

Detection of p-mode Oscillations in HD 35833 with NEID and TESS

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Abstract

We report the results of observations of p-mode oscillations in the G0 subgiant star HD 35833 in both radial velocities and photometry with NEID and TESS, respectively. We achieve separate, robust detections of the oscillation signal with both instruments (radial velocity amplitude $A_{\rm RV} = 1.11 \pm 0.09$ m s⁻¹, photometric amplitude $A_{\rm phot} = 6.42 \pm 0.60$ ppm, frequency of maximum power $\nu_{\rm max} = 595.71 \pm 17.28$ µHz, and mode spacing $\Delta \nu = 36.65 \pm 0.96$ µHz) as well as a nondetection in a TESS sector concurrent with the NEID observations. These data shed light on our ability to mitigate the correlated noise impact of oscillations with radial velocities alone and on the robustness of commonly used asteroseismic scaling relations. The NEID data are used to validate models for the attenuation of oscillation signals for exposure times $t < \nu_{\rm max}^{-1}$, and we compare our results to predictions from theoretical scaling relations and find that the observed amplitudes are weaker than expected by >4 σ , hinting at gaps in the underlying physical models.

Unified Astronomy Thesaurus concepts: Asteroseismology (73); Exoplanet detection methods (489); Radial velocity (1332); Stellar astronomy (1583); Stellar oscillations (161)

Supporting material: data behind figures

1. Introduction

The current generation of extreme precision radial velocity (EPRV) spectrographs, which includes NEID (Schwab et al. 2016), EXPRES (Jurgenson et al. 2016), ESPRESSO (Pepe et al. 2021), and MAROON-X (Seifahrt et al. 2018), is pushing past the 1 m s⁻¹ instrumental noise barrier and paving the way for the detection of 10 cm s⁻¹ RV signals induced by Earthmass exoplanets orbiting Sun-like stars. While we may not achieve this feat with these instruments, they will afford us the opportunity to undertake more detailed studies of one of the largest remaining hurdles in the pursuit of Earth-analog

Original content from this work may be used under the terms of the Creative Commons Attribution 4.0 licence. Any further distribution of this work must maintain attribution to the author(s) and the title of the work, journal citation and DOI. exoplanet detection with radial velocities: intrinsic stellar variability (Fischer et al. 2016; Crass et al. 2021). For even the quietest Sun-like stars, stellar variability manifests at the 1 m s^{-1} level and can easily mask the signals we seek to extract. Without an appropriate means of accounting for these stellar signals, we will fail to detect and characterize sub-m s⁻¹ signals from low-mass, long-period exoplanets in spite of the advantage conferred by recent advances in instrumentation.

Efforts to mitigate the impact of stellar RV variations on exoplanet detection have found success by leveraging our understanding (albeit incomplete) of the underlying physical processes to model these signals rather than naïvely treating them as white noise. It has become commonplace, for example, to use Gaussian processes (GPs) to model stellar activity signals and disentangle them from exoplanet-induced Doppler shifts (e.g., Haywood et al. 2014; Rajpaul et al. 2015; Foreman-Mackey et al. 2017; Gilbertson et al. 2020; Langellier et al. 2021; Zhao et al. 2022). But this strategy is only applicable when it is feasible to obtain many observations on the timescale over which the stellar signals remain correlated. Short-timescale

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Figure 1. Cycles 1 and 3 TESS light curves for HD 35833. We show both the Sector 6 (left) and Sector 32 (right) 2-minute cadence data. From top to bottom: SAP (green), PDCSAP (pink), and custom pixel level decorrelation (PLD)-detrended (blue; Sector 32 only). The photometric analysis herein makes use of the PDCSAP reduction (RMS CDPP \approx 60 ppm) for Sector 6, and the PLD-detrended reduction (RMS CDPP \approx 100 ppm) for Sector 32 detrended light curve is available as the data behind the figure.

(The data used to create this figure are available.)

variations, such as acoustic, pressure-driven "p-mode" oscillations, necessitate a different approach. For Sun-like mainsequence stars, p-mode oscillations produce RV signals that vary with periods of 5–10 minutes. Chaplin et al. (2019) outline a detailed model for how the duration of an exposure affects the residual amplitude of RV signals due to stellar p-mode oscillations, showing that in principle, one can filter oscillation signals to the <10 cm s⁻¹ level with exposure times typical of current EPRV surveys. Using an 8 hr α Cen A RV time series collected by Butler et al. (2004) with the UVES instrument on the VLT, Chaplin et al. (2019) show that the predicted residual amplitude does indeed match the observed RV signal for exposure times longer than typical oscillation periods, validating the general behavior of their model. This model has been used to guide observing strategies for EPRV exoplanet searches (e.g., Blackman et al. 2020; Gupta et al. 2021) and to inform simulations of future surveys (Luhn et al. 2022). Luhn et al. (2022) present a more comprehensive picture of the RV noise contributed by oscillations with a model that describes not only the impact on individual observations, but also correlations across multiple observations. When applying these models to stars other than the Sun, for which we often lack empirical measurements of the oscillation frequencies and amplitudes, we typically rely on scaling relations to predict the asteroseismic parameters from known stellar properties (e.g., Kjeldsen & Bedding 1995). However, these scaling relations are imperfect, as they do not capture the full suite of properties that affect the oscillation signal. If this approach is to be effective on large samples of stars, we must therefore pursue a more rigorous understanding of p-mode oscillations and the radial velocity signals they induce.

In this work, we present the results of both simultaneous and asynchronous observations of p-mode oscillations in the subgiant star HD 35833 in radial velocities and photometry with NEID and the Transiting Exoplanet Survey Satellite (TESS; Ricker et al. 2015), respectively. We describe the observations and data reduction in Section 2. Asteroseismic analyses of these data, detailed in Section 4, yield detections of suppressed²² oscillation signals with NEID and one sector of TESS data, as well as a nondetection in a different TESS sector. The results shed light on our ability to mitigate the correlated noise impact of oscillations with radial velocities alone, and on the accuracy of stellar parameter scaling relations in predicting the frequencies and amplitudes of p-mode oscillations. In Section 5, we discuss potential mechanisms for the observed oscillation amplitude suppression as well as the implications of our results for future RV exoplanet searches and follow-up observations.

2. Observations

2.1. TESS Photometry

TESS observed the subgiant star HD 35833 (HIP 25589, TIC 302423299) during Cycles 1 (Sector 6; 2018 December 11-2019 January 7), 3 (Sector 32; 2020 November 19-2020 December 17), and 4 (Sectors 43, 44, and 45; 2021 September 16-2021 December 2) at a 2-minute cadence. We obtain photometric data for all five sectors from the TESS Science Processing Operations Center (SPOC; Jenkins et al. 2016), and we show the simple aperture photometry (SAP) and pre-search data conditioned simple aperture photometry (PDCSAP) light curves in Figures 1 and 2. The PDCSAP module (Smith et al. 2012; Stumpe et al. 2012, 2014) is designed to correct for long-term instrument systematics, background trends, and other sources of noise while preserving astrophysical signals on shorter timescales. For Sectors 6 and 43-45, we use the PDCSAP reduction for all photometric analyses herein. For Sector 32, however, we see unexpected, high-frequency flux variations in the PDCSAP reduction with amplitudes >500 ppm, nearly an order of magnitude greater than the ~ 60 ppm combined differential photometric precision (CDPP) observed in the other sectors. The absence of this structure in the SAP light curve for Sector 32

 $[\]frac{22}{10}$ In this work, we use the term "suppress" to refer to a partial reduction in oscillation amplitude rather than a total absence of power.



Figure 2. Cycle 4 TESS light curve for HD 35833. We show both the 2-minute cadence data for Sectors 43, 44, and 45. From top to bottom: SAP (green) and PDCSAP (pink). The photometric analysis herein makes use of the PDCSAP reduction (RMS CDPP \approx 60 ppm) for the combined light curve.

suggests the signal is not astrophysical in nature but rather an artifact introduced by the pipeline detrending method.

We reprocess the Sector 32 TESS photometry using the pixel level decorrelation (PLD) toolkit in lightkurve (Lightkurve Collaboration et al. 2018). This method, described in detail by Deming et al. (2015) and Luger et al. (2016, 2018), uses information from nearby pixels to construct a model for the instrumental noise in the region of interest, which can then be subtracted from the pixels in the stellar aperture to produce a systematics-corrected light curve. We extract a 50 pixel×50 pixel cutout from the 10-minute cadence full-frame image (FFI) data centered on HD 35833 (Figure 3), and we calculate the noise model using the lightkurve.PLDCorrector. correct method. We restrict the corrector model to five PCA components so as not to overfit the data, though we note that the result is relatively insensitive to the number of components for ≤ 10 components except at the edges of the light curve. For this correction, we use the default generated background aperture but we manually fix the stellar aperture to be the same as the pipeline aperture from the 2-minute cadence data. There are several known sources in this aperture, but the brightest of these is nearly 7 magnitudes fainter than the target in the Gaia passband (Gaia DR3 3391121978660964992, G=13.601) and the contamination ratio is just 0.00335. We do not expect blending to be a significant source of uncertainty. We interpolate the noise model onto the 2-minute cadence time stamps and subtract it from the SAP flux. Finally, a cubic polynomial is fit to the result to remove any remaining long-term trends.

We also explored detrending the Sector 32 light curve using cotrending basis vectors (CBV) following the methods of Lund et al. (2021). This method did not produce a cleanly detrended light curve for HD 35833, however, so we use the final PLD-corrected light curve (Figure 1; CDPP ≈ 100 ppm) for our analysis of the Sector 32 photometry.

2.2. NEID Radial Velocity Observations

We observed HD 35833 with the NEID spectrograph (Schwab et al. 2016; Robertson et al. 2019) on the WIYN²³ 3.5 m telescope at Kitt Peak National Observatory on the night of 2020 December 12 UT. We observed the star for 5.5 hr at a



Figure 3. TESS FFI cutout used for PLD detrending of the HD 35833 Sector 32 light curve. The stellar aperture is shown in red and the background aperture is shown in gray.

2-minute cadence (90 s exposures), following the target form near-zenith to airmass=2.0. While the intent was to obtain an 8 hr baseline, the observing window was limited by poor weather conditions during the first half of the night. The 2-minute cadence was selected so that we could achieve a high radial velocity precision for each exposure while also finely sampling the stellar oscillations. We observed using the NEID high resolution ($R \sim 113,000$) mode, taking simultaneous etalon calibration frames with the optical density (OD) 1.3 neutral density filter. During the observing sequence, we obtained 166 spectra with a median per-resolution-element signal-to-noise ratio (S/N) of 76 at 550 nm. The S/N of the exposures varies as a result of transparency changes due to intermittent cloud cover, but there is no trend with changing airmass.

The NEID data were processed with version 1.1.2 of the NEID Data Reduction Pipeline²⁴ (DRP). We remove two outliers with S/N < 50 leaving 164 high-quality spectra. We

²³ The WIYN Observatory is a joint facility of the NSF's National Optical-Infrared Astronomy Research Laboratory, Indiana University, the University of Wisconsin–Madison, Pennsylvania State University, the University of Missouri, the University of California–Irvine, and Purdue University.

²⁴ https://neid.ipac.caltech.edu/docs/NEID-DRP/



Figure 4. NEID RV measurements for HD 35833. RVs are calculated using a modified version of the SERVAL template-matching algorithm applied to 1D extracted spectra from the NEID DRP, with low S/N outliers removed.

consider two sources of RVs for our analysis: data products are taken directly from the NEID DRP, which uses the cross-correlation function (CCF; Baranne et al. 1996) method to calculate RVs, and an independent derivation using the template-matching method with a modified version of the SpEctrum Radial Velocity AnaLyser (SERVAL Zechmeister et al. 2018) analysis code, optimized for use for NEID spectra as described in Stefansson et al. (2022). For the SERVAL reduction, we extracted RVs using order indices 69–153 spanning the wavelength region from 398nm to 895nm. The two RV streams are fully consistent with each other, but we elect to use the SERVAL RVs for the remainder of this work as this method yields a better median RV precision (1.05 m s⁻¹) than the CCF-derived RVs from the DRP (1.70 m s⁻¹). The SERVAL radial velocity time series for HD 35833 is shown in Figure 4.

3. Stellar Parameters

The stellar parameters used in this work are taken from the Final Luminosity Age Mass Estimator (FLAME) module of the Gaia Data Release 3 Astrophysical parameters inference system (Apsis) pipeline (Creevey et al. 2022). For bright stars, the FLAME module processes separate sets of inputs, from each of the GSP-Phot and GSP-Spec modules, producing two corresponding sets of stellar parameters (Creevey et al. 2022). We choose to use the FLAME/GSP-Phot results rather than the FLAME/GSP-Spec results for this work, as Fouesneau et al. (2022) and Recio-Blanco et al. (2022) note significant systematic biases in the log g values reported by the GSP-Spec module. FLAME provides values for R_{\star} , M_{\star} , and L_{\star} , from which we also calculate $T_{\rm eff}$ and log g. These parameters and their associated uncertainties are listed in Table 1.

4. Asteroseismic Analysis

4.1. Asteroseismic Scaling Relations for P-mode Oscillations

Several asteroseismic quantities, including the frequencies, amplitudes, and frequency spacing of p-mode oscillations, are expected to scale with fundamental stellar properties. Kjeldsen & Bedding (1995) present a set of theoretically motivated and empirically validated scaling relations for these quantities, which we use here to predict the parameters of the oscillation signals that we observe with TESS and NEID.

The frequency of maximum power for the oscillation power excess scales as

$$\nu_{\rm max} = \nu_{\rm max,\odot} \frac{g/g_{\odot}}{\sqrt{T_{\rm eff}/T_{\rm eff,\odot}}} \tag{1}$$

 Table 1

 Summary of Stellar Parameters for HD 35833

Parameter	Value	Description	Reference
Alternate Ide	entifiers:		
TIC	302423299	TESS Input	Stassun
		Catalog	
Gaia DR3	3391121978660968576	Gaia	Gaia DR3
HIP	25589	HIPPARCOS	HIPPARCOS
		Catalog	
Coordinates	and Parallax:		
$\alpha_{\rm J2016}$	05:28:09.20	Right Ascen- sion (RA)	Gaia DR3
δ_{J2016}	+16:26:19.61	Declination (Dec)	Gaia DR3
ω	14.89 ± 0.02	Parallax (mas)	Gaia DR3
Broadband H	Photometry:		
TESS	6.2426 ± 0.0061	TESS	Stassun
G	6.6975 ± 0.0001	Gaia	Gaia DR3
B_p	7.0194 ± 0.0004	Gaia	Gaia DR3
R_p	6.2005 ± 0.0005	Gaia	Gaia DR3
Derived Stel	lar Parameters:		
$T_{\rm eff}$	5684.2 ± 64.7	Effective Temper-	Gaia DR3
		ature (K)	
L_{\star}	7.147 ± 0.037	Luminosity (L _o)	Gaia DR3
R_{\star}	2.756 ± 0.056	Stellar Radius (R_{\odot})	Gaia DR3
M_{\star}	1.423 ± 0.041	Stellar Mass (M_{\odot})	Gaia DR3
log g	3.710 ± 0.030	Surface Gravity (log (g cm^{-3}))	Gaia DR3

References are: HIPPAROCS (ESA 1997), Stassun (Stassun et al. 2019), Gaia DR3 (Gaia Collaboration et al. 2022).

and $\Delta \nu$, the large oscillation mode spacing, is given by

$$\Delta \nu = \Delta \nu_{\odot} \sqrt{\rho_{\star} / \rho_{\odot}} \tag{2}$$

where the above scaling relations are derived by Kjeldsen & Bedding (1995) and we assume $\nu_{\max,\odot} = 3090 \ \mu$ Hz and $\Delta \nu_{\odot} = 135.1 \ \mu$ Hz for the solar values (Campante et al. 2016). Based on the stellar parameters given in Table 1, we calculate the predicted values of ν_{\max} and $\Delta \nu$ for HD 35833 to be $580.13 \pm 28.91 \ \mu$ Hz and $35.22 \pm 2.36 \ \mu$ HZ, respectively. These values, along with other derived asteroseismic quantities, are shown in Table 2. The quoted uncertainties are propagated from the uncertainties on the stellar parameters, and we do not account for intrinsic scatter in the asteroseismic scaling relations.

The expected amplitudes of the oscillation signals observed with both TESS and NEID can also be calculated using scaling relations from Kjeldsen & Bedding (1995). For TESS, the expected amplitude of radial, l = 0, oscillation modes follows

$$A_{\text{phot}} = A_{\text{phot},\odot} \beta \left(\frac{L_{\star}}{L_{\odot}}\right)^{s} \left(\frac{M_{\odot}}{M_{\star}}\right)^{s} \left(\frac{T_{\text{eff}},\star}{T_{\text{eff},\odot}}\right)^{-2}$$
(3)

where we assume $A_{\text{phot},\odot} = 2.125$ ppm for the TESS bandpass, as calculated by Campante et al. (2016) based on a derivation of the amplitude in the Kepler bandpass by Ballot et al. (2011). We also apply the correction factor β , first introduced by Chaplin et al. (2011a), to account for observed deviation from the above scaling relation for relatively hot stars:

$$\beta = 1 - \exp\left(-\frac{T_{\rm red} - T_{\rm eff}}{1550K}\right) \tag{4}$$

 Table 2

 Predicted and Observed Asteroseismic Parameters for HD 35833

Parameter	Scaling Relations	TESS Sector 6	TESS Sector 32	TESS Sectors 43/44/45	NEID
$\nu_{\rm max}$ ($\mu {\rm Hz}$)	580.13 ± 28.91	582.54 ± 65.42		595.71 ± 17.28	
$\Delta \nu \ (\mu \text{Hz})$	35.22 ± 2.36	33.80 ± 1.17		36.65 ± 0.96	
A _{phot} (ppm)	7.42 ± 0.27	5.65 ± 1.25	<10.90 (99%)	6.42 ± 0.60	
			<7.89 (50%)		
$A_{\rm RV} \ ({\rm m \ s^{-1}})$	2.86 ± 0.10				1.11 ± 0.09

$$T_{\rm red} = 8907 (L_{\star}/L_{\odot})^{-0.093} K.$$
 (5)

We note that this calculation disregards any dependence on stellar activity; this is discussed in further detail in Section 5.1.2. If we assume s = 1 for the mass and luminosity scaling in Equation (3), as in Campante et al. (2016) and Chaplin et al. (2011a), then using the derived stellar parameters listed in Table 1, we find $A_{\text{phot,predicted}} = 7.42 \pm 0.27$ ppm.

It bears emphasizing that Equation (3) is specifically for the amplitudes of the radial oscillation modes, which differ from $A_{\text{max,phot}}$. $A_{\text{max,phot}}$ is defined as the amplitude of a Gaussian fit to the oscillation power excess (with units ppm²/ μ Hz), and it is related to A_{phot} via the equation

$$A_{\rm phot} = \sqrt{\left(A_{\rm max, phot} \frac{\Delta \nu}{c}\right)} \tag{6}$$

where the factor of c, the effective number of oscillation modes per spherical harmonic order, corrects for the contributions of nonradial ($l \neq 0$) modes to the power excess.

For NEID, the expected radial velocity amplitude of the oscillation signal is given by

$$A_{\max, RV} = A_{\max, RV, \odot} \left(\frac{L_{\star}}{L_{\odot}} \right) \left(\frac{M_{\odot}}{M_{\star}} \right).$$
(7)

We adopt $A_{\max,RV,\odot} = 0.19 \text{ m s}^{-1}$ from Chaplin et al. (2019). The radial velocity amplitude for the l = 0 modes is calculated as in Equation (6), using c = 4.09 as derived by Kjeldsen et al. (2008). Here, we find $A_{RV,predicted} = 2.86 \pm 0.10 \text{ m s}^{-1}$. We note that the correction factor c depends on the relative visibilities of different oscillation modes, which may differ for stars of different masses and evolutionary states. The dipolar l = 1 mode in particular has been shown to be weaker in some cases for evolved stars (Stello et al. 2016), which may influence the accuracy of the predicted values of the oscillation amplitudes in this work. We comment on this possibility in Section 5.1.3.

4.2. Photometric Analysis

We conduct a global asteroseismic analysis of the TESS photometry using the pySYD fitting package (Chontos et al. 2021), treating each of the Cycle 1 (Sector 6), Cycle 3 (Sector 32) and Cycle 4 (Sectors 43–45) light curves as independent data sets. The resulting PSDs along with the fitted granulation and oscillation components are shown in Figure 5. We also include the white noise values (9.59 ppm²/ μ Hz, 21.59 ppm²/ μ Hz, and 11.96 ppm²/ μ Hz, for Cycles 1, 3, and 4, respectively), which are calculated as the median power in the range 2000 μ Hz to the Nyquist frequency of 4167 μ Hz. From the Sector 6 light curve, we obtain a significant detection of the oscillation signal with $\nu_{max} = 582.54 \pm 65.42 \,\mu$ Hz and large



Figure 5. Power spectral density for HD 35833 TESS Sector 6 (top), Sector 32 (middle), and the Cycle 4 data (bottom), with fine sampling in light gray and a boxcar-smoothed power spectrum in black. For Sectors 6 and 43–45, we show the individual oscillation (orange), granulation (green), and white noise (gray horizontal line) components as fit by pySYD as well as the sum of these components (blue). For Sector 32, we show the fitted granulation and white noise components, as well as injected oscillation signals with frequencies set by the signal detected in the Cycle 4 data and amplitudes scaled by $1.23 \times$ (dashed line) and $1.70 \times$ (dotted line) for the 50% and 99% detection limits, respectively.

frequency spacing $\Delta \nu = 33.80 \pm 1.17 \,\mu$ Hz, both of which are consistent with the values predicted by theoretical scaling relations (Equations (1) and (2)). The amplitude of the Gaussian fit to the smoothed, background-corrected power excess is $A_{\text{max,phot}} = 2.80 \pm 1.23 \,\text{ppm}^2 / \,\mu$ Hz. As in Huber et al. (2019), we then follow the prescription of Kjeldsen et al. (2008) and

Kjeldsen et al. (2005) to correct for the contributions of nonradial modes and calculate the mean amplitude of the radial, l=0, modes following Equation (6). The value of c used in Section 4.1 is not applicable here, however, as the relative visibilities of radial and nonradial modes differ between photometry and radial velocities. We calculate c = 2.96 for the TESS bandpass, centered on 786.5 nm, by linearly interpolating between the 500 nm and 862 nm values derived by Kjeldsen et al. (2008), and we find that $A_{\rm phot} = 5.65 \pm$ 1.25 ppm for the Sector 6 light curve. We obtain similar results for the Cycle 4 data, with $\nu_{\text{max}} = 595.71 \pm 17.28 \,\mu\text{Hz}$, $\Delta \nu = 36.65 \pm 0.96 \ \mu \text{Hz}$, and $A_{\text{max,phot}} = 3.33 \pm 0.62 \ \text{ppm}^2/$ μ Hz, and we calculate $A_{\text{phot}} = 6.42 \pm 0.60$ ppm. While both data sets had similar photometric precision, the asteroseismic quantities derived from the Cycle 4 light curve are more precise as a consequence of the longer time baseline.

For Sector 32, we do not detect the oscillation power excess with pySYD, likely because of the relatively poor photometric precision of the data (CDPP ≈ 100 ppm for Sector 32 vs. $CDPP \approx 60$ ppm for the other data) and consequently the significantly larger background noise in the power spectrum. We place upper limits on the photometric oscillation amplitude in the Sector 32 light curve via a set of injection recovery tests using the properties of the oscillation signal extracted from the Cycle 4 data. The background-corrected oscillation power excess, with height scaled to simulate a change in A_{phot} , is used to generate an oscillation time series which is then injected into the Sector 32 light curve. The modified light curve is reanalyzed using pySYD to try to recover the injected signal. We achieve a 50% recovery rate for injected signals with $A_{\rm phot} = 7.89 \, \rm ppm \ (1.23 \times the amplitude detected in Cycle 4)$ and a 99% recovery rate for injected signals with $A_{\rm phot} = 10.90 \, \rm ppm \ (1.70 \times the amplitude detected in Cycle 4).$ These results, illustrated in Figure 5, show that we should not expect to detect the oscillation signal in the Sector 32 data if the photometric oscillation amplitude were the same as in the other sectors. In addition, the predicted amplitude is just below the 50% detection threshold, so this result is consistent with expectations from scaling relations.

Though the nondetection in Sector 32 is unsurprising, the observed values of A_{phot} from Sector 6 and Sectors 43–45 differ slightly ($<2\sigma$ discrepancy) from the predicted value. These low amplitudes suggest that the oscillation signal for HD 35833 may be suppressed, or weakened, by some mechanism that is not accounted for in the Equation (3) scaling relation. We discuss this possibility in detail in Section 5.

4.3. Radial Velocity Analysis

The signals of interest for this work vary on timescales of less than an hour. Before proceeding with the RV analysis, we first fit and subtract a slowly varying quadratic trend from the data to remove any long timescale variation, e.g., supergranulation or uncorrected instrument drift. We then run a Lomb-Scargle periodogram analysis (Lomb 1976; Scargle 1982) to search for the oscillation signal, and we detect several peaks near ν_{max} as predicted by the scaling relation in Equation (1) and as measured via the TESS data in Section 4.2 (Figure 6). However, it is clear that the periodogram is not well resolved in frequency space due to the short baseline of the NEID observations, so these data cannot be used to measure the specific frequencies and amplitudes of individual modes. As we illustrate in Figure 6 via simulations of the oscillation time



Figure 6. Lomb-Scargle periodogram of the detrended NEID observations of HD 35833. We use the same frequency limits as in Figure 5 for ease of comparison. We also show the frequency of maximum power for the oscillation signal as detected in TESS Cycles 1 (blue, dotted) and 4 (purple, dashed) and as predicted by scaling relations (red) as vertical lines, as well as the 10%, 5%, and 1% false alarm levels (calculated via the false_alarm_level() method of the astropy.timeseries.LombScargle class) as horizontal lines (black, dotted-dashed). For comparison, we show periodograms for a set of simulated oscillation time series with the same sampling and photon noise properties as the NEID observations (background colored lines).

series, the observed periodogram structure can vary significantly from one 5.5 hr realization to the next. We therefore elect not to fit a Gaussian envelope to the oscillation power excess for the NEID data, and we instead analyze the RVs in the time domain.

4.3.1. Gaussian Processes Conditioning and Decomposition

We perform a GP regression using the RV kernels for granulation and oscillations given in Luhn et al. (2022), based on the photometric-to-RV scalings from Guo et al. (2022). The kernel hyperparameters are set by $T_{\rm eff}$ and log g and as a result, the GP kernels are fully deterministic.

If the structure of the covariance matrix is known, a set of n_1 observations (t_1, y_1) can be used to make predictions over a different sample (t_2, y_2) of length n_2 , as they are jointly Gaussian

$$\begin{bmatrix} \mathbf{y}_1 \\ \mathbf{y}_2 \end{bmatrix} \sim N\left(\begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix}, \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix}\right).$$
(8)

The resulting posterior distribution for μ_2 conditioned on the observations (t_1, y_1) is then centered on the mean

$$\mu_{2|1} = (\Sigma_{11}^{-1} \Sigma_{12})^{l} \mathbf{y}_{1}, \tag{9}$$

with covariance

$$\Sigma_{2|1} = \Sigma_{22} - (\Sigma_{11}^{-1} \Sigma_{12})^{\mathsf{T}} \Sigma_{12}, \qquad (10)$$

and standard deviation

$$\sigma_{2|1} = [\text{Diag}(\Sigma_{2|1})]^{1/2}.$$
 (11)

The previous equations describe standard GP conditioning in the case of a single kernel. In our case, however, we have described the covariance matrix as the sum of two astrophysical GP kernels (granulation and oscillation), e.g.,

$$\Sigma[i, j] = k_a(t_i, t_j) + k_b(t_i, t_j).$$
(12)

In this case, Equation (9) describes the conditioned mean for the sum of two underlying processes. We wish to separate out the mean effect that each of these kernels contributes. We can write Equation (9) more explicitly in this case as

$$\mu_{2|1,a+b} = [\Sigma_{11}^{-1}(\Sigma_{12,a} + \Sigma_{12,b})]^{\mathsf{T}} \mathbf{y}_{1}.$$
 (13)

If we define

$$\mu_{2|1,a} = (\Sigma_{11}^{-1} \Sigma_{12,a})^{\mathsf{T}} \mathbf{y}_1 \tag{14}$$

$$\mu_{2|1,b} = (\Sigma_{11}^{-1} \Sigma_{12,b})^{\mathsf{T}} \mathbf{y}_1, \tag{15}$$

it is true that

$$\mu_{2|1,a+b} = \mu_{2|1,a} + \mu_{2|1,b} \tag{16}$$

where $\mu_{2|1,a}$ represents the conditioned time series described by the kernel $k_a(t_i, t_j)$ and $\mu_{2|1,b}$ represents the conditioned time series described by the kernel $k_b(t_i, t_j)$. In this way, our conditioned mean $\mu_{2|1}$ can be broken up into its component GP kernels. We refer to this as GP *decomposition* to isolate individual components of a multicomponent GP kernel. While the decomposition of the conditioned *means* is relatively trivial, the decomposition of the conditioned *covariance* includes cross terms that cannot be assigned to an individual component.

We wish to focus a moment on our definitions above for $\mu_{2|1,a}$ and $\mu_{2|1,b}$. It is important to note that these are not the same as conditioning the data solely using one kernel or the other, as both kernels are still included in the construction of the covariance matrix Σ_{11} , which also includes the diagonal observation errors (i.e., photon noise). We can show this more explicitly by writing Equation (9) in a slightly different way

$$\mu_{2|1} = \Sigma_{21} \Sigma_{11}^{-1} \mathbf{y}_1. \tag{17}$$

The conditioned mean (Equation (17)) can be expanded to explicitly show each component of our granulation and oscillation GP sum

$$\mu_{2|1,gran+osc} = (\Sigma_{21,gran} + \Sigma_{21,osc}) \\ \times (\Sigma_{11,gran} + \Sigma_{11,osc} + \Sigma_{11,phot})^{-1} \mathbf{y}_{1}, \quad (18)$$

which can be decomposed into a granulation component

$$\mu_{2|1,gran} = (\Sigma_{21,gran})(\Sigma_{11,gran} + \Sigma_{11,osc} + \Sigma_{11,phot})^{-1} \mathbf{y}_1, \quad (19)$$

and an oscillation component

$$\mu_{2|1,osc} = (\Sigma_{21,osc}) \\ \times (\Sigma_{11,gran} + \Sigma_{11,osc} + \Sigma_{11,phot})^{-1} \mathbf{y}_1.$$
(20)

We apply this GP conditioning and decomposition to the detrended NEID time series to isolate the individual oscillation and granulation components of the stellar radial velocity signal. The resulting time series and the radial velocity residuals are shown in Figure 7.

4.3.2. RMS of Decomposed Time Series

One of the main conclusions of the Chaplin et al. (2019) study is that the residual amplitude of a stellar oscillation signal can be predicted if the stellar parameters are known, with appropriate caveats for stochastic mode excitation as discussed in Section 4.3.3. Because the residual amplitude manifests as an RMS when treating oscillations as a source of noise in an RV time series (e.g., for exoplanet surveys), this means that the oscillation contribution to the total RMS can be controlled by adjusting the integration time of one's observations. With the

isolated oscillation RV time series from the GP decomposition for HD 35833, we have the means to explore this prediction.

Here, we vary the integration time, t_i , by binning sets of consecutive exposures, so the integration time is defined as the sum of the on-sky exposure times, t_e , for all exposures in a bin and inter-exposure readout times, t_r . For a bin with M observations, then,

$$t_i = Mt_e + (M - 1)t_r.$$
 (21)

In Figure 8, we show the RV RMS as a function of integration time for the total observed signal and each of the decomposed oscillation and granulation components. We also plot the measured photon noise uncertainties, as well as the sum of the three sources of "noise": photon noise, oscillations, and granulation. We find that the GP decomposition does reliably capture the correlated nature of the oscillation signal, as the observed shape of the oscillation RMS curve agrees with predictions (Figure 9). The total amplitude, however, is somewhat lower than expected, consistent with our analysis of the TESS photometry in Section 4.2 suggesting that the oscillation amplitudes are suppressed.

4.3.3. Radial Velocity Oscillation Amplitudes

Chaplin et al. (2019) lay out a detailed explanation of how one can predict the residual amplitudes of stellar p-mode oscillation signals for radial velocity observations given knowledge of the stellar properties. In the Chaplin et al. (2019) model, oscillation amplitudes are attenuated by the factor

$$\eta(\nu) = \operatorname{sinc}(\pi\nu t_e) \tag{22}$$

where ν is the oscillation frequency and t_e is the exposure time of the observation. This transfer function is accurate when calculating the amplitude attenuation for a single continuous exposure, or, equivalently, a sequence of back-to-back, uninterrupted exposures. But in practice, most radial velocity instruments cannot achieve a 100% duty cycle when taking multiple exposures; individual exposures will be separated by a nonzero readout time, during which the stellar oscillation signal is not being observed. In this case, the transfer function must be scaled by a more complicated function of the exposure time, readout time, and total number of exposures, N

$$\eta(\nu) = \frac{1}{N} \operatorname{sinc}(\pi \nu t_e) \sum_{j=0}^{N} e^{\pi i \nu (2j-N+1)(t_e+t_r)}.$$
 (23)

We show the full derivation of this equation in the Appendix.

The predicted residual oscillation amplitude for our NEID observations, shown in Figure 9, is calculated following the procedure outlined in Chaplin et al. (2019) but with the generalized form of the transfer function (Equation (23)) used in place of the simple sinc function (Equation (22)). Because the net oscillation signal is produced by many simultaneously beating modes that are stochastically excited and damped, this prediction only represents the median case. We expect the residual amplitude that we measure with NEID to deviate from this prediction somewhat, as it will depend on the exact interference pattern of the oscillations during the short, 5.5 hr observing window. To illustrate the expected distribution of the median case, we simulate an oscillation time series 10,000 times, copying the NEID observing cadence and exposure times, and randomizing the phases of the modes each time. We



Figure 7. Detrended NEID RV time series for HD 35833 (top) and residuals (bottom). The individual oscillation and granulation signals, separated using the GP decomposition method described in Section 4.3.1, are shown as black dashed (oscillation) and dotted (granulation) lines. The sum of these signals is shown as a solid blue line. The residual signal is calculated by interpolating the oscillation and granulation signals to the observation time stamps and subtracting these from the detrended RVs.

(The data used to create this figure are available.)

then bin the data and calculate residual amplitudes for each realization as described in Section 4.3.2, and we compute the 16–84, 2.3–97.7, and 0.2–99.8 percentile ranges to place 1σ , 2σ , and 3σ constraints on the residual amplitude we expect to observe in each bin.

As we show in Figure 9, the residual amplitude of the observed oscillation signal is significantly lower than what we predict, falling outside of the 3σ contours for most integration times $\leq \nu_{\text{max}}^{-1}$. As with the photometric data, these results suggest that the oscillations were suppressed.

We fit for the reduction in amplitude by taking the ratios of the predicted and observed residual amplitude curves in the range 2 minutes \leq integration time \leq 30 minutes; we do not expect the data to be sensitive to variations at the \ll 1 m s⁻¹ level even when binning dozens of observations. The mean of the ratios, weighted by the uncertainty on the oscillation RMS, is 2.58 in this range. That is, the observed residual amplitude is 2.58 times lower than the predicted value. Applying this scale factor to our model for the radial velocity amplitude yields $A_{\rm RV} = 1.11 \pm 0.09$ m s⁻¹. We recognize that this is an inferred value, based on the observed residual amplitude of the time series rather than the measured amplitudes of individual modes, and that we should not expect to precisely recover the true value of $A_{\rm RV}$ given the short baseline of the NEID time series. But the discrepancy between the inferred RV oscillation amplitude and the predicted value is nevertheless significant, as we illustrate in Figure 9.

5. Discussion

5.1. Causes of Amplitude Suppression

The oscillation amplitudes detected in both the TESS light curves and the NEID RV time series are weaker than anticipated. This discrepancy is marginal in the photometry, at $< 2\sigma$, but significant in the RV residual amplitude data. To better understand the underlying cause, we first consider the precision and accuracy of the stellar parameters. We use these parameters as inputs to the scaling relations from which the predicted amplitudes are derived; if the inputs are inaccurate it is possible that the observed oscillation amplitudes are not inconsistent with the theoretical scaling relations. But if the stellar parameters are reasonably accurate, this points to shortcomings of the scaling relations, or of our understanding of the physics of p-mode oscillations, as the reason for the observed amplitude discrepancy.

5.1.1. Accuracy of Stellar Parameters

We reparameterize Equations (3) and (5) in terms of T_{eff} and log g to reduce the number of free parameters in the amplitude



Figure 8. RV RMS as a function of integration time for NEID observations of HD 35833. We show the total observed RMS (red points), the photon noise as calculated from the information content of the spectra (black dashed line), the RMS of the decomposed oscillation (black dotted–dashed line) and granulation (black dotted line) signals, and the sum of the three sources of RV "noise" (blue solid line). The shape of the oscillation signal is consistent with predictions from the Chaplin et al. (2019) model.

scaling relations

$$A_{\rm phot} = A_{\rm phot,\odot} \beta \left(\frac{g}{g_{\odot}}\right)^{-1} \left(\frac{T_{\rm eff},\star}{T_{\rm eff,\odot}}\right)^2$$
(24)

and

$$T_{\rm red} = 8907 \left(\frac{R_{\star}}{R_{\odot}}\right)^{-0.186} \left(\frac{T_{\rm eff}, \star}{T_{\rm eff,\odot}}\right)^{-0.372} K$$
(25)

where β still depends on T_{red} as in Equation (4). The scaling relation for the RV oscillation amplitude given in Equation (7) can similarly be reparameterized as

$$A_{\max, RV} = A_{\max, RV, \odot} \left(\frac{g}{g_{\odot}}\right)^{-1} \left(\frac{T_{\text{eff}}, \star}{T_{\text{eff}, \odot}}\right)^{4}.$$
 (26)

We calculate the predicted values of A_{phot} and A_{RV} for values of $T_{\rm eff}$ and $\log g$ on the ranges 4500 K < $T_{\rm eff}$ < 7000 K and $3.2 < \log g < 4.2$, holding the only remaining free parameter, R_{\star} , fixed to the value given in Table 1. We show the relative difference between the predicted and observed values of A_{phot} in Figure 10 alongside an equivalent plot for ν_{max} . We also show the relative difference between the predicted and observed values of $\Delta \nu$ as a function of stellar mass and radius. Whereas the observed values of ν_{max} and $\Delta \nu$ agree well with predictions, the amplitudes do not. The RV amplitude, in particular, is not consistent with $< 10\sigma$ for any reasonable set of stellar parameters. Furthermore, it is readily apparent that a change in log g by more than ~ 0.1 dex would be incompatible with the observed value of ν_{max} . Uncertainties in the stellar parameters are therefore rather unlikely to be responsible for the apparent suppression of the oscillation amplitudes.

5.1.2. Magnetic Suppression of p-mode Oscillations

Detailed helioseismic studies have shown that the frequencies and amplitudes of solar p-mode oscillations vary with



Figure 9. Predicted (orange) and observed, decomposed (black) residual amplitude for the decomposed p-mode oscillation signal in the NEID RVs for HD 35833. While the shapes are the same, the amplitude of the observed signal is suppressed relative to predictions by a factor of 2.58 (blue). We also indicate the predicted value of $\nu_{\rm max}^{-1}$ as a gray vertical line.

the solar activity cycle (Chaplin et al. 2000; Jiménez-Reyes et al. 2003), with higher levels of activity correlating with weaker oscillation amplitudes. And in more recent years, studies of individual stars (e.g., García et al. 2010; Bonanno et al. 2019) and large stellar samples (Chaplin et al. 2011b; Bonanno et al. 2014; Mathur et al. 2019) have revealed this same anticorrelation in stars other than the Sun. These findings suggest that stellar activity can lead to the suppression of p-mode oscillations, a prediction that is corroborated by solar magnetoconvection models (Cattaneo et al. 2003) and empirical results from spatially resolved observations of oscillations in the vicinity of sunspots (Braun et al. 1987, 1988).

We use the Ca II H & K S-index as a proxy for chromospheric activity to assess whether magnetic suppression is the mechanism responsible for the weaker-than-expected RV oscillation signal observed for HD 35833. As Bonanno et al. (2014) show, high S-index values are correlated with smaller oscillation amplitudes in the Kepler bandpass; we expect the same to be true for TESS, of course, and for NEID as well, as photometric and radial velocity oscillation signals are simply different manifestations of the same physical process. However, measurements of the S-index for HD 35833 show that this star is not magnetically active. Isaacson & Fischer (2010) calculate S-index values for HD 35833 for 14 nights of Lick observatory data collected across 4.5 yr from 2004 to 2009 as part of the California Planet Search survey; we show these measurements in Figure 11. The median S-index for the Lick time series is 0.16, and the values range from 0.14–0.19. This is consistent with the values reported by Isaacson & Fischer (2010) for very quiet subgiants with similar B - V color, indicating that HD 35833 was not exhibiting high levels of activity at the time.

We also consider present-day activity levels for HD 35833, as this is the more pertinent metric in the context of the NEID observations. The S-indices for the NEID time series are calculated using a modified version of the NEID DRP, in which the flux in the line cores and reference regions is determined using a weighted mean of multiple overlapping orders rather than a single order. The median S-index for the NEID observations of HD 35833 is 0.11, with only a modest variation



Figure 10. The relative difference between the predicted and observed values of ν_{max} , $\Delta\nu$, A_{phot} , and A_{RV} for grids of possible stellar parameters for HD 35833. The dark swaths indicate regions of parameter space for which the predicted values would be consistent with observations. The T_{eff} and log g values (M_{\star} and R_{\star} values for $\Delta\nu$) given in Table 1 are shown as a white circle (black for A_{RV}), with error bars for the formal uncertainties. We note that the observed values are consistent with predictions for ν_{max} and $\Delta\nu$, and $<5\sigma$ discrepant for A_{phot} . But there is no reasonable set of stellar parameters that is simultaneously compatible with the aforementioned values as well as with the observed value of A_{RV} .



Figure 11. Ca II H & K S-index measurements for HD 35833 from Lick observations (Isaacson & Fischer 2010). The median value of 0.16 shows that activity-induced suppression cannot explain the observed p-mode oscillation amplitudes.

throughout the night (Figure 12). While these values are not calibrated to the same scale as the Isaacson & Fischer (2010) sample and thus cannot be directly compared, we confirm that the star remained quiet by noting the absence of emission features in the Ca II H & K line cores in the NEID spectra. This is consistent with our expectation that activity levels for subgiants should not change significantly on decade-long time-scales (Wright 2004). HD 35833 was not exhibiting high levels



Figure 12. Ca II H & K S-index measurements for HD 35833 from the NEID time series.

of activity on the night it was observed with NEID, and the oscillations were not magnetically suppressed.

5.1.3. Other Causes of Amplitude Suppression

Mathur et al. (2019) explore the possibility that metallicity may be responsible for a reduction in oscillation amplitudes in a sample of Kepler stars for which no strong magnetic activity signals are detected. A dependence on metallicity was first suggested by Houdek et al. (1999), who point to the impact of opacity on convective velocities and find that lower metallicity stars are indeed expected to have weaker oscillation amplitudes. This finding is supported by more detailed 3D hydrodynamical simulations by Samadi et al. (2010). However, the results of the Mathur et al. (2019) analysis are inconclusive; while the stars without detected oscillations have subsolar metallicities on average, a significant number of them have [Fe/H] > 0. With the lack of observational evidence confirming the expected effect of metallicity, we refrain from considering it as the source of oscillation suppression in the case of HD 35833.

In using the correction factor c from Kjeldsen et al. (2008) to calculate the expected oscillation amplitudes in Section 4.1, we implicitly assume that the relative visibilities of different modes in HD 35833 are similar to those of the Sun. However, multiple studies have shown that the visibilities of the dipolar (l=1)oscillation modes may be significantly reduced or absent altogether for some evolved stars (Mosser et al. 2012; García et al. 2014; Stello et al. 2016). If these modes are not present in HD 35833, the overall oscillation amplitudes will be weaker than the predicted values given in Table 2. To determine whether this might explain the observed RV amplitude suppression, we recalculate the predicted residual amplitude signal for the NEID time series in the extreme case that the visibility $S_1 = 0$. We find that while the ratio between the predicted and observed residual amplitudes is not as drastic in this case ($A_{\rm RV}$. $predicted/A_{RV,observed} = 1.97$), the discrepancy is still significant at the $>5\sigma$ level, indicating that missing dipolar modes cannot be the sole cause of the observed suppression.

5.2. Implications for Radial Velocity Exoplanet Detection

The agreement between the predicted and observed shapes of the oscillation RV RMS for HD 35833 (Figure 9) empirically validates the Chaplin et al. (2019) residual amplitude model for integration times $t_i < \nu_{\text{max}}^{-1}$ for the first time. This complements the Chaplin et al. (2019) analysis of α Cen A RVs, for which the residual amplitudes show good agreement with the model at integration times $t_i > \nu_{\text{max}}^{-1}$. In addition, our analysis in Section 4.3.1 highlights the useful application of the Luhn et al. (2022) Gaussian process kernels in fitting and removing correlated RV signals. These results support the use of the Chaplin et al. (2019) attenuation model and the Luhn et al. (2022) correlated noise model to inform exposure times and observing strategies for EPRV exoplanet surveys.

At the same time, we must recognize that there are gaps in our current understanding of the physical processes governing p-mode oscillations, and that we have yet to understand these signals at the 10 cm s⁻¹ level. The significant discrepancy between the residual RV oscillation amplitudes we detect and the residual amplitudes predicted by scaling relations is cause for concern, as it indicates there may be additional physics not captured by these relations. Though this mechanism suppressed the oscillations in HD 35833, it may in other cases excite stronger oscillations, making it more difficult to detect exoplanets with small semiamplitudes. In addition, the TESS and NEID data for HD 35833 reveal a very different ratio between photometric and RV amplitudes than scaling relations predict, suggesting that we should be more cautious about using photometric data to inform estimates of the RV noise contribution from oscillations. Inaccurate or imprecise models of stellar variability will also inhibit robust mass and orbit measurements for any low-amplitude exoplanet signals that are detected. Even with a reliable strategy in place for mitigating the noise contribution of p-mode oscillations, we must continue to undertake detailed studies of these signals if we aim to push the limits of RV exoplanet detection and characterization.

6. Conclusion

We report separate detections of p-mode oscillations in the subgiant star HD 35833 with NEID and TESS. Independent analyses of both data sets show that the oscillation amplitudes are significantly weaker than expected from scaling relations, yet we are unable to link this finding to any known suppression mechanisms. The NEID data also validate the Chaplin et al. (2019) amplitude attenuation model for RV observations with exposure times $> \nu_{\text{max}}^{-1}$, resolving a critical unknown in the design of current and future EPRV exoplanet surveys.

We failed to detect oscillations simultaneously in RVs and photometry, because the elevated white noise floor in the Sector 32 TESS data precluded the detection of the p-mode signal. While these data were not wholly uninformative, as upper limits on the oscillation amplitude in Sector 32 are consistent with predicted amplitudes, future analyses in this same vein would still be valuable. If a tight relation between photometric and RV oscillation amplitudes can be identified, observers can develop more informed RV observing strategies by using long-baseline photometric data (e.g., from TESS or its successor) to more accurately predict the strength of the oscillation signal in RVs. We also invite readers to make use of the data sets presented herein for future studies of oscillations involving larger ensembles of stars.

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Facilities: TESS, WIYN (NEID), Gaia.

Software: astropy (Astropy Collaboration et al. 2018), barycorrpy (Kanodia & Wright 2018), lightkurve (Lightkurve Collaboration et al. 2018), matplotlib (Hunter 2007), numpy (Harris et al. 2020), scipy (Oliphant 2007), SERVAL (Zechmeister et al. 2018)

Appendix Residual Radial Velocity Amplitudes for Sequences of Exposures

Here, we describe the analytic calculation of the residual RV amplitude of a stellar oscillation signal for a sequence of exposures separated in time. We begin with the simple case of a single exposure, and then generalize this to multiple exposures. For a single, continuous observation of duration t_e , the attenuation of the oscillation signal, or the transfer function, is given by the Fourier transform of a rectangular pulse centered on zero with width t_e and unit height (Figure 13).

$$\eta(\nu) = \frac{1}{t_e} \int_{-\frac{1}{2}t_e}^{\frac{1}{2}t_e} e^{-2\pi i\nu t} dt$$
$$= \frac{1}{\pi \nu t_e} \frac{e^{\pi i\nu t_e} - e^{-\pi i\nu t_e}}{2i}$$
$$= \frac{1}{\pi \nu t_e} \sin(\pi \nu t_e)$$
$$= \operatorname{sinc}(\pi \nu t_e).$$
(A1)

In this case, the transfer function is equivalent to a simple sinc function. For a sequence of observations each with exposure time t_e and separated by readout time t_r , the transfer function is the Fourier transform of a sequence of rectangular pulses collectively centered on zero (Figure 14). To calculate the Fourier transform, we parameterize the phase offsets using the variable

$$k = 2j - N + 1 \tag{A2}$$

where N is the total number of observations and j is the index of a given observation. The transfer function is then

$$\eta(\nu) = \frac{1}{Nt_e} \sum_{j=0}^{N} \int_{-\frac{1}{2}t_e + \frac{k}{2}(t_e + t_r)}^{\frac{1}{2}t_e + \frac{k}{2}(t_e + t_r)} e^{-2\pi i \nu t} dt$$
$$= \frac{1}{Nt_e} \sum_{j=0}^{N} \frac{1}{\pi \nu} \frac{e^{\pi i \nu t_e} - e^{-\pi i \nu t_e}}{2i} e^{k\pi i \nu t_e} e^{k\pi i \nu t_e} e^{k\pi i \nu t_r}$$
$$= \frac{1}{N} \operatorname{sinc}(\pi \nu t_e) \sum_{j=0}^{N} e^{\pi i \nu k (t_e + t_r)}.$$
(A3)

Substituting j and N back in for k using Equation (A2), this yields

$$\eta(\nu) = \frac{1}{N} \operatorname{sinc}(\pi \nu t_e) \sum_{j=0}^{N} e^{\pi i \nu (2j - N + 1)(t_e + t_r)}.$$
 (A4)



Figure 13. The transfer function for a single finite exposure.



Figure 14. The transfer function for a sequence of six exposures separated by nonzero readout time.

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