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A LOCAL EARTHQUAKE CODA MAGNITUDE AND ITS RELATION
TO DURATION, MOMENT $M_o$, AND LOCAL RICHTER MAGNITUDE $M_L$

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ABSTRACT

A relationship was found between the seismic moment, $M_o$, of shallow local earthquakes and the total duration of the signal, $t$, in seconds, measured from the earthquake origin time, assuming that the end of the coda is composed of backscattering surface waves due to lateral heterogeneity in the shallow crust following Aki. Using the linear relationship between the logarithm of $M_o$ and the local Richter magnitude $M_L$, we obtain a relationship between $M_L$ and $t$, of the form: $M_L = a_0 + a_1 \log t + a_2 t^{1/3} + f(t)$, where $a_0$, $a_1$, $a_2$ are constants depending on $Q$ and geometric spreading, and $f(t)$ is a function of the instrument response and a (weak) function of the scattering process. This relationship is different from the empirical one generally used $M_L = a_0 + a_1 \log \tau + a_2 (\log \tau)^2 + a_3 \Delta$, where $\tau$ is the duration measured from the first P arrival time and $\Delta$ is epicentral distance in km. In the theoretical relationship, the dependence on epicentral distance is implicit in $t$. This relationship has been used to calculate a coda magnitude $M_C$, which has been compared to $M_L$ for Southern California earthquakes which occurred during the period from 1972 to 1975. This comparison has been made independently at six stations of the CIT network. At all stations, a good linear fit ($M_L = C_0 + C_1 M_C$) has been obtained. The standard errors range from .2 to .3 and the correlation coefficients from .80 to .90. Once station gain is accounted for, station correction terms are less than 0.17 magnitude unit when comparing $M_L$ and $M_C$. $M_C$ estimation is not limited to a duration measurement but can utilize the entire earthquake coda. The use of several measurements per seismogram produces a more robust estimate of earthquake size.
INTRODUCTION

With the advent of computerized local earthquake analysis systems using large seismographic networks, new methods of earthquake magnitude determinations are needed. Current methods of magnitude determination for local earthquakes, estimations that require the amplitude of body waves or the duration of the coda, are not generally useful for computerized systems. In larger earthquakes, body wave amplitudes can exceed the linear range of the recording system and, in some automated systems, recording can cease before the end of the coda. The intent of this paper is to develop a method of magnitude estimation that can utilize any part of the coda, some of which will certainly be preserved in the linear range of computer recording systems. It is shown how this coda magnitude method relates to other estimates of earthquake size by the use of coda duration, moment $M_o$, and local Richter magnitude $M_L$.

An empirical correlation between the magnitude of earthquakes, $M$, and the duration of the recorded signal, $\tau$, has consistently been observed. Linear relationships between $M$, $\log \tau$ (in seconds), the epicentral distance, $\Delta$ (in km), and the focal depth, $h$ (in km), have been empirically determined for various geographical areas, using multiple regression analysis techniques.

The first study was made by Bisztricsany (1958), who found a linear relationship between the magnitude of teleseisms (in the range of 5 to 8), the logarithm of the duration of the surface wave train, and the epicentral distance. Later on, this method has been applied to local earthquakes (of epicentral distances less than 100 km), with shallow focus (focal depth less than 60 km), and the duration $\tau$ has been defined as the total
length of the record, instead of the length of the surface wave train only. In chronological order, such studies have been due to Solov'ev (1965) for Sakhalin Island; Tsunura (1967) for Japanese earthquakes recorded at the network of the Wakayama Microearthquake Observatory, and whose magnitudes were determined by the Japan Meteorological Agency; Crosson (1972) for earthquakes in the Puget Sound Region, using magnitudes calculated from Wood-Anderson records, Lee et al., (1972) for central California earthquakes with Richter magnitude $M_L$; Real and Teng (1973) for Southern California also with $M_L$; Hermann (1975) for the central United States and body wave magnitude determinations, Bakun and Lindh (1976) for the Oroville, California, region, and $M_L$. Similar studies have been made for the Mica Array, Canada, the Santiago station in Chile, the Alaska network and the University of Utah seismograph stations, U.S.A., etc. ...(see Lee and Wetmiller, 1976).

For shallow earthquakes, focal depths are much smaller than epicentral distances and much less accurately determined. As a result, it is harder to investigate the dependence of duration on focal depth than on epicentral distance. No such dependence has ever been found in a conclusive way.

The local magnitude $M_L$, introduced by Richter (1935), is given by:

$$M_L = \log A - \log A_0 (\Delta, \text{km})$$

where $A$ is the maximum amplitude in mm on the Wood-Anderson seismogram for an earthquake at distance $\Delta$ km, and $A_0$ is that for a particular earthquake selected as standard. (All logarithms used here are common logarithms to the base 10.) The magnitudes determined by this method are usually above 2.5, so that, for smaller earthquakes, an equivalent magnitude has to be used. It is obtained by reading the maximum amplitude recorded by a short-period high gain seismometer and
conversion into the amplitude that would have been recorded if a Wood-Anderson instrument had been used.

When plotting the local magnitude $M_L$ versus the logarithm of duration, a slight curvature is consistently observed (Lee et al., 1972, Figure 4, Real and Teng, 1973, Figures 5 and 6; Bakun and Lindh, 1976, Figure 4a; Lee and Wetmiller, 1976, page 23). An example for the station MWC is shown in Figure 1 from Real and Teng (1973). The slope increases with increasing magnitude. Consequently, a better fit of the data is obtained by introducing a quadratic term in $\log \tau$ in the magnitude --- duration relation. Therefore, the general form of the empirical relation between magnitude and duration is given by:

$$M = a_0 + a_1 \log \tau + a_2 (\log \tau)^2 + a_3 \Delta,$$  

where $a_0, a_1, a_2$ and $a_3$ are constants. Reviews of existing formulas have been given by Real and Teng (1973) and Bakun and Lindh (1976). The standard deviations are in the range of .1 to .2, the estimated standard errors, of .2 ~ .3. The dependence on epicentral distance is small and always negligible for distances less than 200 km (e.g., Tsumura, 1967; Crosson, 1972).

As established by Tsumura (1967), the errors in the duration magnitude due to subjective errors in reading durations do not exceed ± .3 magnitude unit. The main source of observational error is the uncertainty in determining the end of the earthquake signal. According to Real and Teng (1973), it is fixed with an uncertainty of about 20%. It results in an approximately constant uncertainty of $\log \tau$ over the range of magnitude of interest, .15 unit of magnitude. In the same study, it has been established that there is no azimuthal dependence, as long as a sufficient number of events covering a large area and recorded on a long period of
time (say, a few years) is analyzed, and that the effect of variations in microseismic noise is insignificant in Southern California.

**Theoretical Relationship Between M and Duration**

In this paper, we try to find a theoretical justification for the empirical relationships existing between local magnitude, duration of the signal, and epicentral distance. Aki (1969) has proposed a theory in which the coda of local earthquakes is composed of backscattered waves due to lateral heterogeneity in the crust. He has observed that the power spectra of coda waves at a given time measured from the earthquake origin time ("lapse time") appear to be nearly independent of the epicentral distance and of the nature of direct wave path between station and epicenter, as can be seen by comparing the records of the same event at different stations (Aki, 1969; Aki and Chouet, 1975). Aki (1969) has constructed his model for coda waves on the following three assumptions:

1. The distribution of the scatterers is two-dimensional over the earth's surface, random, and uniform.
2. The primary waves and the secondary waves are surface waves of the same kind, and a constant group velocity independent of frequency is assumed for them, for the sake of simplicity.
3. The distance between station and scatterer is on the same order as the distance between hypocenter and scatterer and is much larger than the hypocentral distance. He makes use of the Born approximation by neglecting multiple scattering. He obtains the following relation:

\[
M_0 \left[2N(r_0)\right]^{1/2} |\phi_o(f_o,r_0)| = t^{1/2} e^{\left(\pi f_p t/Q\right)}
\]

\[
\left(-\frac{1}{Q} \frac{dt}{df_p}\right)^{1/4} A(t) \frac{1}{\sqrt{8}} I(f_p)^{-1}
\]

(2)
where \( M_0 \) is the seismic moment, \( N(r_0) \) is the number of scatterers within a radius of \( r_0 \), \( |\phi_o(f_p|r_0)| \) is the absolute value of the Fourier Transform of displacement due to secondary waves generated at a scatterer at a distance \( r_0 \) by a source of unit moment located at the same distance from the scatterer, \( A(t) \) is the average peak-to-peak amplitude on a seismogram estimated from a short time sample around a given absolute lapse time \( t \), \( I(f_p) \) is the instrument response, and \( Q \) is an apparent quality factor, which is assumed constant and includes the effects of intrinsic attenuation of the wave medium and scattering process. The peak frequency \( f_p \) was obtained by measuring the wave period directly on the seismogram at time \( t \). For his data (aftershocks of the Parkfield earthquake, June 28, 1966), Aki found the following empirical formula for the peak frequency \( f_p \) versus the corresponding lapse time \( t \):

\[
\frac{t}{100} = \left( \frac{f_p}{0.82} \right)^{-1.5}
\]  

(3)

where \( t \) is in seconds and \( f_p \) is in cps, in the range \( 30 \leq t \leq 1000 \). This relation seems to be independent of earthquake size, epicentral distance, and nature of the direct path between station and epicenter. In the following, we will assume that the same dispersion relation applies to earthquakes studied here. Let \( B(f_p) \equiv [2N(r_o)]^{1/2}|\phi_o(f_p|r_o)| \). The values of \( B(f_p) \) as a function of peak frequency have been empirically determined by Aki and are given in his Table 3 (Aki, 1969). Using (3), (2) becomes:

\[
M_o = B(f_p)^{-1} t^{11/12} c (\pi 17.67 t^{1/3}/Q) \frac{1}{5.24} Q^{-14} \cdot A(t) \left| I(f_p) \right|^{-1}
\]

(4)
This formula is strictly valid only in the range of frequencies which are lower than the corner frequency of the particular event under consideration. The effect of the radiation pattern has been neglected, which is a reasonable assumption because the scattered energy should represent a broad range of source azimuths. A linear relationship between the logarithm of seismic moment $M_o$ and the local magnitude $M_L$ for a given geographical area has been consistently established in California over a wide range of magnitudes ($0 < M_L < 6$) (e.g., Aki, 1969; Wyss and Brune, 1968; Thatcher and Hanks, 1973; Johnson and McEvilly, 1974; Bakun and Lindh, 1976). No systematic dependence of this $M_o$-$M_L$ relationship on either epicenter location or focal depth has ever been found. Uncertainties in the determination of magnitudes and variation in the geology at the recording site introduce an uncertainty factor of 2 or 3 in relating moment to magnitude (Wyss and Brune, 1968; Johnson and McEvilly, 1974).

We will express this relationship in the following way:

$$\log M_o = a + b M_L'$$

(5)

where $a$ and $b$ are constants. If we designate $M_L'$ as the coda magnitude $M_C$, (4) can then be rewritten as:

$$a + b M_C = -0.72 + \log A(t) + \log B(f_p)^{-1}$$

$$- \frac{1}{4} \log Q + \frac{11}{12} \log t + \frac{\pi}{Q} 17.67 \log e^{1/3} + \log |I(f_p)|^{-1}.$$  

(6)

Then we divide the term corresponding to the instrument effect into two parts: the gain $G$, and the normalized amplitude spectrum $R(f_p)$. We now have the relation for coda magnitude $M_C$:

$$M_C = \frac{1}{b} \left( -a - 0.72 + \log A(t) + \log R(f_p)^{-1} + \log B(f_p)^{-1} - \frac{1}{4} \log Q + \frac{11}{12} \log t + \frac{24.11}{Q} t^{1/3} \right).$$

(7)
Two different definitions of signal duration $t$ are commonly used. The time is always measured from the onset of the first P arrival. The difference occurs in the way of defining the end of the coda. It has been defined as the time when the signal returns to the noise level present before the arrival of the P wave (e.g., Crosson, 1972) (Criterion 1) or as the time when the trace amplitude (peak-to-peak) falls below a given level on some reading device (e.g., Lee et al., 1972) (Criterion 2).

If $t$ in expression (7) is the time when the coda ends, measured from the earthquake origin time, we have:

$$t = \frac{A}{\alpha} + \tau$$

where $\alpha$ is nearly a constant depending on the P wave velocity. With Criterion 2, $A$ in expression (7) is a constant. With Criterion 1, $A$ represents the average amplitude of the background noise. It is a function of the recording station, but the ratio $\frac{A}{G}$ is nearly constant for all stations in a given area, since an increase in the magnification of the instrument results in an increase in the average noise level.

Hermann (1975) has studied the influence of different cut-off levels for defining duration for earthquakes with magnitudes between 2.5 and 4.5 on the magnitude-duration relation. Duration measurements made at two different levels at each station yielded relationships differing from one another by only an additive constant and, within the limits of accuracy, the linear relation between magnitude and $\log \tau$ were equivalent at the two levels, as expected. In the following, we will consider $A$ as a constant.

(7) can be rewritten in the following way, when using (8):

$$M_C = \frac{1}{b} \left[ -a - 0.72 + \log \frac{A}{G} + \log R(f_p)^{-1} + \log B(f_p)^{-1} - \frac{1}{4} \log Q + \frac{11}{12} \log \left( \frac{A}{\alpha} + \tau \right) + \frac{24.11}{Q} \left( \frac{A}{\alpha} + \tau \right)^{1/3} \right].$$
In agreement with the empirical relationship (1) between $M_L$, $\tau$, and $A$, $M_L$ is found to be an increasing function of both $\tau$ and $A$, but $M_C$ in (9) is more complicated. In addition to the logarithmic term in duration, there is a power term; and the effect of the gain of the instrument has already been mentioned. For a given area, $Q$ has been considered as constant. Aki (1969) fixed it at 200. The term $\log R(f_p)^{-1}$ depends on the instrument type. As it is an implicit function of $\tau$ and $A$ (through (3) and (8)), the instrument is likely to influence all the coefficients in the empirical formula (1). This result is consistent with the observations of Bakun and Lindh (1976): They plotted $M_L$ vs. $\log \tau$ for three different instruments and found a different intercept and a different slope in each case (their Figure 4b). The USGS and ORV instruments, which have about the same normalized response below 5 Hz, gave about the same slope, whereas the ORV-LP instrument which has a normalized response very different from the two others gave a different slope (their Figure 2).
DATA ANALYSIS

In order to further check the duration-$M_L$ relationship, we now compare local magnitudes $M_L$ with $M_C$ from the theoretical expression (9) for events in Southern California recorded during the period from 1972 to 1975. The following stations of the telemetered seismic network operated by the Seismological Laboratory of the California Institute of Technology have been used: Isabella (ISA), Goldstone (GSC), Saddleback Butte (SBB), Mount Wilson (MWC), Twentynine Palms (TPC), and Palomar (PLM). Their locations in the Southern California Seismograph Network are shown in Figure 2. The signal was recorded by short-period vertical Benioff seismometers and registered on drum recorders revolving at a speed of 60 mm/min.

Durations $\tau$ were read using Criterion 2, the average peak-to-peak amplitude threshold $A$ being approximately 0.1 cm. The gain, $G$, was initially assumed to be the one at the station PAS, $G_{PAS}$, of $10^5$. The effect of different gains is discussed below. The coefficient $a$ of expression (7) was taken as $1.62$ sec/km, $Q$ fixed at 200 following Aki (1969) and the values of $B(f_p)$ were taken from Aki's Table 3, once $f_p$ had been determined from $\tau$ and $A$ using relations (3) and (7). The constants $a$ and $b$ in the moment-magnitude relation (5) that we used were also determined by Aki (1969) in his study of aftershocks of the Parkfield earthquake, for the magnitude range of 3 to 5.5 he found:

$$\log M_o = 15.8 + 1.5 M_L.$$  (10)

In this way, we have calculated a coda magnitude $M_C$ for each event at each of the six stations, as expressed by:

$$M_C = + 11.07 - \frac{1}{1.5} \log G_{PAS} + \frac{1}{1.5} \log R(f_p)^{-1}$$

$$+ \frac{1}{1.5} \log B(f_p)^{-1} + 0.61 \log (aA+\tau) + 0.08(aA+\tau)^{1/3}$$  (11)
and shown in Table 1 (for \( t < 60 \text{ sec} \), values of \( B(f_p) \) used in Table 1 are different from Aki, as discussed below). \( R(f_p) \) is the normalized instrument response of the short period Benioff seismometer. Figure 3 shows the correlation between \( M_L \) and \( M_C \). \( M_L \) ranged from 2.6 to 5.5. Below 2.6, most values of \( M_L \) have been obtained from short period seismograms instead of Wood-Anderson records and they are omitted.

The results of a least-square fit of the data for each station for \( M_L \geq 2.6 \) are given in Table 2, including also the geographical coordinates of the stations, the period of time during which the data have been obtained, the number of events used in the analysis, the standard error of estimate and the correlation coefficient. The standard errors are \( .2 \sim .3 \) which is comparable to what is usually obtained by fitting such data with an empirical relationship of the form of expression (1). The correlation coefficients range from .80 for MWC where there is much scatter in the data, particularly at low magnitudes, to .90 for ISA. The number of events is on the order of 300 \( \sim 400 \), depending on the station, and the data have been obtained between 1973 and 1975 for most stations, so that the azimuthal coverage is good and the events have a broad distribution in size, epicentral distance and source mechanism.
The relationship between $M_C$ and $M_L$ is put in a convenient form when it is written as:

$$M_L = C_0 + C_1 M_C$$

and these parameters are shown in Table 3. It can be seen that the slope $C_1$ is systematically too high taking values between 1.22 and 1.55, the average value being 1.4. We now evaluate what parameters of the theory can affect the slope.

Aki and Chouet have also proposed a second model for coda waves in which the seismic energy transfer was considered as a diffusion process. With the diffusion theory, the geometrical spreading term $t^{1/2}$ in equation (2) is replaced by $t^{3/4}$, i.e., a power intermediate to 1 (corresponding to body waves) and 1/2 (surface waves). It is easily seen that this would change only the coefficient of log $(aA + t)$ in the expression giving the coda magnitude (11), from .61 to .78. The diffusion theory applies only to the parts of the energy scattered in the backward direction. Like the single-scattering theory that we have used, it is a simplification, and probably, the actual power of the geometrical spreading term lies between the values of 1/2 (single-scattered surface waves) and 3/4 (diffusion process) and the calculated slopes are shown in Figure 4. The effect of incorporating a diffusion process theory would reduce the observed slopes from 1.4 to about 1.3, which is not a significant difference. The coda magnitude $M_C$ for each event recorded at each station has been calculated from expression (11). This expression makes use of the empirical dispersion law (3) for coda waves found by Aki for the Parkfield region (Aki, 1969). A priori, this law may not apply to our set of data and a
more rigorous treatment would necessitate a derivation of a similar law for the data used in this study. Any discrepancy between the actual dispersion law and the one we used may yield a slope different from 1 in the $M_L$ vs. $M_C$ curve and even to a curvature of this line. In order to get $M_C$, we also used the values of $B(f_p)$ tabulated by Aki (1969). These values were obtained by fitting the data in his Figure 3. There are not many data for peak frequencies between 0.2 and 1 cps, which is the range of frequency corresponding to the durations of most events with magnitudes between 2.6 and 5.5. Therefore, the function $B(f_p)$ is not well constrained and may be subject to reevaluation. However, the two most likely candidates to change the slope of the $M_L$ vs. $M_C$ correlations are either the $M_o$ vs. $M_L$ relationship or $Q$.

**$M_o$-$M_L$ Relationship**

Wyss and Brune (1968) have Fourier analyzed surface waves from 13 earthquakes in the Parkfield region and have found the following relationship between seismic moment $M_o$ and Richter magnitude $M_L$:

$$\log M_o = 1.4 \cdot M_L + 17.0,$$

for $3.2 < M_L < 5.5$. $M_o$ of Southern California earthquakes calculated by Thatcher and Hanks (1973) from long-period S wave spectral levels yielded the relation:

$$\log M_o = 1.5 \cdot M_L + 16.0,$$

for $M_L$ between 2.5 and 6. These two studies give a coefficient $b$ of 1.4 to 1.5, similar to the one found by Aki (equation 10). However, more recent studies in Central California have yielded a coefficient $b$ closer to 1. Johnson and McEvilly (1974) have estimated seismic moments
from the low-frequency levels of the spectra for 13 earthquakes with magnitudes between 2.4 and 5.1 located near the San Andreas fault and found the following relation:

$$\log M_o = (1.16 \pm 0.06)M_L + (17.60 \pm 0.28).$$

And for earthquakes near Oroville, Bakun and Lindh (1976) have found:

$$\log M_o = (1.21 \pm 0.03)M_L + (17.02 \pm 0.07)$$

for $0 < M_L < 6$. The value of 1.5 for the coefficient of $M_L$ in the moment-magnitude relation may be more appropriate for Southern California than the value of 1.2 found for Central California. However, a coefficient $b$ closer to 1 would explain the average slope $c_1$ higher than 1 found for our $M_L$ vs. $M_C$ curves. In fact if the actual value of $b$ was 1.2 instead of 1.5, this slope would be reduced from 1.4 to 1.1.

**Effective Q**

In our calculation of coda magnitude $M_C$, we have used a $Q$ of 200 (Aki, 1969). However, a lower value may be more realistic, since values of 50-200 at frequencies on the order of 1 Hz have been found for the $Q$ of coda waves in Central California and Western Japan (Aki and Chouet, 1975). At these frequencies (0.2 Hz - 2 Hz) which are dominant at the end of the coda of local earthquakes, the coda is probably made of backscattering surface waves from heterogeneities in the shallow, low $Q$ crust. If the actual average apparent $Q$ of coda waves of local earthquakes in Southern California was 70, a value favored by Aki and Chouet (1975), instead of 200, the observed slope $c_1$ of $M_L$ vs. $M_C$ relationship (12) would be reduced from 1.4 to about 1.1 as seen in Figure 5.

In order to investigate the influence of gain of the instrument and/or geology at the station site on the coda magnitude $M_C$, we now
consider the residuals $\Delta M = M_C - M_L$ at the various stations for a
well-sampled interval of magnitude, chosen around $M_L = 3.1$, given in
Table 3. If $G_{STA}$ is the gain at a station STA, by comparing (9) and
(11) we get:

$$M_C = \frac{1}{1.5} \log \frac{G_{STA}}{G_{PAS}} + M_L. \quad (13)$$

The values of the gain corrections ($AG = \frac{1}{1.5} \log \frac{G_{STA}}{G_{PAS}}$) are also
given in Table 3, as well as the values of $\Delta M - AG$. It is clear from
this table that the major contribution to the discrepancy between $M_L$
and $M_C$ is due to the effect of the gain at each particular station.

A coda excitation factor, depending on the surface geology of the station
site, (Aki, 1969) should have the same effect on the magnitude residuals
as the gain of the instrument. This factor can be 5 to 8 times larger
on sediments than on granite (Aki, 1969), which corresponds to a rela-
tive change of .5 to .6 in $\Delta M$. $M_C$ is overestimated at ISA by 0.17
and underestimated at SBB by 0.13, with the other stations intermediate
to these values. These corrections may be insignificant considering the
scatter of the data. All of the stations used in this study are sited
directly on crystalline bedrock (granite or metamorphic), with the excep-
tion of Palomar, which has a few meters of alluvium between the instrument
and the granitic bedrock. All are essentially on the ground surface
except for Isabella, which is about 70 meters in a tunnel in bedrock. ISA
has an anomalously high residual, $\Delta M - AG$, of .17 magnitude unit (Table 3)
which may be due to its particular setting in the Southern Sierra mountains.
Note that the excitation factor should more generally be a function of fre-
quency, so that it could not only shift the $M_L$ vs. $M_C$ curve but also distort
it.
We now compare the theoretical magnitude $M_C$ from equation (9) with the empirical curve of our Figure 1 found by Real and Teng (1973) for durations at station MWC: $M_T = 1.37 + 0.41 \log (T)^2$ (the distance term is neglected). If we assume the same parameters as used above in calculation of $M_C$, except for an allowance of variance of $Q$ for a good fit, we get the theoretical curve shown in Figure 1, which is calculated with a $Q$ of 500. (Because preliminary results have shown no significant variation in the scattering property term $B(f_p)$, it is assigned a constant value of $4.32 \times 10^{-25}$ which is of course dependent on $Q$. This has only minor effects on the shape of the low-magnitude end of the curve.) In the empirical relationship of Real and Teng, the curvature comes from a $(\log T)^2$ term whereas in the $M_C$ curve it comes from a $T^{1/3}$ term. It is interesting to note that expansion of $T^a$ in a series is:

$$T^a = 1 + \frac{2.3a}{1!} \log T + \frac{(2.3a)^2}{2!} (\log T)^2 + \ldots$$

Thus, the next term after $\log T$ is a $(\log T)^2$ term in parallel with Real and Teng's empirical fit. There is however a discrepancy between the duration data of Real and Teng and that analyzed here. It is represented by the fact that the former is better fit with a high $Q$ of about 500 whereas the latter requires a lower $Q$ of about 70. At the present this difference is not understood. The data of Real and Teng were gathered from 1969 to 1971 events and the data here are from 1972 through 1975. This presents the possibility of a change of $Q$ between the intervals. Alternatively, although the reading methods were supposedly the same, it is possible that variances in definition of the "end" of the coda for large and small events can occur between readers. Whether these or other
causes are responsible is a subject that deserves further study.

Because the $M_C$ magnitude is a function of instrument response, it is interesting to compare the theoretical $M_C$ with duration data of Bakun and Lindh (1977). They found a different duration magnitude relation for the instruments ORV, a Benioff seismometer, and ORV-LP, a long period instrument with peak gain at about 0.2 Hz. Figure 6 shows $M_C$ calculated with a $Q$ of 70 compared with the empirical piecewise linear fits of Bakun and Lindh to ORV and ORV-LP data. (A constant value of $B(f_p) = 6.24 \times 10^{-23}$ was used for the $M_C$ calculation.) The Benioff ORV data are fit well except at the break in the two linear empirical curves. But more importantly, the ORV-LP data show markedly shorter durations than ORV below $M_L = 5.5$ in agreement with the theoretical curves. The theoretical $M_C$ curve is not defined below about $M_L = 4.4$. This implies that, because of the drop off of ORV-LP instrument response with higher frequency above 6 Hz combined with smaller earthquake amplitudes, scattered energy drops below the noise level.

The development of the theory leading to the expression for $M_C$ in equation (9) was based on the assumption that the peak frequency $f_p$ was below the characteristic or "corner" frequency of the seismic source. This assumption begins to be compromised at $M_L$ values significantly below 2 because here the characteristic frequency is about 2 Hz (see, for example, Johnson and McEvilly, 1974). Further, since the evidence here indicates that the duration magnitude scale is a subset of the general coda magnitude $M_C$ in equation (7), it follows that extrapolation of duration magnitude scales to levels significantly below $M_L = 2$ is also suspect.

A further consideration that compromises the extension of duration and coda magnitude scales to small earthquakes is the evidence that
the later part of the coda is mainly composed of backscattered waves. Certainly this is not true at the time of the direct S-wave arrival. Depending on the length of what one considers "S-wave" on the seismogram, it is reasonable to suppose that within 5 to 10 seconds of the S-wave arrival we are dealing more with direct S-wave energy. For both this, and the characteristic frequency problem discussed above, extrapolations of duration and coda magnitude scales below about $M_L = 1.5$ should be carefully reevaluated.

In Table 1 are summarized the $M_C$ values calculated for the coda at $t$, the time from origin time, for 0.1 cm RMS amplitude at a station with $1 \times 10^5$ gain. $M_C$ values for Q's of 500, used in Figure 1, 200, used in Figure 3, and 70, used in Figure 6, are given to illustrate the effect of variable Q. In these tables, the scattering property $B(f_p)$ is assumed constant and adjusted such that $M_C$ values at $t = 60$ sec are equal. Again, the assumption that $B(f_p)$ is constant instead of variable as in Aki (1969) significantly affects only magnitudes below $M_L = 2.6$ and would result in a reduction of the lower values by about 0.3 magnitude units if the Aki values were used. Study of this parameter will be done in a future paper. An important point to note here is that Q can be an important factor in relating the coda magnitudes of small earthquakes to large earthquakes. The difference in coda magnitude can be as large as one unit over an $M_L$ range of 3 to 5 for published data in the literature. This result makes the calibration of individual stations for Q an important consideration in the establishment of a coda magnitude scale.

The use of $M_C$ values in Table 1 for any part of the coda involves only an additive term for the coda amplitude at the corresponding time.
as can be seen in equation (7). Different instrument gain requires an added constant in the same equation. Use of a different instrument response requires the removal of the Benioff response term in Table 1 or equation (7), \( \log R(f_p)^{-1} \), and addition of the equivalent term corresponding to the desired instrument response.
CONCLUSIONS

Backscattering surface waves due to lateral heterogeneity in the shallow crust appear to be a good explanation for the late part of the coda of local earthquakes and duration times. The scattering theory measures the moment $M_o$ and provides a physical basis for the observed empirical correlation between local Richter magnitude $M_L$ and total duration of the coda.

With the use of the scattering theory, we derive a coda magnitude $M_C$ that can be used at any part of the coda and, in particular, where durations are measured. We compare $M_L$ to the coda magnitude $M_C$ calculated from duration data at six stations of the CIT network. For a linear fit $M_L = C_0 + C_1 M_C$, we find slopes $C_1$ systematically higher than one with the use of our initial assumptions based on the parameters of Aki (1969). This high value of $C_1$ implies one, or more likely a combination, of the following factors: the coda waves may have a geometrical spreading factor intermediate between surface waves and body waves; the coefficient $b$ of 1.5 in the $M_L - M_o$ relationship may be too high; the apparent Q of 200 that is used is too high and a value of 70 may be more appropriate.

For the six bedrock stations analyzed here, a correction for station gain within the theoretical framework of $M_C$ removes all significant site correction factors from the duration data. The remaining site correction on "coda excitation" terms deviate less than 0.17 magnitude units from the average.

Comparison of the coda magnitude $M_C$ with the duration data of Real and Teng (1973) shows that the theoretical $M_C$ matches the curvature of the data that led Real and Teng to add a higher order term to their
fit of $M_L$ versus log $t$. However, a high $Q$ of 500 in the $M_C$ relation was necessary to fit their data. Comparison of $M_C$ with the duration data of Bakun and Lindh (1977) shows the effect of different instrument response on the duration data. With a $Q$ of 70, the $M_C$ calculation fits their short-period Benioff data from ORV and simultaneously fits the widely divergent long period ORV-LP data for magnitudes above $M_L = 4.4$.

The results here strongly support the theory that the coda is composed of scattered energy. However, that same theory implies that scattering assumptions do not hold in the early part of the S-wave coda. For this reason, it appears that the extrapolation of duration and coda magnitude scales below $M_L = 1.5$ should be carefully reevaluated.

The coda magnitude $M_C$ can be used either with coda duration data or with any portion of the coda that is not "close" to the direct S-wave arrival. The latter feature is important for the usefulness of $M_C$ in computerized seismic data systems where parts of the coda may exceed the system's linear range, preventing the calculation of a body wave magnitude, or may be cut off before the time where duration would normally be measured.

Most importantly, the coda magnitude $M_C$ has a physical basis that relates moment $M_o$, Richter magnitude $M_L$, duration, and the general shape of the coda to the station gain, instrument response, and effective $Q$. There is evidence here that $Q$ can vary significantly from station to station which may require a calibration of individual stations. The study of regional variation of $Q$ and other parameters that affect the theory of scattered coda energy will be the subject of future investigations.
ACKNOWLEDGEMENTS

The authors wish to thank Don Anderson, Keiiti Aki, Carl Johnson, Karen McNally and Bernard Minster for critically reviewing the manuscript and providing several helpful suggestions. The data used in this paper is due to many individuals at the Seismological Laboratory who maintain the network and process the data and their efforts are greatly appreciated. This work was supported by JPL/NASA contract No. 64598, NASA Grant No. NSG-5224, and U.S.G.S. contract No. 14-08-0001-16719. Contribution No. 3062, Division of Geological and Planetary Sciences, California Institute of Technology, Pasadena, California 91125.
REFERENCES


TABLE I

CALCULATION OF THE CODA MAGNITUDE $M_C$

<table>
<thead>
<tr>
<th>$t$ (sec)</th>
<th>$f_p$ (Hz)</th>
<th>$\log R(f_p)^{-1}$</th>
<th>$M_C$</th>
</tr>
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<tr>
<td></td>
<td></td>
<td></td>
<td>$Q = 500$</td>
</tr>
<tr>
<td>5</td>
<td>6.04</td>
<td>-4.98</td>
<td>1.82</td>
</tr>
<tr>
<td>7</td>
<td>4.83</td>
<td>-5.00</td>
<td>1.91</td>
</tr>
<tr>
<td>10</td>
<td>3.81</td>
<td>-4.97</td>
<td>2.03</td>
</tr>
<tr>
<td>12</td>
<td>3.37</td>
<td>-4.94</td>
<td>2.10</td>
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<td>15</td>
<td>2.90</td>
<td>-4.96</td>
<td>2.19</td>
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<td>20</td>
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<tr>
<td>60</td>
<td>1.15</td>
<td>-4.38</td>
<td>2.95</td>
</tr>
<tr>
<td>70</td>
<td>1.04</td>
<td>-4.31</td>
<td>3.05</td>
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<td>80</td>
<td>0.95</td>
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<td>3.14</td>
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<td>90</td>
<td>0.88</td>
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<td>3.23</td>
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<td>100</td>
<td>0.82</td>
<td>-4.08</td>
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<td>120</td>
<td>0.73</td>
<td>-3.99</td>
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<td>150</td>
<td>0.63</td>
<td>-3.82</td>
<td>3.62</td>
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<tr>
<td>200</td>
<td>0.52</td>
<td>-3.59</td>
<td>3.86</td>
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<tr>
<td>300</td>
<td>0.39</td>
<td>-3.28</td>
<td>4.20</td>
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<tr>
<td>400</td>
<td>0.33</td>
<td>-3.04</td>
<td>4.46</td>
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<tr>
<td>500</td>
<td>0.28</td>
<td>-2.89</td>
<td>4.64</td>
</tr>
<tr>
<td>600</td>
<td>0.25</td>
<td>-2.73</td>
<td>4.81</td>
</tr>
<tr>
<td>700</td>
<td>0.22</td>
<td>-2.60</td>
<td>4.95</td>
</tr>
</tbody>
</table>


**TABLE 2**

FIT OF $M_C = b_o + b_1 M_L$

<table>
<thead>
<tr>
<th>Station</th>
<th>Latitude</th>
<th>Longitude</th>
<th>Period</th>
<th>Number of Events With $M_L \geq 2.6$</th>
<th>$b_o$</th>
<th>$b_1$</th>
<th>Correlation Coefficient</th>
<th>Std Error of Estimate</th>
</tr>
</thead>
<tbody>
<tr>
<td>GSC</td>
<td>35°18.1' N</td>
<td>116°48.3' W</td>
<td>1973-1975</td>
<td>308</td>
<td>1.11</td>
<td>0.75 ± 0.03</td>
<td>0.86</td>
<td>0.23</td>
</tr>
<tr>
<td>ISA</td>
<td>35°39.8'</td>
<td>118°28.4'</td>
<td>1973-1975</td>
<td>288</td>
<td>1.27</td>
<td>0.73 ± 0.02</td>
<td>0.90</td>
<td>0.18</td>
</tr>
<tr>
<td>MWC</td>
<td>34°13.4'</td>
<td>118°03.5'</td>
<td>1972-1975</td>
<td>346</td>
<td>1.14</td>
<td>0.72 ± 0.03</td>
<td>0.80</td>
<td>0.26</td>
</tr>
<tr>
<td>PLM</td>
<td>33°21.2'</td>
<td>116°51.7'</td>
<td>1973-1975</td>
<td>402</td>
<td>1.58</td>
<td>0.64 ± 0.02</td>
<td>0.81</td>
<td>0.22</td>
</tr>
<tr>
<td>SBB</td>
<td>34°41.0'</td>
<td>117°49.5'</td>
<td>1974-1975</td>
<td>292</td>
<td>1.44</td>
<td>0.69 ± 0.03</td>
<td>0.85</td>
<td>0.21</td>
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<tr>
<td>TPC</td>
<td>34°06.4'</td>
<td>116°02.9'</td>
<td>1973-1975</td>
<td>298</td>
<td>0.96</td>
<td>0.82 ± 0.03</td>
<td>0.83</td>
<td>0.27</td>
</tr>
</tbody>
</table>
### Table 3

**Fit of** $M_L = C_0 + C_1 M_C$

<table>
<thead>
<tr>
<th>Station</th>
<th>Gain</th>
<th>$\Delta G$</th>
<th>$C_0$</th>
<th>$C_1$</th>
<th>$\Delta M$</th>
<th>$\Delta M - \Delta G$</th>
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<tbody>
<tr>
<td>GSC</td>
<td>4.80 E5</td>
<td>.46</td>
<td>-1.48</td>
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<td>.34</td>
<td>- .12</td>
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<tr>
<td>ISA</td>
<td>2.40</td>
<td>.26</td>
<td>-1.73</td>
<td>1.37</td>
<td>.43</td>
<td>+ .17</td>
</tr>
<tr>
<td>MWC</td>
<td>2.37</td>
<td>.25</td>
<td>-1.58</td>
<td>1.38</td>
<td>.27</td>
<td>+ .02</td>
</tr>
<tr>
<td>PLM</td>
<td>4.83</td>
<td>.46</td>
<td>-2.45</td>
<td>1.55</td>
<td>.46</td>
<td>+ .00</td>
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<tr>
<td>SBB</td>
<td>8.00</td>
<td>.61</td>
<td>-2.10</td>
<td>1.46</td>
<td>.48</td>
<td>- .13</td>
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<tr>
<td>TPC</td>
<td>3.76</td>
<td>.39</td>
<td>-1.17</td>
<td>1.22</td>
<td>.40</td>
<td>+ .01</td>
</tr>
</tbody>
</table>

* $G_{PAS} = 1.00 \text{ E}5$
FIGURES

Figure 1. Local Richter magnitude \( M_L \) versus log \( \tau \) (duration) for station MWC. Dashed line is empirical fit of Real and Teng (1973) to their data. Continuous line is the calculation of \( M_C \) using parameters described in the text. A small vertical offset is introduced to separate the curves.

Figure 2. Locations of the stations GSC, ISA, MWC, PLM, SBB, and TPC within the Southern California Seismographic Network.

Figure 3. Correlation between \( M_L \) and \( M_C \) for the stations GSC, ISA, MWC, PLM, SBB, and TPC. The parameters for the linear least square fits are given in Table 2.

Figure 4. The change in slope of \( M_C \) due to different geometrical spreading terms (following Aki and Chouet, 1977) for (a) \( \tau^1 \), backscattering body waves, (b) \( \tau^{3/4} \), diffusion process, and (c) \( \tau^{1/2} \), backscattering surface waves.

Figure 5. The change in slope of \( M_C \) due to different values of \( Q \). A value close to 70 fits the data of Figures 3 and 6. A \( Q \) of 500 fits the data of Figure 1.

Figure 6. Local Richter magnitude \( M_L \) versus log \( \tau \) (duration) for stations ORV and ORV-LP. Dashed line and squares are linear fits to data and data for ORV and ORV-LP, respectively, from Bakun and Lindh (1977). Solid lines are the theoretical coda magnitude \( M_C \) calculated with \( Q = 70 \) for ORV and ORV-LP to illustrate the effect of different instrument response. \( M_C \) for ORV-LP is not defined below \( M_L = 4.4 \) for this particular value of \( Q \).
Theoretical $M_C$

Observed (Real and Teng, 1973)

Fig. 1
Fig. 3
(a) Backscattering surface waves
(b) Diffusion process
(c) Backscattering body waves
Theoretical

Observed

ORV

ORV-LP

Fig. 6