   ERACC#.6*(AUP-ALD)*S
   KX#10.5+X(J,1)*S/EFF
   KX#MINO(MAXO(KX,1),61)
C
C *****
C15 * PLOT THE UNSCRAMLED SPECTRUM. *
C *****
C
   SAVE#PLOT(KX)
   PLOT(KX)#SYM1
   WRITE(6,105)XRUN(1),XRS,ERACC,ENERGY,(PLOT(L),L#1,KX)
40 PLOT(KX)#SAVE
105 FORMAT(1X,F12.5,F7.2,F7.3,F7.2,61A1)
   GO TO 3
45 STOP
   END

```

REBIN1 - EFN SOURCE STATEMENT - IFN(S) -

```

C #####
C # SUBROUTINE TO MAKE LINEARITY CORRECTION. #
C # THIS VERSION USES P. HEATH 5TH ORDER POLYNOMIAL FOR #
C # THE 3 IN. X 3 IN. NA-I(TL) DETECTOR. #
C #####
C
      SUBROUTINE REBIN(N,G,ZO,CHFT)
      DIMENSION SUM(200),B(200)
      COMMON B
      DO 5 K#1,200
      5 SUM(K)#0.

C
C *****
C * TO FIRM A RUNNING SUM. *
C *****
C
      DO 10 J#1,N
      10 I#N-J+1
      SUM(I+1)#SUM(I+2)+B(I)
      SUM(1)#SUM(2)+B(1)
      DO 20 J#1,N

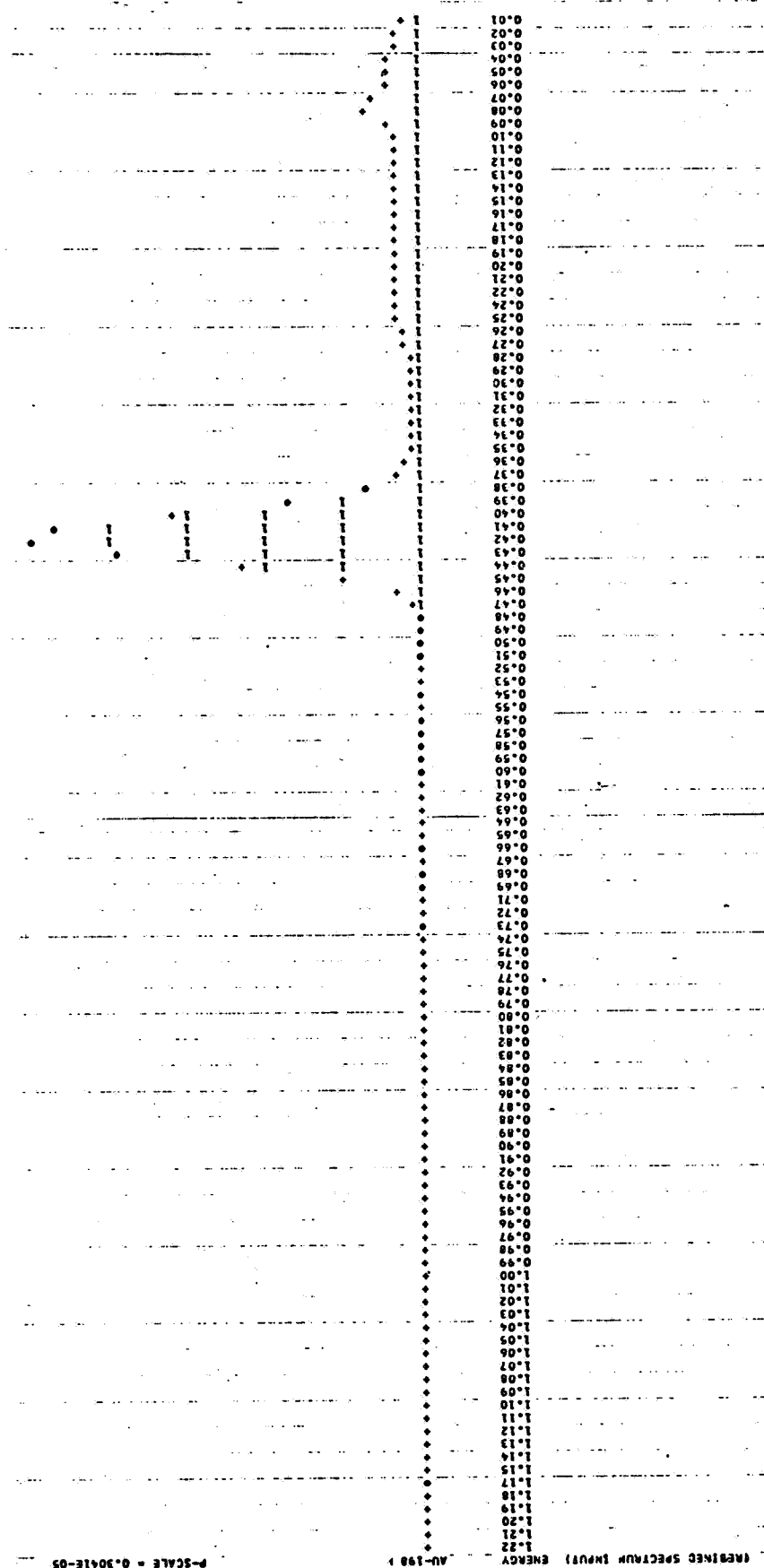
C
C *****
C * CALCULATE THE ENERGIES AT THE UPPER AND LOWER EDGES *
C * OF A CHANNEL AND CONVERT THEM TO CHANNELS BY THE *
C * NON-LINEAR ENERGY TO CHANNEL RELATIONSHIP. *
C *****
C
      EUP#G*(FLOAT(J)-ZO+.5)
      CUP#.40715+103.08984*EUP-6.60938*EUP**2+2.55859*EUP**3
      1 -0.52246*EUP**4+0.041382*EUP**5
      ELO#FUP-G
      CLN#.40715+103.08984*ELO-6.60938*ELO**2+2.55859*ELO**3
      1 -0.52246*ELO**4+0.041382*ELO**5
C
      CORRECTION FOR BINED CHANNELS.
      CUP#CUP/CHFT
      CLN#CLN/CHFT

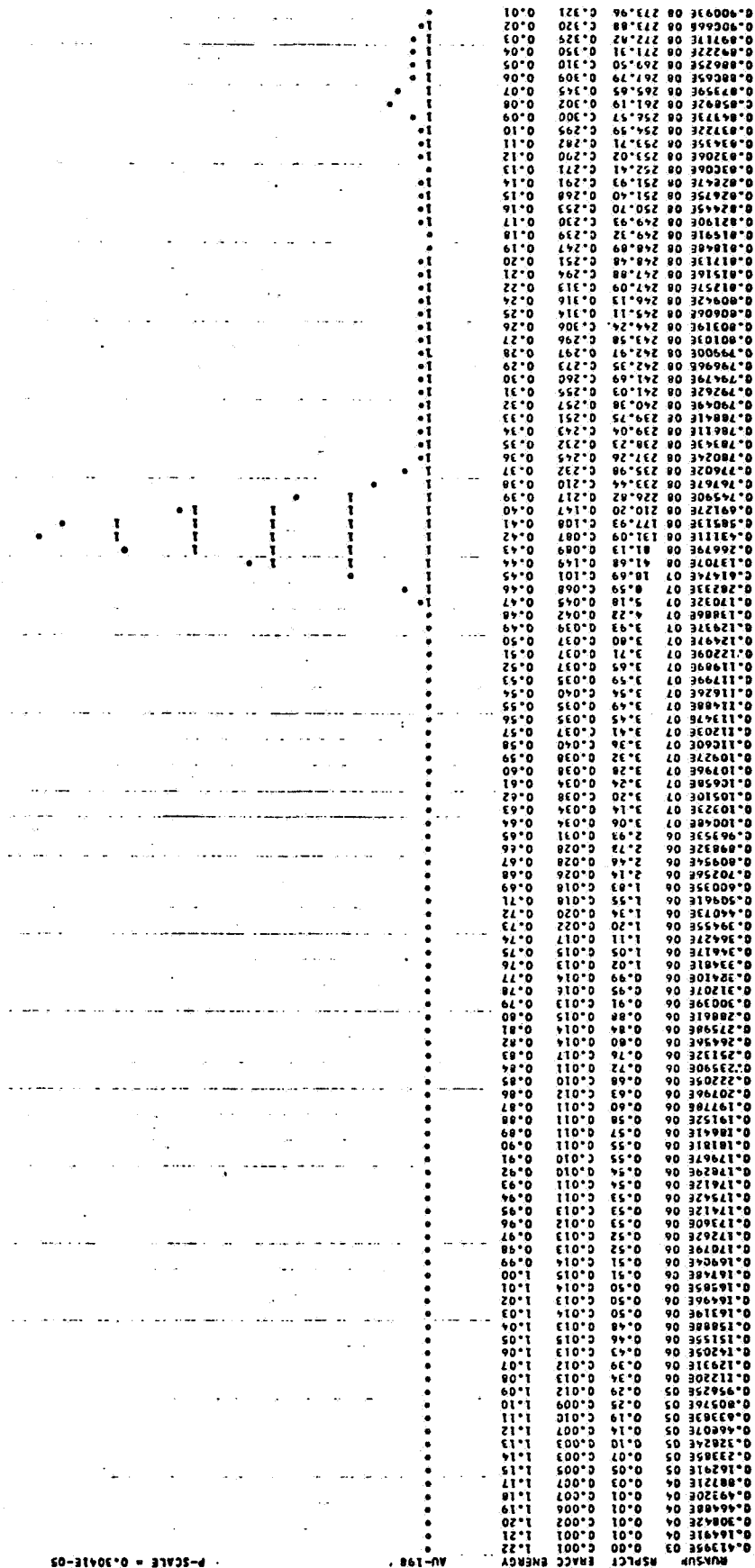
C
C *****
C * EVALUATE THE CORRESPONDING RUN SUMS AT THE UPPER AND *
C * LOWER EDGES OF THE REVISED CHANNELS AND OBTAIN THE *
C * DIFFERENCES TO GIVE THE REBINED COUNTS. *
C *****
C
      NUP#CUP
      NLO#CLN
      FUP#CUP-FLOAT(NUP)
      FLO#CLN-FLOAT(NLO)
      VALUP#(1.-FUP)*SUM(NUP+1)+FUP*SUM(NUP+2)
      VALLO#(1.-FLO)*SUM(NLO+1)+FLO*SUM(NLO+2)
      20 B(J)#VALLO-VALUP
      WRITE(6,110)
      110 FORMAT(16H REBIN COMPLETED)
      RETURN
      END

```

App. III. Results from High Speed SEGO

Graphs of pulse-height distributions and unfolded spectra produced by High Speed Sego follow. These results were obtained by applying the code to the data of R. Heath obtained with a 3 x 3 in. NaI(Tl) spectrometer.





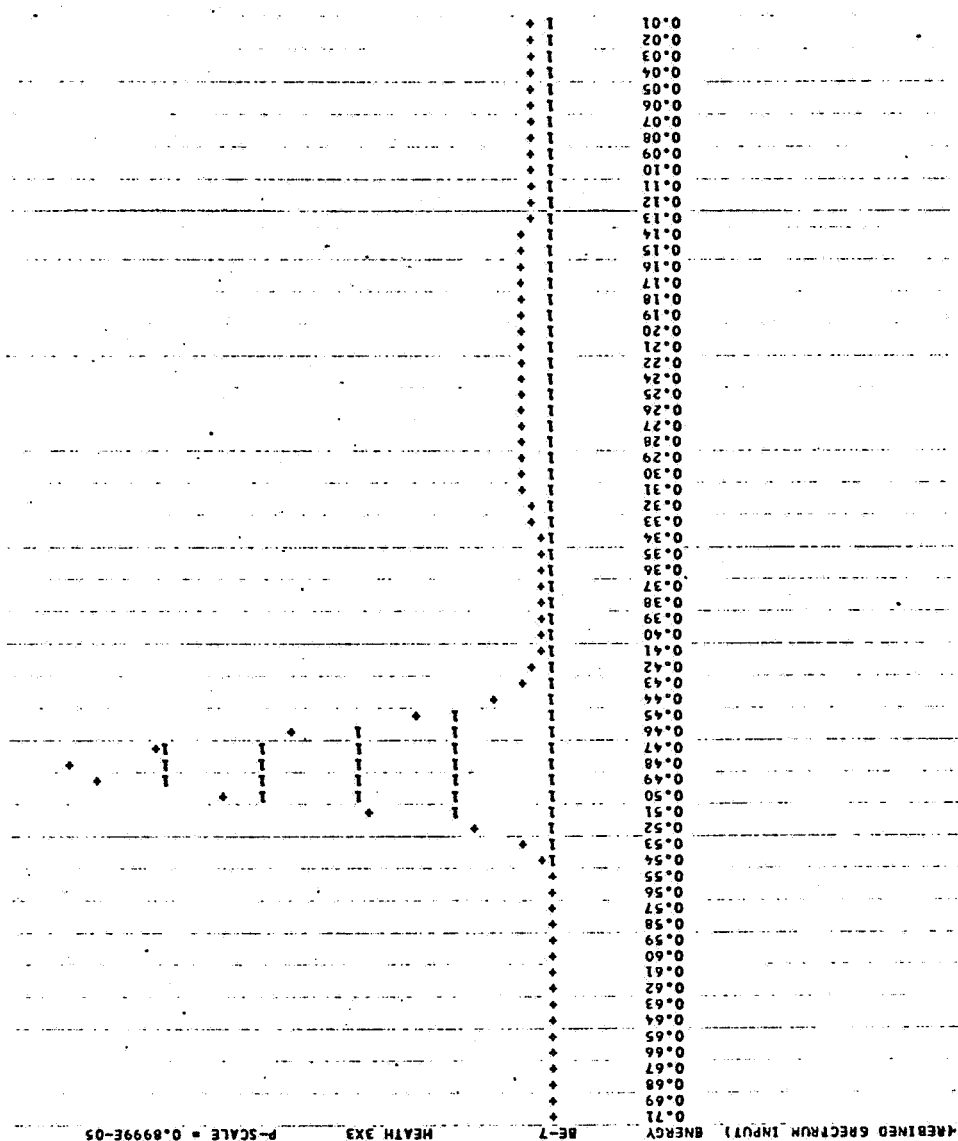


Fig. III-3. Corrected Pulse-Height Distribution for ${}^7\text{Be}$.

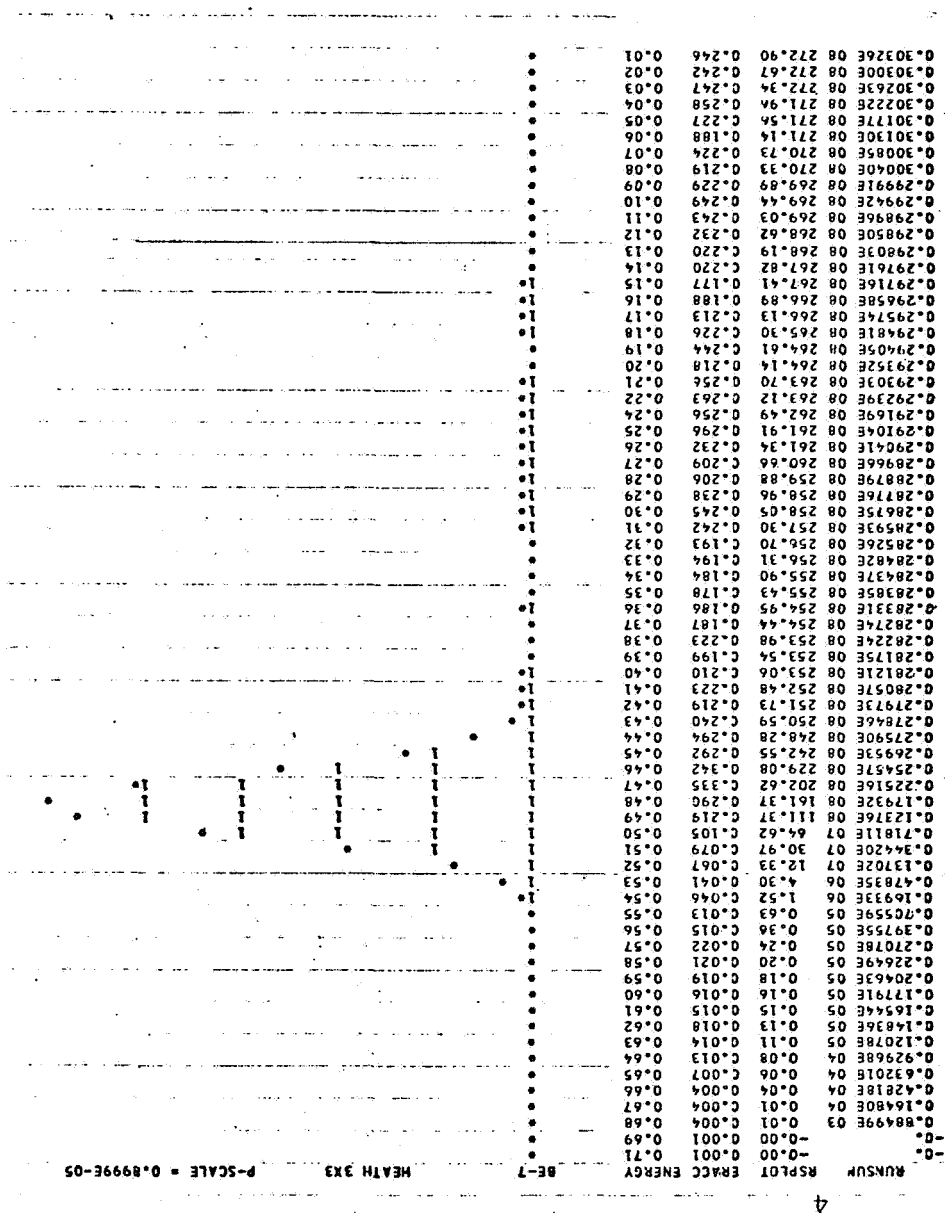


Fig. III-4. Unfolded Gamma Spectrum for ^7Be .

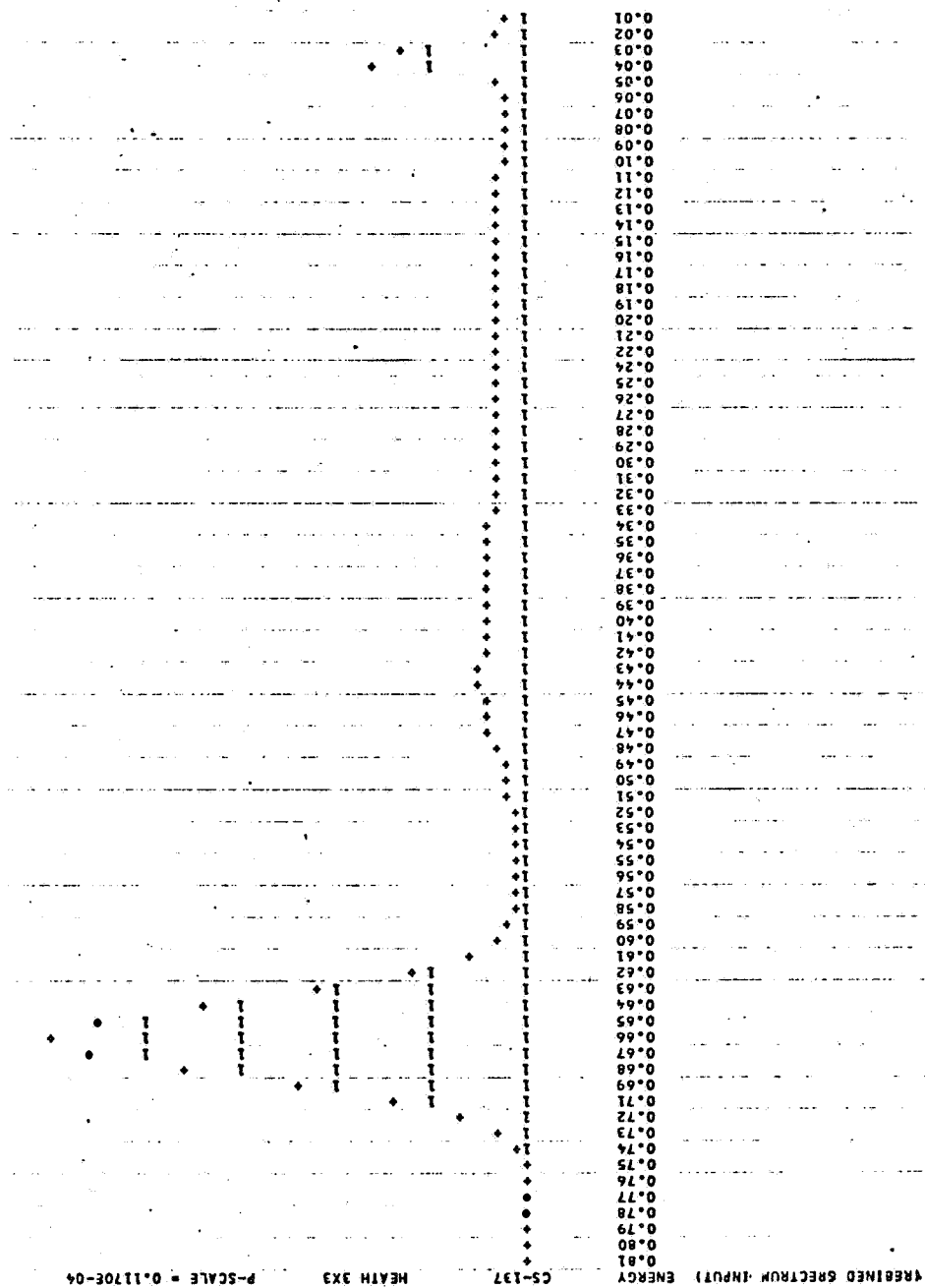


Fig. III-5. Corrected Pulse-Height Distribution for ^{137}Cs .

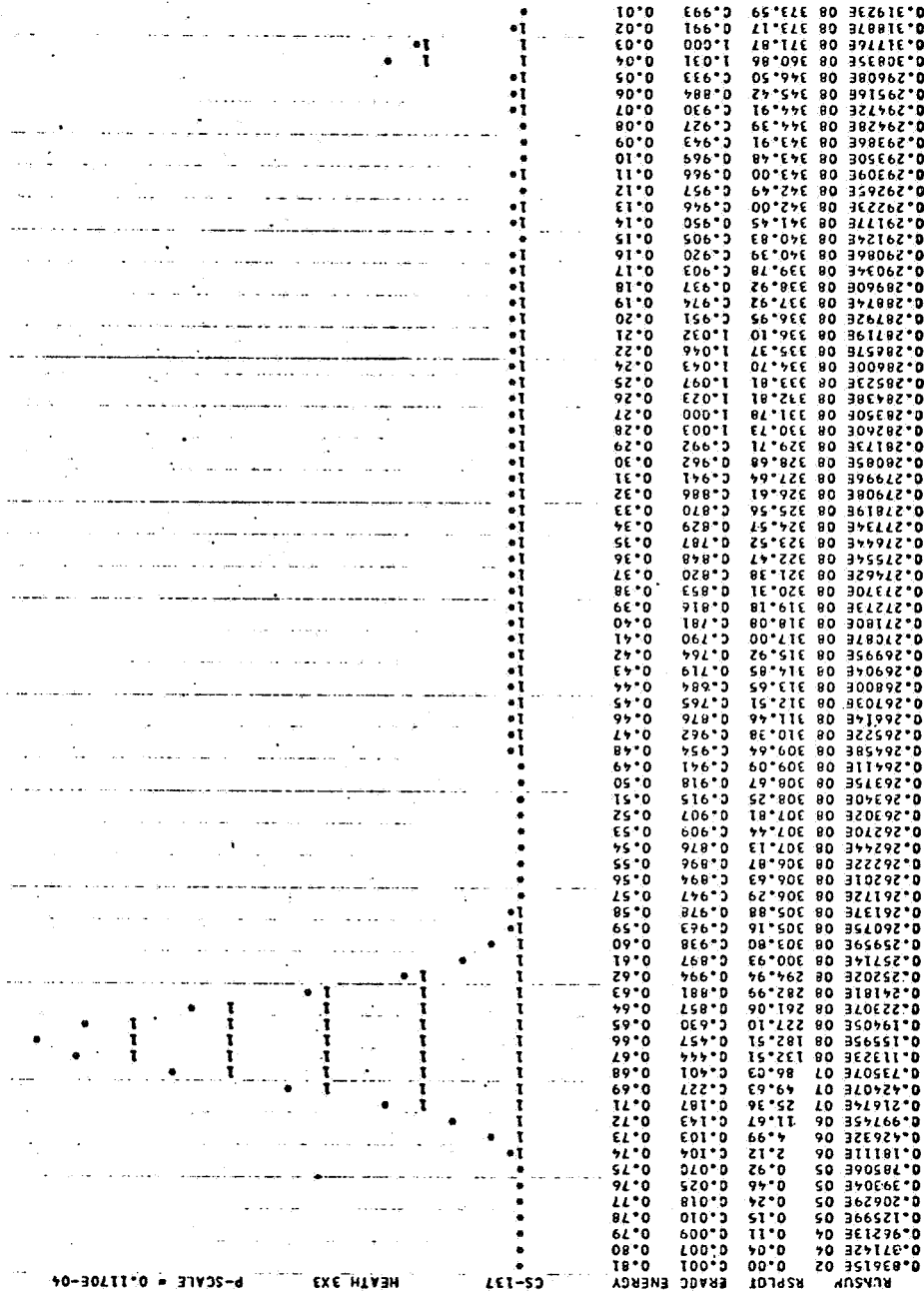


Fig. III-6. Unfolded Gamma Spectrum for ^{137}Cs .

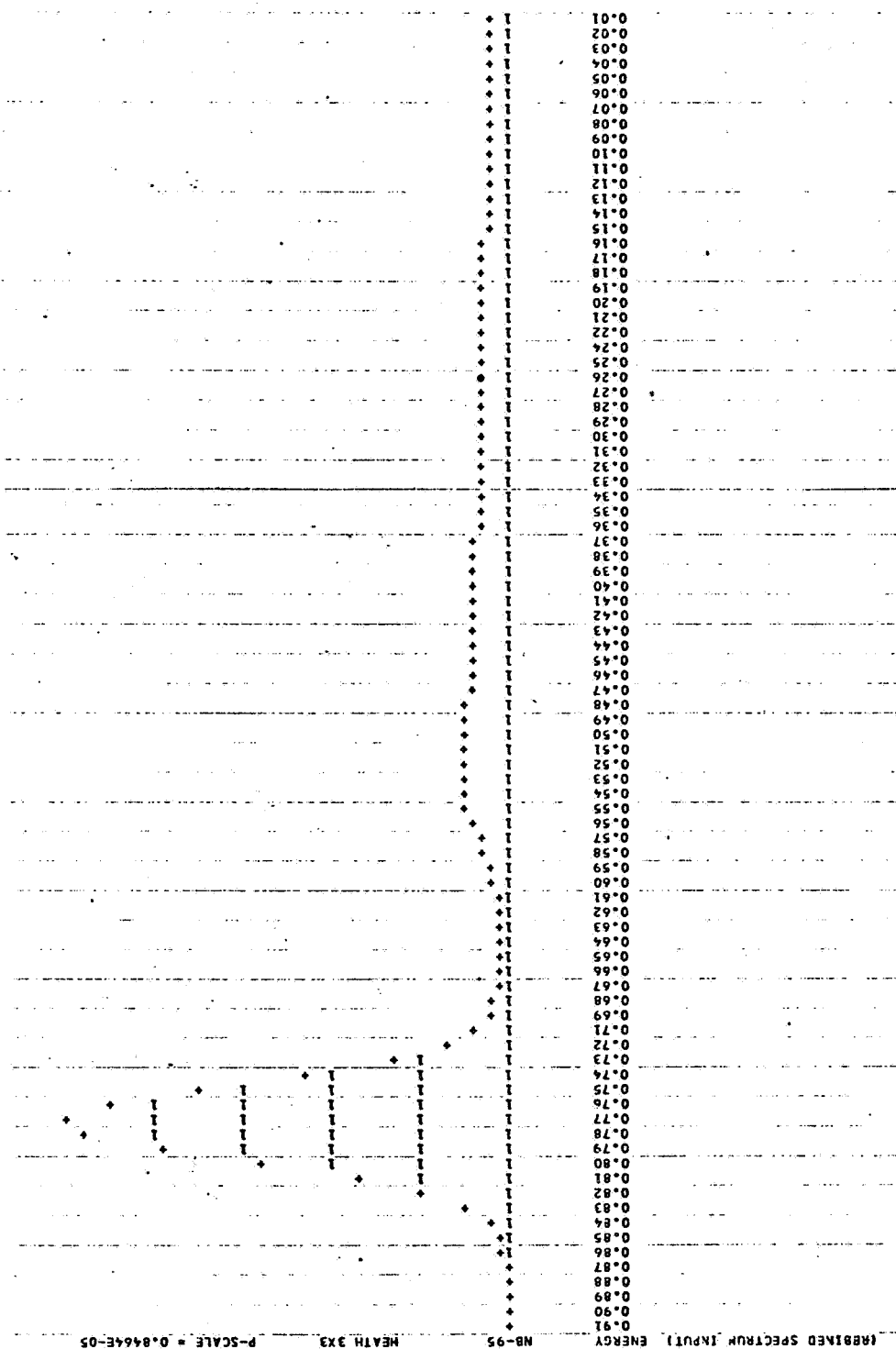


Fig. III-7. Corrected Pulse-Height Distribution for ^{95}Nb .

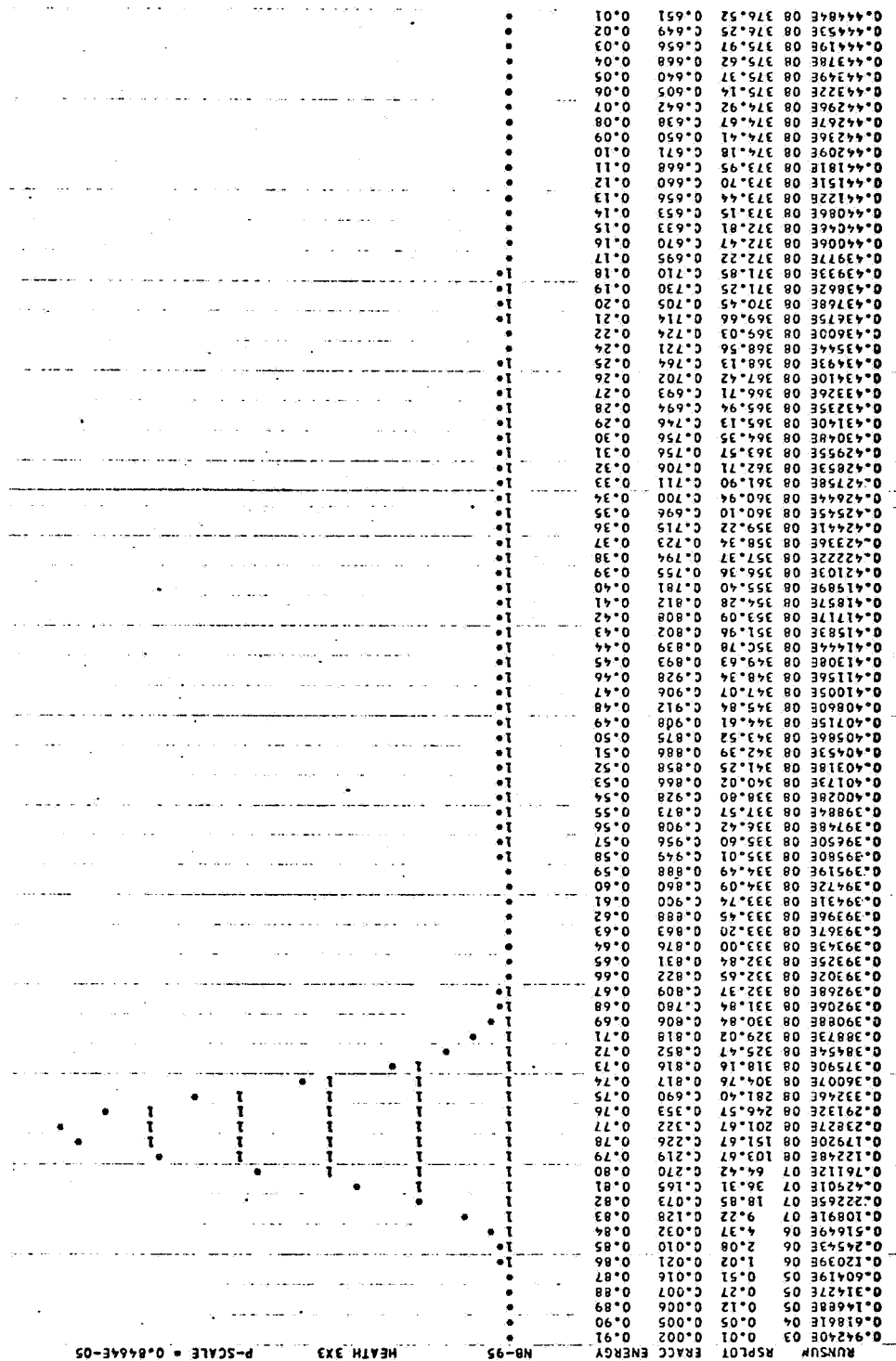


Fig. III-8. Unfolded Gamma Spectrum for ⁹⁵Nb.

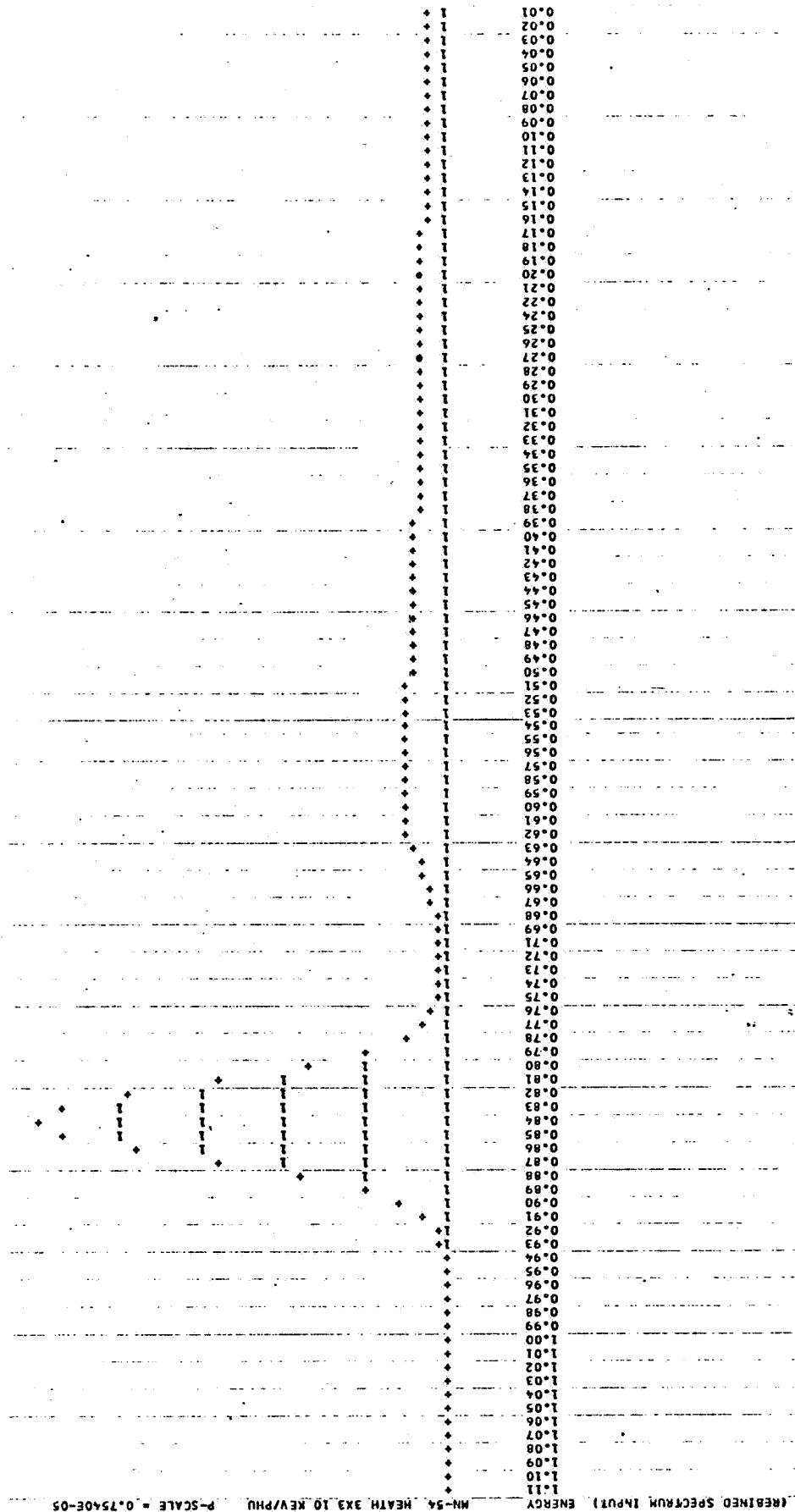
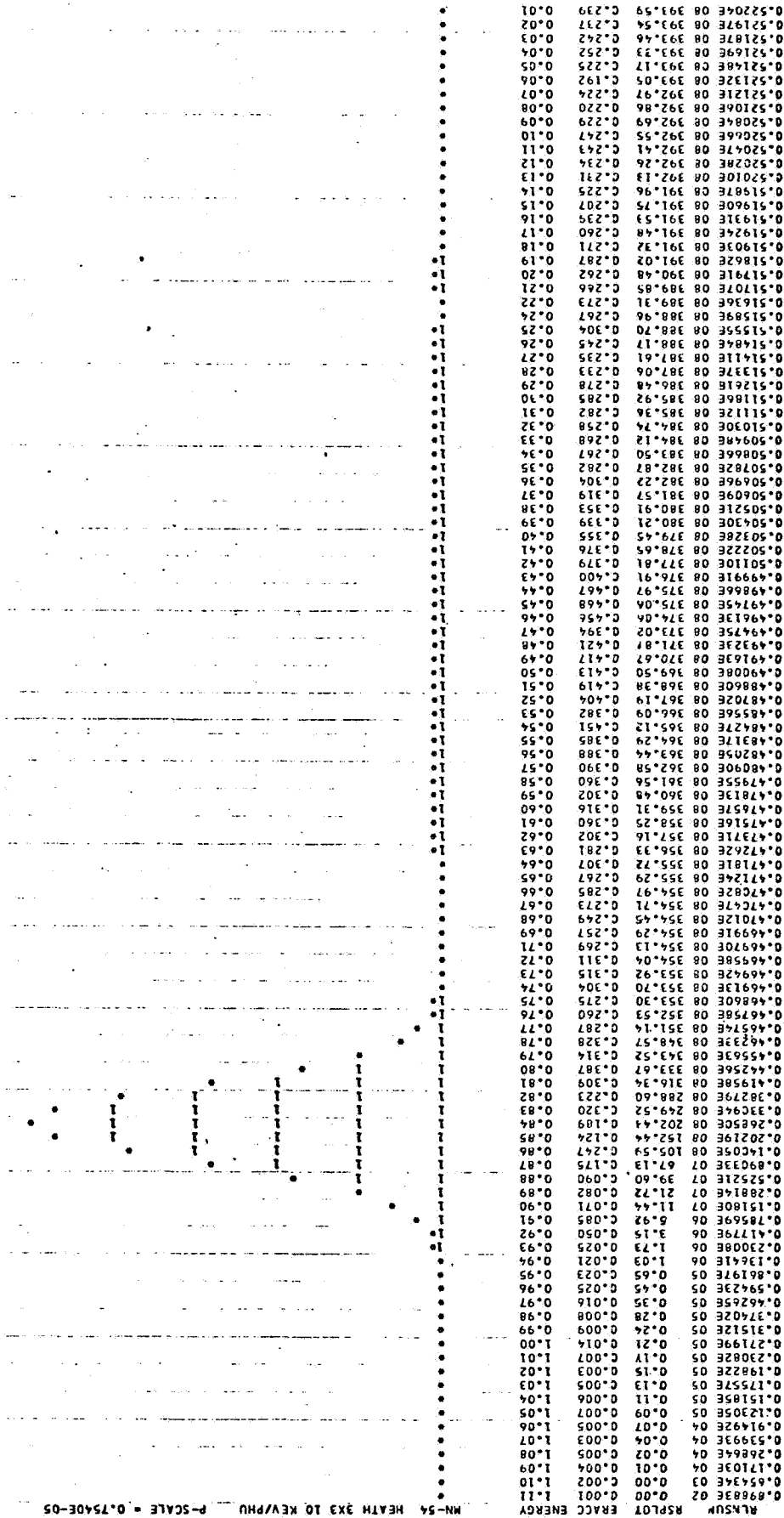


Fig. III-9. Corrected Pulse-Height Distribution for ^{54}Mn .



MN-54 HEATH 3X3 10 KEV/PHU P-SCALE = 0.7540E-05

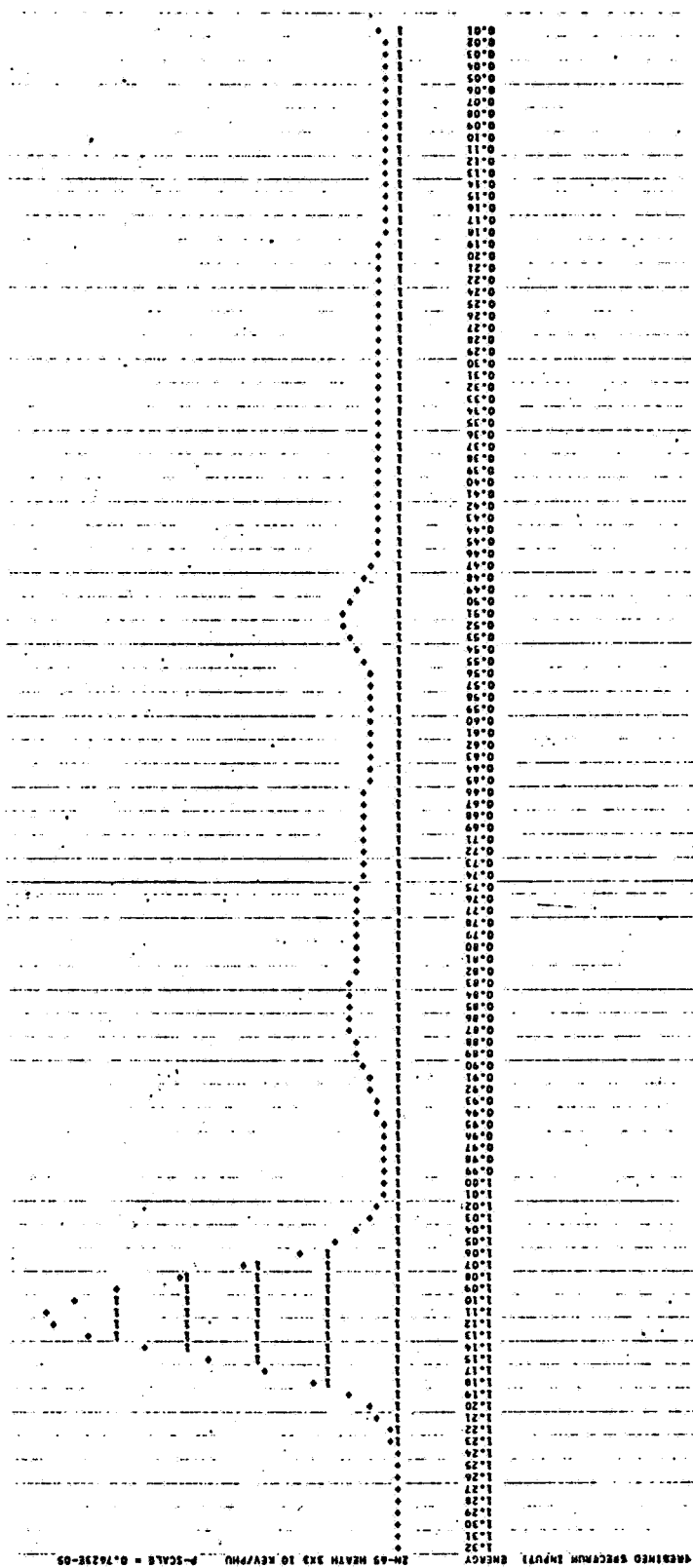


Fig. III-11. Corrected Pulse-Height Distribution for ^{65}Zn .

Fig. III-12. Unfolded Gamma Spectrum for ^{65}Zn .

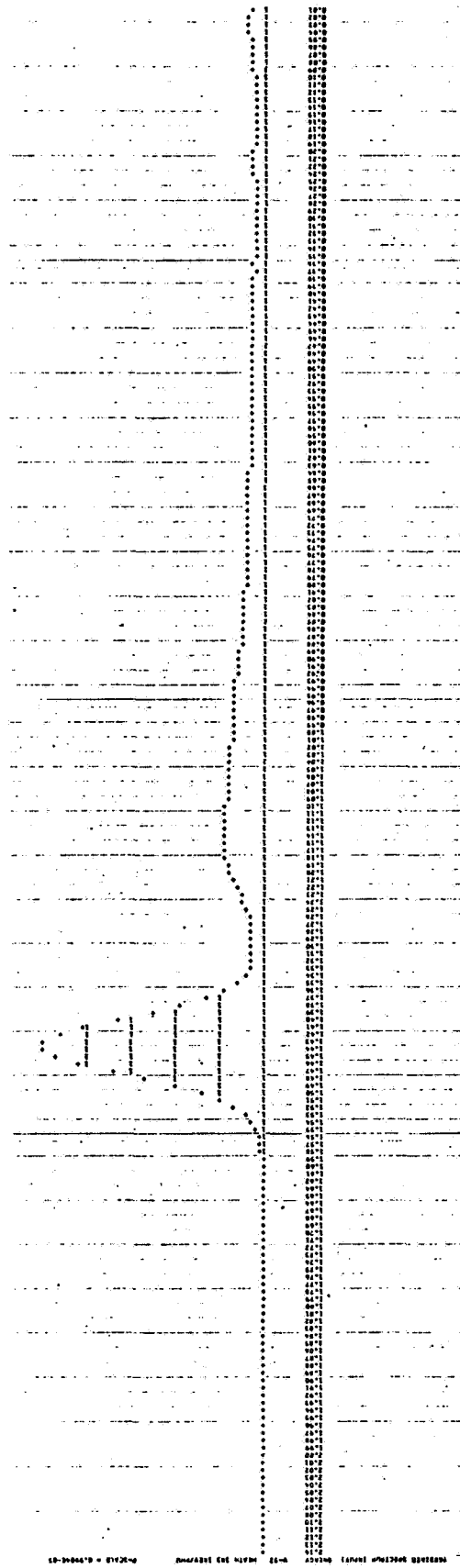


Fig. III-13. Corrected Pulse-Height Distribution for 52V.

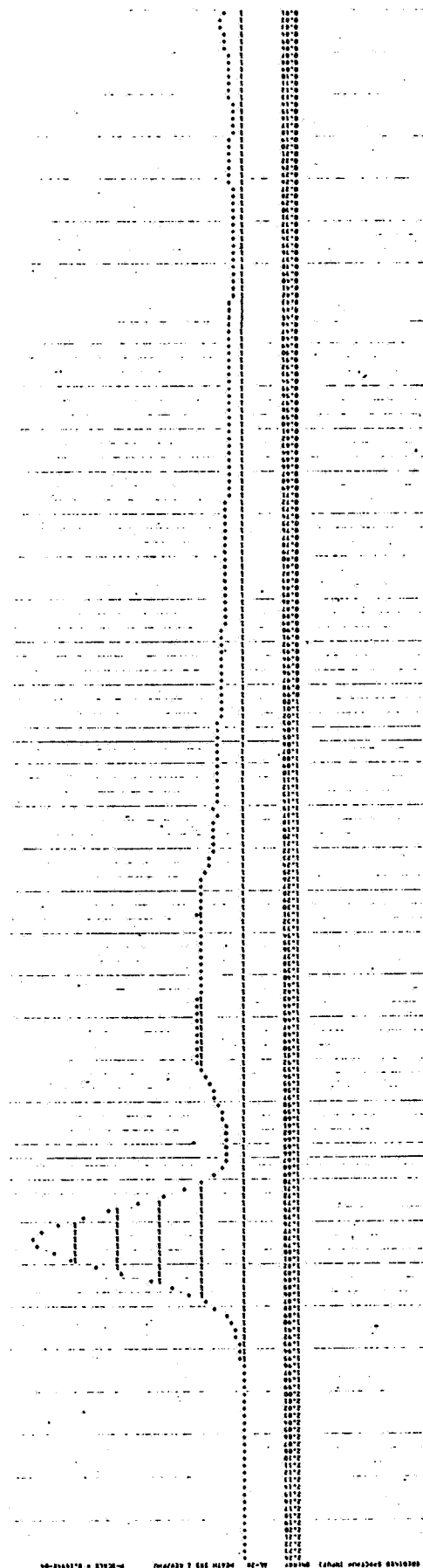


Fig. III-15. Corrected Pulse-Height Distribution for ^{28}Al .

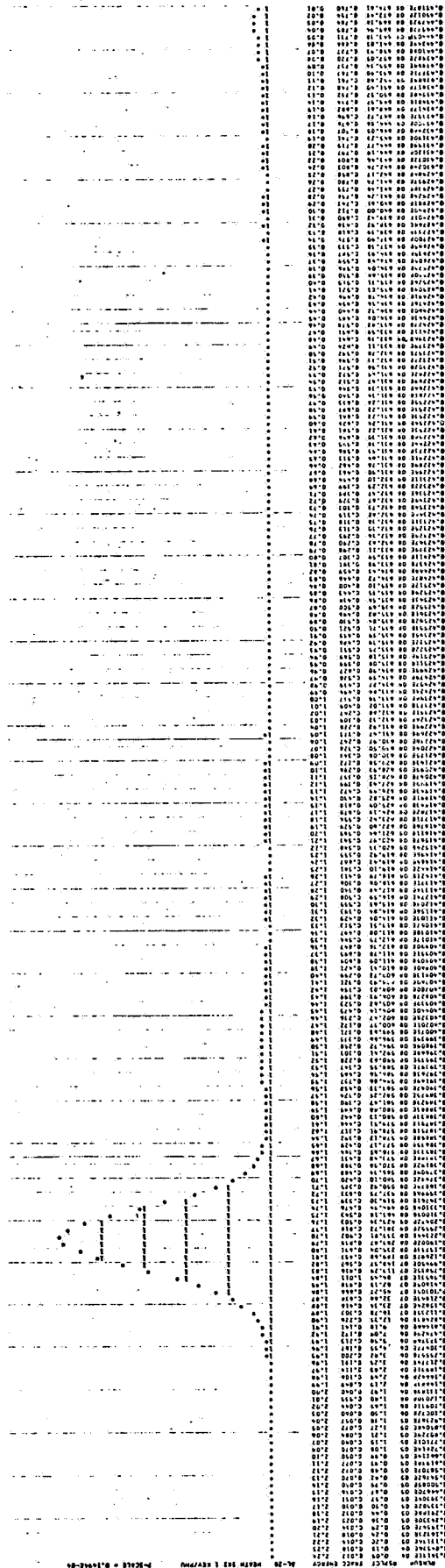


Fig. III-16. Unfolded Gamma Spectrum for ^{28}Al .

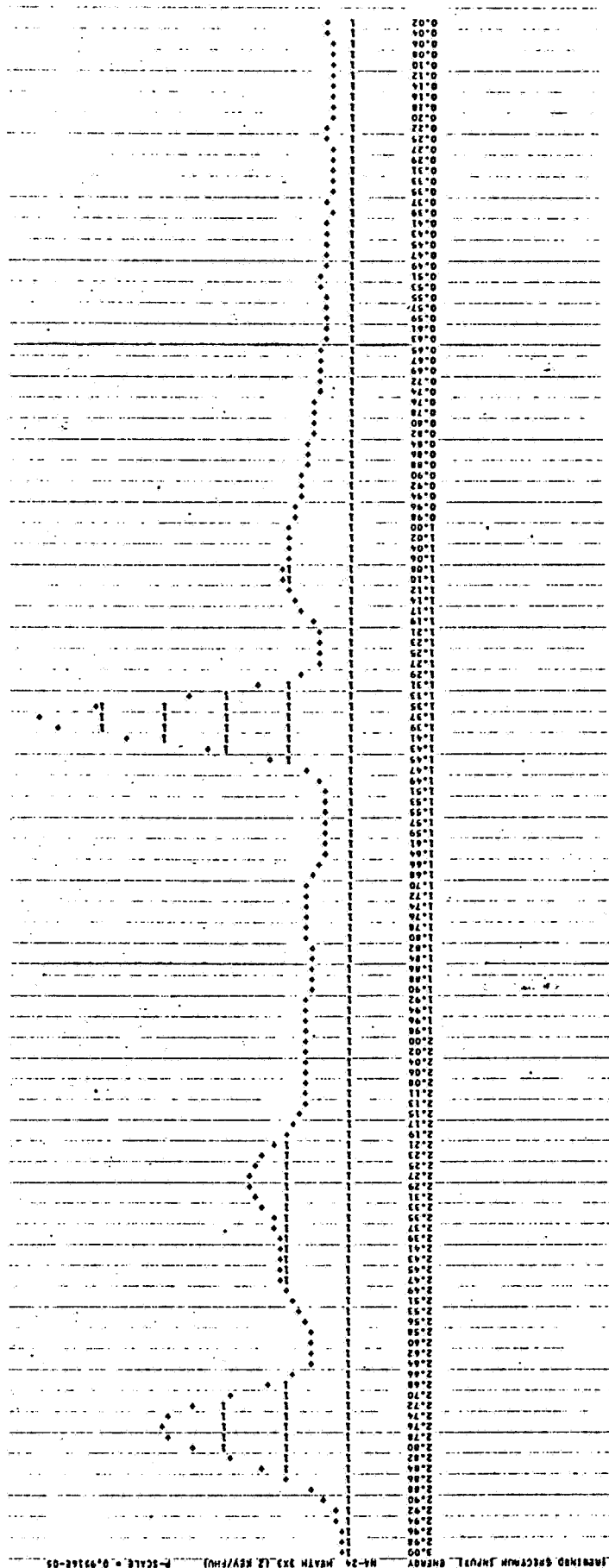


Fig. III-17. Corrected Pulse-Height Distribution for ^{24}Na .

Fig. III-18. Unfolded Gamma Spectrum for ^{24}Na .

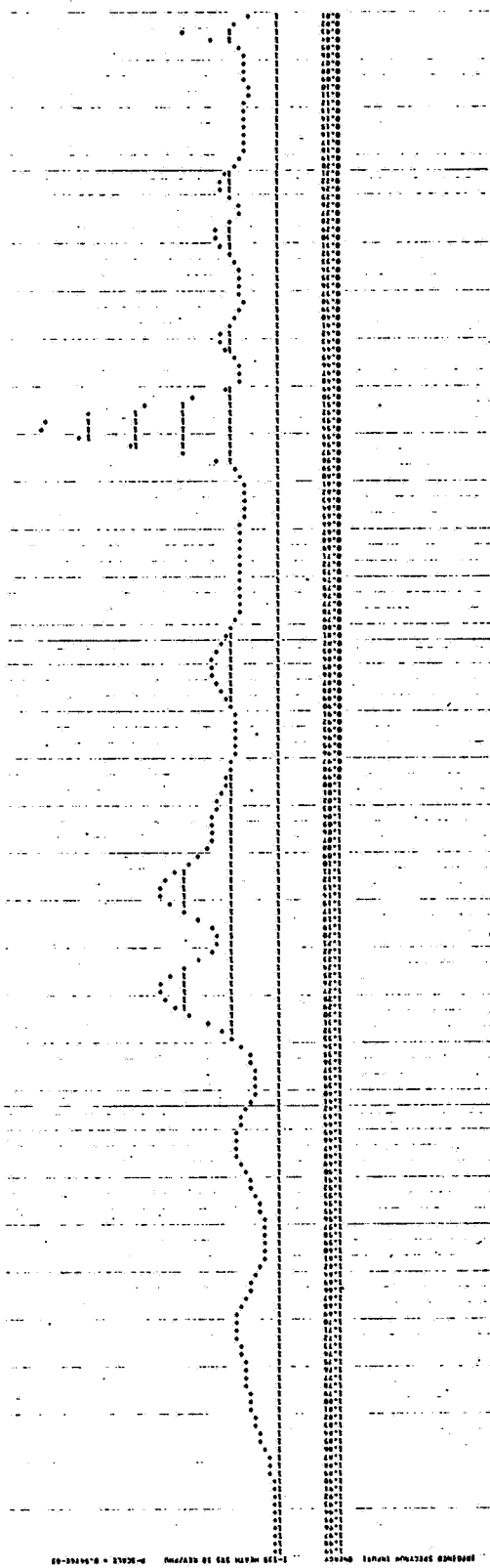


Fig. III-19. Corrected Pulse-Height Distribution for ^{135}I .

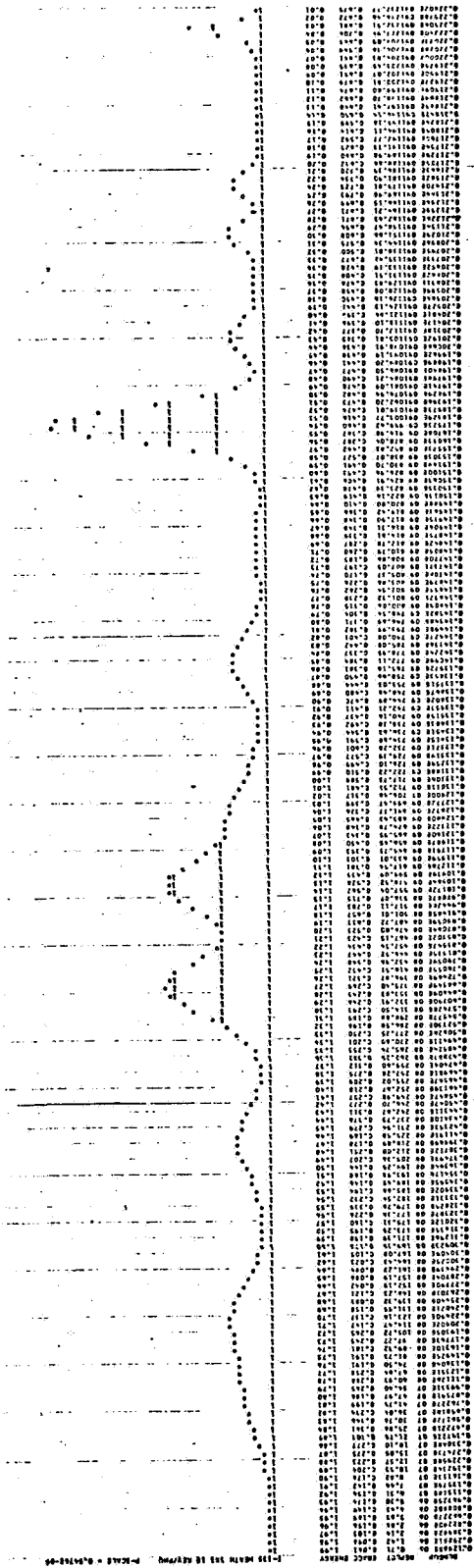


Fig. III-20. Unfolded Gamma Spectrum for ^{135}I .

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