

1     **Transfer of Satellite Rainfall Uncertainty from Gauged to**  
2     **Ungauged Regions at Regional and Seasonal Timescales**

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## ABSTRACT

Hydrologists and other users need to know the uncertainty of the satellite rainfall data sets across the range of time/space scales over the whole domain of the data set. Here, ‘uncertainty’ refers to the general concept of the ‘deviation’ of an estimate from the reference (or ground truth) where the deviation may be defined in multiple ways. This uncertainty information can provide insight to the user on the realistic limits of utility, such as hydrologic predictability, that can be achieved with these satellite rainfall data sets. However, satellite rainfall uncertainty estimation requires ground validation (GV) precipitation data. On the other hand, satellite data will be most useful over regions that lack GV data, for example developing countries. This paper addresses the open issues for developing an appropriate uncertainty transfer scheme that can routinely estimate various uncertainty metrics across the globe by leveraging a combination of spatially-dense GV data and temporally sparse surrogate (or proxy) GV data, such as the Tropical Rainfall Measuring Mission (TRMM) Precipitation Radar and the Global Precipitation Measurement (GPM) mission Dual-Frequency Precipitation Radar. The TRMM Multi-satellite Precipitation Analysis (TMPA) products over the US spanning a record of 6 years are used as a representative example of satellite rainfall. It is shown that there exists a quantifiable spatial structure in the uncertainty of satellite data for spatial interpolation. Probabilistic analysis of sampling offered by the existing constellation of passive microwave sensors indicate that transfer of uncertainty for hydrologic applications may be effective at daily time scales or higher during the GPM era. Finally, a commonly used spatial interpolation technique (kriging), that leverages the spatial correlation of estimation uncertainty, is assessed at climatologic, seasonal, monthly and weekly timescales. It is found that the effectiveness of kriging is sensitive to the type of uncertainty metric, time scale of transfer and the density of GV data within the transfer domain. Transfer accuracy is lowest at weekly timescales with the error doubling from monthly to weekly.

59    However, at very low GV data density ( $<20\%$  of the domain), the transfer accuracy is too low to  
60    show any distinction as a function of the timescale of transfer.

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62    **Keywords:** Satellite precipitation, uncertainty, transfer, spatial interpolation, GPM.

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## 1.0 INTRODUCTION

Precipitation is arguably one of the most important components of the water cycle over land. One study shows that almost 70-80% of the variability in the terrestrial water cycle can be explained from the spatio-temporal variability observed in precipitation over land (Syed et al., 2004). Existing missions such as the Tropical Rainfall Measurement Mission (TRMM) provide vital precipitation information for water cycle studies (Huffman et al., 2007). Furthermore, planned missions such as the Global Precipitation Measurement (GPM) mission will provide a global hydrologic remote sensing observatory to advance the use of precipitation sensing technologies in scientific inquiry into hydrologic processes (Krajewski et al., 2006). With the global and more frequent precipitation observational capability planned for GPM, such precipitation measuring satellite missions permit us to refine knowledge from physical and hydrologic models that can then be converted to local and global strategies for water resources management (Voisin et al., 2008; Hossain et al., 2007). *[Hereafter, because our focus is on liquid precipitation, the term 'rainfall' will be used as shorthand for 'precipitation' for convenience]*

However, a crucial challenge in advancing satellite rainfall-based surface hydrologic prediction, is the need to bridge the scale incongruity between overland hydrologic processes that evolve at small scales (i.e.,  $< 1$  hour and  $< 5$  km) and operational satellite precipitation datasets that will always be restricted to coarser scales from passive microwave sensors (i.e.,  $> 1$  hour and  $> 5$  km; Hossain and Lettenmaier, 2006). There are two paths that have historically been followed as a response to this scale incongruity: 1) apply satellite rainfall data available at the native scale for hydrologic prediction (e.g., Harris and Hossain, 2008; Su et al., 2008; Voisin et al., 2008); and 2) apply spatial and spatio-temporal disaggregation (or downscaling) techniques



to resolve satellite rainfall data at the required smaller space-time scales for hydrologic prediction (e.g., Forman et al., 2009; Bindlish and Barros, 2000). Each option leads to non-negligible uncertainty in hydrologic simulation. In the first case, the major source of this uncertainty is due to the algorithmic and sampling uncertainty (for passive microwave-PMW sensors) of satellite rainfall data at the native scale. In the second case, the primary source of uncertainty is due to the statistical disaggregation technique that further propagates the native scale uncertainty to sub-grid uncertainty in ways that are not well understood (see for example, Rahman et al., 2009). Either way, hydrologists and other users, need to know the uncertainty of the satellite rainfall data sets across the range of time/space scales over the whole domain of the data set. This uncertainty can provide insightful information to the user on the realistic limits of utility that can be achieved with satellite rainfall data sets, for example for hydrologic predictability (Hong et al., 2006), on which we will focus in this paper. While representing the uncertainty structure of satellite rainfall as a function of scale against quality-controlled ground validation datasets remains a critical research problem for GPM, therein lies a paradox. Satellite rainfall uncertainty estimation requires ground validation (GV) precipitation data. On the other hand, satellite data will be most useful over ungauged regions in the developing world (Tang and Hossain, 2009).

*In-situ* rainfall information from rain gauge networks is generally considered the standard choice for GV data (Villarini and Krajewski, 2007; Habib et al., 2004; McCollum et al., 2002). Such data is often referred to as ‘reference’ or ‘truth’. However, *in-situ* gauges are point measurements and unless there exists a dense network to adequately capture the space-time variability of rainfall process, its use for validating areal-averaged satellite rainfall data for surface hydrologic processes remains questionable (Ciach and Krajewski, 1999). The work of

Gebremichael et al. (2003) clearly demonstrates the sensitivity of satellite rainfall uncertainty estimation to gauge density. Thus, in most regions across the globe without adequate *in-situ* rain gauge coverage, the uncertainty associated with satellite data have been parameterized to sampling configuration of the sensors at this stage (Li et al., 1998; Huffman, 1997). Some examples of this parameterization are the Global Precipitation Climatology Project (GPCP; Huffman, 2005; Huffman et al., 1997) dataset and the TRMM Multi-satellite Precipitation Algorithm (TMPA; Huffman et al., 2007) that now provide an estimate of the Root Mean-Squared Uncertainty (RMSE) of the satellite rainfall estimates on the basis of sampling pattern and the period of rainfall accumulation of interest to the user.

While such parameterized methodologies for estimating uncertainty have been useful in providing users with a level of confidence associated with satellite rainfall estimates, such uncertainty is essentially a standard deviation measure of sampling uncertainty. Many of these uncertainty methodologies are based on the conceptual argument that uncertainty (i.e, standard deviation,  $\sigma_E$ ) can be related directly or inversely to observation interval ( $\Delta t$ ), observation period (T), spatial averaging area (A), and rain rate (R):

$$\sigma_E = f\left(\frac{1}{R}, \frac{1}{A}, \frac{\Delta t}{T}, parameters\right) \quad (1)$$

as expressed by Steiner et al. (2003), among others. In many cases, the functional form of this ‘predicted’ uncertainty is not benchmarked to the realities of the ground observations and hence may not provide a reasonable assessment in indicating the expected reliability for water cycle studies (Gebremichael et al., 2010). Recently, several other parameterized methodologies have evolved based on data assimilation approaches (e.g. Kalman filtering in GsMAP satellite product

of Ushio et al., 2009). In these approaches, an estimate of uncertainty that is available is essentially related to the methodology of the filtering technique and does not necessarily indicate the actual level of agreement with GV rainfall data. In some instances, however, the uncertainty is estimated by comparing the output of a wide-coverage technique (such as infrared-IR advected PMW) to a more localised but higher accuracy product (such as PMW only; Ushio et al., 2009).

There now exists a sufficient body of knowledge on uncertainty metrics and models that we should consider a transition to a more hydrologically-relevant framework in anticipation of the satellite data-rich scenario of GPM. Although existing uncertainty metrics and uncertainty models represent an important first step, most treat uncertainty as a single measure representative for a large space and time domain. This uni-dimensional uncertainty measure is invariably the standard deviation of uncertainty (e.g. Eqn. 1). However, a satellite rainfall product with an uncertainty standard deviation ( $\sigma_E$ ) of X mm/hr over a large space-time domain can be represented by a multiplicity of distinct spatio-temporal patterns of rainfall, each having a distinct response in surface hydrology (see for example, Lee and Anagnostou, 2004).

What is therefore needed now for advancing the hydrological application of GPM is a practical methodology that can routinely ‘transfer’ a set of hydrologically-relevant uncertainty metrics from locations/regions having GV-based values to ungauged regions for improving water cycle studies or water resources management. Here, ‘transfer’ is akin to spatial interpolation at non-sampled locations (grid boxes) using measurements from sampled but sparse locations (grid boxes). Figure 1 provides a conceptual rendition of this idea of ‘transfer’ of uncertainty based on the concept of spatial interpolation (*taken from Ling and Hossain, 2009*).

This paper analyzes the open issues for developing an appropriate uncertainty transfer scheme that can routinely estimate various uncertainty metrics across the globe by leveraging a

combination of spatially-dense GV data and temporally sparse surrogate (or proxy) GV data from sources such as the TRMM-like PR sensor anticipated during the GPM era. The TRMM Multi-satellite Precipitation Analysis (TMPA) products 3B42RT and 3B41RT (Huffman et al., 2007) over the US spanning a record of 6 years are used as a representative example of satellite rainfall. The paper presents a probabilistic analysis of sampling offered by the existing constellation of precipitation-relevant satellite PMW sensors in order to understand the current and expected spatial coverage during the GPM era. A commonly used spatial interpolation technique (kriging), that leverages the spatial correlation of rainfall estimation uncertainty, is then investigated for its effectiveness. This effectiveness is cast in the context of the expected sparseness in GV data expected from TRMM and GPM missions. Finally, important issues needing closure are summarized on the basis of our investigation of transfer of satellite rainfall uncertainty from GV to non-GV regions. To avoid confusion among readers, hereafter, the terms ‘uncertainty’ or ‘uncertainty metric’ will be used to define the quality indices of the satellite rainfall estimate derived at GV locations (such as bias, root mean squared error, probability of detection). The terms ‘error’ or ‘transfer error’ will be used specifically to define the quality of the transfer (spatial interpolation) process of uncertainty metrics at non-GV locations.

## **2.0 SPATIAL CORRELATION OF SATELLITE RAINFALL UNCERTAINTY**

The very first requirement for an effective transfer (spatial interpolation) scheme is the presence of a quantifiable spatial structure (or spatial correlation) in the variable being transferred. Therefore, we first investigated the presence of spatial correlation of satellite rainfall uncertainty. First, in order to minimize the error of the GV rainfall data, we used the National Center for Environmental Prediction’s (NCEP) 4 km Stage IV NEXRAD rainfall data that is adjusted to precipitation gages and conveniently available as a quality-controlled data mosaic

over the U.S. (Lin and Mitchell, 2005; Fulton et al., 1998). TMPA's near real-time satellite rainfall data-products from PMW-calibrated Infrared (IR) and merged PMW-IR estimates (labeled 3B41RT and 3B42RT, respectively; Huffman et al., 2007) were used as the satellite rainfall data. The data for GV and satellite rainfall data spanned 6 years from 2002 to 2007. The NEXRAD Stage IV GV rainfall data were first remapped to 0.25° 3-hourly resolution for consistency with the native scale of the satellite rainfall products. 3B41RT data were also remapped at the 3-hourly time scale. After a thorough quality assessment and quality control (QA/QC), the datasets were organized by season and various regions for the years 2002-2007.

In order to study how the satellite rainfall uncertainty is spatially dependent (or correlated), Tang and Hossain (2009) derived the spatial correlograms for each uncertainty metric using the TMPA dataset described above. Herein, the correlation length (CL), where the autocorrelation dropped to 1/e (e-folding distance), was first computed. Next, the empirical semi-variograms were derived and then idealized as exponential semi-variogram functions,

$$\gamma(h) = c_0 + c(1 - e^{-h/a}) \quad (2)$$

where  $\gamma(h)$  is the semi-variance at spatial lag 'h',  $c_0$  represents the nugget variance (i.e., the minimum variability observed or the 'noise' level at a separation distance of 0);  $c$  is the sill variance (when spatial lag is infinite); and  $a$  is the correlation length. Figure 2 provides a summary of the 'climatologic' correlation length (e-folding distance) by season for various uncertainty metrics of the satellite rainfall products such as Probability of Detection (POD) for rain, POD for no rain, false alarm ratio (FAR), root mean squared error (RMSE), and bias. Herein, 'climatologic' refers to the mean error derived from the entire 6 year of data. Appendix one provides the mathematical formulation for the error metrics.

Figure 2 clearly demonstrates that, at the climatologic (long-term) time scale, satellite rainfall uncertainty can have distinct spatial organization that can be leveraged for spatial interpolation. The correlation lengths for a given uncertainty metric as a function of season appear to be at least 3-5 (0.25°) TMPA grid boxes long. As a rule of thumb, this indicates that the transfer of error from sampled locations may be effective up to 4 grid boxes (~ 100 km) away. Another interesting feature that is revealed in this figure is the significantly higher correlation lengths (and spatial organization) observed for 3B41RT than 3B42RT. This can be traced to the sources of the specific satellite estimates: 3B41RT is uniformly computed using a calibration of infrared (IR) brightness temperatures to a combined PMW estimate. The statistics of the uncertainty are spatially very homogenous since they originate from a single probability distribution at regional scales. On the other hand, 3B42RT uses a variety of PMW rainfall estimates with gaps filled during a 3-hour sampling period with the 3B41RT estimate ‘as is’. This fill-in causes the 3B42RT data to draw on two different probability distributions in space for uncertainty statistics (IR and PMW); the increased spatial heterogeneity in the uncertainty structure leads to shorter correlation length. This analysis shows that any uncertainty transfer scheme should benefit from improvements in the 3B42RT product to make it statistically more homogenous in space.

### **3.0 SPATIAL COVERAGE OFFERED BY CURRENT CONSTELLATION OF PMW SENSORS**

Having observed a distinct spatial organization of uncertainty, we also need to understand the space/time dimension that is implicit in the concept of real-time uncertainty ‘transfer’ over non-GV regions. The space dimension pertains to the regions with spatially sparse GV data due to inadequate *in-situ* gauge data (such as that shown in Figure 1), which are, of course, recorded

at fixed positions. The time dimension pertains to the temporally sparse case of using the most accurate rainfall source currently available from space, such as the orbiting TRMM PR as ‘proxy’-GV data, over regions where there is no ground-based GV data. Depending on how we define GV data, there can be several types of GV ‘voids’ where uncertainty information will be need to be estimated for GPM. For example, if we rely on the ‘conventional’ ground source for GV data, voids will be represented by large and stationary regions having little or no instrumentation. On the other hand, if a ‘proxy’ for GV is defined from orbiting sensors, such as the TRMM PR, or even a highly accurate PMW sensor, then voids will be numerous grid boxes dynamically changing in location with each satellite orbit.

The left panel of Figure 3 shows the probability of a 3B42RT grid box ( $0.25^\circ$ ) having a conical-scanning PMW overpass (comprising either TMI, SSMI or AMSR) in 3, 6 and 24 hour windows. On the right panels of Figure 3, the probability of a 3B42RT grid box having a TRMM Microwave Imager (TMI) scan is shown. These probability maps were created using a 100 day period of any PMW sensor) from the 2007- 2008 period. It is clear from the maps that the spatio-temporal dynamics of the location of PMW scans is strongly sensitive to the accumulation periods of hydrologic relevance and one that must be investigated carefully in order to identify how an uncertainty transfer scheme may work using proxy-GV data.

At time scales of 3-6 hours, there are vast regions lacking conventional surface GV data in the tropics of Africa, Asia and South America where the probability of having a PMW scan is less than 50% (Figure 3). This makes the estimation of uncertainty through transfer from GV regions at these locations more important for hydrologic applications. While GPM may improve the coverage of PMW scans, such large voids with a low probability for a PMW scan will still remain over these regions due the continued dependence on polar-orbiting sensors. Since gauge-



based GV is sparse for these tropical regions at hydrologic scales, proxy GV data from space-borne sensors (such as that expected from the GPM Dual Frequency Precipitation Radar) may be one of the few ways to explore if ‘transfer’ of uncertainty is realistic. For the higher latitude land regions (which comprise mostly the industrialized world with reasonably gauged fixed-location GV instrumentation), the uncertainty could be transferred from the stationary GV regions. The right panels of Figure 3 also show that the probability of having a TMI scan also happens to be lowest (0.1-0.2 in 3 hours) over the tropical regions. However, over a 24 hour time period there is considerably higher probability of having such a scan (~0.7-0.8). This implies that the practicable timescale for transferring uncertainty metrics over the tropics from a sun-asynchronous PMW sensor is at least 24 hours.

#### 4.0 TRANSFER OF UNCERTAINTY BY SPATIAL INTERPOLATION

Tang and Hossain (2009) recently showed that most uncertainty metrics (such as bias and POD) are amenable to ‘transfer’ from gauged to ungauged locations using spatial interpolation at climatologic (six-year average) timescales. The method of ordinary kriging (OK) was used for testing the ‘transfer’ of uncertainty metrics. OK is the most common (and one of the simplest) spatial interpolation estimator used to find the best linear unbiased estimate of a second-order stationary random field with an unknown constant mean as follows:

$$\hat{Z}(x_0) = \sum_{i=1}^n \lambda_i Z(x_i) \quad (3)$$

where  $\hat{Z}(x_0)$  = kriging estimate at location  $x_0$ ;  $Z(x_i)$  = sampled value at location  $x_i$ ;  $\lambda_i$  = weighting factor for  $Z(x_i)$  (summing to one over all  $i$ ), and  $n$  is the number of sampled (known) locations. Kriging methods have already been used for spatial interpolation of precipitation from point gauge data with considerable success (see for example, Seo et al., 1990; Krajewski, 1987).



Using the same six-year database of high resolution satellite rainfall data from TMPA over the central US, the OK method was applied to assess the effectiveness of transfer of uncertainty metrics from GV to non-GV grid boxes, using correlation as the main assessment metric. Assuming that only 50% of the region (i.e., grid boxes) was gauged (i.e., having access to GV data), OK was implemented to estimate uncertainty metrics at the other 50% of the (non-GV) region. Selection of ‘GV’ grid boxes was random and hence each kriging realization was repeated 10 times in a Monte Carlo (MC) fashion to derive an average scenario. The semi-variogram and correlation lengths were computed on the basis of the 50% of the assumed ‘available’ data. Spatial correlograms for each uncertainty metric were derived and the correlation length (CL), where the autocorrelation dropped to  $1/e$  (e-folding distance), was computed. The empirical semi-variograms were derived and then idealized as exponential semi-variogram functions.

Tang and Hossain (2009) showed that the transfer of uncertainty metrics using kriging did not lead to wholesale changes in the pattern of the uncertainty field when compared to the true climatologic uncertainty field (see upper left and upper right panels of Figure 4a). Overall, their assessment indicated that ‘transfer’ uncertainty metrics from a gauged to an ungauged location through spatial interpolation has merit for selected uncertainty metrics. In Figure 4b, the histograms for ‘kriging error’ and actual uncertainty (over ungauged grid boxes) demonstrate the accuracy of the transfer method for FAR. Here, the ‘kriging error’ refers to the difference between ‘kriged uncertainty’ and ‘actual uncertainty’, whereas the ‘actual uncertainty’ is the ‘measured uncertainty’. In other words, the ‘kriging error’ is the estimation uncertainty while the ‘actual uncertainty’ is the true dataset uncertainty. The actual GV-based uncertainty (i.e., FAR in this case) is shown in pink while the black line represents the histogram for kriging error. The

histograms for kriging error are considerably lower, by almost an order of magnitude, compared to the actual GV-based uncertainty and are almost unbiased.

However, a point to note is that the work of Tang and Hossain (2009) demonstrated the utility of transfer only at the climatologic time scales with a high degree of GV coverage (50%). Also, at the climatologic scales, the spatial structure of uncertainty can be expected to be well defined and reasonably homogenous (longer correlation lengths of uncertainty that lead to high accuracy for kriging; see Figure 2). Furthermore, the use of the correlation measure may not necessary reflect the most rigorous assessment of accuracy for the transfer of error metrics. For example, there may be high correlation even with large systematic bias in the ‘kriged’ error metric at non-GV grid boxes. In this study, we therefore explored the effectiveness of kriging at seasonal (and lower) timescales and modeled how the effectiveness of transfer is impacted by GV data coverage. We also assessed the accuracy of transfer using marginal and non-correlation type measures.

Figure 5 shows how GV coverage (as randomly located grid boxes over a region) impacts the accuracy of kriging-based transfer of uncertainty over the grid boxes lacking GV data for two different time scales (climatologic and seasonal in the left and right panels, respectively). The exercise was performed in a manner similar to Tang and Hossain (2009) over the central United States. The GV coverage was systematically varied from 10% to 90% and the effectiveness of kriging of uncertainty metric at locations lacking GV data was assessed using the correlation measure with in-situ (sampled) uncertainty metric. For the seasonal case, the summer months of June-July-August in 2007 over the central US was chosen as an example and one seasonal variogram was modeled.

The most striking feature of the GV-density study is that the effectiveness of an uncertainty transfer scheme, specifically kriging in this example, worsens considerably at low GV coverage (correlation dropping to under 0.7) as time scales shorten. Qualitatively, this result is expected, and clearly indicates that if a transfer scheme for estimating uncertainty metrics is finer than seasonal scale (ranging from 3-6 hourly to weekly-monthly), the effectiveness for uncertainty transfer would intuitively worsen further with kriging. A similar assessment can be made from Figure 3 on the potential of kriging using dynamically located a sun-asynchronous PMW scans (such as TMI) as proxy-GV data. At 3-6 hours, the probability of a grid box being scanned by a sun-asynchronous TMI ranges from 0.1 to 0.4. In other words, this is equivalent to a large region having a fixed GV coverage of 10-40%. Naturally therefore, the effectiveness of OK over the tropics using proxy-GV data in the GPM era may probably not be any better at timescales shorter than a day.

In order to demonstrate a more rigorous level of accuracy of the interpolation scheme beyond the correlation measure, we reviewed our kriging simulations for two more scenarios: 1) transfer of uncertainty metric at monthly time scales and 2) transfer of uncertainty metric at weekly time scales. For each scenario, we performed a more in-depth assessment for the months of summer and weeks of June of 2007. We used mean error, in place of correlation measure, to assess the accuracy of the transfer, using the following error definitions:

$$\text{Error} = (\text{Interpolated Uncertainty Metric} - \text{Actual Uncertainty Metric}) / \text{Actual Uncertainty Metric} \quad (4)$$

$$\text{Mean Error} = \text{Mean of Error (as defined above in Eqn 4) over non-GV grid boxes} \quad (5)$$

$$\text{Std. Dev of Error} = \text{Standard Deviation of Error (as defined above) over non-GV grid boxes} \quad (6)$$

Tables 1a and 1b summarize the assessment of OK method using mean relative error (Eqn 5) as the main assessment metric for transfer of uncertainty metrics – BIAS, RMSE, POD and FAR. For each uncertainty metric both mean and standard deviation of error of transfer is shown as measures of accuracy and precision, respectively. It is observed that, unlike correlation measure, the mean and standard deviation of error reveal a somewhat different picture on the utility of OK method. The error metric BIAS has the lowest accuracy ranging from 50% error (at 10% missing GV gridboxes) to 100% error (at 90% missing GV gridboxes) for monthly time scale. For weekly timescale, the mean error ranges from 80% (at 10% missing GV grid boxes) to 120% (at 90% missing GV grid boxes). On the other hand, POD, followed by RMSE and FAR, have the highest accuracy for transfer of error metrics at both timescales according the mean relative error measure. As expected, the precision of the kriging based transfer scheme degrades at shorter timescales. At very low GV coverage (<20%) the standard deviation of transfer error is high (>100%), indicating poor performance of the OK method regardless of the timescale at which the uncertainty metrics are transferred.

## **5.0 CONCLUSION: THE CURRENT OPEN ISSUES ON UNCERTAINTY TRANSFER**

In light of the impact of GV coverage and the timescale on effectiveness of uncertainty transfer, we now need to consider the following: 1) explore other techniques for transfer that are more sophisticated than ordinary kriging; and 2) understand how we can leverage the methodological error estimate that is routinely available from uncertainty models such as Huffman (1997) and Kalman filtering techniques that many satellite rainfall data algorithms use (such as, Ushio et al. 2009). In our interpolation method, the spatial structure of rainfall has not been used alongside that of estimation uncertainty. Because the two (rainfall and its estimation

uncertainty) are related, it may be worthwhile to pursue co-kriging type conditional interpolation schemes that leverage existing information on the satellite rainfall distribution as an extra constraint.

Also, for spatial interpolation methods, we should keep in mind that traditional geostatistical tools are pattern filling methods based on the spatial correlation exhibited by two points in space separated by a lag  $h$ . The variogram computed using this two-point geostatistical approach may simplify the spatial patterns manifested by the complex precipitation systems and surface emissivities that dictate the accuracy of satellite rainfall products at hydrologic time scales over land. For the case of spatial interpolation of ground water contamination, it has recently been demonstrated that the use of a highly non-linear pattern learning technique in the form of an artificial neural network (ANN) can yield significantly superior results under the same set of constraints when compared to ordinary kriging method (Chowdhury et al., 2009). Thus, the use of non-linear mapping techniques are worth an investigation.

Another aspect to keep in mind is the nature of use of each uncertainty metrics. Different users will have naturally different needs. Hydrologist users engaged in flash flood or monsoonal flood forecasting will probably be more interested in the  $POD_{RAIN}$  (to understand the accuracy in estimating peak flow), FAR (to minimize false alarms in flood warnings), and BIAS (to minimize under/over estimation in river stage) for each grid box (see Harris and Hossain, 2008 and Hossain and Anagnostou, 2004). Hydrologists engaged in continuous simulation based on soil moisture accounting for drought monitoring and water management would probably focus more on  $POD_{NORAIN}$  (to minimize uncertainty in underestimating the soil wetness and evapotranspiration) for each grid box. On the other hand, crop yield and famine forecasters would like to focus more on the seasonal bias over a large agricultural zone during the growing

season as the important indicator of reliability of a satellite rainfall product (personal communication with Dr. Chris Funk of University of California Santa Barbara).

In summary, developing an uncertainty transfer scheme that is amenable to operational implementation for estimation of uncertainty metrics for satellite rainfall data over regions lacking surface GV data is a necessary requirement for current and future satellite precipitation missions to advance their hydrologic potential. Hydrologist users around the world need to have a clear understanding of the pros and cons of applying satellite rainfall data for terrestrial hydrologic applications at a given scale if the benefit of these missions is to be maximized. One way of facilitating the understanding is through the routine provision of various measures of uncertainty that are of hydrologic relevance. If this uncertainty information is provided alongside the global and more frequent precipitation observational capability planned in GPM, it will permit us to refine knowledge from physical and hydrologic models that can then be converted to local and global strategies for water resources management. Work is currently undergoing to address some of the open issues discussed above and we hope to report them in the near future.

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## APPENDIX: FORMULATION OF UNCERTAINTY METRICS

Consider the 2x2 contingency table of hits and misses associated with satellite rainfall estimates:

TABLE A.1

$$\text{Probability of Detection for Rain (POD}_{\text{RAIN}}\text{): } \frac{N_A}{N_A + N_B} \quad (\text{A.1})$$

$$\text{Probability of Detection for No Rain (POD}_{\text{NORAIN}}\text{): } \frac{N_D}{N_D + N_C} \quad (\text{A.2})$$

$$\text{False Alarm Ratio (FAR): } \frac{N_B}{N_B + N_A} \quad (\text{A.3})$$

The  $\text{POD}_{\text{RAIN}}$  essentially defines how often a satellite rainfall estimate is likely to correctly detect gridboxes as rainy according to the reference or ground validation data. Similarly,  $\text{POD}_{\text{NORAIN}}$  defines how often a satellite rainfall estimate is likely to correctly detect a non-rainy grid box as non-rainy according to the ground validation data.

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**Figure 1.** Conceptual rendition of the idea of ‘transfer’ of uncertainty information from a gaged (GV) location to an ungauged (non-GV) location. Upper panel depicts the notion of ‘uncertainty’ of satellite rainfall data (in this case, the scalar deviation of magnitudes is termed ‘uncertainty’ although there are many other types of uncertainty). Lower panel depicts how the known uncertainty (derived from GV sites shown in black in the middle panel) would be ‘transferred’ to the non-GV (ungauged) sites shown in blue (right most panel). *Reprinted from Tang and Hossain, 2009.*

**Figure 2.** Correlation length of uncertainty metrics at climatologic timescales for 3B41RT (upper panel) and 3B42RT (lower panel) shown as a function of season. Note the distance unit is 0.25 degree grid boxes (~ 25 km). The vertical bars are shown in order from left to right as ‘Bias’, ‘RMSE’, ‘POD rain’, ‘POD no-rain’, ‘FAR’. *(Taken from Tang and Hossain, 2009).*

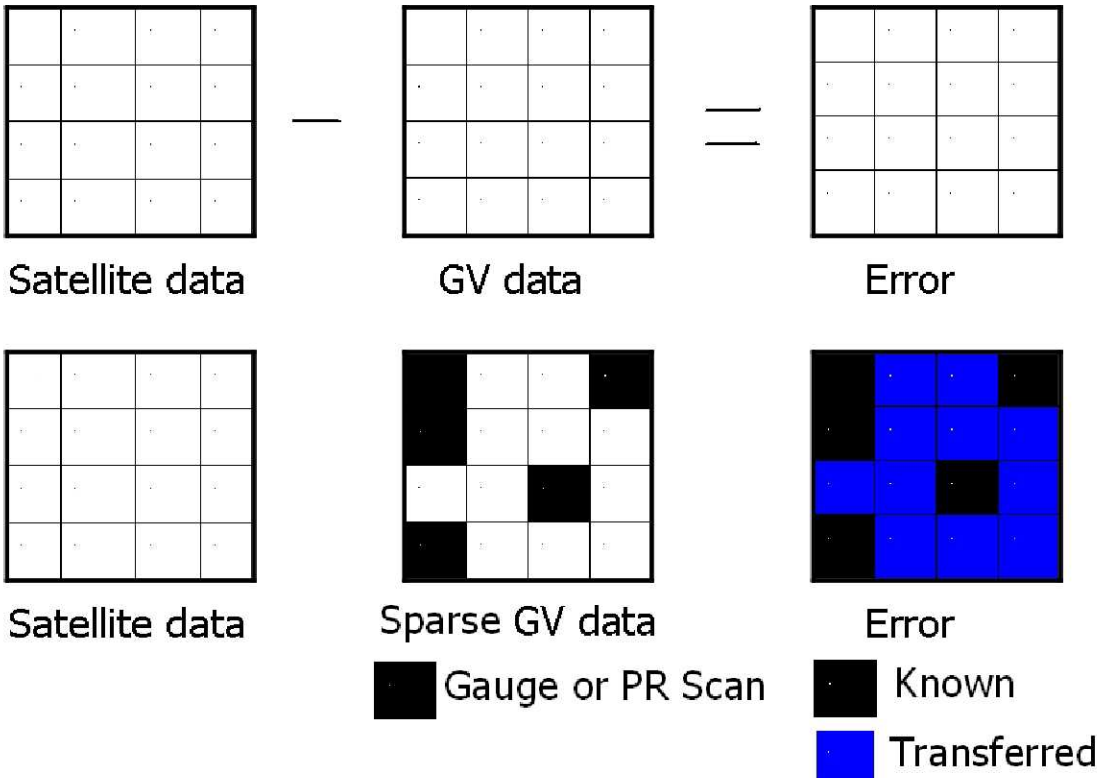
**Figure 3.** Left panels: probability of a 3B42RT 0.25° grid box having a PMW scan from either TMI, SSMIs or AMSR for 3 hour (upper), 6 hour (middle), and 24 hour periods (bottom). Right panels: Same as left panels but only for TMI.

**Figure 4a.** Transfer of Bias of 3B41RT from gauged to ungauged locations. Upper left-most panel shows the true field of uncertainty in bias based on 6 years of data. The lower left-most panel is the randomly selected 50% of the region for computation of the empirical variogram and correlation length. The lower middle panel shows the other 50% of the region that was assumed have no GV. Lower right panel shows the estimation of the bias at the non-GV grid boxes using ordinary kriging.

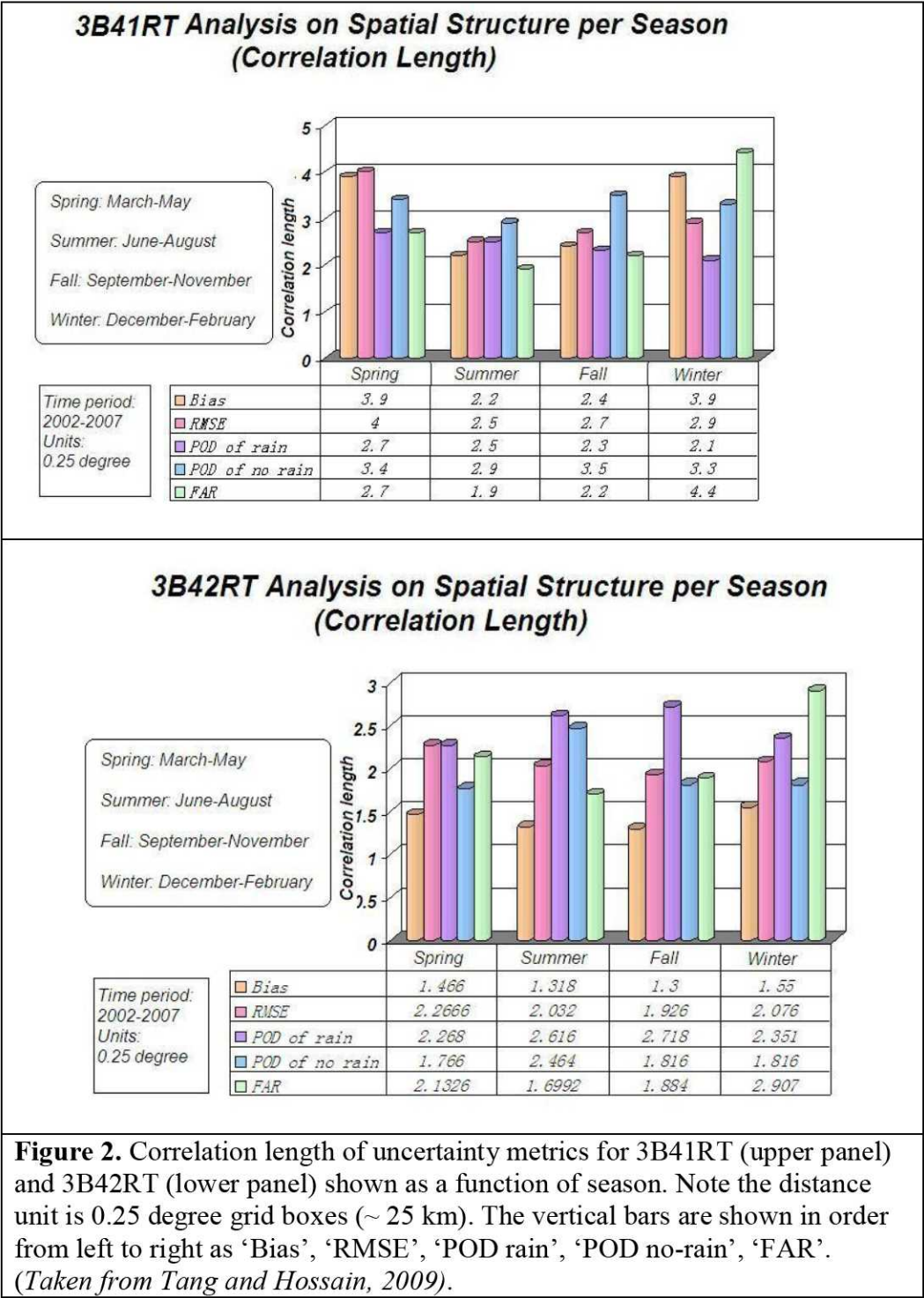
**Figure 4b.** Histograms of kriging error and actual error for false alarm ratio (FAR) over ungauged gridboxes. Here kriging error (shown in black) is defined as the difference between transferred (or kriged) FAR and the actual FAR derived from GV data. The actual GV-based FAR is shown in pink.

**Figure 4b.** Histograms of kriging error and actual error for false alarm ratio (FAR) over ungauged gridboxes. Here kriging error (shown in black) is defined as the difference between transferred (or kriged) FAR and the actual FAR derived from GV data. The actual GV-based FAR is shown in pink.

**Figure 5.** Impact of GV coverage (or sparseness) on the effectiveness of uncertainty metric transfer by ordinary kriging at climatologic scale (upper panel) and seasonal scale (lower panel) for the central US. . Computed with TMPA data collected for a 100 day period (May-August) in 2007-2008.

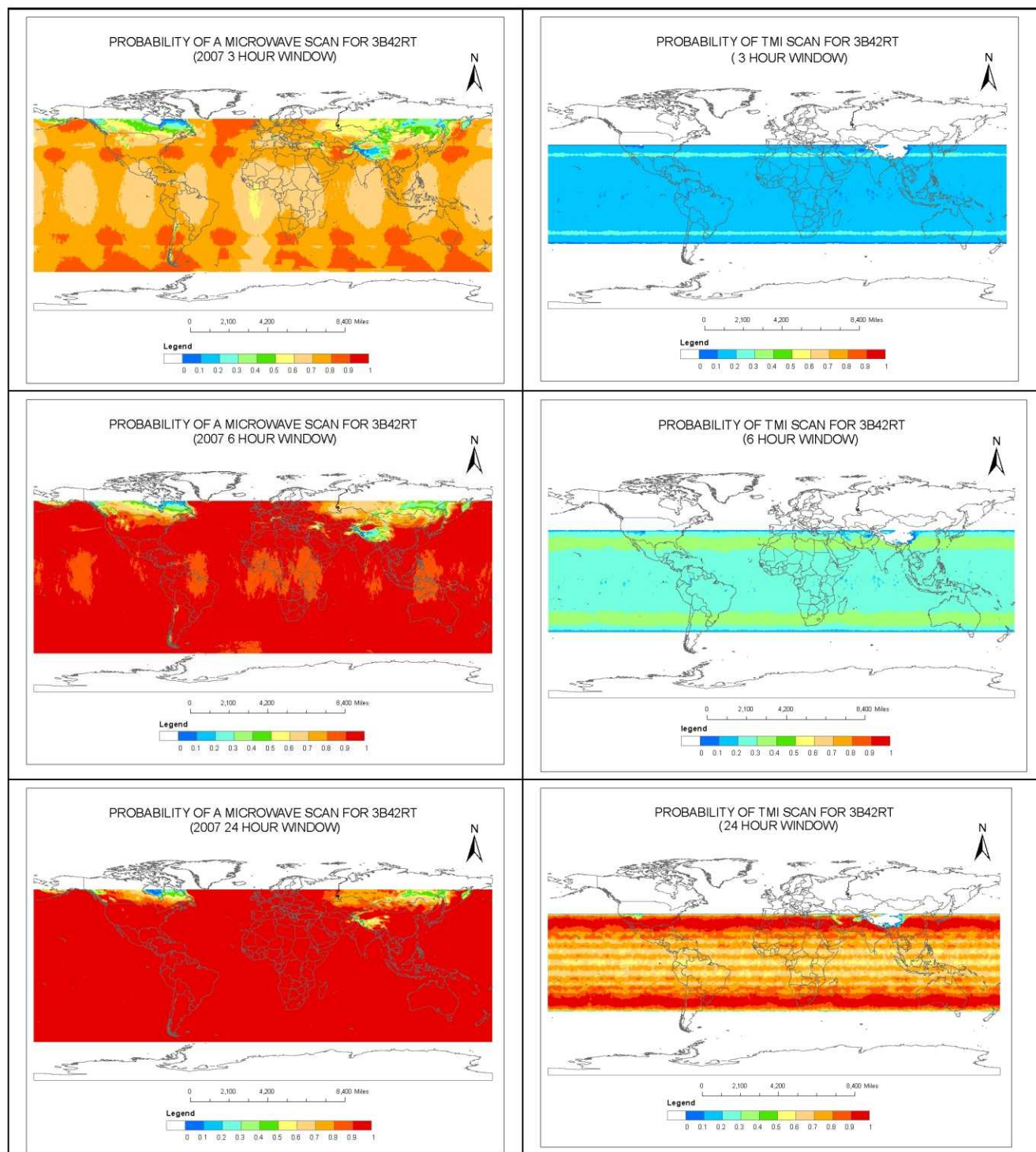


**Figure 1.** Conceptual rendition of the idea of ‘transfer’ of uncertainty information from a gaged (GV) location to an ungaged (non-GV) location. Upper panel depicts the notion of ‘uncertainty’ of satellite rainfall data (in this case, the scalar deviation of magnitudes is termed ‘uncertainty’ although there are many other types of uncertainty). Lower panel depicts how the known uncertainty (derived from GV sites shown in black in the middle panel) would be ‘transferred’ to the non-GV (ungaged) sites shown in blue (right most panel). *Reprinted from Tang and Hossain, 2009.*



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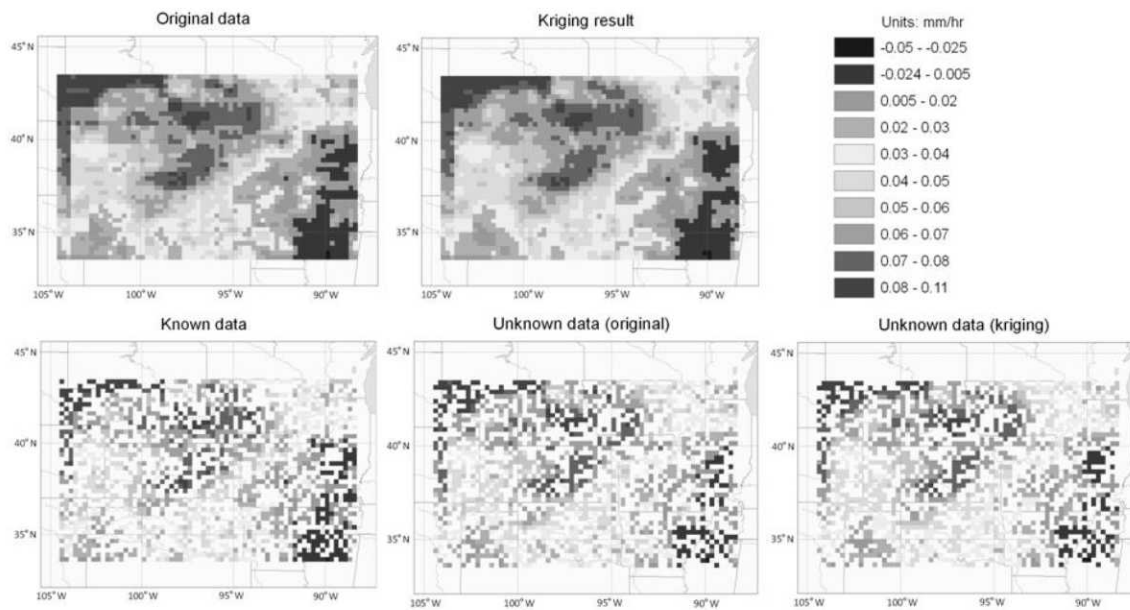


**Figure 3.** Left panel: probability of a 3B42RT 0.25° grid box having a PMW scan from either TMI, SSMIs or AMSR for 3 hour (upper), 6 hour middle) and 24 hour periods (bottom). Right panel: Same as left panel but only for TMI. These probability maps were created using a 100 day period of any PMW sensor from the 2007- 2008 period.



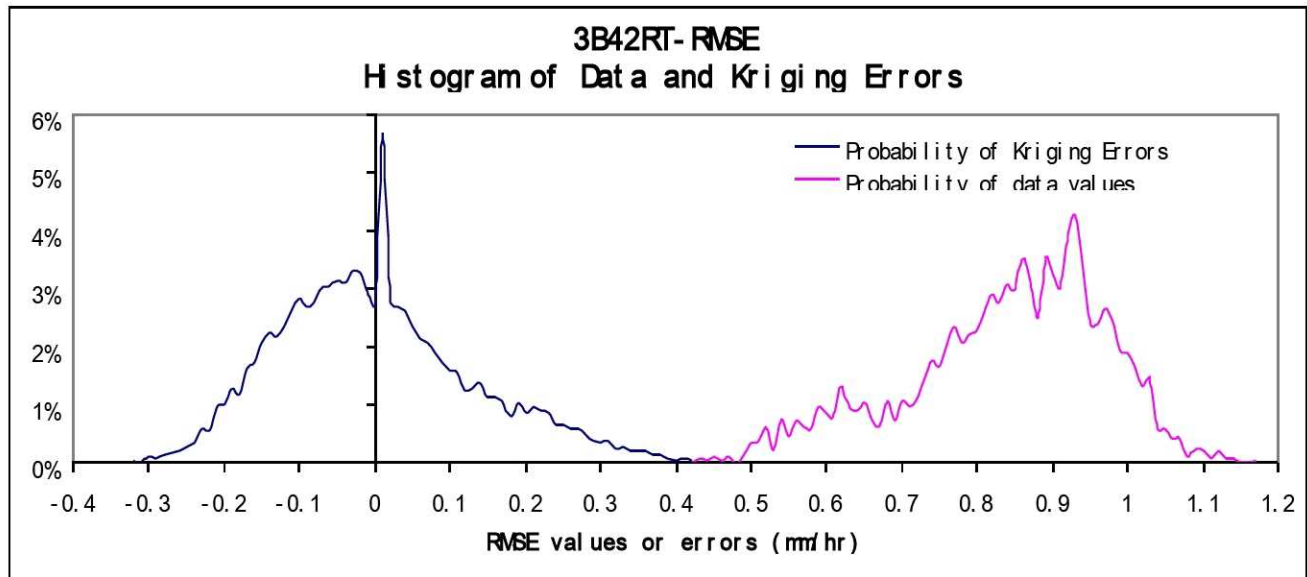
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### 3B41RT BIAS



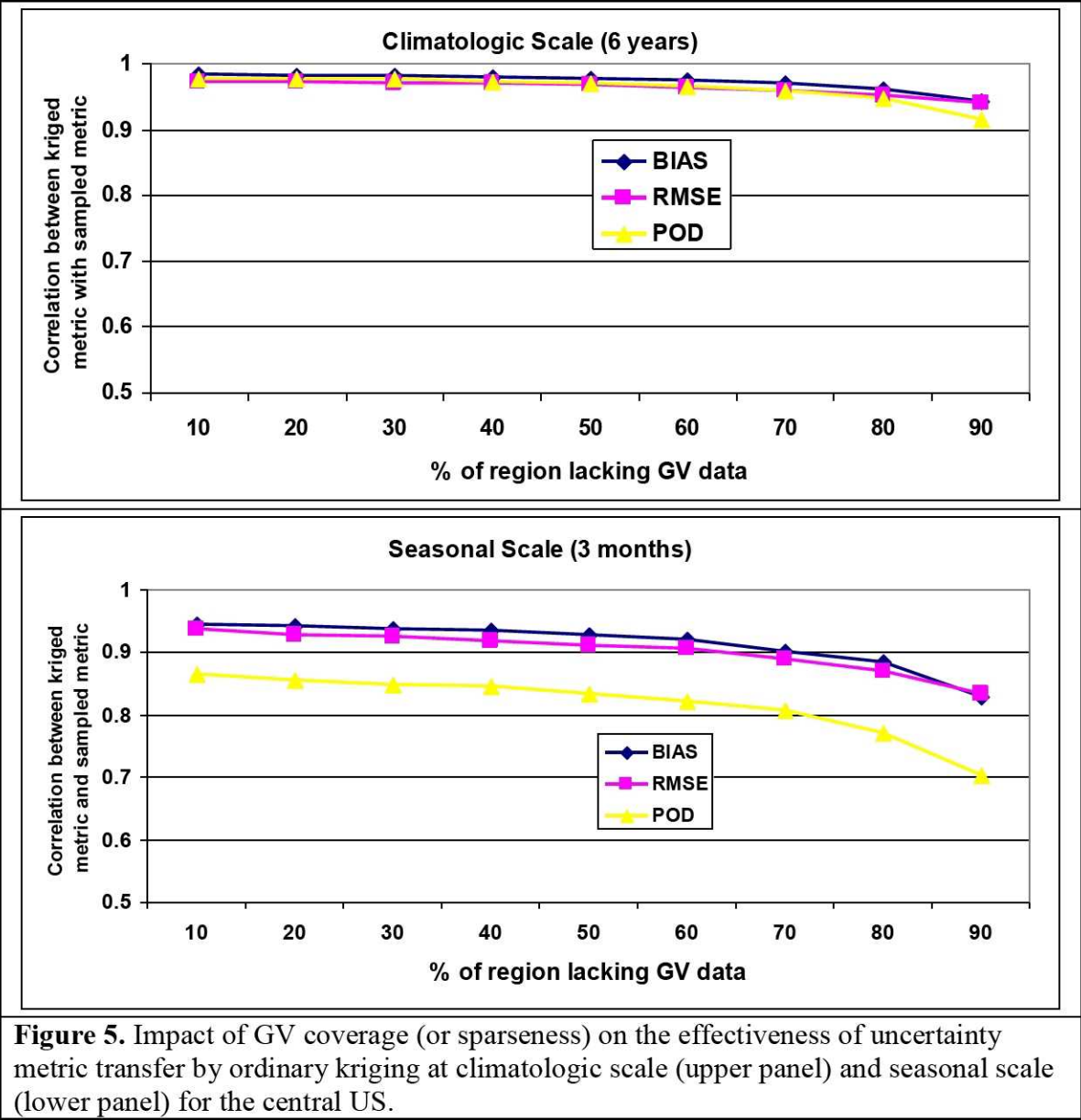
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**Figure 4a.** Transfer of Bias of 3B41RT from gauged to ungauged locations. Upper leftmost panel shows the true field of uncertainty on bias based on 6 years of data. The lower left most panel is the randomly selected 50% of the region for computation of the empirical variogram and correlation length. The lower middle panel shows the other 50% of the region that was assumed to be non-GV grid boxes. Lower right panel shows the estimation of the bias at the non-GV grid boxes using ordinary kriging.



**Figure 4b.** Histograms of kriging error and actual uncertainty for false alarm ratio (FAR) over ungauged gridboxes. Here kriging error (shown in black) is defined as the difference between transferred (or kriged) FAR and the actual FAR derived from GV data. The actual GV-based FAR is shown in pink.

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**Table 1a.** Assessment of the transfer of uncertainty metrics at monthly time scales (for summer months of June-July-August).

MONTH	% of region lacking GV data	BIAS		RMSE		POD		FAR	
		Mean Error <sup>1</sup>	Std. Dev of Error <sup>2</sup>	Mean Error	Std. Dev Of Error	Mean Error	Std. Dev of Error	Mean Error	Std. Dev of Error
June									
	10	0.53	0.78	0.19	0.18	0.12	0.10	0.22	0.27
	20	0.64	0.86	0.22	0.29	0.13	0.13	0.23	0.25
	30	0.60	0.93	0.21	0.24	0.13	0.12	0.24	0.31
	40	0.66	1.00	0.22	0.24	0.13	0.14	0.22	0.26
	50	0.68	1.08	0.23	0.20	0.14	0.16	0.24	0.37
	60	0.73	1.12	0.25	0.27	0.14	0.16	0.25	0.33
	70	0.75	1.11	0.26	0.25	0.15	0.15	0.26	0.35
	80	0.85	1.24	0.27	0.30	0.16	0.16	0.28	0.40
	90	1.00	1.43	0.31	0.29	0.19	0.23	0.31	0.47
July		Mean Error	Std. Dev of Error	Mean Error	Std. Dev Of Error	Mean Error	Std. Dev of Error	Mean Error	Std. Dev of Error
	10	0.54	0.71	0.20	0.20	0.16	0.15	0.32	0.43
	20	0.64	0.98	0.22	0.22	0.17	0.18	0.27	0.31
	30	0.67	0.98	0.22	0.21	0.16	0.19	0.28	0.32
	40	0.61	0.85	0.23	0.24	0.18	0.22	0.31	0.41
	50	0.67	1.01	0.25	0.24	0.18	0.21	0.31	0.39
	60	0.72	1.09	0.26	0.27	0.17	0.19	0.32	0.40
	70	0.80	1.13	0.26	0.24	0.19	0.24	0.33	0.44
	80	0.87	1.32	0.30	0.34	0.21	0.25	0.35	0.48
	90	0.99	1.46	0.33	0.34	0.24	0.30	0.39	0.51
August		Mean Error	Std. Dev of Error	Mean Error	Std. Dev Of Error	Mean Error	Std. Dev of Error	Mean Error	Std. Dev of Error
	10	0.47	0.79	0.21	0.22	0.14	0.17	0.28	0.47
	20	0.56	0.84	0.24	0.32	0.16	0.19	0.28	0.41
	30	0.52	0.78	0.23	0.28	0.15	0.17	0.28	0.35
	40	0.62	1.00	0.24	0.30	0.15	0.17	0.27	0.34
	50	0.62	0.95	0.25	0.34	0.16	0.21	0.27	0.35
	60	0.70	1.09	0.26	0.30	0.17	0.21	0.27	0.34
	70	0.69	1.05	0.30	0.36	0.17	0.20	0.29	0.40
	80	0.78	1.16	0.36	0.49	0.18	0.20	0.31	0.43
	90	0.83	1.19	0.35	0.42	0.21	0.23	0.33	0.50

Error = (Interpolated Uncertainty Metric - Actual Uncertainty Metric)/Actual Uncertainty Metric

<sup>1</sup>Mean Error= Mean of Error (as defined above) over non-GV grid boxes

<sup>2</sup>Std. Dev of Error = Standard Deviation of Error (as defined above) over non-GV grid boxes

624 **Table 1b.** Assessment of transfer of uncertainty metric at weekly time scales (for June weeks).

Week (of June)	% of region lacking GV data	BIAS		RMSE		POD		FAR	
1 <sup>st</sup> Week		Mean Error <sup>1</sup>	Std. Dev of Error <sup>2</sup>	Mean Error	Std. Dev Of Error	Mean Error	Std. Dev of Error	Mean Error	Std. Dev of Error
	10	0.78	1.13	0.45	0.73	0.27	0.29	0.35	0.32
	20	0.88	1.17	0.46	0.68	0.28	0.29	0.35	0.32
	30	0.87	1.16	0.41	0.53	0.31	0.34	0.36	0.33
	40	0.86	1.15	0.48	0.76	0.33	0.35	0.35	0.31
	50	0.97	1.25	0.41	0.54	0.33	0.35	0.34	0.30
	60	0.93	1.15	0.53	0.83	0.33	0.36	0.38	0.34
	70	1.05	1.35	0.52	0.80	0.34	0.36	0.38	0.32
	80	0.98	1.15	0.56	0.81	0.37	0.38	0.38	0.33
	90	1.15	1.37	0.74	1.06	0.39	0.38	0.40	0.38
2 <sup>nd</sup> Week		Mean Error	Std. Dev of Error	Mean Error	Std. Dev Of Error	Mean Error	Std. Dev of Error	Mean Error	Std. Dev of Error
	10	0.85	1.23	0.39	0.77	0.23	0.25	0.35	0.31
	20	0.81	1.15	0.37	0.64	0.25	0.27	0.34	0.32
	30	0.82	1.20	0.47	0.78	0.26	0.29	0.37	0.37
	40	0.80	1.19	0.47	0.77	0.27	0.31	0.37	0.37
	50	0.90	1.33	0.47	0.75	0.26	0.27	0.35	0.33
	60	0.94	1.28	0.48	0.78	0.28	0.31	0.36	0.32
	70	0.94	1.29	0.51	0.84	0.29	0.32	0.38	0.36
	80	1.08	1.45	0.61	0.96	0.29	0.33	0.38	0.38
	90	1.11	1.46	0.68	1.06	0.31	0.38	0.40	0.37
3 <sup>rd</sup> Week		Mean Error	Std. Dev of Error	Mean Error	Std. Dev Of Error	Mean Error	Std. Dev of Error	Mean Error	Std. Dev of Error
	10	0.76	1.04	0.42	0.56	0.28	0.33	0.31	0.28
	20	0.80	1.25	0.50	0.77	0.28	0.29	0.32	0.29
	30	0.79	1.25	0.54	0.91	0.31	0.34	0.33	0.30
	40	0.82	1.19	0.52	0.80	0.32	0.37	0.32	0.29
	50	0.83	1.18	0.55	0.93	0.31	0.33	0.32	0.32
	60	0.86	1.26	0.52	0.75	0.32	0.36	0.33	0.32
	70	0.94	1.32	0.58	0.90	0.32	0.34	0.33	0.33
	80	0.98	1.30	0.60	0.88	0.35	0.38	0.34	0.35
	90	1.09	1.43	0.73	1.16	0.39	0.44	0.39	0.38

Error = (Interpolated Uncertainty Metric - Actual Uncertainty Metric)/Actual Uncertainty Metric

<sup>1</sup>Mean Error= Mean of Error (as defined above) over non-GV grid boxes

<sup>2</sup>Std. Dev of Error = Standard Deviation of Error (as defined above) over non-GV grid boxes

**Table A.1.** Contingency Table (a HIT is defined when both satellite and GV agree on the type of event detected; a MISS is when there is disagreement between satellite and GV detected events).

		Truth/Reference	
		<i>Rainy Gridboxes</i>	<i>Non-rainy Gridboxes</i>
Satellite Estimates	Rainy Gridboxes	N <sub>A</sub> (HIT)	N <sub>B</sub> (MISS)
	Non-rainy Gridboxes	N <sub>C</sub> (MISS)	N <sub>D</sub> (HIT)