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# A STATISTICAL ANALYSIS OF ELEVATED TEMPERATURE GRAVIMETRIC CYCLIC 

# OXIDATION DATA OF 36 Ni - AND Co-BASE SUPERALLOYS BASED ON AN <br> OXIDATION ATTACK PARAMETER 

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## SUMMARY

A large body of high temperature cyclic oxidation data generated from tests at NASA Lewis Research Center involving gravimetric/time values for 36 Ni and Co-base superalloys was reduced to a single attack parameter, $\mathrm{K}_{\mathrm{a}}$, for each run. This $\mathrm{K}_{\mathrm{a}}$ value was used to rank the cyclic oxidation resistance of each alloy at 1000,1100 , and $1150^{\circ} \mathrm{C}$. These $\mathrm{K}_{\mathrm{a}}$ values were also used to derive an estimating equation using multiple linear regression involving $\log _{10} K_{a}$ as a function of alloy chemistry and test temperature. This estimating equation has a high degree of fit and could be used to predict cyclic oxidation behavior for similar alloys and to design an optimum high strength Ni-base superalloy with maximum high temperature cyclic oxidation resistance. The critical alloy elements found to be beneficial were $\mathrm{Al}, \mathrm{Cr}$ and Ta .

## INTRODUCTION

Cyclic oxidation data in the form of specific weight change/time values and x-ray diffraction results for retained scales as well as spalled oxide(s) has been collected in two recent NASA reports (refs. 1 and 2). These reports covered 36 high-temperature Ni - or Co-base superalloy turbine alloys (table I). These alloys were tested in standard NASA Lewis cyclic oxidation test rigs which have been described in detail in reference 3. Most of the samples tested in these studies were run in a standard mode of a 1.0 hr exposure in the hot zone and then automatically lifted out of the furnace for a minimum of 20 min . This standard cycle was repeated continuously with the sample removed at selected intervals for intermittent weighing to generate the specific weight ( $\Delta W / A$ ) versus time curves. X-ray diffraction analysis was performed at selected intervals as well. In most cases the standard 1 hr cyclic tests for these alloys were 100 hr at $1150^{\circ} \mathrm{C}, 200 \mathrm{hr}$ at $1100^{\circ} \mathrm{C}$, or 500 hr at $1000^{\circ} \mathrm{C}$.

Most of these alloys, particularly at the higher test temperatures, showed an eventual sample specific weight loss due to scale spalling as the sample cools between heating cycles - more than offsetting the oxygen pickup during scale formation at the exposure temperature. The shape of these $\Delta W / A$ versus time curves closely resemble classic paralinear kinetic behavior (refs. 4 to 6).

This gravimetric cyclic oxidation data can be converted into a single attack parameter, $\mathrm{K}_{\mathrm{a}}$ (see below) to rank the oxidation resistance at a given temperature. The higher this $\mathrm{K}_{\mathrm{a}}$ value the poorer the resistance. Based on analysis of a large body of data generated by this laboratory, $K_{a}$ values are ranked as follows (ref. 7):

$$
\begin{aligned}
& \mathrm{K}_{\mathrm{a}} \leq 0.20 \text { excellent } \\
& 0.20 \text { to } 0.5 \text { good } \\
& 0.50 \text { to } 1.0 \text { fair } \\
& 1.0 \text { to } 5.0 \text { poor } \\
& >5.0 \text { catastrophic }
\end{aligned}
$$

The goals of this investigation are to derive the attack parameter, $\mathrm{K}_{\mathrm{a}}$ for each individual alloy sample tested using the suitable model equation; compare the derived $\mathrm{K}_{\mathrm{a}}$ values at 1000,1100 , and $1150^{\circ} \mathrm{C}$ to rank the oxidation resistance of alloys; and thirdly, to attempt by regression analysis to derive an estimating equation for $K_{a}$ (or more realistically $\log _{10} K_{a}$ ) as a function of test temperature and alloy composition. If the third goal is feasible the estimating equation will be used to estimate $K_{a}$ for an alloy not included in this study and finally predict an optimum alloy composition for an alloy of this type.

## ESTIMATING CORROSION ATTACK

All the specific weight change/time data and related kinetics are based on the simple mass balance equation at any time, $t$ :

$$
\begin{equation*}
\Delta W / A=W_{r}-W_{m} \tag{1}
\end{equation*}
$$

where $\Delta W / A$ is the sample's specific weight change value which is plotted against time in these type of handbook figures; $W_{r}$ is the specific weight of the retained scale, and $W_{m}$ is the accumulated specific weight of all the metal converted to oxide up to that time regardless whether the metal is still in the retained scale, or lost by any other process (e.g., scale spalling, and/or scale vaporization and/or scale erosion). This $W_{m}$ value is the critical parameter in any corrosion process and always increases monotonically with time. The problem in any corrosion study is to somehow estimate $W_{m}$ preferably as a function of time.

In most corrosion studies a test sample is run for a given time, removed from test and descaled and the thickness change measured. This value can be directly converted to a $\mathrm{W}_{\mathrm{m}}$ value provided there is no significant alloy element concentration gradient or grain boundary penetration in the alloy. This is not a very practical method in high temperature oxidation studies since it effectively destroys the sample and is a difficult measurement to make particularly for complex alloys. An even more complex extension of this approach is to metallographically mount a cross section of the test sample and determine not only thickness change but any grain boundary attack. Special etching techniques or electron microprobe analysis can then be used to determine any diffusional effects. However, it would be more practical if some nondestructive technique to measure thickness change of the sample as a function of time could be developed, with these more complex and time consuming analysis serving to provide verification.

Another approach is to focus on the $W_{F}$ value. Since it is assumed that the $\Delta W / A$ value can be derived for any time by simply weighing the sample at that time then if $W_{r}$ can be determined then the $W_{m}$ values can be readily solved using equation (1) for a series of times. For two limiting cases $W_{r}$ presents no particular problem. In the first case typical of most high temperature isothermal studies no scale loss occurs. So the $W_{r}$ value at any time is simply the $\Delta W / A$ value multiplied by a stoichiometric oxide constant (refs. 8 and 9 ). For example, in an isothermal parabolic oxidation process after time, $t$ :

$$
W_{m}=b k_{p}^{1 / 2} t^{1 / 2}-k p^{1 / 2} t^{1 / 2}
$$

or

$$
\begin{equation*}
W_{m}=k_{p}^{1 / 2} t^{1 / 2}(b-1) \tag{2}
\end{equation*}
$$

where $k_{p}$ is the parabolic scaling constant and $b$ is the stoichiometric constant based on the composition of the scale.

In the other limiting case where the scale spalls to essentially bare metal, occasionally found in cyclic oxidation, equation (1) reverts to

$$
\begin{equation*}
-W_{m}=\sim \Delta W / A \tag{3}
\end{equation*}
$$

where $\Delta W / A$ values are negative. This has been observed, for example, in burner rig oxidation studies where an insignificant amount of oxide remains (refs. 10 to 14).

There have been attempts at this laboratory and elsewhere to measure $W_{r}$ directly using some physical method (e.g., $\beta$-back scatter, ultrasonic, or microwave technique). So far, however, no method has proven practical. Therefore, an indirect means of estimating $W_{m}$ as a function of time must be found to analyze the large body of cyclic oxidation data.

One approach is to attempt to model the scaling/scale loss process using differential equations based on parabolic scale growth, occurring simultaneously with a linear scale loss. This model has been solved using the mass balance approach and requires only the constants $\mathbf{k}_{\mathrm{p}}, \mathbf{k}_{\mathrm{p}}$, and the stoichiometric constant for the scale formed to be able to determine $\Delta W / A, W_{r}$, and most importantly $W_{m}$ for any time $t$ (refs. 4 to 6). But since $k_{p}$ and particularly $k_{\ell}$ are not generally known, Barrett and Presler (ref. 9) derived a computer program to analyze paralinear behavior and determine $\Delta W / A, W_{r}$, and $W_{m}$ values along with the $k_{p}$ and $k_{g}$ values as a function of time using just two sets $\Delta W / A$, time inputs, and a stoichiometric constant. This program has been used successfully to analyze isothermal oxidation of chromia forming alloys where scale vaporization is significant (ref. 9). Attempts have also been made to use this COREST program to analyze cyclic oxidation behavior of the type of $\Delta \mathrm{W} / \mathrm{A}$ with time curves shown in the two turbine alloy reports but its success had been limited (refs. 14 and 15) but it is useful as a first approximation.

A more successful approach has been to actually model the cyclic oxidation process, cycle by cycle, on a computer. Any scale growth process, usually a parabolic rate constant, can be used as input. The nature of the spalling process should also be known. For chromia or alumina forming alloys it appears the rate of spalling is a fixed percent of the oxide thickness (ref. 16). As in the other methods the stoichiometric constants can usually be estimated quite easily. This computer program termed COSP (ref. 17) generates the $\Delta W / A, W_{r}$, and $W_{m}$ versus time just as in COREST. This approach has been fairly successful with the more simple type heater alloys but has been more difficult to use in analyzing the cyclic oxidation behavior of more complex alloys like high temperature superalloys.

Another approach which has proven successful is to fit the specific weight change/time data to a simple quasi-paralinear equation by multiple linear regression:

$$
\begin{equation*}
\Delta W / A=k_{1}^{1 / 2} t^{1 / 2} \pm k_{2} t \pm \sigma \tag{4}
\end{equation*}
$$

Here $k_{1}{ }^{1 / 2}$ and $k_{2}$ are constants analogous to the scale growth and scale spalling constants and $\sigma$ is the standard error of estimate. If the fit is good enough (usually $\mathrm{R}^{2}>0.90$ ) and $\mathrm{k}_{1}{ }^{1 / 2}$ is significant and positive and $k_{2}$ is statistically significant then the attack parameter $K_{a}$ is defined as:

$$
\begin{equation*}
K_{a}=\left(k_{1}^{1 / 2}+10\left|k_{2}\right|\right) \tag{5}
\end{equation*}
$$

or

If $k_{1}^{1 / 2}$ is either not significant or negative and $k_{2}$ is significant then $K_{a}$ is defined as

$$
\begin{equation*}
K_{a}=20\left|\mathrm{k}_{2}\right| \tag{6}
\end{equation*}
$$

The rational behind these $\mathrm{K}_{\mathrm{a}}$ derivations are discussed in references 7, 16, and 18 to 22. It has been shown that these $K_{a}$ values are valid as estimators of oxidation resistance and are well correlated with both thickness change measurements and $W_{m}$ estimates derived by both the COREST and COSP computer programs discussed above. This $K_{a}$ estimation technique has the advantage that if the specific weight change/time data is in a computer data base for a given run the data can be automatically processed for a regression fit according to equation (4) and $K_{a}$ computed according to equations (5) or (6) depending on the significance and sign of the coefficients $k_{1}{ }^{1 / 2}$ and $k_{2}$. By this process fairly irregular kinetics can be evaluated. This $K_{a}$ approach was chosen to analyze the large number of runs for the complex superalloys referred to in this report.

## Derivation of $K_{a}$ Values from the Cyclic Oxidation Data

A total of 323 runs based on the 36 alloys listed in table $I$ of $\Delta W / A$ versus time data were individually analyzed according to equation (4) by multiple linear regression. This approach leading to $K_{a}$ values for each run is detailed in Appendix $A$.

After discarding 8 outliers as described in the appendix a total of 315 valid $K_{a}$ values were available to rank the alloys. These valid $K_{a}$ values can be compared at each test temperature for each alloy as a. series of bar graphs. For ease of description the 36 alloys tested were divided into two distinct groups and plotted in figures $1(\mathrm{a})$ to (c) and figures $2(\mathrm{a})$ to (c). In the first grouping, all Ni-base, the alloys were essentially alumina/aluminate scale formers. These alloys, 15 in number, contained 5 to $6 \mathrm{wt} \% \mathrm{Al}$ and a minimum of $5 \mathrm{wt} \% \mathrm{Cr}$. The second grouping, containing both Ni - and Co -base alloys, were either $\mathrm{Cr}_{2} \mathrm{O}_{3}$ /chromite or possibly MO scale formers. This group of 21 alloys contained either less than $5 \mathrm{wt} \%$ Al with Cr of $9 \mathrm{wt} \%$ or greater and were basically the $\mathrm{Cr}_{2} \mathrm{O}_{3} /$ chromite scale formers. Or else they had quite high Al levels but no Cr and tended to form NiO as the surface oxide in spite of the high Al levels.

These two sets of alloys are plotted as a series of bar graphs in order of increasing Al content at the three test temperatures.

The coordinates are $\mathrm{K}_{\mathrm{a}}$ values plotted on a log based scale. Also indicated are the rankings from excellent to catastrophic. The top of each bar is the maximum $\mathrm{K}_{\mathrm{a}}$ value derived for that alloy at the given temperature. Any horizontal lines below the top represent replicates. This gives an indication of the scatter for each alloy. As expected, oxidation resistance decreases with an increase in test temperature and the number of alloys showing excellent to good oxidation resistance (i.e., $\mathrm{K}_{\mathrm{a}} \leq 0.2$ or $\leq 0.5$ ) decreases with increasing temperature as well. Although these plots are quite informative they tend to be somewhat pessimistic because they focus more on maximum values than on average values. Based on these plots three alloys, all $\mathrm{Al}_{2} \mathrm{O}_{3} /$ aluminate formers, have the best oxidation resistance. In decreasing order of resistance they are: (1) TRW-R, (2) B-1900, and (3) NASA-TRW-VIA.

## Modeling Oxidation Attack, $K_{\mathrm{a}}$ as a f (Alloy Chemistry, Temperature)

In an earlier study (ref. 22) at this laboratory the derived oxidation attack parameter in the form of $\log _{10} \mathrm{~K}_{\mathrm{a}}$ was used to study systematic variations in $\mathrm{Co}, \mathrm{Ta}, \mathrm{Al}, \mathrm{Cr}$, and Mo in a prototype Ni-base turbine alloy. The basic alloy content was $\mathrm{Ni}-1 \mathrm{wt} \% \mathrm{Ti}-2 \mathrm{wt} \% \mathrm{~W}-1 \mathrm{wt} \% \mathrm{Nb}-0.1 \mathrm{Zr}-0.12 \mathrm{C}-0.01 \mathrm{~B}$. The alloy had five target levels each of $\mathrm{Al}(3.25,4,4.75,5.50$, and 6.25$) ; \operatorname{Cr}(6,9,12,15$, and 18$) ; \operatorname{Co}(0,5,10,15$, and 20); Mo ( $0,1,2,3$, and 4 ); and $\mathrm{Ta}(0,2,4,6$, and 8 ) all in weight percent. This series of alloys represented a $2^{5}$ composite statistically designed experiment representing a total of 43 individual alloys. The samples were tested for $2001-\mathrm{hr}$ cycles at $1100^{\circ} \mathrm{C}$ to derive the $\mathrm{K}_{\mathrm{a}}$ values as described above. This design along with a suitable number of replicates enabled a second degree estimating equation to be derived by multiple linear regression as a function of the five composition variables.

This same basic approach was to be used to analyze statistically the 36 alloys with the valid 315 derived $\mathrm{K}_{\mathrm{a}}$ values of this study. This analysis differs signifcantly from the above mentioned $\mathbf{2}^{5}$ statistically designed study as follows:
(1) It includes both Ni and Co-base alloys although the preponderance are Ni-base.
(2) There are 13 compositional variables as shown in table $\mathrm{I}-\mathrm{Cr}, \mathrm{Al}, \mathrm{Ti}, \mathrm{Mo}, \mathrm{W}, \mathrm{Nb}, \mathrm{Ta}$, $\mathrm{C}, \mathrm{B}, \mathrm{Zr}, \mathrm{Hf}, \mathrm{V}$, and Re.
(3) The alloys were tested at two, three, or even four different temperatures.
(4) The compositions were essentially random (i.e., the alloy compositions were not systematically varied).
(5) An additional temperature term of the form $X_{i}=1 / T_{k}{ }^{\circ}$ is required as well.

In addition the following simplyfing assumptions were made:
(A) Nominal alloy chemistries will be used even if multiple heats of the same alloys were tested.
(B) A fourteenth composition variable was added and was defined as the $\mathrm{Co}+\mathrm{Fe}$ content in the Ni base alloys or the $\mathrm{Ni}+\mathrm{Fe}$ content in the Co base alloys.
(C) The minor Cu content in the Mar-M-246 alloy was not included.

Note there were a number of replicate runs. In multiple regression analysis this allows the pure error variance to be separated from the residual error variance so the significance of the model may be tested ${ }^{\text {- }}$ with the lack of fit variance. This approach will be shown for the ultimate model derived in this analysis.

Initially only a first order model will be considered (i.e., the independent variables will be first degree only or linear - $\mathrm{x}_{1}, \mathrm{x}_{2}, \ldots$ ) using the basic 15 terms. Assume the model:

$$
\begin{align*}
& \log \mathrm{K}_{\mathrm{a}}=\mathrm{a}+\mathrm{b}_{1} \mathrm{C}_{\mathrm{r}}+\mathrm{b}_{2} \mathrm{Al}+\mathrm{b}_{3} \mathrm{NiCo}+\mathrm{b}_{4} \mathrm{Ti}+\mathrm{b}_{5} \mathrm{Mo}+\mathrm{b}_{6} \mathrm{~W}+\mathrm{b}_{7} \mathrm{Nb}+\mathrm{b}_{8} \mathrm{Ta}  \tag{7}\\
& +\mathrm{b}_{9} \mathrm{C}+\mathrm{b}_{10} \mathrm{~B}+\mathrm{b}_{11} \mathrm{Zr}+\mathrm{b}_{12} \mathrm{Hf}+\mathrm{b}_{13} \mathrm{~V}+\mathrm{b}_{14} \mathrm{Re}+\mathrm{b}_{15}\left(1 /\left(\text { temp }+273{ }^{\circ} \mathrm{C}\right)\right) \pm \sigma
\end{align*}
$$

The multiple regression analysis stepwise procedure was used ${ }^{1}$ which rejected any of the 15 terms not significant to the 0.15 level. The final estimating equation involved 11 significant terms with a suprisingly high $\mathrm{R}^{2}$ value of just over 80 percent. The lack of fit (L.O.F) variance is highly significant implying as expected the model is not adequate. The summary table for this analysis is shown in Appendix C.

The next step is to build a model involving both first and second order terms. In most cases a second order equation is sufficient to model most estimating processes of this type. Thus the model equation would be of the form

$$
\begin{equation*}
\log K_{a}=a_{1}+b_{1} x_{1}+b_{2.2} x_{1}^{2}+b_{1.2} x_{1} x_{2}+b_{2} x_{2}+\ldots b_{15.15} x_{15}^{2} \tag{8}
\end{equation*}
$$

For $x_{i}=15$ this would involve a possible 135 terms which would not be practical to run in a stepwise multiple regression analysis. Instead a series of subsets of $x_{i}, x_{i}^{2}, x_{i} x_{i} \ldots$ terms were used involving 20 to 25 of the 135 possible terms. The significant terms were then accumulated, A total of 23 likely terms were then used to derive a final estimating equation. A rejection level of $\alpha=0.15$ was again used.

Table II summarizes this analysis. Including the coefficients for the final 14 term equation (9), fourteen of the 23 terms were found to be significant. These coefficients along with the intercept are listed in this table along with their significance levels. This technique also generated the predicted values for each sample run as well as $\log \mathrm{K}_{\mathrm{a}}$ values for any of the 36 alloys not tested at 1000 or $1100^{\circ} \mathrm{C}$.

Table III is an analysis of variance (ANOVA) summary table to partition the variability (i.e., sum of squares) to test the goodness of fit of the 14 term model equation. This is possible because of the large number of replicate terms which represent pure error. This enables the residual error found in regression analysis to be separated into pure error and lack of fit. The F - ratio of $\mathrm{MS}_{\mathrm{L} . \mathrm{O}, \mathrm{F}}$ to $\mathrm{MS}_{\text {error }}$ is roughly 1.26. Thus the L.O.F term is not significant to the $\alpha=0.05$ level. This indicates the model estimating equation is adequate for predictive purposes. The $\mathbf{R}^{2}$ value is close to 0.85 which is quite high for this type of estimation. Even if a better model estimating equation could be found involving more of the 135 possible second order terms or involving even higher order terms or possibly other variables not included in the model only an $\mathbf{R}^{2}$ value of 0.886 could have been achieved because of the pure replicate error. On this basis the estimated equation explains just over 95 percent of the possible variability that could be modeled.

Figures 3(a) to (c) and figures 4(a) to (c) show the derived $\mathrm{K}_{\mathrm{a}}$ estimates from the 14 term estimating equation on a $\log _{10}$ bar graph scale for each alloy at 1000,1100 , and $1150{ }^{\circ} \mathrm{C}$ for the two alloy groupings. These values are listed in tables IV and V. Also shown on the same bar graphs are the

[^0]average observed $\mathrm{K}_{\mathrm{a}}$ values ${ }^{2}$ for each alloy for ready comparison. At $1000^{\circ} \mathrm{C}$ only 11 of the 36 alloys were tested, so 25 alloys represent just the predicted values. At $1100^{\circ} \mathrm{C} 34$ of the 36 alloys were tested, while at $1150{ }^{\circ} \mathrm{C}$ all 36 alloys were run. In general the mean and predicted values fall in or near the same rating category. The overall agreement between the predicted and average $\mathrm{K}_{\mathrm{a}}$ values appear good.

Figure 5 shows a plot of the regression standard residuals plotted against the predicted values for all the 315 runs. The random nature of the residuals are a good indicator of the validity and unbias nature of the regression equation. A scatter diagram of the predicted $\log K_{a}$ values ploteed against the $\log$ of their observed values is shown in figure. 6. The data was fitted by simple linear regression and gives a resultant diagonal straight line with a slope near unity. Also shown are the + or -2.5 standard deviation lines which would include 95 percent of the data points. This is a further validation of the 14 term regression equation to estimate $\log \mathrm{K}_{\mathrm{a}}$ values.

A further check on efficency of the estimating equation is how well it predicts $K_{a}$ values for a similar alloy not included in the original 36 alloy data base. The alloy chosen was NASAIR-100 which has a nominal composition in weight percent of $\mathrm{Ni}-9 \mathrm{Cr}-5.8 \mathrm{Al}-0.5 \mathrm{Co}-10.5 \mathrm{~W}-3.3 \mathrm{Ta}-1.2 \mathrm{Ti}-1 \mathrm{Mo}-0.03 \mathrm{max}$ $\mathrm{Zr}-0.006 \mathrm{C}-0.002 \mathrm{~B}$. Two samples were tested for 1001 hr cycles at $1150{ }^{\circ} \mathrm{C}$. Also a single sample was tested at $1200{ }^{\circ} \mathrm{C}$ even though this was outside the temperature test range by $50^{\circ} \mathrm{C}$. Table VI summarizes the $K_{a}$ derivations for these cyclic runs. From the estimated $\log K_{a}$ values from the 14 term estimating equation (9) and the derived $\log K_{a}$ values from the computed $K_{a}$ values derived from the oxidation rate constants. The agreement appears quite good. At $1150^{\circ} \mathrm{C}$ both actual $\log \mathrm{K}_{\mathrm{a}}$ values are within $1-1 / 2$ sigma units, while at $1200^{\circ} \mathrm{C}$ the values are within one sigma unit of each other. This leads further credence as to the validity of the 14 term estimating equation as well as the overall approach.

## Implications for Alloy Chemistry From The Model Estimating Equation

The final 14 term estimating equation (9) summarized in table III has certain obvious implications from the alloy chemistry standpoint. There are only three terms with beneficial negative coefficients which lower the $\mathrm{K}_{\mathrm{a}}$ estimates. These improve the cyclic oxidation resistance of this type of Ni-based or Co-based superalloy. Both Al and Cr improve the resistance and so does Ta as long as Al is present. Alloy elements which are neutral (i.e., have no effect) on the cyclic oxidation resistance at least within the alloy ranges (i.e., sample space) of the 36 alloys tested are C, B, and Zr. This also applies to Co in Ni -based or Ni in Co-based alloys.

This leaves $\mathrm{Ti}, \mathrm{Hf}, \mathrm{V}, \mathrm{Re}, \mathrm{Nb}, \mathrm{Mo}$, and W to be evaluated from the coefficients. Nb is the most obvious element to omit since has a positive interaction with $\mathrm{Ti}, \mathrm{Ta}$ and Hf. This then allows 1.0 -percent Hf to be alloyed since it is neutral without Nb . Rhenium and V should also be eliminated. Tungsten, Mo, and Ti should probably also be dropped since they are all involved with positive terms. However, since around 1.0 -percent Ti is usually alloyed to this type of Ni-base superalloy for reasons other than oxidation resistance it should be fixed at roughly 1 percent. One percent Hf could be added also as long as Nb is not present.

This could lead to a typical prototype turbine alloy of $\mathrm{Ni}-10 \mathrm{Co}-0.9 \mathrm{Ti}-1 \mathrm{Hf}-0.1 \mathrm{C}-0.015 \mathrm{~B}-0.1 \mathrm{Zr}$ with $\mathrm{XAl}-\mathrm{YCr}-\mathrm{ZTa}$. It is then possible to use the estimating equation to optimize the composition within certain alloy constraints. If Mo and W are required for any reason they should be kept as low as possible.

[^1]This is assumed to be a Group I alloy-a basic alumina/aluminate former which has an Al content constrained between 5 and $6 \mathrm{wt} \%$. The Cr contents for this type of alloy that varies between 5 and $13 \mathrm{wt} \%$ while Ta when present ranges between 2 and $9 \mathrm{wt} \%$. The role of Cr in helping to stabilize the protective alumina/aluminate scale in heater alloys and Ta in forming the tri-rutile oxide $\mathrm{Ni}(\mathrm{Ta}) \mathrm{O}_{4}$ which also confers protection in more complex alumia/aluminate forming alloys have been discussed elsewhere (refs. 8 and 7). This statistical analysis tends to confirm these earlier conclusions. The optimum contents of $\mathrm{Al}, \mathrm{Cr}$, and Ta were determined using the above constraints and generating a series of contour plots from the 14 term estimating equation at $1100^{\circ} \mathrm{C}$. A factor was added ( $2.5 \times 0.352155$ ) to give a 95 percent confidence interval so that the alloy would have excellent cyclic oxidation resistance (i.e., $\log \mathrm{K}_{\mathrm{a}} \leqq-0.7$ ). The criterion chosen was such that the total $\mathrm{Cr}+\mathrm{Al}+\mathrm{Ta}$ content would be at a minimum. On this basis the composition for the "best" cyclic oxidation resistance should be $6 \mathrm{Al}-5 \mathrm{Cr}-$ 8.6 Ta . Thus a typical ideal alloy should be $\mathrm{Ni}-10 \mathrm{Co}-6 \mathrm{Al}-5 \mathrm{Cr}-8.6 \mathrm{Ta}-0.9 \mathrm{Ti}-1 \mathrm{Hf}-0.15 \mathrm{C}-0.015-0.05 \mathrm{Zr}$. This high strength superalloy would satisfy all the compositional constraints of a group I alumina/aluminate forming alloy with good cyclic oxidation resistance and contain no deletereous alloy additions implicit from the 14 term estimating equation.

## SUMMARY OF RESULTS

As a result of statistical analysis of 323 cyclic oxidation runs in static air for 36 Ni - and Co - base high strength superalloys in the 1000 to $1150{ }^{\circ} \mathrm{C}$ range using an oxidation attack parameter, $\mathrm{K}_{\mathrm{a}}$ derived from $\Delta W / A$, time data the following results were obtained:
(1) Using multiple linear regression analysis with $\log K_{a}$ as the dependent variable a second degree estimating equation can be derived as a function of nominal alloy composition and test temperature based on $315 \mathrm{~K}_{\mathrm{a}}$ values with a high degree of fit.
(2) The derived 14 term estimating equation has an $R^{2}$ value of close to 85 percent and the numerous replicate runs show the maximum possible $R^{2}$ would be close to 89 percent due to 11 percent pure error and only 4 percent lack of fit. This indicates this particular 14 term model is adequate and can be used to predict oxidation results and design alloys with a high degree of confidence.
(3) Based on the coefficients of the regression equation Cr and Al are considered beneficial, and Ta is beneficial when Al is present. Nb is deleterious when $\mathrm{Ta}, \mathrm{Ti}$, and Hf are present and should be omitted. Mo and W should be at a minimum since they adversely affect Al and Cr , respectively. $\mathrm{Re}, \mathrm{V}$, and Ti should not be alloyed if possible. Ni in Co-base alloys and Co in Ni-base alloys appear innocuous as does $\mathrm{C}, \mathrm{B}$, and Zr within the range of their nominal compositions of the 36 alloys studied.
(4) The same estimating equation appeared equally valid for either Ni - or Co-base alloys and for both alumina/aluminate formers or chromia/chromite formers.
(5) Of the 36 alloys studied (see table I) the five best all group I alumina/aluminate formers can be ranked as follows from best to worse (low $\mathrm{K}_{\mathrm{a}}$ to high) based on the estimating equation computed at $1100^{\circ} \mathrm{C}$ :
(a) B-1900
(b) B-1900 $+\mathrm{H}_{\mathrm{f}}$
(c) NASA-TRW-VIA
(d) TRW-R
(e) TAZ 8A
(6) The estimating equation was used to calculate $K_{a}$ values for NASAIR- 100 a related alloy and compared to $K_{a}$ values derived from cyclic oxidation tests at 1150 and $1200^{\circ} \mathrm{C}$. The actual and derived $K_{a}$ 's agreed well within the 95 percent confidence interval.
(7) An optimum Ni-base alloy with maximum possible cyclic oxidation resistance along with a minimum total alloy content with good mechanical properties was designed using both the $\log \mathrm{K}_{\mathrm{a}} 14$ term estimating equation and the compositional constraints implicit in table I. This alloy in weight percent was the alumia/aluminate former alloy:
$\mathrm{Ni}-10 \mathrm{Co}-5 \mathrm{Cr}-6 \mathrm{Al}-8.6 \mathrm{Ta}-0.9 \mathrm{Ti}-0.15 \mathrm{C}-0.015 \mathrm{~B}-0.05 \mathrm{Zr}$.

## CONCLUSIONS

1. A cyclic oxidation attack parameter, $\mathrm{K}_{\mathrm{a}}$ derived from gravimetric/time data which has proven useful in the past to quantitatively rank cyclic oxidation resistance for a number of heater type alloys was successfully to evaluate the cyclic oxidation resistance of a large number of complex Ni - and Co-base high strength superalloys.
2. Using $\log _{10} \mathrm{~K}_{\mathrm{a}}$ as the dependent variable an estimating equation involving alloy chemistry and test temperature was derived from the experimentally derived $K_{a}$ values using multiple linear regression. This allowed the oxidation resistance of the alloys studied as well as similar alloys to be successfully predicted and ranked.
3. The estimating equation can be used to design comparable alloys based on alloy composition and test temperature.

## APPENDIX A - DERIVATION OF INDIVIDUAL $K_{a}$ VALUES

A total of $323^{1}$ runs based on the 36 alloys in table I of the $\Delta W / A$ versus time data from references 1 and 2 were individually analyzed according to equation (4), by multiple linear regression.

$$
\Delta W / A=k_{1}^{1 / 2} t^{1 / 2}+k_{2} t \pm \text { S.E.E. }
$$

Where $\Delta W / A$ is the specific weight change at any time, $t$ in hours, $k_{1}^{1 / 2}$ is a growth constant that when squared is analogous with the parabolic scaling constant, $k_{p}$; and $k_{2}$ is a linear coefficient and S.E.E. is the standard error of estimate on the $\Delta W / A$ estimates. The significance level for each coefficient is tested to the 10 percent significance level. If both are significant and $k_{1}^{1 / 2}$ is postivie then an attack parameter, $\mathrm{K}_{\mathrm{a}}$ is defined as:

$$
K_{a}=\left(k_{1}^{1 / 2}+10\left|k_{2}\right|\right)
$$

But if $\mathrm{k}_{1}^{1 / 2}$ is either negative or not significant then $\mathrm{K}_{\mathrm{a}}$ is re-defined as

$$
\mathrm{K}_{\mathrm{a}}=20\left|\mathrm{k}_{\mathbf{2}}\right|
$$

The other limiting case is when there is no linear component such as spalling, scale vaporization, excessive scale growth etc., $\mathrm{K}_{\mathrm{a}}$ reduces to simply $\mathrm{K}_{\mathrm{a}}=\mathbf{k}_{\mathrm{l}}^{1 / 2}$ or for diffusion controlled scaling $K_{a}=k_{p}^{1 / 2}$. Here $k_{p}$ is the conventional isothermal parabolic scaling constant.

The runs analyzed ranged in temperatures from 1000 to $1150^{\circ} \mathrm{C}$. The times analyzed were at $1000^{\circ} \mathrm{C}$ were $500 \mathrm{hr}, 1100^{\circ} \mathrm{C}-200 \mathrm{hr}$ and $1150^{\circ} \mathrm{C}-100 \mathrm{hr}$. The times may be shorter if the specific weight charges are extreme ( $>100 \mathrm{mg} / \mathrm{cm}^{2}$ ) usually with associated massive scale spall.

The total of 323 cyclic oxidation sample runs involving 36 alloys were analyzed as described above using regression analysis on the specific weight change/time data. $\mathrm{K}_{\mathrm{a}}$ values were then computed from the appropriate $\mathrm{k}^{1 / 2}$ and/or $\mathrm{k}_{2}$ constants. Table A-I summarizes the class of $\mathrm{K}_{\mathrm{a}}$ values derived for each alloy at each temperature. There were 20 runs at $1000^{\circ} \mathrm{C}, 128$ at $1100^{\circ} \mathrm{C}$ and 172 at $1150^{\circ} \mathrm{C}$. There were also three runs at $1093{ }^{\circ} \mathrm{C}\left(2000^{\circ} \mathrm{F}\right)$. An examination of these $323 \mathrm{Ka}_{\mathrm{a}}$ values led to dropping 8 of these values. Seven were inferred to be statistical outliers (runs 204-3, 336-4, 472-6, 324-4, 656-1, 657-4, and 664-6). In addition run 481-6 was dropped because its $\Delta W / A$ values were positive but gave too poor a fit to any of the standard model equations to drive $K_{\mathbf{a}}$.

The individual $\mathrm{K}_{\mathrm{a}}$ values are listed in table A-II. Of the 315 valid runs 231 follow the type I paralinear model the remaining 84 are of the type III type showing a linear weight loss. In general the individual regression fits are quite good to models I or III with $R^{2}$ values usually well over 90 percent. Of the 315 valid runs, 25 had $R^{2}$ values under 90 percent. Of these, 16 had $R^{2}$ values in 80 to 90 percent range, 5 in the 70 to 80 percent range, 3 in the 60 to 70 percent range, and 1 in the 50 to 60 percent

[^2]range. In the overall analysis, however, these three values with the lowest $\mathrm{R}^{2}$ model fits in the 50 to 70 percent range were not even close to being statistical outliers so they were retained for the overall analysis. These valid $K_{a}$ values can then be used for further comparison and analyses.

## APPENDIX B - SUPPLEMENTAL CYCLIC OXIDATION PLOTSS

Figures B-1 to B-28 show the additional 28 alloy runs not included in references 1 and $\mathbf{2}$. The $\mathrm{Ka}_{\mathrm{a}}$ values were derived as described in the body of the text. The test cycles were 1 hr in static air.

## APPENDIX C - BASIC LINEAR OXIDATION MODEL

A summary of the simplest linear model involving 11 significant terms of the original 15 first order terms listed in the main body of the text are shown in tables C-I and C-II. A reasonable $\mathbf{R}^{2}$ is derived as indicated in table C-I. However, table C-II indicates the residual sum of squares when partitioned into true error (i.e. replicate) and lack of fit error the simplest model is not adequate. This led to the more complex final model which included second degree terms.

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table I.-NOMINAL ALLOY COMPOSITION FOR HIGH-TEMPERATURE TURBINE alloys

| Alloy | Compoaltion, w1\% |  |  |  |  |  |  |  |  |  |  |  |  | Commente |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Ni | Co | Cr | Al | Ti | Mo | w | cb | T. | $c$ | B | Zr | Hf |  |
| Alloy 625 | Belance | $\cdots$ | 22.5 | 0.2 | 0.2 | 9.0 | $\cdots$ | See Commente | Seen Comments | 0.05 | $\cdots$ | $\cdots$ | $\cdots$ | $\mathrm{Cb}+\mathrm{Ta}=3.66^{1}$ |
| Alloy 718 |  | $\cdots$ | 19.0 | . 5 | . 8 | 3.05 | ..... | Soe Commeats | Soe Commente | 04 | 0.005 | ...... | $\cdots$ | With 18.5 wt\% $\mathrm{Fs} ; \mathrm{Cb}+\mathrm{Ta}=5.30^{2}$ |
| Astroloy |  | 15.0 | 15.0 | 4.4 | 3.5 | 5.25 | 4.0 | $\cdots$ | --------... | . 06 | . 09 | 0.06 | ----- | Slimilar to U-700 |
| B. 1900 |  | 10.0 | 8.0 | 6.0 | 1.0 | 6.0 | ${ }^{1} 1$ | 0.1 | 4.3 | 1 | . 015 | . 08 | $\cdots$ |  |
|  |  | 10.0 15.0 | 8.0 19.0 | 6.0 5.5 | 1.0 5.5 | 6.0 3.0 | ${ }^{1}$ | 1 | 4.3 | ${ }_{18} 18$ | . 015 | .08 .05 |  |  |
| (1N.100 ${ }_{\text {IN. } 713 \mathrm{LC}}$ |  | 15.0 | 10.0 12.0 | 5.5 5.9 | 6.5 | 3.0 4.5 | $\cdots$ | 2.0 | $\cdots$ | .18 .05 | .015 .010 | .05 .10 | $\cdots$ |  |
| IN-738 |  | 8.5 | 16.0 | 3.4 | 3.4 | 1.75 | 2.6 | . 9 | 1.75 | ${ }^{17}$ | . 010 | . 10 | ---- | ${ }^{*}{ }^{\text {c ineluden }}$ one hot-worked alloy with $\sim \mathrm{a} \mathrm{Cb}$ |
| IN-792 |  | 9.0 | 12.7 | 3.2 | 4.2 | 2.0 | 3.9 |  | 3.9 | . 21 | . 02 | . 10 | . 75 | 4 |
| IN-939 |  | 19.0 | . 22.0 | 2.0 | 3.6 | --- | 2.0 | 1.0 | 1.5 | . 15 | . 01 | . 10 | --- |  |
| MAR-M-200 |  | 10.0 | 2.0 | 5.0 | 2.0 | $\cdots$ | 12.5 | 2.7 | $\cdots$ | .15 18 | . 015 | . 05 | $\cdots$ | 2 |
| MAR-M-200 + Ht |  | 10.0 | 0.0 | 5.0 | 20 | $\overline{35}$ | ${ }_{5}^{11.5}$ | 1.0 2.7 | $\cdots$ | . 15 | ${ }^{.015}$ | . 05 | 1.5 | 2 |
| MAR-M-246 |  | 11 | 11 | 5.0 | 1.5 | 2.5 | ${ }_{6}$. | 2.7 | 2.0 | . 18 | . 0.01 | . 01 | $\cdots$ | With 0.1 wt\% Cu |
| MAR-M-247 |  | 10.0 | 8.2 | 5.5 | 1.0 | . 6 | 10.0 |  | 3.0 | . 16 | . 02 | . 09 | 1.5 |  |
| MAR-M-421 |  | 0.6 | 15.8 | 4.3 | 1.8 | 20 | 3.8 | 2.0 | - | . 16 | . 015 | ${ }^{205}$ | --- |  |
| NASA-TRW-VIA |  | 7.5 | 6.1 | 5.4 | 1.0 | 2.0 | 5.5 | . 5 | 9.0 | . 13 | . 02 | . 13 | 40 | With 0.5 wt\% Rem; ${ }^{\text {c }}$ |
| Nimonic 115 |  | 14.0 | 14.8 | 4.9 | 4.0 | 3.5 | - | $\cdots$ | - | . 16 | . 015 | . 001 | $\cdots$ |  |
| NX-188 |  | --3.0. | $<10$ | 8.0 | . | 18.0 | $\ldots$ | $\cdots$ | $\square$ | 04 | -- | $\cdots$ | $\cdots$ | 3 |
| Reme 41 |  | 110 9.6 | 19.0 | 1.5 | 3.1 | 10.0 | $\cdots$ | $\cdots$ | $\square$ | 109 | . 01 | -- | - |  |
|  |  | 9.6 10.0 | 14.0 9.0 | 3.0 | 5.0 4.0 | 4.0 2.0 | 7.0 | $\cdots$ | 3.8 | . 17 | . 015 | . 03 | $\cdots$ |  |
| Rene 125 |  | 10.0 | 9.0 | 8.0 | 2.5 | 2.0 | 7.0 | - ---7-...... | 3.8 | . 10 | . 02 | . 05 | 1.50 | ${ }_{4}$ |
| R-150-SX |  | 12.0 | 5.0 | 5.5 | --- | 1.0 | 5.0 | ------ | 6.0 | --- | --.-. | -- | $\cdots$ | With 3.0 wt Re and 2.2 wt V |
| TAZ-8A |  | $\cdots$ | 6.0 | 6.0 | $\cdots$ | 4.0 | 10 | 2.5 | 8.0 | . 125 | . 004 | 1.0 | $\cdots$ |  |
| TRWW R |  | 8.0 | 8.0 | 5.3 | . | 3.0 | 4.0 | . 3 | 6.0 | 05 | . 015 | . 12 | 1.00 |  |
| TRW-1800 |  | ${ }^{6}$ | 13.0 | 6.0 | . 6 | - | 9.0 | 1.5 | $\cdots$ | .08 |  |  | ---. |  |
| U-520 U.700 |  | 12.0 12.5 | 19.0 | 2.0 | 3.0 | 80 | 1.0 | ------ | $\cdots$ | . 06 | ${ }^{.005}$ | $\ldots$ | - |  |
| U-700 |  | 12.6 | 18.0 | 4.3 | 3.5 | 4.5 | -- | $\cdots$ | - | . 07 | . 03 |  |  | verying Co brules |
| U-710 |  | 15.0 | 18.0 | 2.5 | 5.0 | 3.0 | 1.5 | $\cdots$ | $\cdots$ | . 10 | . 012 | $\cdots$ | $\cdots$ |  |
| U. 720 |  | 15.0 | 18.0 | 2.5 | 5.0 | 30 | 1.2 | $\cdots$ | ------ | . 04 | . 03 | . 03 | -- |  |
| Weapeloy |  | 13.5 | 19.5 | 1.3 6.5 | $\stackrel{3}{3.0}$ | 6.3 18.5 | $\cdots$ | $\underline{\square}$ |  | . 15 | .006 | . 06 | $\cdots$ |  |
| MAR-M-609 | 10.0 | Balama | 23.5 | ---. | . 2 | -- | 7.0 | -- | 3.5 | . 00 | $\ldots$ | .50 | $\ldots$ |  |
|  | 10.0. | Belance | ${ }_{26.5}^{21.0}$ | - | - | - | 11.0 | 2.0 | $\cdots$ | . 48 | - | -- | - | Wuth 2.0 wtx Fe* |

[^3][^4]TABLE II.-MULTIPLE REGRESSION* RESULTS
FOR LOG ${ }_{10}$ Ka AS A FUNCTION OF ALLOY COMPOSITION IN wt\%, AND OF ABSOLUTE TEST TEMPERATURE IN $1 / T_{K}$ BASED ON AN INITIAL SELECTION OF 23 MOST LIKELY

1st AND 2nd ORDER REACTIONS.
NUMBER DATA VALUES $\mathbf{n}=315$

| $\mathbf{Z i}=23, \mathbf{Z f}=14$ |  |  |
| :---: | :---: | :---: |
| Significant terms, Z | Coefficient | t-statistic |
| Al-Ta | -0.03008490 | -7.365 |
| $1 / \mathrm{T}_{\mathrm{K}}$ | -28733.83015 | -11.020 |
| $\mathrm{Al}^{\mathbf{2}}$ | -. 05162169 | -9.088 |
| Al.V | +. 16395511 | 7.053 |
| Cr | -. 71873828 | -5.241 |
| Nb -Ta | +.05346153 | 7.115 |
| $\mathrm{Cr} \cdot\left(1 / \mathrm{T}_{\mathrm{K}}\right.$ ) | +924.75130 | 4.850 |
| Ti.Ta ${ }^{\text {m }}$ | +. 01932161 | 2.432 |
| $\mathrm{Cr} . \mathrm{W}$ | +.003726623 | 5.878 |
| Al. Mo | +. 01273215 | 6.960 |
| $\mathrm{Ti} \cdot \mathrm{Nb}$ | +. 08140372 | 4.089 |
| $\cdots \mathrm{Nb} \cdot \mathrm{Hf}$ | +. 24155034 | 2.930 |
| Ti | +. 08344541 | 2.890 |
| Re | +. 21293029 | 1.739 |
| $a_{0}$, intercept | 22.75638644 |  |

$\mathrm{R}^{\mathbf{3}}=84.43 \% \quad$ S.E.E. $=0.352155 \quad \mathrm{Zi}=23$ $\mathrm{Al} \cdot \mathrm{Ta}, 1 / \mathrm{T}_{\mathrm{K}}, \mathrm{Al}^{2}, \mathrm{Al} \cdot \mathrm{V}, \mathrm{Cr}_{\mathrm{r}}, \mathrm{Nb} \cdot \mathrm{Ta}, \mathrm{Cr} \cdot\left(1 / \mathrm{T}_{\mathrm{K}}\right)$, $\mathrm{Ti} \cdot \mathrm{Ta}, \mathrm{Cr} \cdot \mathrm{W}, \mathrm{Al} \cdot \mathrm{Mo}, \mathrm{Ti} \cdot \mathrm{Nb}, \mathrm{Nb} \cdot \mathrm{Hf}, \mathrm{Ti}, \mathrm{Re}, \mathrm{Al}, \mathrm{Mo}$, $\mathrm{Nb}, \mathrm{Ta}, \mathrm{C}, \mathrm{Zr}, \mathrm{Hf}, \mathrm{Cr}^{2}, \mathrm{Ti} \cdot \mathrm{Z}_{\mathrm{r}}$
"Stepwise Regreassion-Variables are added one at a time atarting with the most significant, the F-atatistic for a variable must be sigaificant to 0.15 . After a variable is added, however, the atepwise method looks at all the variables already in the model and deletes any that does not produce an $F$-statistic significant to the 0.15 level.

TABLE III.-ANALYSIS OF VARIANCE (ANOVA)
SUMMARY FOR $n=315$ DATA SET; $\mathbf{Z f}=14$
SHOWING SOURCES OF VARIATION INCLUDING LACK
OF FIT OF THE ESTIMATING EQUATION

| Source | Degrees of freedom, d. 1 | Sum of square: | Mean squaren |
| :---: | :---: | :---: | :---: |
| Model | 14 | 201.68573 | 14.40469511 |
| Residual | 300 | 37.20395146 | . 12401317 |
| Lack of fit | (67) | (9.8844261) | (.14762875) |
| Replication | (233) | (27.319525) | (.11725118) |
| Total | 314 | 238.86988 |  |

$$
F-\text { Ratio }=\frac{M S(\text { LOF })}{\text { MS(REPS) }}=\frac{0.14752875}{0.11725118}=1.258^{\mathrm{a}}
$$

> The lack of fit term appears not be be significent
> since the $F-$ Ratio for $(1-a)$ where $a=0.95=1.658$ which exceeds the MS(LOF)/MS(REPS) ratio derived in this tudy. Therefore this model is considered satiafactory.

TABLE IV.-GROUP I ALLOYS - ALUMINA/ALUMINATE SCALE FORMERS COMPARISON OF PREDICTED Ka's FROM LOG Ka ESTIMATES FOR COEFFICIENTS LISTED IN TABLE II TO THE AVERAGE* OF THE OBSERVED Ka's FOR EACH ALLOY AT EACH TEST TEMPERATURE

| Alloy | Wt\% ${ }^{\text {b }}$ |  |  | $1000{ }^{\circ} \mathrm{C}$ |  | $1100^{\circ} \mathrm{C}$ |  | $1150^{\circ} \mathrm{C}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Al | Cr | Ta | Average Ka | $\begin{aligned} & \text { Predicted } \\ & \mathrm{Ka} \end{aligned}$ | Average Ka | $\begin{gathered} \text { Prodicted } \\ K a \end{gathered}$ | Average Ka | $\begin{gathered} \text { Predicted } \\ \mathrm{Ks} \\ \hline \end{gathered}$ |
| MAR-M-200 | 5.0 | 9.0 | $\ldots$ |  | 0.9752 | 7.2548 | 14.3609 | 88.3329 | 47.7780 |
| MAR-M-200 + Hf | 5.0 | 9.0 | .... |  | 1.0993 | 16.9870 | 16.1768 | 58.2881 | 63.8568 |
| MAR-M-211 | ¢. 0 | 9.0 | .... |  | . 7883 | 13.2160 | 11.6007 | 24.1583 | 38.6218 |
| MAR-M-246 | 5.0 | 11.0 | 2.0 |  | . 0726 | 1.6534 | . 8376 | 18.0767 | 2.5006 |
| Rent - 125 | 5.0 | 9.0 | 3.8 |  | . 1400 | 1.9005 | 2.0602 | 9.7719 | 6.8580 |
| TRW-R | 5.3 | 8.0 | 6.0 | 0.0555 | . 0323 | . 1063 | . 5365 | . 8302 | 1.8863 |
| NASA-TRW-VIA | 5.4 | 6.1 | 9.0 |  | . 0169 | . 3155 | . 3533 | 1.3698 | 1,3776 |
| IN-100 | 5.5 | 10.0 | .... |  | 1.8657 | 14.0391 | 24.3067 | 83.0398 | 76.6307 |
| MAR-M-247 | 5.5 | 8.2 | 3.0 | . 0525 | . 0477 | . 5022 | . 7743 | 4.3845 | 2.6928 |
| R-150-SX | 5.5 | 5.0 | 6.0 | 3.5375 | 2.8480 | 45.0103 | 68.2400 | 314.856 | 282.519 |
| IN-713 LC | 5.9 | 12.0 | -... |  | . 0924 | . 7146 | . 9439 | 1.2619 | 2.6685 |
| B-1900 | 6.0 | 8.0 | 4.3 | . 0532 | . 0187 | . 1839 | .3100 | 1.3843 | 1.0898 |
| B-1900 + Hf | 6.0 | 8.0 | 4.3 |  | . 0197 | . 4228 | . 3277 | 1.0774 | 1.1532 |
| TAZ - 8A | 6.0 | 6.0 | 8.0 | . 0972 | . 0252 | . 4243 | . 5244 | 2.2900 | 2.0634 |
| TRW - 1800 | 6.0 | 13.0 | ...- |  | . 0968 | . 7309 | . 8746 | 3.6902 | 2.3416 |

${ }^{\text {a }}$ Observed $\mathrm{Ka}^{\prime}$ s are based on the antilog of the average of the Log Ka values for each alloy at each test temperatuare.
${ }^{\mathrm{b}} \mathrm{Al}_{1} \mathrm{Cr}$, and Ta are the key elements in improving cyclic oxidation renintance.

TABLE V.-GROUP II ALLOYS - CHROMIA/CHROMITE AND NiO SCALE FORMERS - COMPARISON OF PREDICTED Ka's FROM Log Ka ESTIMATES FOR COEFFICIENTS LISTED IN TABLE II TO THE AVERAGEa OF THE OBSERVED Ka's FOR EACH ALLOY AT EACH TEST TEMPERATURE

| Alloy | W \% $^{\text {b }}$ |  |  | $1000{ }^{\circ} \mathrm{C}$ |  | $1100^{\circ} \mathrm{C}$ |  | $1150{ }^{\circ} \mathrm{C}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Al | Cr | Ta | Avorage Ka | $\begin{gathered} \text { Predicted } \\ \mathrm{Ka} \end{gathered}$ | Average Ka | $\begin{gathered} \text { Predicted } \\ \mathrm{Ka} \end{gathered}$ | Average Ka | $\begin{gathered} \text { Predicted } \\ \mathrm{Ka} \\ \hline \end{gathered}$ |
| MAR-M-509 | 0 | 23.5 | 3.5 |  | 10.2035 | 25.2623 | 25.6668 | 46.5804 | 38.7764 |
| WI-52 | 0 | 21.0 | .... |  | 16.1108 | 33.6529 | 54.9552 | 116.887 | 95.1412 |
| X-40 | 0 | 25.5 | .... |  | 12.4060 | 35.5703 | 24.4580 | 27.6292 | 33.1348 |
| Alloy 625 | 0.2 | 22.5 | 1.9 |  | 3.9692 | 28.7163 | 11.2780 | 36.4196 | 17.9926 |
| Alloy 718 | 0.5 | 19.0 | 3.3 |  | 8.3100 | 28.5671 | 36.1671 | 43.3921 | 69.8240 |
| Waspaloy | 1.3 | 19.5 | .... | 4.7380 | 3.7067 | 5.7051 | 15.1791 | 23.1244 | 28.5170 |
| René 41 | 1.5 | 19.0 | $\cdots$ |  | 4.6173 |  | 20.0954 | 33.0520 | 38.7982 |
| IN-939 | 2.0 | 22.0 | 1.5 |  | 9.9811 | 32.5843 | 30.1413 | 55.3798 | 49.4148 |
| U-520 | 2.0 | 19.0 | ---- |  | 3.9657 | 31.6500 | 17.2593 | 55.9731 | 33.3208 |
| U.710 | 2.5 | 18.0 | ...- |  | 4.1103 | 33.7545 | 20.2068 | 48.908 | 41.1959 |
| U-720 | 2.5 | 18.0 | .... | 6.3587 | 3.9242 | 32.3348 | 19.2918 | 41.5751 | 39.3306 |
| René 80 | 3.0 | 14.0 | ---- |  | 2.4992 | 37.3206 | 20.0015 | 60.3715 | 50.7086 |
| IN-792 | 3.2 | 12.7 | 3.9 |  | 2.0481 | 21.9872 | 19.2034 | 49.8747 | 52.2593 |
| IN-738 | 3.4 | 16.0 | 1.8 | 1.6985 | 3.1246 | 27.3451 | 19.5987 | 37.0869 | 44.5570 |
| MAR-M-421 | 4.3 | 15.8 | ---- |  | 1.3436 | 9.5308 | 8.6353 | 34.9361 | 19.8471 |
| René 120 | 4.3 | 9.0 | 3.8 |  | . 6020 | 6.8484 | 8.8588 | 14.9107 | 24.4930 |
| U-700 | 4.3 | 15.0 | .-.- | 1.1562 | . 7657 | 3.6784 | 5.4247 | 21.2444 | 13.0235 |
| Astroloy | 4.4 | 15.0 | ...- |  | 1.2896 | 3.2373 | 9.1370 | 61.7246 | 21.9361 |
| Nimonic 115 | 4.9 | 14.6 | --.- | . 3982 | . 4071 |  | 3.0284 | 1.6397 | 7.4309 |
| WAZ-20 | 6.5 | -. | ---- |  | . 3425 | 20.0738 | 15.0883 | 82.7178 | 82.0313 |
| NX-188 | 8.0 | ---. | ---- |  | . 0518 | 3.4403 | 2.2817 | 7.7592 | 12.4050 |

'Observed $\mathrm{Ka}^{\prime} \mathrm{a}$ are based on the antilog of the average of the Log Ka values for each alloy at each teat temperatuare.
${ }^{\mathrm{b}} \mathrm{Al}, \mathrm{Cr}$, and Ta are the key elementa in inproving cyclic oxidation resistance.

TABLE VI.-COMPARISON OF OBSERVED AND PREDICTED Ka VALUES FOR A TYPICAL TURBINE ALLOY Ni-BASE NASAIR-100(Ni-9Cr-5.75Al-1.2Ti-1Mo-3.30Ta-10.5W-.03Zr) TESTED IN CYCLIC OXIDATION FOR ONE HR EXPOSURE CYCLES IN STATIC AIR AT 1150 AND $1200^{\circ} \mathrm{C}$

| Run | Test <br> temperature | Teat time, <br> hrs | $\Delta W / A$ fasl, <br> $\mathrm{mg} / \mathrm{cm}^{2}$ | Ka <br> observed | Log Ks <br> observed | Log Ka <br> predicted ${ }^{2}$ | Standard <br> deviation, $\sigma$ | Deviation <br> $\sigma$-units ${ }^{b}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $44-1$ | $1150^{\circ} \mathrm{C}$ | 100 | -33.54 | 5.8137 | 0.7645 | 0.2684 | 0.3522 | 1.408 |
| $44-3$ | $1150^{\circ} \mathrm{C}$ | 100 | -38.75 | 6.9583 | 0.7751 | 0.2685 | 0.3522 | 1.438 |
| $42-1$ | $1200^{\circ} \mathrm{C}$ | 30 | -48.14 | 12.2041 | 1.0865 | 0.7554 | 0.3522 | 0.940 |

[^5]TABLE A-I.-CLASSIFICATION OF OBSERVED KE VALUES DERIVED FROM INDIVIDUAL $\triangle W / A$ VERSUS TIME VALUES FOR EACH ALLOY RUN FOR A TOTAL OF 323 RUNS INCLUDING EIGHT PROBABLE OUTLIERS

| Alloy | Number of samples tested at |  |  | Observed Ks, type |  | Number of outiler(a)" and reason(a) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $1000^{\circ} \mathrm{C}$ | $1100^{\circ} \mathrm{C}$ | $1150{ }^{\circ} \mathrm{C}$ | Paralinear | Linear |  |
| Alloy 625 | 0 | 1 | 1 | 2 | 0 | 0 |
| Alloy 718 | 0 | 1 | 1 | 2 | 0 | $0$ |
| Astroloy | 0 | 1 | 1 | 1 | 1 | 0 |
| B. 1900 | 1 | 8 | 30 | 23 | 16 | $10=4.457$ |
| $\mathrm{B}-1900$ + Hf | 0 | 3 | 3 | 0 | 6 | $10=3.798$ |
| IN-100 | 0 | $3^{1}$ | 13 | 11 | 5 |  |
| IN-713 LC | 0 | 1 | 2 | 0 | 3 | 0 |
| IN.738 | 1 | 10 | 5 | 16 | 0 |  |
| IN. $792{ }^{\text {b }}$ | 0 | 8 | 11 | 18 | 0 | $\begin{gathered} 1 \text { o }=-3.972, \text { approximate parabolic } \\ \mathbf{R}^{4}=0.998 \end{gathered}$ |
| IN-939 | 0 | 1 | 1 | 2 | 0 | 0 |
| MAR-M-200 | 0 | 3 | 4 | 5 | 2 |  |
| MAR-M-200 + Hf | 0 | 6 | 8 | 12 | 2 | 0 |
| MAR-M-211 | 0 | 3 | 3 | 5 | 1 | $2 \sigma=-3.175, \theta=-3.677$ |
| MAR-M-246 | 0 | 1 | 1 | 2 | 0 | 0 0 2.78 |
| MAR-M-247 | 2 | 5 | 5 | 9 | 3 | $1 \pm-2.785$ |
| MAR-M-421 | 0 | 1 | 1 | 2 | 0 | $0$ |
| NASA-TRW-VIA | 0 | 6 | 13 | 15 | 4 | 0 |
| Nimonic 115 | 2 | 1 | 1 | 2 | 2 | $10 \pm-3.13 i$ |
| NX-188 | 0 | 2 | 3 | 4 | 1 | 0 |
| Rene-41 | 0 | 0 | 3 | 3 | 0 | 0 |
| Rene-60 | 0 | 2 | 3 | 5 | 0 | 0 |
| René-120 | 0 | 1 | 2 | 3 | 0 | 0 |
| René-125 | 0 | 3 | 2 | 4 | 1 | 0 |
| R-150-SX | 2 | 1 | 1 | 3 30 | 1 | 0 |
| TAZ-8A | 1 | 11 | 11 | 30 | 3 |  |
| TRW-R | 1 | 2 | 2 | 1 | 4 | 0 |
| TRW-1800 | 0 | 1 | 1 | 1 | 1 | 0 |
| U. 520 | 0 | 1 | 1 | 2 | 0 | 0 |
| U.700 | 5 | 27 | 12 | 21 | 23 | 0 |
| U.710 | 0 | 1 | 1 | 2 | 0 | 0 |
| U-720 | 2 | 1 | 1 | $1{ }^{4}$ | 0 | $\mathrm{l}^{0}$ |
| Waspaloy | 3 | 5 | 5 | 12 | 1 | $1^{\text {c }}$ |
| WAZ-20 | 0 | 2 | 3 | 3 | 2 |  |
| MAR-M-509 | 0 | ${ }_{2}{ }^{\text {d }}$ | 3 | ${ }^{6}$ | 0 | 0 |
| WI-52 X-40 | 0 | $2^{\text {d }}$ 1 | 7 | 3 | 0 | 0 0 |
|  |  |  |  |  |  |  |
| Total | 20 | 128 | 172 | 230 | 89 | 8 |

${ }^{2}$ An additional IN-100 sample tested at $1093^{\circ} \mathrm{C}$. paraliaear behavior.
${ }^{\text {b }}$ One IN-792 sample showed almont pure parabolic behavior but was deemed an outlier.
${ }^{\circ}$ One Waspalloy sample ( $481-6$ ) teated for 200 I . hr cycles at $1100^{\circ} \mathrm{c}$ gave tuch a poot nt to any of 3 posilble models-
paralinear, linear or parabolic that it was automatically considered an outliet.
${ }^{\text {d }}$ Two additinal WI-52 samples tested at $1093^{\circ} \mathrm{C}$, paralinear behavior.
"Based on the model:
$\log \mathrm{Ka}=\mathrm{a} \cdot \mathrm{CoNi}+\mathrm{b} \cdot \mathrm{Ti}+\mathrm{c} \cdot \mathrm{Mo}+\mathrm{d} \cdot \mathrm{W} \pm \mathrm{e} \cdot \mathrm{Nb}+\mathrm{f} \cdot \mathrm{Ta}+\mathrm{g} \cdot \mathrm{C}+\mathrm{h} \cdot \mathrm{B}+\mathrm{l} \cdot \mathrm{Zr}+\mathrm{j} \cdot \mathrm{HI}+\mathrm{k} \cdot \mathrm{V}+\mathrm{l} \cdot \mathrm{Al} \cdot \mathrm{Cr}_{\mathrm{t}}+\mathrm{m} \cdot \mathrm{Al}^{\mathbf{2}}+\mathrm{n} \cdot \mathrm{Cr}^{2}+\mathbf{o} \cdot \mathbf{1 / T} \mathrm{T}_{\mathrm{K}}+\mathrm{p} \cdot \mathrm{Cr}+\mathrm{q} \cdot \mathrm{Re} \pm \sigma$ if $\sigma> \pm 2.5$ the sample is dropped as an outlier.

TABLE A-II.-INDIVIDUAL Ka Values and associated specific weight change data for
EACH ALLOY SAMPLE RUN, $\mathrm{n}=315$

| Alloy | Test temperature, ${ }^{\circ} \mathrm{C}$ | Run number | Test time, hr | Model type | $k_{1}^{1 / 2}$ | $\mathbf{k}_{2}$ | Ka | $\mathrm{R}^{\mathbf{2}}$ | Final $\Delta W / A$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Alloy 625 | 1100 | 351-4 | 200 | Paralinear | 7.99315 | -2.07222 | 28.7154 | 0.998 | -293.20 |
| Alloy 625 | 1150 | 352.4 | 100 |  | 7.69380 | -2.87258 | 36.4196 | . 999 | -208.10 |
| Alloy 718 | 1100 | 351-3 | 200 |  | 8.17729 | -2.03898 | 28.5671 | . 998 | -284.60 |
| Alloy 718 | 1150 | 352-3 | 100 |  | 8.67148 | -3.47206 | 43.3921 | . 999 | -255.70 |
| Astroloy | 1100 | 473-3 | 200 | $\checkmark$ | 1.21721 | -. 20201 | 3.2373 | . 928 | -30.25 |
| Astroloy | 1150 | 472-3 | 100 | Linear | -........- | -3.08623 | 61.7246 | . 992 | -318.80 |
| B-1900 | 1000 | 471-3 | 500 | Paralinear | . 03803 | -. 00151 | . 0531 | . 926 | +. 19 |
|  | 1100 | 103-3 | 200 | Paralinear | . 07635 | -. 01044 | . 1808 | . 978 | -. 97 |
|  |  | 103-4 |  | Paralinear | . 08866 | -. 01469 | . 2356 | . 951 | -1.56 |
|  |  | 186-6 |  | Linear | -...--.--* | -. 01597 | . 3193 | . 873 | -2.52 |
|  |  | 190-5 |  | Linear | --...-** | -. 00843 | . 1686 | . 832 | $-1.20$ |
|  |  | 276-6 |  | Paralinear | . 04583 | -. 00840 | . 1298 | . 983 | $-.97$ |
|  |  | 324-2 |  |  | . 06368 | -. 01226 | . 1863 | . 972 | $-1.40$ |
|  | $t$ | 327-1 | $\checkmark$ |  | . 03604 | -. 00924 | . 1284 | . 983 | -1.21 |
|  | 11.50 | 41-1 | 100 |  | . 06418 | -. 03528 | . 4169 | . 999 | -2.87 |
|  |  | 78-1 |  |  | . 58862 | -. 24889 | 3.0775 | . 994 | -19.91 |
|  |  | 78-2 |  | $\checkmark$ | . 65950 | -. 25321 | 3.1916 | . 995 | -19.59 |
|  |  | 95-1 |  | Linear | --------* | -. 05565 | 1.1130 | . 995 | -5.56 |
|  |  | 95-2 |  | Linear | -----.--- | -. 05231 | 1.0462 | . 995 | -5.05 |
|  |  | 101.3 |  | Paralinear | . 18539 | $-.04590$ | . 6444 | . 996 | -2.62 |
|  |  | 101.6 |  | Linear | -...-.-.-- | -. 04207 | . 8414 | . 988 | -3.97 |
|  |  | 107.4 |  | Linear | -------- | -. 06512 | 1.3025 | . 995 | -6.80 |
|  |  | 107.5 |  | Paralinear | . 40414 | -. 13133 | 1.7174 | . 997 | -9.46 |
|  |  | 123-1 |  |  | . 65939 | -. 16387 | 2.1981 | . 986 | -12.11 |
|  |  | 123-2 |  |  | . 72746 | $-.20699$ | 2.7974 | . 982 | -15.16 |
|  |  | 123-3 |  |  | . 57362 | -. 13841 | 1.9577 | . 985 | -9.12 |
|  |  | 123.4 |  |  | . 15333 | -. 05408 | . 6941 | . 981 | -4.35 |
|  |  | 123-5 |  |  | . 32815 | -. 12461 | 1.5743 | . 989 | -10.10 |
|  |  | 123-6 |  | $t$ | . 52619 | -. 12212 | 1.7474 | . 978 | -7.93 |
|  |  | 128-1 |  | Linear | --------* | -. 07332 | 1.4665 | .995 | -7.08 |
|  |  | 128-2 |  | Linear | -----.-.- | $-.05824$ | 1.1648 | . 999 | -5.62 |
|  |  | 130-1 |  | Paralinear | . 71171 | $-.16798$ | 2.3915 | . 976 | -11.14 |
|  |  | 130-2 |  |  | 2.32699 | -. 49507 | 7.2777 | . 981 | -28.76 |
|  |  | 130-3 |  |  | . 77096 | -. 20694 | 2.8404 | . 987 | -14.43 |
|  |  | 130-4 |  |  | . 21995 | -. 07800 | 1.0000 | . 999 | -5.66 |
|  |  | 130-5 |  |  | . 07632 | -. 06096 | . 6860 | . 998 | -5.42 |
|  |  | $130-6$ |  | . | . 42854 | $-.15200$ | 1.9485 | . 990 | -11.99 |
|  |  | 146-5 |  | Linear | --------- | -. 04454 | . 8908 | . 986 | -4.25 |
|  |  | 204-4 |  |  | --------- | -. 07026 | 1.4053 | . 947 | -6.13 |
|  |  | 221-1 |  |  | --.------ | -. 05004 | 1.0008 | . 990 | -4.75 |
|  |  | 221-5 |  |  | -- | -. 07562 | 1.5125 | . 997 | -7.31 |
|  |  | 321.2 |  |  | ---------* | -. 05778 | 1.1557 | . 995 | -5.50 |
|  |  | 328.1 |  |  | --------* | -. 03415 | . 6830 | . 989 | -3.27 |
|  | $\checkmark$ | $337-4$ | $\downarrow$ |  | -- | -. 03844 | . 7688 | . 994 | -4.11 |
| B-1900 + Hf | 1100 | 190-4 | 200 |  | -------- | -. 01208 | . 2416 | . 902 | -1.94 |
|  | 1100 | 326-3 | 200 |  | ---..--- | $-.08729$ | 1.7458 | . 983 | -1.65 |
|  | 1100 | 475.1 | 200 |  | -.......-- | $-.00896$ | . 1791 | . 959 | -1.44 |
|  | 1150 | 323-3 | 100 |  | -...-...-- | -. 0437 | . 874 | . 967 | -3.85 |
|  | 1150 | 474.1 | 100 | $\downarrow$ | -........- | $-.0664$ | 1.327 | . 978 | -7.80 |
| IN-100 | 1093 | 100-1 | 100 | Paralinear | 6.9924 | -2.1500 | 28.493 | . 999 | -148.10 |

TABLE A-II.-Coptinued.

| Alloy | Teat temperature, ${ }^{-} \mathrm{C}$ | Run number | Teat time, hr | Model type | $k_{1}^{1 / 2}$ | $\mathbf{k}_{2}$ | Ks | $\mathrm{R}^{\mathbf{2}}$ | $\begin{aligned} & \text { Final } \\ & \Delta W / A \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| IN-100 | 1100 | $393-1$ | 200 | Paralinear | 1.0514 | -0.1415 | 2.466 | 0.985 | $-15.25$ |
|  | 1100 | 413-4 | 75 | Linear | *-"-*** | -6.3439 | 126.878 | . 999 | -462.4 |
|  | 1100 | 469-1 | 200 | Linear | -7...... | -. 4421 | 8.842 | . 983 | -63.34 |
|  | 1150 | 41-6 | 100 | Paralinear | 5.3939 | -1.0188 | 15.582 | . 955 | -56.20 |
|  | 1 | 95.3 | 75 | Paralinear | 21.2371 | -6.4656 | 85.793 | . 897 | -306.0 |
|  |  | 95-6 | 75 | Linear | -->......- | -5.2591 | 105.183 | . 999 | -385.0 |
|  |  | 105-1 | 90 | Paralinear | 4.1880 | -7.9080 | 83.268 | . 999 | -652.7 |
|  |  | 105-2 | 100 |  | 13.8025 | -7.9190 | 92.992 | . 999 | $-635.2$ |
|  |  | 127-1 |  |  | 41.1241 | -7.9930 | 121.054 | . 983 | -417.6 |
|  |  | 127-2 |  |  | 14.9888 | $-10.7823$ | 122.812 | . 899 | -827.8 |
|  |  | 127-3 |  |  | 18.2168 | -3.8912 | 57.129 | . 089 | -220.2 |
|  |  | 127-4 |  |  | 16.6828 | -3.4270 | 50.952 | . 984 | -191.0 |
|  |  | 127.5 |  |  | 28.0016 | -5.2697 | 80.699 | . 988 | -277.2 |
|  |  | 127.6 | - | $\downarrow$ | 23.6953 | -4.3537 | 67.232 | . 953 | -231.9 |
|  |  | 414.4 | 60 | Linear | -п*- | -7.3451 | 146.902 | . 999 | -438.2 |
|  | $\downarrow$ | 470-1 | 45 |  | --**-*** | -11.8857 | 237.714 | . 999 | - 821.9 |
| IN-713 LC | 1100 | 473-5 | 200 |  | - | -. 0357 | . 715 | . 997 | -6.20 |
| IN-713 LC | 1150 | 41-4 | 100 |  | -...-...-. | $-.1386$ | 2.772 | . 993 | -12.98 |
| IN-713 LC | 1150 | 472-5 | 100 | ¢ |  | -. 0287 | . 575 | .966 | -2.52 |
| IN-738 | 1000 | 674-3 | 500 | Paralinear | 1.0279 | -. 0870 | 1.698 | . 953 | $-12.65$ |
|  | 1100 | 324-1 | 200 |  | 9.4313 | -1.3867 | 23.298 | . 976 | -55.81 |
|  |  | 413-2 |  |  | 11.9709 | -1.6564 | 28.535 | . 964 | -182.4 |
|  |  | 469-6 |  |  | 3.4517 | -.6258 | 8.710 | . 997 | -95.13 |
|  |  | 659-1 |  |  | 13.2680 | -2.6506 | 39.774 | . 998 | -338.4 |
|  |  | 663-2 |  |  | 13.5724 | -3.1061 | 44.634 | . 988 | -183.3 |
|  |  | 664-2 |  |  | 11.9394 | -1.8196 | 30.135 | . 992 | -199.50 |
|  |  | 679-4 |  |  | 7.5608 | -2.3929 | 31.490 | . 999 | -363.60 |
|  |  | 679-5 |  |  | 11.8385 | -1.9061 | 30.900 | . 994 | -215.80 |
|  |  | 680-4 |  |  | 7.2468 | -2.2193 | 29.440 | . 998 | -332.70 |
|  | $\downarrow$ | 680-5 | 6 |  | 4.6310 | -2.1713 | 26.344 | . 996 | -357.9 |
|  | 1150 | 41-2 | 100 |  | 8.5420 | -1.8366 | 26.908 | . 965 | -112.6 |
|  |  | 321-1 |  |  | 9.9065 | -2.2040 | 31.946 | . 976 | -134.1 |
|  |  | 414-2 |  |  | 11.7430 | -2.6738 | 38.481 | . 983 | -160.8 |
|  |  | 470-6 |  |  | 13.4443 | -2.9574 | 43.018 | . 986 | -170.6 |
|  | - | 658-1 | , |  | 5.5606 | -4.3745 | 49.305 | . 999 | -371.9 |
| IN-792 | 1100 | 310-2 | 200 |  | 9.0621 | -1.4302 | 23.364 | . 991 | -161.9 |
|  |  | 376-2 |  |  | 9.7766 | -1.6063 | 25.841 | . 995 | $-184.5$ |
|  |  | 326-5 |  |  | 9.5023 | -1.4302 | 23.804 | . 990 | -156.30 |
|  |  | 336-5 |  |  | 8.9258 | -1.6386 | 24.312 | . 995 | -184.4 |
|  |  | 411-6 |  |  | . 0792 | -1.2207 | 12.287 | . 973 | -148.8 |
|  |  | 469-4 |  |  | 10.0727 | -1.5685 | 25.757 | . 983 | -183.6 |
|  | $\checkmark$ | 657-5 | , |  | 9.4552 | -1.3011 | 22.466 | . 965 | -144.3 |
|  | 1150 | 323-2 | 100 |  | 13.6102 | -3.2478 | 46.088 | . 994 | -192.1 |
|  | 115 | 323-5 |  |  | 13.4964 | -3.3138 | 46.634 | . 996 | -196.2 |
|  |  | 337.5 |  |  | 12.4612 | -3.2332 | 44.794 | . 995 | -205.0 |
|  |  | 412.6 |  |  | 13.2188 | -3.3625 | 46.844 | . 993 | -208.5 |
|  |  | 425.4 |  |  | 13.8841 | -3.7341 | 51.225 | . 998 | $-233.2$ |
|  |  | 425-5 |  |  | 14.3983 | -4.1172 | 55.570 | . 998 | -264.5 |
|  |  | 426-4 |  |  | 14.1476 | -3.7264 | 51.411 | . 998 | -229.6 |
|  | $\downarrow$ | 426-5 | $\checkmark$ | $t$ | 13.1177 | -3.8563 | 51.681 | . 998 | -251.4 |

TABLE A-II.-Continued.

| Alloy | Test temperature, ${ }^{\circ} \mathrm{C}$ | Run number | Test <br> time, hr | Model type | $k_{1}^{1 / 2}$ | $k_{2}$ | $\mathbf{K a}$ | $\mathrm{R}^{2}$ | $\begin{aligned} & \text { Final } \\ & \Delta W / A \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| IN-792 | 1150 | 428-4 | 100 | Paralinear | 13.9847 | -3.6630 | 50.615 | 0.998 | -225.0 |
| IN-792 | 1150 | 428-5 | 100 |  | 17.1251 | -4.4736 | 61.861 | . 997 | -273.8 |
| IN-792 | 1150 | 470-4 | 100 |  | 12.3665 | -3.2108 | 44.474 | . 992 | -203.4 |
| IN-939 | 1100 | 327-3 | 200 |  | 12.3857 | -2.0199 | 32.584 | . 996 | -227.6 |
| IN-939 | 1150 | 328-3 | 100 |  | 15.8826 | -3.9472 | 55.380 | . 996 | -233.2 |
| MAR-M-200 | 1100 | 310-3 | 200 |  | 1.7693 | -. 3701 | 5.470 | . 994 | -52.16 |
|  | 1100 | 391.1 | 200 |  | 1.1751 | -. 3713 | 4.888 | . 999 | -58.06 |
|  | 1100 | 391.2 | 200 | - | 6.1989 | -. 9083 | 14.281 | . 989 | -50.55 |
|  | 1150 | 225-1 | 75 | Linear | ------*-* | -5.0986 | 101.972 | . 998 | -369.2 |
|  | 11 | 225-2 | 75 | Linear | --........ | -5.0528 | 101.056 | . 999 | -368.2 |
|  |  | 392-1 | 100 | Paralinear | 11.2087 | -2.6179 | 37.388 | . 984 | -165.2 |
|  | $\downarrow$ | 392-2 | 100 |  | 16.4969 | -4.0094 | 56.591 | . 994 | -243.3 |
| MAR-M-200 + Hi | 1100 | 310-4 | 200 |  | 5.7798 | -. 8618 | 14.398 | . 994 | -95.85 |
|  |  | 310-5 |  |  | 6.3588 | -. 8557 | 14.916 | . 974 | -94.95 |
|  |  | 391.3 |  |  | 7.5777 | -1.0607 | 18.185 | . 984 | -115.7 |
|  |  | 391-4 |  |  | 10.9500 | -1.3809 | 24.758 | . 944 | -35.11 |
|  |  | 391-5 |  |  | 6.4031 | $-.8243$ | 14.646 | . 957 | -90.17 |
|  | . | 391-6 | $\downarrow$ |  | 7.1013 | $-.9866$ | 16.967 | . 983 | -107.0 |
|  | 1150 | 225-3 | 100 |  | 4.6373 | -4.2870 | 47.507 | . 999 | -380.3 |
|  |  | 225-4 |  | $\downarrow$ | 5.2434 | -1.0455 | 15.698 | . 982 | -58.81 |
|  |  | 225.5 |  | Linear | --.------ | -3.9051 | 78.102 | . 999 | -385.0 |
|  |  | 225.6 |  | Linear | ----..--- | -4.4577 | 89.153 | . 999 | -439.9 |
|  |  | 392-3 |  | Paralinear | 22.2491 | -5.1485 | 73.734 | . 997 | -295.0 |
|  |  | 392-4 |  |  | 20.9305 | -5.2509 | 73.439 | . 998 | -313.7 |
|  |  | 392-5 |  |  | 23.0734 | -4.6244 | 69.317 | . 985 | -242.8 |
|  | $\checkmark$ | 392-6 | - |  | 21.4160 | -4.6941 | 68.357 | . 993 | -261.4 |
| MAR-M-211 | 1100 | 324-4 | 115 |  | 51.5721 | -9.4149 | 145.721 | . 983 | -524.9 |
|  | 1100 | 473-6 | 200 |  | . 3227 | -. 08759 | 1.199 | . 989 | -14.62 |
|  | 1150 | 321-4 | 100 |  | 1.4392 | $-.39142$ | 5.353 | . 979 | -27.93 |
|  | 1150 | 478-1 | 100 |  | 32.1709 | -7.6849 | 109.019 | . 995 | $-452.8$ |
| MAR-M-246 | 1100 | 325-3 | 200 |  | . 2656 | -. 1288 | 1.653 | . 994 | -24.44 |
| MAR-M-246 | 1150 | 322-3 | 100 |  | 5.0692 | -1.3008 | 18.077 | .975 | -92.89 |
| MAR-M-247 | 1000 | 452-5 | 500 |  | . 0471 | -. 0012 | . 059 | . 991 | +.46 |
|  | 1000 | 480-3 | 500 | $\downarrow$ | . 0343 | -. 0012 | . 046 | . 954 | +.24 |
|  | 1100 | 453-5 | 200 | Linear | ------- | -. 0280 | . 560 | . 993 | -5.30 |
|  | $1$ | 481-3 |  | Paralinear | . 0789 | -. 0320 | . 399 | . 997 | -4.92 |
|  |  | 657.1 |  | Paralinear | . 2228 | $-.0334$ | . 556 | . 998 | -3.50 |
|  |  | 657-2 |  | Paralinear | . 1964 | -,0259 | . 456 | . 994 | -2.52 |
|  | $\downarrow$ | 657-3 | + | Linear | --.---* | -. 0282 | . 564 | . 979 | -4.86 |
|  | 1150 | 454-5 | 100 | Paralinear | . 4067 | -. 2250 | 2.657 | . 996 | -19.46 |
|  |  | 482-3 |  | Paralinear | 1.1464 | -. 4054 | 5.200 | . 995 | -30.86 |
|  |  | 656-2 |  | Paralinear | 2.9041 | -. 6259 | 9.163 | . 995 | -35.68 |
|  | $\downarrow$ | 656-3 | $\cdots$ | Linear | -...... | -. 1459 | 2.919 | . 973 | -14.21 |
| MAR-M-421 | 1100 | 325-1 | 200 | Paralinear | 3.8911 | -. 5640 | 9.531 | . 944 | -74.11 |
| MAR-M-421 | 1160 | 322-1 | 100 |  | 12.0706 | -2.2866 | 34.936 | . 940 | $-128.7$ |
| NASA-TRW-VIA | 1100 | 103-1 | 200 |  | . 2144 | -. 0198 | . 412 | . 988 | -. 94 |
|  | , | 103-2 | 1 |  | . 1933 | -. 0174 | . 367 | . 982 | -. 77 |
|  |  | 103-6 |  |  | . 1118 | -. 0111 | . 223 | . 874 | -. 54 |
|  |  | 190-6 |  |  | . 0528 | -. 0154 | . 207 | . 992 | -2.32 |
|  | $\checkmark$ | 473-4 | $\downarrow$ | $\cdots$ | . 1981 | -. 0258 | . 456 | . 939 | -1.88 |

TABLE A-II.-Coptinued.

| Alloy | Teat temperature, ${ }^{\circ} \mathrm{C}$ | Run number | Test time, hr | Model type | $\mathrm{F}_{1}^{1 / 2}$ | $k_{3}$ | Ka | $R^{2}$ | $\begin{aligned} & \text { Final } \\ & \Delta W / A \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| NASA-TRW-VIA | 1100 | 659-6 | 200 | Paralinear | 0.0991 | -0.0212 | 0.311 | 0.963 | -2.41 |
|  | $1150$ | 41-3 | 100 |  | . 4364 | -. 0838 | 1.274 | . 991 | -3.87 |
|  | 1 | 78-6 |  |  | . 4543 | -. 1233 | 1.687 | . 991 | -8.27 |
|  |  | 101-4 |  |  | . 4176 | $-.0750$ | 1.168 | . 999 | -3.26 |
|  |  | 105-5 |  |  | . 1437 | -. 0564 | . 708 | . 997 | -4.13 |
|  |  | 129-1 |  | Linear <br> Linear Paralipear | ****************) | -. 0367 | . 734 | . 995 | -3.81-3.77 |
|  |  | 129-2 |  |  | -...--- | -. 0382 | . 763 | . 983 |  |
|  |  | 129-3 |  |  | 0.8529 | $-.1894$ | 2.747 | . 996 | -11.01 |
|  |  | 129-4 |  | Paralipear | . 6891 | -. 1357 | 1.946 | . 999 | -7.68 |
|  |  | 129-5 |  |  | . 1439 | -. 0533 | . 677 | . 992 | -4.20-15.81 |
|  |  | 129-6 |  |  | 1.3212 | -. 2894 | 4.215 | . 999 |  |
|  |  | 204-5 |  | Linear | --.---* | $-.0877$ | 1.754 | . 998 | -8.96 |
|  |  | 472-4 |  | Paraliaear | . 2585 | -. 1205 | 1.463 | . 998 | -9.82 |
|  | $\checkmark$ | 658-6 | $\checkmark$ | Linear | ......- | -. 0768 | 1.636 | . 996 | -7.35 |
| Nimonic 115 | 1000 | 675-4 | 500 | Paralinear | . 1930 | -. 0124 | . 317 | . 825 | $-1.47$ |
| Nimonic 115 | 1000 | 675-5 | 500 | Paralinear | . 3220 | -. 0178 | . 500 | . 542 | -4.60 |
| Nimonic 115 | 1150 | 663-6 | 100 | Linear | -----** | -. 0820 | 1.640 | . 962 | -7.24 |
| NX-188 | 1100 | 393-2 | 200 | Paralinear | . 8623 | -. 2386 | 3.248 | . 990 | -39.06 |
|  | 1100 | 413-3 | 200 |  | . 5128 | $-.3131$ | 3.644 | . 997 | -58.45 -37.87 |
|  | 1150 | 102-3 | 100 |  | . 5371 | -.4188 | 4.725 | . 999 | -37.87 |
|  |  | 102-6 | 1 |  | 2.1865 | -. 8314 | 10.500 | . 997 | -61.88 |
| 1 |  | 414-3 |  | Linear Paralinear | 10.2068 | -. 4708 | 9.416 | . 998 | -48.39 -156.4 |
| Rend 41 |  | 100-5 | - |  |  | -2.6283 | 36.480 | . 988 | $-156.4$ |
| Rene 41 |  | 137-3 | , |  | 8.2779 | -2.4160 | 32.438 | . 995 | $-150.9$ |
| Rend 41 | $t$ | 137-6 |  |  | 8.5318 | -2.1973 | 30.505 | . 998 | -130.4 |
| Rend 80 | 1100 | 232-3 | 200 |  | 10.6738 | -2.9182 | 39.856 | . 999 | -426.4 |
|  | 1100 | 659-2 | 200 |  | 13.7574 | -2.1189 | 34.946 | . 993 | -234.3 |
|  | 1150 | 108-3 | 100 |  | $14.0964$ | $-5.2980$ | 67.077 | $.999$ | $-380.0$ |
|  | 1150 | 108-6 | 100 |  | $\begin{array}{r} 12.8470 \\ 6.8785 \end{array}$ | $\begin{aligned} & -5.0996 \\ & -4.4707 \end{aligned}$ | 63.838 | $\begin{aligned} & .999 \\ & .999 \end{aligned}$ | $\begin{aligned} & -373.9 \\ & -370.6 \end{aligned}$ |
|  | 1150 | 658-2 | 100 |  |  |  | 51.386 |  |  |
| René 120 | 1100 | 232-6 | 200 |  | $\begin{aligned} & 6.8785 \\ & 2.9870 \end{aligned}$ | $\begin{array}{r} -4.4707 \\ -.3887 \end{array}$ | 6.854 | . 984 | $-38.57$ |
| Rend 120 | 1150 | 108-4 | 100 |  | 4.6318 | $\begin{array}{r} -.3887 \\ -1.0219 \end{array}$ | $\begin{aligned} & 14.851 \\ & 14.970 \end{aligned}$ | . 996 | -57.63 |
| Rend 120 | 1150 | 108-5 | 100 | $\downarrow$ | $\begin{aligned} & 4.9019 \\ & 1.4998 \end{aligned}$ | $\begin{array}{r} -1.0219 \\ -1.0068 \end{array}$ |  | . 994 | -53.30 |
| René 125 | 1100 | 325-4 | 200 |  |  | -. 1942 | 3.442 | . 959 | -20.97 |
|  | 1100 | 659-3 | 200 | Limear Paralinear | --....- | $-.0190$ | . 380 | $\begin{aligned} & .997 \\ & .967 \end{aligned}$ | -3.92 |
|  | 1100 | 659-3 | 200 |  | 2.1047 | $\begin{aligned} & -.3141 \\ & -.6214 \end{aligned}$ | $\begin{aligned} & 5.246 \\ & 9.304 \end{aligned}$ | $.967$ | $-38.76$ |
|  | 1150 | 322-4 | 100 | Paralinear | 3.0903 |  |  | $.981$ | $\begin{aligned} & -34.69 \\ & -52.21 \end{aligned}$ |
| , | 1150 | 658-4 | 100500 |  | 2.7092 | $-.7554$ | $\begin{array}{r} 9.304 \\ 10.263 \end{array}$ | . 990 | -52.21 -16.78 |
| R-160-SX | 1000 | 615-3 |  |  | . 6394 | $-.0514$ | $\begin{array}{r} 10.263 \\ 1.153 \end{array}$ |  | $-16.78$ |
| $1$ | 1000 | 678-6 | 500 |  | 5.8266 | -. 5024 | 10.850 | . 922 | $-148.1$ |
|  | 1100 | 614-3 | 160 | Linear <br> Paralinear | 3.7768$\qquad$ | -4.1233 | 45.010 | . 999 | -598.4 |
| $\downarrow$ | 1150 | 613-3 | 45 |  |  | -15.7428 | 314.856 | . 993 | $-667.0$ |
| TAZ-8A | 1000 | 471-6 | 500 |  | . 0851 | -. 0012 | . 097 | . 994 | $+1.40$ |
|  | 1100 | 232-2 | 200 |  | . 7243 | -. 0823 | 1.547 | . 955 | +7.40 +2.95 |
|  |  | 324-3 | , |  | . 3521 | -. 0095 | . 447 | . 999 | +2.95 |
|  |  | 413-1 |  |  | . 1172 | -. 0094 | . 211 | . 648 | -. 06 |
|  |  | 413-6 |  |  | . 4530 | -. 0235 | . 688 | . 992 | +1.31 |
|  |  | 469-2 |  |  | . 3932 | -. 0184 | . 678 | . 998 | +1.84 |
|  |  | 473-2 |  |  | . 1203 | -. 0063 | . 184 | . 981 | +.43 |
| $\square$ | $\downarrow$ | 657.8 | $\dagger$ | $t$ | . 0401 | -. 0046 | . 086 | . 748 | $-.19$ |

TABLE A-II.-Continued.


TABLE A-II.-Continued.

| Alloy | $\begin{gathered} \text { Test } \\ \text { temperature, } \\ { }^{\circ} \mathrm{C} \\ \hline \end{gathered}$ | $\begin{gathered} \text { Run } \\ \text { number } \end{gathered}$ | Tent time, hr | Model type | $\mathrm{ki}^{1 / 2}$ | $k_{3}$ | Ka | $\mathbf{R}^{\mathbf{2}}$ | $\begin{gathered} \text { Final } \\ \Delta W / \mathbf{A} \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| U-700 | 1100 | 655-5 | 200 | Paralinear | 7.3772 | -0.9441 | 16.819 | 0.914 | -111.2 |
|  |  | 656-6 | 200 |  | 6.7959 | -1.8583 | 25.379 | . 999 | -271.3 |
|  |  | 679-1 | 200 |  | .... | -. 1278 | 2.656 | . 958 | -32.13 |
|  |  | 679-2 | 100 | Linear | ........ | -. 0938 | 1.876 | . 936 | -19.72 |
|  |  | $680-1$ |  |  | ........ | -. 1230 | 2.460 | . 973 | -23.97 |
|  | $\downarrow$ | 680-2 |  |  | ........ | -. 0854 | 1.708 | . 940 | -15.05 |
|  | 1150 | 321-6 |  |  | .......... | $-.4705$ | . 9.411 | . 941 | -60.56 |
|  |  | 323-6 |  |  |  | -.4260-3.9220 | 8.520 | . 964 | -51.14 |
|  |  | 423-5 |  | Paralinear | 16.4.... |  | 55.680 | . 992 |  |
|  |  | 438-1 |  |  | 2.0961 | -. 6998 | 9.095 | . 958 | -230.7 -58.16 |
|  |  | 438-2 |  |  | 15.2100 | -3.9650 | 54.860 | . 995 | -243.2 |
|  |  | 449-6 |  |  | 15.0393 | -3.0999 | 46.038 | . 970 | -174.8 |
|  |  | 454-1 |  | $\downarrow$ | 2.2690 | -. 6897 | 9.168 | . 938 | -57.77 |
|  |  | 470-5 |  | Linear <br> Linear <br> Paralinear | ....... | -. 3743 | 7.486 | . 960 | -45.27 |
|  |  | 478.6 |  |  | - | -. 5407 | 10.814 | . 926 | -71.45 |
|  |  | 654-4 |  |  | 14.7616 | -3.6417 | 51.179 | . 992 | -217.5 |
|  |  | 654-5 |  |  | 14.1388 | -3.5586 | 40.725 | . 995 | -214.4 |
|  | $\downarrow$ | 654-6 | $t$ |  | 9.3159 | -3.4349 | 43.665 | . 999 | -246.5 |
| U.710 | 1100 | 324-5 |  |  | 11.6597 | -2.2095 | 33.755 | . 997 | -270.2 |
| U. 710 | 1150 | 321-5 | 200 100 |  | 9.4443 | -3.9464 | 48.908 | . 999 | -294.1 |
| U-720 | 1000 | 674-6 | 500 |  | 2.9558 | -. 2821 | 8.777 | . 978 | -77.59 |
|  | 1000 | 675-6 | 500 |  | 3.5686 | -. 3431 | 7.000 | . 973 | $\begin{gathered} -93.57 \\ -313.5 \end{gathered}$ |
|  | 1100 | 655-3 | 200 |  | 9.5565 | $\begin{aligned} & -2.2778 \\ & -3.6764 \end{aligned}$ | $\begin{aligned} & 32.335 \\ & 41.575 \end{aligned}$ | . 999 |  |
| $\downarrow$ | 1150 | 654-3 | 100 |  | 4.8115 |  |  | . 999 | $\begin{aligned} & -313.5 \\ & -313.4 \end{aligned}$ |
| Waspaloy | 1000 | 436-6 | 500 |  | 3.6677 | -. 2862 | $6.530$$5.511$ | . 854 | $\begin{aligned} & -313.4 \\ & -75.73 \end{aligned}$ |
|  | 1000 | 480-6 |  |  | 3.0613 | -. 2450 |  | . 900 | $-57.36$ |
|  | 1000 | 615-5 | 500 500 |  | 1.7020 | $\begin{array}{r} -.1253 \\ -1.9120 \end{array}$ | 2.955 | . 667 | $\begin{gathered} -44.80 \\ -248.5 \end{gathered}$ |
|  | 1100 | 393-5 |  |  | 9.0460 |  | $\begin{array}{r} 28.166 \\ 3.760 \end{array}$ | . 999 |  |
|  |  | 437.6 | 200 |  | 1.6630 | -. 2097 |  | .930.968 | $\begin{array}{r} -248.5 \\ -23.91 \end{array}$ |
|  |  | 473-1 |  |  | 1.7550 | -. 2465 | $\cdot 4.220$ |  | $-30.25$ |
|  |  | 614-5 | 1 |  | 1.0993 | $\begin{array}{r} -.1271 \\ -2.9190 \end{array}$ | 2.371 | . 869 | -14.48-165.2 |
|  | 1150 | 438.6 | 100 |  | 14.1400 |  | $\begin{aligned} & 43.330 \\ & 56.035 \end{aligned}$ | $\begin{array}{r} .985 \\ .989 \end{array}$ |  |
|  |  | 470-2 | 100 |  | 11.7414 | -4.4294 |  |  | $\begin{array}{r} -165.2 \\ -318.9 \end{array}$ |
|  |  | 472-1 |  |  | 4.082 | -. 7967 | $\begin{array}{r} 12.049 \\ 9.660 \end{array}$ | . 992 | -41.06 |
|  |  | 482-6 |  |  | 3.5827 | -. 6077 |  | ,980 | $\begin{gathered} -27.53 \\ -226.7 \end{gathered}$ |
| , | - | 613-5 | 1 |  | 19.2421 | -. 4156 | 23.398 | . 995 |  |
| WAZ-20 | 1100 | 232-5 | 200 | $\square$ | 3.6298 | $\begin{aligned} & -1.0868 \\ & -1.8428 \end{aligned}$ | $\begin{aligned} & 14.498 \\ & 27.794 \end{aligned}$ | . 991 | $-155.5$ |
|  | 1100 | 413.5 | 200100 |  | 9.3657 |  |  | $\begin{array}{r} .999 \\ .999 \end{array}$ | $\begin{aligned} & -240.6 \\ & -568.3 \end{aligned}$ |
|  | 1150 | 102-4 |  | Linear <br> Linear Paralinear | --.---- | $\begin{aligned} & -6.8479 \\ & -5.2145 \end{aligned}$ | $\begin{array}{r} 27.794 \\ 116.958 \end{array}$ |  |  |
|  | 1150 | 102-5 | 100 |  | --....- |  | $\begin{array}{r} 104.291 \\ 46.400 \end{array}$ | .999.999 | $\begin{aligned} & -505.3 \\ & -322.5 \end{aligned}$ |
| $\downarrow$ | 1150 | 414-5 | 100 |  | 6.8340 | -3.9566 |  |  |  |
| MAR-M-509 | 1100 | 310-1 | 200 |  | 9.3085 | $\begin{aligned} & -1.3218 \\ & -1.7972 \end{aligned}$ | 22.524 | . 990 | $-137.1$ |
|  | 1100 | 326-4 | 200 |  | 10.3614 |  | 28.333 | . 999 | -211.2 |
|  | 1150 | 102-1 | 100 |  | 9.6676 | -1.8752 | 28.420 | . 987 | -97.87 |
|  | 1150 | 102-2 |  |  | 17.2012 | -3.3339 | 80.540 | . 981 | -177.5 |
| , | 1150 | 323-4 |  |  | 21.8875 | -4.8477 | 70.364 | . 996 | -265.2 |
| W1.52 | 1093 | 120-1 |  |  | 9.7830 | -4.3592 | 53.375 | . 999 | -327.5 |
|  | 1093 | 120-2 |  | $\checkmark$ | 2.3512 | -3.8348 | 40.699 | . 998 | -346.4 |
|  | 1100 | 393-3 | 200 | Linear | .....-- | -3.7798 | 75.596 | . 998 | -579.6 |
| $\downarrow$ | 1100 | 469-3 | 200 | Paralinear | 14.5889 | -. 0392 | 14.981 | . 989 | -559.0 |

TABLE A-II.-Concluded.

| Alloy | Test temperature, ${ }^{\circ} \mathrm{C}$ | $\begin{gathered} \text { Run } \\ \text { number } \end{gathered}$ | Test time, hr | Model type | $k_{1}{ }^{1 / 2}$ | $\mathbf{k}_{2}$ | Ka | $\mathrm{R}^{\mathbf{2}}$ | $\begin{gathered} \text { Final } \\ \Delta W / A \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| W-152 | 1150 | 99.1 | 100 | Linear | ....... | -6.2106 | 124.212 | 0.999 | -608.2 |
|  |  | 99.2 |  | Linear | --.-..- | -6.8008 | 136.016 | . 999 | -663.3 |
|  |  | 105-4 |  | Paralinear | 9.9264 | -7.6869 | 86.795 | . 999 | -650.6 |
|  |  | 105-5 | , | Paralinear | 8.8308 | -7.2813 | 81.644 | . 999 | -623.7 |
|  |  | 128-4 | 75 | Linear | --.-..- | $-5.8658$ | 117.317 | . 982 | -387.4 |
|  |  | 128-5 | 75 | Linear | ...... | -5.8844 | 117.689 | . 993 | -419.0 |
|  |  | 470-3 | 45 | Linear | ..... | -9.0172 | 180.344 | . 999 | -405.7 |
| X-40 | $\begin{aligned} & 1100 \\ & 1150 \end{aligned}$ | 393-4 | 200 | Paralinear | $\begin{aligned} & 15.2770 \\ & 15.5662 \end{aligned}$ | -2.0293 | 35.570 | .971.994 | $\begin{array}{r} -206.3 \\ -186.0 \end{array}$ |
|  |  | 95-4 | 100 |  |  | $\begin{array}{r} -2.3280 \\ -.3388 \end{array}$ | $\begin{array}{r} 38.846 \\ 5.106 \end{array}$ |  |  |
|  |  | 95-5 |  |  | 1.7174 |  |  | $\begin{array}{r} .816 \\ .983 \end{array}$ | $\begin{gathered} -186.0 \\ -25.44 \end{gathered}$ |
|  |  | 105-3 |  |  | 11.4739 | -2.2589 | $\begin{array}{r} 5.106 \\ 34.063 \end{array}$ |  | $\begin{gathered} -25.44 \\ -121.8 \end{gathered}$ |
|  |  | 105-6 | $\checkmark$ |  | 10.8964 | -2.1528 | 32.424 | . 989 | $\begin{gathered} -113.6 \\ -42.54 \end{gathered}$ |
|  |  | 128-3 | 45 |  | 5.5343 | -1.6776 | 22.310 | . 950 |  |
|  | , | 128-6 | 100 | , | 15.5855 | $\begin{aligned} & -3.4410 \\ & -3.5114 \end{aligned}$ | $\begin{aligned} & 49.995 \\ & 50.303 \end{aligned}$ | $\begin{array}{r} .995 \\ .996 \end{array}$ | $\begin{gathered} -42.54 \\ -188.6 \end{gathered}$ |
|  | $\downarrow$ | 146-3 | 100 | $\downarrow$ | 15.1885 |  |  |  | $\begin{aligned} & -188.6 \\ & -197.9 \end{aligned}$ |

TABLE C-I-MULTIPLE REGRESSION" RESULTS FOR LOG ${ }_{10} \mathrm{Ka}$ AS A FUNCTION OF ALLOY COMPOSITION IN wt , AND OF ABSOLUTE TEST TEMPERATURE IN $1 / T_{k}$ BASED ON AN INITIAL SELECTION OF $151^{\text {"* }}$ ORDER VARIABLES.

NUMBER OF DATA VALUES $n=315$.

| $\left[Z_{i}=15, z_{\text {f }}=11\right]$ |  |  |
| :---: | :---: | :---: |
| Significant tē̈ms, Z | Coefficient | t-statistic |
| Ta | -0.15488235 | -9.666 |
| $1 / \mathrm{T}_{\mathrm{K}}$ | -17305.08365 | -16.606 |
| $\mathrm{Al}^{\mathbf{K}}$ | -0.33925047 | .7.333 |
| Cr | -0.08308178 | -4.459 |
| Ti | +0.26407575 | 11.464 |
| Nb | +0.24172264 | 4.789 |
| C | +1.99987840 | 5.810 |
| Re | +0.87295039 | 9.593 |
| Zr | +0.37654324 | 2.416 |
| Mo | +0.04526628 | 2.995 |
| Hf | +0.17781791 | 2.309 |
| $\mathrm{a}_{0}$, intercept | 14.77171564 |  |
| $\mathrm{R}^{2}=80.04 \%$ | E.E. $=0.396669$ | $Z_{i}=15$ |

$\mathrm{Co} / \mathrm{Ni}, \mathrm{Cr}, \mathrm{Al}, \mathrm{Ti}, \mathrm{Mo}, \mathrm{W}, \mathrm{Nb}, \mathrm{Ta}, \mathrm{C}, \mathrm{B}, \mathrm{Zr}, \mathrm{Hf}, \mathrm{V}, \mathrm{Re}, 1 / \mathrm{T}_{\mathrm{k}}$
"Stepwise regression-variables are added one at a time starting with the most significant, the F -otatistic for a variable must be significant to 0.15. Alter a variable is added, however, the stepwise method looks at all the variablen already in the model and deletea any that does not produce an F-statistic significant to the 0.15 level.

TABLE C-II-ANALYSIS OF VARIANCE (ANOVA) SUMMARY
FOR $n=315$ DATA SET; $\mathrm{E}_{\mathrm{f}}=11$ SHOWING SOURCES
OF VARIATION INCLUDING LACK OF FIT
OF THE ESTIMATING EQUATION

| Source | Degrees of freedom, d.f. | Sum of squares | Mean squarea |
| :---: | :---: | :---: | :---: |
| Model | 11 | 191.19376 | 17.3812505 |
| Residual | 303 | 47.67592788 | 0.1573463 |
| Lack of fit | (70) | (20.35603) $\}$ | (0.29080575) |
| Replication | (233) | (27.319525) | (0.11725118) |
| Total | 314 | 238.86968 |  |

$$
F-\text { ratio }=\frac{M S(\text { LOF })}{M S(\text { REPS })}=\frac{0.29080675}{0.11725118}=2.480^{\circ}
$$

[^6]

Figure 1.-Observed oxidation attack parameters - Ka's for Group I alumina/aluminate scale alloy formers tested at 1000,1100 and $1150{ }^{\circ} \mathrm{C}$ respectively (multiple horizontal lines indicate replicates).


Figure 1.-Continued.



Figure 2.-Observed oxidation attack parameters - Ka's for Group 11 chromia/chromite or NiO scale alloy formers tested at 1000, 1100 and $1150^{\circ} \mathrm{C}$ respectively (multiple horizontal lines indicate replicates).


Figure 2.-Continued.


Figure 2.-Concluded.


Figure 3.-Comparison of the average observed and the predicted oxidation attack parameters, Ka's, for Group I alumina/aluminate scale alloy formers at 1000, 1100, and $1150^{\circ} \mathrm{C}$ respectively.




Figure 4.-Comparison of the average observed and the predicted oxidation attack parameters, Ka's, for Group II chromia/chromite or NiO scale alloy formers at 1000,1100 and $1150^{\circ} \mathrm{C}$ respectively.




Figure 5.-Standardized residual values va. 315 predicted log Ka values derived from a 14 term regression estimating equation involving alloy composition and temperature for 36 high strength Ni - and Co - base superalloys.


Figure 6.-Predicted log Ka values derived from a 14 term estimating equation involving alloy composition and temperature va. observed log Ka values for 36 high strength Ni - and Co-base superalloys (the straight lines on thls plot represent a simple linear regression fit of this data with $\pm 2.5$ standard deviation limits).


Figure B-1.-B-1900, $1150^{\circ} \mathrm{C}$, run 78-1.


Figure B-3.-W-52, $1150^{\circ} \mathrm{C}$, run 99-1.


Figure B-5.-WI-52, $1150^{\circ} \mathrm{C}$, run 128-4.


Figure B-2.-IN-100, $1150^{\circ} \mathrm{C}$, nun 95-6.


Figure B-4.-B-1900, $1150{ }^{\circ} \mathrm{C}$, run 123-3.


Figure B-6.-NASA-TRW-VIA, $1150^{\circ} \mathrm{C}$, run 129-2.


Figure B-7.-René-41, $1150^{\circ} \mathrm{C}$, run 137-3.


Figure B-9.-B-1900 $+\mathrm{Ht}, 1150^{\circ} \mathrm{C}$, run 204-3.


Figure B-11.-TAZ-8A, $1100^{\circ} \mathrm{C}$. nun 232-2.


Figure B-8.-René-41, $1150^{\circ} \mathrm{C}$, run 137-6.


Figure B-10.-B-1900, $1150^{\circ} \mathrm{C}$, run 221-1.


Figure B-12.-WAZ-20, $1100^{\circ} \mathrm{C}$, run 232-5.


Figure B-13.-Alloy $625,1150^{\circ} \mathrm{C}$, run 352-4.


Figure B-15.-U-700, $1100^{\circ} \mathrm{C}$, run 422-5.


Figure B-17.-U-700, $1000^{\circ} \mathrm{C}$, run 424-5.


Figure B-14.-Alloy $625,1150^{\circ} \mathrm{C}$, run 352-4.


Figure B-16.-U-700, $1150^{\circ} \mathrm{C}$, run 423-5.


Figure B-18.-IN-792, $1150^{\circ} \mathrm{C}$, run 428-4.


Figure B-19.-Waspaloy, $1000^{\circ} \mathrm{C}$, nun 436-6.


Figure B-21.-U-700, $1100^{\circ} \mathrm{C}$, run 448-6.


Figure B-20.-U-700, $1000^{\circ} \mathrm{C}$, run 447-6.


Figure B-22.-U-700, $1150^{\circ} \mathrm{C}$, nun 449-6.


Figure B-23.-MAR-M-247, $1150^{\circ} \mathrm{C}$, run 454-5.


Figure B-24.-TAZ-8A, $1150^{\circ} \mathrm{C}$, run 472-2.


Figure B-25.-MAR-M-211, $1150^{\circ} \mathrm{C}$, run 478-1.


Figure B-27.-IN-792, $1100^{\circ} \mathrm{C}$, run 657-5.


Figure B-26.-TAZ-8A, $1150^{\circ} \mathrm{C}$, run 656-2.


Figure B-28.-René-125, $1100^{\circ} \mathrm{C}$, run 659-3.



[^0]:    ${ }^{1}$ The SAS statistical computer package (version 5) for the VM main frame operating system was used for all data analysis in this study.

[^1]:    ${ }^{2}$ The average $K_{a}$ 's are defined as the antilog of the average of the $\log K_{a}$ values for each alloy at each temperature.

[^2]:    ${ }^{1}$ Included also are 28 runs not listed in references 1 and 2, but plotted in the Appednix B of this report.

[^3]:    

    | 3 |
    | :--- |
    | 3 |
    | 8 |
    | 1 |
    | 3 |
    | 3 |
    | 3 |
    | 3 |

[^4]:    

[^5]:    ${ }^{2}$ Based on the derived estimating equation, see table II.
    b $(\log \mathrm{Ka}$ observed $-\log \mathrm{Ka}$ predicted)
    Standard deviation

[^6]:    ${ }^{\text {a }}$ The lack of fit term appears to be significant since the F-ratio for $(1-a)$ where $a=0.95=1.658$ which does not exceed the MS(LOF)/MS(REPS) ratio derived for this first order model. Therefore this model is not considered satisfactory.

